

# Machine Learning 10-601

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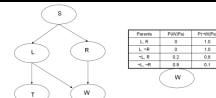
## Today:

- Inference in graphical models
- Learning graphical models

## Readings:

- Bishop chapter 8

## Bayesian Networks Definition



A Bayes network represents the joint probability distribution over a collection of random variables

A Bayes network is a directed acyclic graph and a set of conditional probability distributions (CPD's)

- Each node denotes a random variable
- Edges denote dependencies
- For each node  $X_i$  its CPD defines  $P(X_i | Pa(X_i))$
- The joint distribution over all variables is defined to be

$$P(X_1 \dots X_n) = \prod_i P(X_i | Pa(X_i))$$

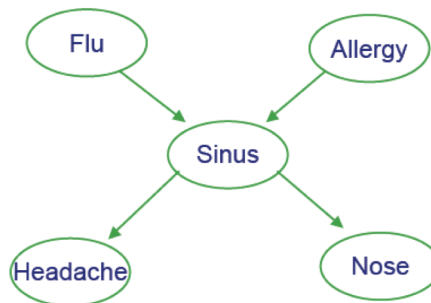
$Pa(X)$  = immediate parents of  $X$  in the graph

## Inference in Bayes Nets

- In general, intractable (NP-complete)
- For certain cases, tractable

## Example

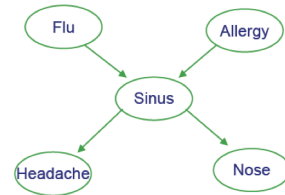
- Flu and Allergies both cause Sinus problems
- Sinus problems cause Headaches and runny Nose



## Prob. of joint assignment: easy

Suppose we are interested in joint assignment  $\langle F=f, A=a, S=s, H=h, N=n \rangle$

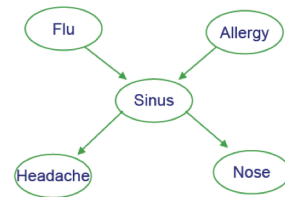
What is  $P(f,a,s,h,n)$ ?



let's use  $p(a,b)$  as shorthand for  $p(A=a, B=b)$

## Marginal probabilities $P(X_i)$ : not so easy

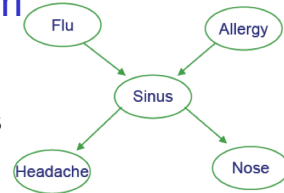
- How do we calculate  $P(N=n)$  ?



let's use  $p(a,b)$  as shorthand for  $p(A=a, B=b)$

## Generating a random sample from joint distribution: easy

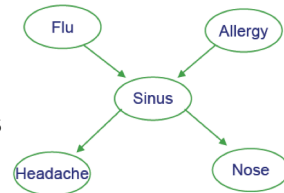
How can we generate random samples drawn according to  $P(F,A,S,H,N)$ ?



let's use  $p(a,b)$  as shorthand for  $p(A=a, B=b)$

## Generating a sample from joint distribution: easy

How can we generate random samples drawn according to  $P(F,A,S,H,N)$ ?



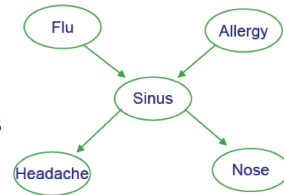
To generate a random sample for roots of network ( F or A ):

1. let  $\theta = P(F=1)$  # look up from CPD
2.  $r$  = random number drawn uniformly between 0 and 1
3. if  $r < \theta$  then output 1, else 0

let's use  $p(a,b)$  as shorthand for  $p(A=a, B=b)$

## Generating a sample from joint distribution: easy

How can we generate random samples drawn according to  $P(F,A,S,H,N)$ ?



To generate a random sample for roots of network ( F or A ):

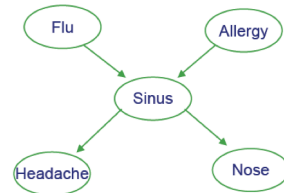
1. let  $\theta = P(F=1)$  # look up from CPD
2.  $r$  = random number drawn uniformly between 0 and 1
3. if  $r < \theta$  then output 1, else 0

To generate a random sample for S, given F,A:

1. let  $\theta = P(S=1|F=f,A=a)$  # look up from CPD
2.  $r$  = random number drawn uniformly between 0 and 1
3. if  $r < \theta$  then output 1, else 0

## Generating a sample from joint distribution: easy

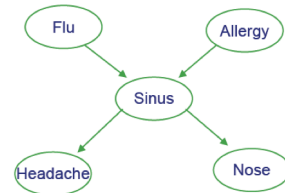
Note we can estimate marginals like  $P(N=n)$  by generating many samples from joint distribution, then count the fraction of samples for which  $N=n$



Similarly, for anything else we care about, calculate its maximum likelihood estimate from the sample  $P(F=1|H=1, N=0)$

→ weak but general method for estimating any probability term...

