

# Lecture 1

You are presented with a challenge: estimate the probability that a thumbtack lands on its head.  $R^T \triangleright^H$

You have only 6 draws of this thumbtack:

$\triangleright \curvearrowright \triangleright \curvearrowright \triangleright$  i.e. HTHHTH

What is the probability of heads? It seems like  $\frac{4}{6} = \frac{2}{3}$  is a good answer. We will see here why.

## Maximum Likelihood Estimation (MLE)

Let's restate the problem:

We assume the thumbtack flip follows a Bernoulli distribution with parameter  $\theta$

$$P(X) = \begin{cases} \theta & \text{if } x=H \\ 1-\theta & \text{if } x=T \end{cases}$$

Now we write the probability of the data  $D$  given  $\theta$ , also called the likelihood of the data:

$$\begin{aligned} P(D|\theta) &= P(X_1, X_2, X_3, X_4, X_5, X_6 | \theta) \\ &= P(X_1|\theta) P(X_2|\theta) \dots P(X_6|\theta) \quad \leftarrow \text{because the flips are IID} \\ &= \theta^4 (1-\theta)^2 \\ &= \theta^{\alpha_H} (1-\theta)^{\alpha_T} \quad \text{where } \alpha_H = \# \text{ of heads} \\ & \quad \alpha_T = \# \text{ of tails} \end{aligned}$$

MLE consist in picking the value of  $\theta$  that maximizes the likelihood.

$$\theta_{MLE} = \arg \max_{\theta} \theta^{\alpha_H} (1-\theta)^{\alpha_T}$$

It is convenient to compute the log likelihood (LL)

$$\begin{aligned} LL &= \ln \theta^{\alpha_H} (1-\theta)^{\alpha_T} = \ln \theta^{\alpha_H} + \ln (1-\theta)^{\alpha_T} \\ &= \alpha_H \ln \theta + \alpha_T \ln (1-\theta) \end{aligned}$$

Because log is monotonous;  
 $\arg \max_{\theta} \log P(D|\theta) = \arg \max_{\theta} P(D|\theta)$

↳ Before you optimize you need to verify that the objective function is concave

Solve for  $\theta_{MLE}$ :

$$\frac{\partial LL}{\partial \theta} = 0$$

$$\frac{\alpha_H}{\theta_{MLE}} - \frac{\alpha_T}{1-\theta_{MLE}} = 0$$

$$\theta_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}$$

↳ this corresponds to the intuitive answer  $\frac{2}{3}$

There are many ways to show that a function is concave. Here we will show that the second derivative is always  $< 0$ . Revising convexity (concavity) properties will be useful later in the course.

$$\frac{\partial^2 LL}{\partial \theta^2} = \frac{\partial}{\partial \theta} \left( \frac{\alpha_H}{\theta} - \frac{\alpha_T}{1-\theta} \right)$$

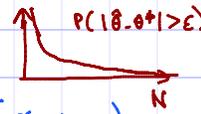
$$= \frac{-\alpha_H}{\theta^2} - \frac{\alpha_T}{(1-\theta)^2} < 0$$

How confident can we be of our answer? Are 6 flips enough? Would we feel more confident if we had 1000 flips? How much more confident?

We can use Hoeffding's inequality adapted for Bernoulli variables to compute a minimum sample size for an error of at most  $\epsilon$ .

$$P(|\hat{\theta} - \theta^*| \geq \epsilon) \leq 2e^{-2N\epsilon^2}$$

$$N = \alpha_H + \alpha_T, \quad \hat{\theta} = \alpha_H / (\alpha_H + \alpha_T)$$



### Maximum a posteriori estimation (MAP)

Now assume that instead of a thumbtack, for which it's hard to guess the probability of heads, you were tasked with finding the probability of a coin falling on heads.

Suddenly, 4 heads and 2 tails don't seem enough to say that the coin is biased (a biased coin has  $\theta \neq 0.5$ ). We have a strong prior on  $\theta$  being 0.5 or very close to 0.5 for a normal coin.

However if we had observed 4000 Heads and 2000 Tails, we would think the coin is biased. This can be modeled as a Bayesian estimation problem where our prior belief about the coin is expressed as a prior distribution  $p(\theta)$ .

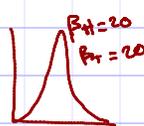
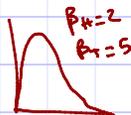
For this problem, a convenient distribution is the Beta distribution. It allows us to express different beliefs about  $\theta$ , the parameter of a Bernoulli distribution.

$$\theta \sim \text{Beta}(\beta_H, \beta_T)$$

$$p(\theta) = \frac{\theta^{\beta_H-1} (1-\theta)^{\beta_T-1}}{B(\beta_H, \beta_T)}$$

Here we call the two parameters  $\beta_H$  &  $\beta_T$  because they effectively act as additional head & tail counts in the posterior distribution

normalizing constant



$\beta_H, \beta_T$  affect the prior distribution

We compute the posterior distribution:

$$P(\theta|D) = \frac{P(D|\theta)P(\theta)}{P(D)} \quad (\text{Bayes rule})$$

$$\propto P(D|\theta)P(\theta) \propto \theta^{\alpha_H} (1-\theta)^{\alpha_T} \theta^{\beta_H-1} (1-\theta)^{\beta_T-1} = \theta^{\alpha_H+\beta_H-1} (1-\theta)^{\alpha_T+\beta_T-1}$$

$$\text{We find } \theta_{\text{MAP}} = \underset{\theta}{\text{argmax}} \theta^{\alpha_H+\beta_H-1} (1-\theta)^{\alpha_T+\beta_T-1}$$

To solve this, we repeat similar steps to above.

$$\theta_{\text{MAP}} = \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H + \alpha_T + \beta_T - 2}$$

as  $N \rightarrow \infty$  the effect of the prior washes out for small  $N$ , the prior can have a big effect.