## Generating a sample from joint distribution: easy



We can easily sample  $P(F,A,S,H,N)$

Can we use this to get  $P(F,A,S,H | N)$ ?

Directly sample  $P(F,A,S,H | N)$ ?

## Gibbs Sampling:

Goal: Directly sample conditional distributions

$$P(X_1, \dots, X_n | X_{n+1}, \dots, X_m)$$

Approach:

- start with the fixed observed  $X_{n+1}, \dots, X_m$  plus arbitrary initial values for unobserved  $X_1^{(0)}, \dots, X_n^{(0)}$
- iterate for  $s=0$  to a big number:

$$X_1^{s+1} \sim P(X_1 | X_2^s, X_3^s \dots X_n^s, X_{n+1}, \dots, X_m)$$

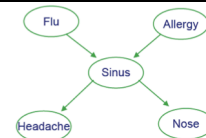
$$X_2^{s+1} \sim P(X_2 | X_1^{s+1}, X_3^s \dots X_n^s, X_{n+1}, \dots, X_m)$$

...

$$X_n^{s+1} \sim P(X_n | X_1^{s+1}, X_2^{s+1}, \dots, X_{n-1}^{s+1}, X_{n+1}, \dots, X_m)$$

Eventually (after burn-in), the collection of samples will constitute a sample of the true  $P(X_1, \dots, X_n | X_{n+1}, \dots, X_m)$

\* but often use every 100th sample, since iters not independent



## Gibbs Sampling:

Approach:

- start with arbitrary initial values for  $X_1^{(0)}, \dots, X_n^{(0)}$  (and observed  $X_{n+1}, \dots, X_m$ )

- iterate for  $s=0$  to a big number:

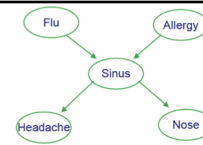
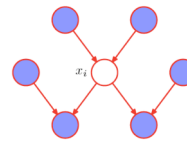
$$X_1^{s+1} \sim P(X_1 | X_2^s, X_3^s \dots X_n^s, X_{n+1}, \dots, X_m)$$

$$X_2^{s+1} \sim P(X_2 | X_1^{s+1}, X_3^s \dots X_n^s, X_{n+1}, \dots, X_m)$$

...

$$X_n^{s+1} \sim P(X_n | X_1^{s+1}, X_2^{s+1}, \dots, X_{n-1}^{s+1}, X_{n+1}, \dots, X_m)$$

Only need Markov Blanket at each step!



Gibbs is special case of Markov Chain Monte Carlo method

## Prob. of marginals: not so easy

But sometimes the structure of the network allows us to be clever → avoid exponential work

eg., chain

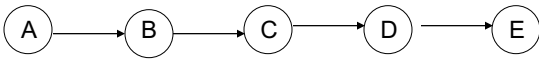


what is  $P(C=1 | B=b, D=d)$ ?

what is  $P(C=1)$  ?

## Variable Elimination example

But sometimes the structure of the network allows us to be clever → avoid exponential work

eg., chain   
what is  $P(C=1)$  ?

## Inference in Bayes Nets

- In general, intractable (NP-complete)
- For certain cases, tractable
  - Assigning probability to fully observed set of variables
  - Or if just one variable unobserved
  - Or for singly connected graphs (ie., no undirected loops)
    - Variable elimination
- Can often use Monte Carlo methods
  - Generate many samples, then count up the results
  - Gibbs sampling (example of Markov Chain Monte Carlo)
- Many other approaches
  - Variational methods for tractable approximate solutions
  - Junction tree, Belief propagation, ...

see Graphical Models course 10-708



## Learning Bayes Nets from Data

### Learning of Bayes Nets

- Four categories of learning problems
  - Graph structure may be known/unknown
  - Variable values may be fully observed / partly unobserved
- Easy case: learn parameters when graph structure is *known*, and training data is *fully observed*
- Interesting case: graph *known*, data *partly observed*
- Gruesome case: graph structure *unknown*, data *partly unobserved*

## Learning CPTs from Fully Observed Data

- Example: Consider learning the parameter

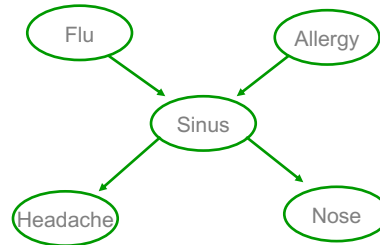
$$\theta_{s|ij} \equiv P(S = 1 | F = i, A = j)$$

- MLE (Max Likelihood Estimate) is

$$\theta_{s|ij} = \frac{\sum_{k=1}^K \delta(f_k = i, a_k = j, s_k = 1)}{\sum_{k=1}^K \delta(f_k = i, a_k = j)}$$

$k^{\text{th}}$  training example

$\delta(X) = 1$  if  $X$  is true  
0 otherwise



- Remember why?

let's use  $a_k$  to represent value of  $A$  on the  $k$ th example

## MLE estimate of $\theta_{s|ij}$ from fully observed data

- Maximum likelihood estimate  
 $\theta \leftarrow \arg \max_{\theta} \log P(\text{data} | \theta)$

- Our case:

$$P(\text{data} | \theta) = \prod_{k=1}^K P(f_k, a_k, s_k, h_k, n_k)$$

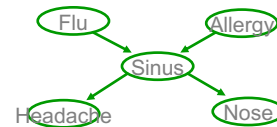
$$P(\text{data} | \theta) = \prod_{k=1}^K P(f_k) P(a_k) P(s_k | f_k a_k) P(h_k | s_k) P(n_k | s_k)$$

$$\log P(\text{data} | \theta) = \sum_{k=1}^K \log P(f_k) + \log P(a_k) + \log P(s_k | f_k a_k) + \log P(h_k | s_k) + \log P(n_k | s_k)$$

$$\frac{\partial \log P(\text{data} | \theta)}{\partial \theta_{s|ij}} = \sum_{k=1}^K \frac{\partial \log P(s_k | f_k a_k)}{\partial \theta_{s|ij}}$$

$$\theta_{s|ij} = \frac{\sum_{k=1}^K \delta(f_k = i, a_k = j, s_k = 1)}{\sum_{k=1}^K \delta(f_k = i, a_k = j)}$$

let's use  $a_k$  to represent value of  $A$  on the  $k$ th example

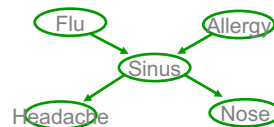


## MLE for $\theta_{s|ij} = P(S = 1|F = i, A = j)$ from fully observed data

- Maximum likelihood estimate

$$\theta \leftarrow \arg \max_{\theta} \log P(\text{data}|\theta)$$

$$\theta_{s|ij} = \frac{\sum_{k=1}^K \delta(f_k = i, a_k = j, s_k = 1)}{\sum_{k=1}^K \delta(f_k = i, a_k = j)}$$



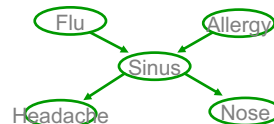
like flipping coin  $\sum_{k=1}^K \delta(f_k = i, a_k = j)$  times to see  
how often  $s_k = 1$

## MAP for $\theta_{s|ij} = P(S = 1|F = i, A = j)$ from fully observed data

- Maximum likelihood estimate

$$\theta \leftarrow \arg \max_{\theta} \log P(\text{data}|\theta)$$

$$\theta_{s|ij} = \frac{\sum_{k=1}^K \delta(f_k = i, a_k = j, s_k = 1)}{\sum_{k=1}^K \delta(f_k = i, a_k = j)}$$



- MAP estimate

$$\theta \leftarrow \arg \max_{\theta} \log P(\theta|\text{data}) = \arg \max_{\theta} \log [P(\text{data}|\theta)P(\theta)]$$

If assume prior  $P(\theta_{s|ij}) = \text{Beta}(\beta_1, \beta_0) = \frac{1}{B(\beta_1, \beta_0)} \theta_{s|ij}^{\beta_1-1} (1 - \theta_{s|ij})^{\beta_0-1}$

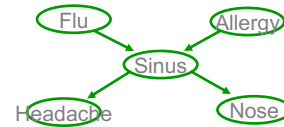
$$\theta_{s|ij} = \frac{(\beta_1 - 1) + \sum_{k=1}^K \delta(f_k = i, a_k = j, s_k = 1)}{(\beta_1 - 1) + (\beta_0 - 1) + \sum_{k=1}^K \delta(f_k = i, a_k = j)}$$

like coin flipping, including hallucinated examples

## Estimate $\theta$ from partly observed data

- What if FAHN observed, but not S?
- Can't calculate MLE

$$\theta \leftarrow \arg \max_{\theta} \log \prod_k P(f_k, a_k, s_k, h_k, n_k | \theta)$$

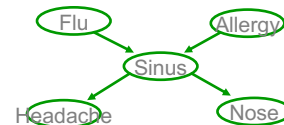


- Let  $X$  be all *observed* variable values (over all examples)
- Let  $Z$  be all *unobserved* variable values
- Can't calculate MLE:
 
$$\theta \leftarrow \arg \max_{\theta} \log P(X, Z | \theta)$$
- WHAT TO DO?

## Estimate $\theta$ from partly observed data

- What if FAHN observed, but not S?
- Can't calculate MLE

$$\theta \leftarrow \arg \max_{\theta} \log \prod_k P(f_k, a_k, s_k, h_k, n_k | \theta)$$



- Let  $X$  be all *observed* variable values (over all examples)
- Let  $Z$  be all *unobserved* variable values
- Can't calculate MLE:

$$\theta \leftarrow \arg \max_{\theta} \log P(X, Z | \theta)$$

- EM seeks\* the estimate:

$$\theta \leftarrow \arg \max_{\theta} E_{Z|X, \theta} [\log P(X, Z | \theta)]$$

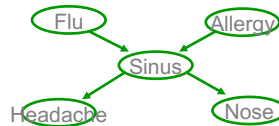
\* EM guaranteed to find local maximum

## Expected value

$$E_{P(X)}[f(X)] = \sum_x P(X = x) f(x)$$

- EM seeks estimate:

$$\theta \leftarrow \arg \max_{\theta} E_{Z|X,\theta}[\log P(X, Z|\theta)]$$



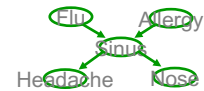
- here, observed  $X=\{F,A,H,N\}$ , unobserved  $Z=\{S\}$

$$\log P(X, Z|\theta) = \sum_{k=1}^K \log P(f_k) + \log P(a_k) + \log P(s_k|f_k a_k) + \log P(h_k|s_k) + \log P(n_k|s_k)$$

$$E_{P(Z|X,\theta)} \log P(X, Z|\theta) = \sum_{k=1}^K \sum_{i=0}^1 P(s_k = i|f_k, a_k, h_k, n_k) [\log P(f_k) + \log P(a_k) + \log P(s_k|f_k a_k) + \log P(h_k|s_k) + \log P(n_k|s_k)]$$

let's use  $a_k$  to represent value of  $A$  on the  $k$ th example

## EM Algorithm - Informally



EM is a general procedure for learning from partly observed data

Given observed variables  $X$ , unobserved  $Z$  ( $X=\{F,A,H,N\}$ ,  $Z=\{S\}$ )

Begin with arbitrary choice for parameters  $\theta$

Iterate until convergence:

- E Step: use  $X$ ,  $\theta$  to estimate the unobserved  $Z$  values
- M Step: use  $X$  values and estimated  $Z$  values to derive a better  $\theta$

Guaranteed to find local maximum.

Each iteration increases  $E_{P(Z|X,\theta)}[\log P(X, Z|\theta')]$

## EM Algorithm - Precisely

EM is a general procedure for learning from partly observed data

Given observed variables  $X$ , unobserved  $Z$  ( $X=\{F,A,H,N\}$ ,  $Z=\{S\}$ )

Define  $Q(\theta'|\theta) = E_{P(Z|X,\theta)}[\log P(X, Z|\theta')]$

*current* *M step new*

Iterate until convergence:

- E Step: Use  $X$  and current  $\theta$  to calculate  $P(Z|X,\theta)$
- M Step: Replace current  $\theta$  by  

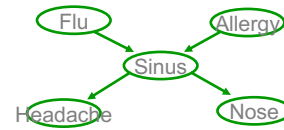
$$\theta \leftarrow \arg \max_{\theta'} Q(\theta'|\theta)$$

Guaranteed to find local maximum.

Each iteration increases  $E_{P(Z|X,\theta)}[\log P(X, Z|\theta')]$

## E Step: Use $X, \theta$ , to Calculate $P(Z|X, \theta)$

observed  $X=\{F,A,H,N\}$ ,  
unobserved  $Z=\{S\}$



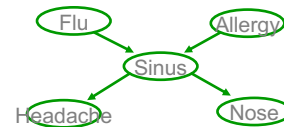
How? Bayes net inference problem.

$$P(S_k = 1 | f_k a_k h_k n_k, \theta) =$$

let's use  $a_k$  to represent value of A on the kth example

## E Step: Use $X, \theta$ , to Calculate $P(Z|X, \theta)$

observed  $X=\{F,A,H,N\}$ ,  
unobserved  $Z=\{S\}$



How? Bayes net inference problem.

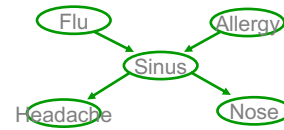
$$P(S_k = 1 | f_k a_k h_k n_k, \theta) =$$

$$P(S_k = 1 | f_k a_k h_k n_k, \theta) = \frac{P(S_k = 1, f_k a_k h_k n_k | \theta)}{P(S_k = 1, f_k a_k h_k n_k | \theta) + P(S_k = 0, f_k a_k h_k n_k | \theta)}$$

let's use  $a_k$  to represent value of A on the kth example

## EM and estimating $\theta_{s|ij}$

observed  $X = \{F, A, H, N\}$ , unobserved  $Z = \{S\}$



E step: Calculate  $P(Z_k|X_k; \theta)$  for each training example,  $k$

$$P(S_k = 1 | f_k a_k h_k n_k, \theta) = E[s_k] = \frac{P(S_k = 1, f_k a_k h_k n_k | \theta)}{P(S_k = 1, f_k a_k h_k n_k | \theta) + P(S_k = 0, f_k a_k h_k n_k | \theta)}$$

M step: update all relevant parameters. For example:

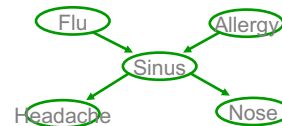
$$\theta_{s|ij} \leftarrow \frac{\sum_{k=1}^K \delta(f_k = i, a_k = j) E[s_k]}{\sum_{k=1}^K \delta(f_k = i, a_k = j)}$$

$$\text{Recall MLE was: } \theta_{s|ij} = \frac{\sum_{k=1}^K \delta(f_k = i, a_k = j, s_k = 1)}{\sum_{k=1}^K \delta(f_k = i, a_k = j)}$$

## EM and estimating $\theta$

More generally,

Given observed set  $X$ , unobserved set  $Z$  of boolean values



E step: Calculate for each training example,  $k$

the expected value of each unobserved variable in each training example

M step:

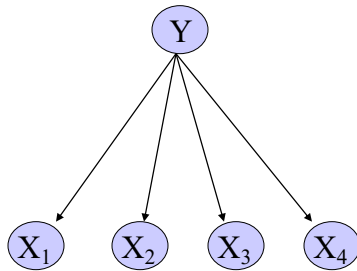
Calculate  $\theta$  similar to MLE estimates, but replacing each count by its expected count

$$\delta(Z = 1) \rightarrow E_{Z|X, \theta}[Z] \quad \delta(Z = 0) \rightarrow (1 - E_{Z|X, \theta}[Z])$$



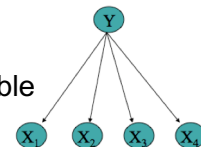
## Using Unlabeled Data to Help Train Naïve Bayes Classifier

Learn  $P(Y|X)$



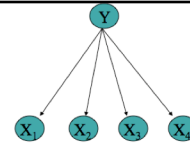
Y	X1	X2	X3	X4
1	0	0	1	1
0	0	1	0	0
0	0	0	1	0
?	0	1	1	0
?	0	1	0	1

E step: Calculate for each training example,  $k$   
the expected value of each unobserved variable



Y	X1	X2	X3	X4
1	0	0	1	1
0	0	1	0	0
0	0	0	1	0
?	0	1	1	0
?	0	1	0	1

## EM and estimating $\theta$



Given observed set  $X$ , unobserved set  $Y$  of boolean values

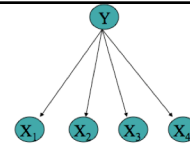
E step: Calculate for each training example,  $k$   
the expected value of each unobserved variable  $Y$

$$E_{P(Y|X_1 \dots X_N)}[y(k)] = P(y(k) = 1 | x_1(k), \dots, x_N(k); \theta) = \frac{P(y(k) = 1) \prod_i P(x_i(k) | y(k) = 1)}{\sum_{j=0}^1 P(y(k) = j) \prod_i P(x_i(k) | y(k) = j)}$$

M step: Calculate estimates similar to MLE, but  
replacing each count by its expected count

let's use  $y(k)$  to indicate value of  $Y$  on  $k$ th example

## EM and estimating $\theta$



Given observed set  $X$ , unobserved set  $Y$  of boolean values

E step: Calculate for each training example,  $k$   
the expected value of each unobserved variable  $Y$

$$E_{P(Y|X_1 \dots X_N)}[y(k)] = P(y(k) = 1 | x_1(k), \dots, x_N(k); \theta) = \frac{P(y(k) = 1) \prod_i P(x_i(k) | y(k) = 1)}{\sum_{j=0}^1 P(y(k) = j) \prod_i P(x_i(k) | y(k) = j)}$$

M step: Calculate estimates similar to MLE, but  
replacing each count by its expected count

$$\theta_{ij|m} = \hat{P}(X_i = j | Y = m) = \frac{\sum_k P(y(k) = m | x_1(k) \dots x_N(k)) \delta(x_i(k) = j)}{\sum_k P(y(k) = m | x_1(k) \dots x_N(k))}$$

$$\text{MLE would be: } \hat{P}(X_i = j | Y = m) = \frac{\sum_k \delta((y(k) = m) \wedge (x_i(k) = j))}{\sum_k \delta(y(k) = m)}$$

- 
- **Inputs:** Collections  $\mathcal{D}^l$  of labeled documents and  $\mathcal{D}^u$  of unlabeled documents.
  - Build an initial naive Bayes classifier,  $\hat{\theta}$ , from the labeled documents,  $\mathcal{D}^l$ , only. Use maximum a posteriori parameter estimation to find  $\hat{\theta} = \arg \max_{\theta} P(\mathcal{D}|\theta)P(\theta)$  (see Equations 5 and 6).
  - Loop while classifier parameters improve, as measured by the change in  $l_c(\theta|\mathcal{D}; \mathbf{z})$  (the complete log probability of the labeled and unlabeled data)
    - **(E-step)** Use the current classifier,  $\hat{\theta}$ , to estimate component membership of each unlabeled document, *i.e.*, the probability that each mixture component (and class) generated each document,  $P(c_j|d_i;\hat{\theta})$  (see Equation 7).
    - **(M-step)** Re-estimate the classifier,  $\hat{\theta}$ , given the estimated component membership of each document. Use maximum a posteriori parameter estimation to find  $\hat{\theta} = \arg \max_{\theta} P(\mathcal{D}|\theta)P(\theta)$  (see Equations 5 and 6).
  - **Output:** A classifier,  $\hat{\theta}$ , that takes an unlabeled document and predicts a class label.

From [Nigam et al., 2000]



## Experimental Evaluation

- Newsgroup postings
  - 20 newsgroups, 1000/group
- Web page classification
  - student, faculty, course, project
  - 4199 web pages
- Reuters newswire articles
  - 12,902 articles
  - 90 topics categories

## 20 Newsgroups

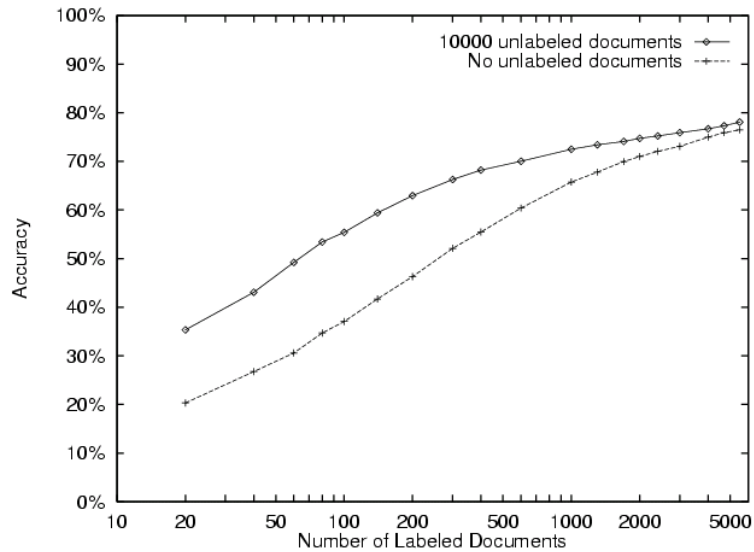


Table 3. Lists of the words most predictive of the **course** class in the WebKB data set, as they change over iterations of EM for a specific trial. By the second iteration of EM, many common **course**-related words appear. The symbol *D* indicates an arbitrary digit.

Iteration 0		Iteration 1	Iteration 2
intelligence	word <i>w</i> ranked by $P(w Y=\text{course})$ $/P(w Y \neq \text{course})$	<i>DD</i>	<i>D</i>
<i>DD</i>		<i>D</i>	<i>DD</i>
artificial	Using one labeled example per class	lecture	lecture
understanding		cc	cc
<i>DDw</i>		<i>D*</i>	<i>DD:DD</i>
dist		<i>DD:DD</i>	due
identical		handout	<i>D*</i>
rus		due	homework
arrange		problem	assignment
games		set	handout
dartmouth		tay	set
natural		<i>DDam</i>	hw
cognitive		yurttas	exam
logic		homework	problem
proving		kfoury	<i>DDam</i>
prolog		sec	postscript
knowledge		postscript	solution
human		exam	quiz
representation		solution	chapter
field		assaf	ascii

## What you should know about EM

- For learning from partly unobserved data
- MLE of  $\theta = \arg \max_{\theta} \log P(\text{data}|\theta)$
- EM estimate:  $\theta = \arg \max_{\theta} E_{Z|X,\theta}[\log P(X, Z|\theta)]$   
Where X is observed part of data, Z is unobserved
- EM for training Bayes networks
- Recall EM for Gaussian Mixture Models
- Can also derive your own EM algorithm for your own problem
  - write out expression for  $E_{Z|X,\theta}[\log P(X, Z|\theta)]$
  - E step: for each training example  $X^k$ , calculate  $P(Z^k | X^k, \theta)$
  - M step: chose new  $\theta$  to maximize  $E_{Z|X,\theta}[\log P(X, Z|\theta)]$

## Learning Bayes Net Structure

## How can we learn Bayes Net graph structure?

In general case, open problem

- can require lots of data (else high risk of overfitting)
- can use Bayesian priors, or other kinds of prior assumptions about graph structure to constrain search

One key result:

- Chow-Liu algorithm: finds “best” tree-structured network
- What’s best?
  - suppose  $P(\mathbf{X})$  is true distribution,  $T(\mathbf{X})$  is our tree-structured network, where  $\mathbf{X} = \langle X_1, \dots, X_n \rangle$
  - Chow-Liu minimizes Kullback-Leibler divergence:

$$KL(P(\mathbf{X}) \parallel T(\mathbf{X})) \equiv \sum_k P(\mathbf{X} = k) \log \frac{P(\mathbf{X} = k)}{T(\mathbf{X} = k)}$$

## Kullback-Leibler Divergence

- $KL(P(X) \parallel T(X))$  is a measure of the difference between distribution  $P(X)$  and  $T(X)$

$$KL(P(\mathbf{X}) \parallel T(\mathbf{X})) \equiv \sum_k P(\mathbf{X} = k) \log \frac{P(\mathbf{X} = k)}{T(\mathbf{X} = k)}$$

- It is asymmetric, always greater or equal to 0
- It is 0 iff  $P(X)=T(X)$

$$\begin{aligned} KL(P(X) \parallel T(X)) &= \sum_k P(X = k) \log P(X = k) - \sum_k P(X = k) \log T(X = k) \\ &= -H(P) + H(P, T) \end{aligned}$$

$$\text{where cross entropy } H(P, T) = \sum_k -P(X = k) \log T(X = k)$$

## Chow-Liu Algorithm

Key result: To minimize  $KL(P \parallel T)$  over possible tree networks  $T$  representing true  $P$ , it suffices to find the tree network  $T$  that maximizes the sum of mutual informations over its edges

Mutual information for an edge between variable  $A$  and  $B$ :

$$I(A, B) = \sum_a \sum_b P(a, b) \log \frac{P(a, b)}{P(a)P(b)}$$

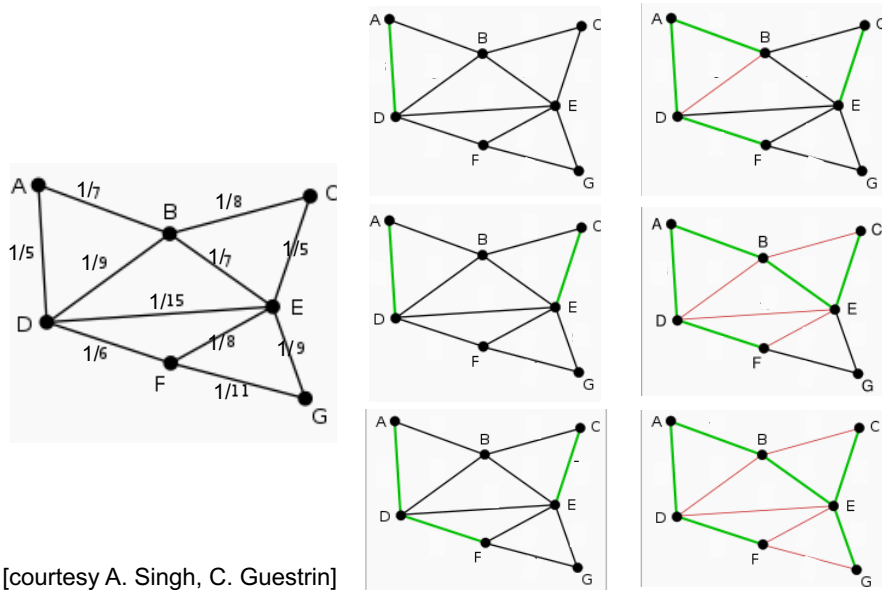
This works because for tree networks with nodes  $\mathbf{X} \equiv \langle X_1 \dots X_n \rangle$

$$\begin{aligned} KL(P(\mathbf{X}) \parallel T(\mathbf{X})) &\equiv \sum_k P(\mathbf{X} = k) \log \frac{P(\mathbf{X} = k)}{T(\mathbf{X} = k)} \\ &= - \sum_i I(X_i, Pa(X_i)) + \sum_i H(X_i) - H(X_1 \dots X_n) \end{aligned}$$

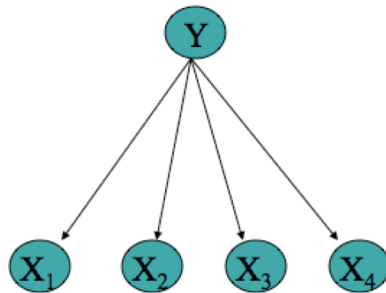
## Chow-Liu Algorithm

1. for each pair of variables  $A, B$ , use data to estimate  $P(A, B)$ ,  $P(A)$ , and  $P(B)$
2. for each pair  $A, B$  calculate mutual information
$$I(A, B) = \sum_a \sum_b P(a, b) \log \frac{P(a, b)}{P(a)P(b)}$$
3. calculate the maximum spanning tree over the set of variables, using edge weights  $I(A, B)$   
(given  $N$  vars, this costs only  $O(N^2)$  time)
4. add arrows to edges to form a directed-acyclic graph
5. learn the CPD's for this graph

## Chow-Liu algorithm example Greedy Algorithm to find Max-Spanning Tree



## Tree Augmented Naïve Bayes



[Nir Friedman et al., 1997]



## Bayes Nets – What You Should Know

- Representation
  - Bayes nets represent joint distribution as a DAG + Conditional Distributions
- Inference
  - NP-hard in general
  - For some graphs, closed form inference is feasible
  - Approximate methods too, e.g., Monte Carlo methods, ...
- Learning
  - Easy for known graph, fully observed data (MLE's, MAP est.)
  - EM for partly observed data, known graph
  - Learning graph structure: Chow-Liu for tree-structured networks
  - Hardest when graph unknown, data incompletely observed