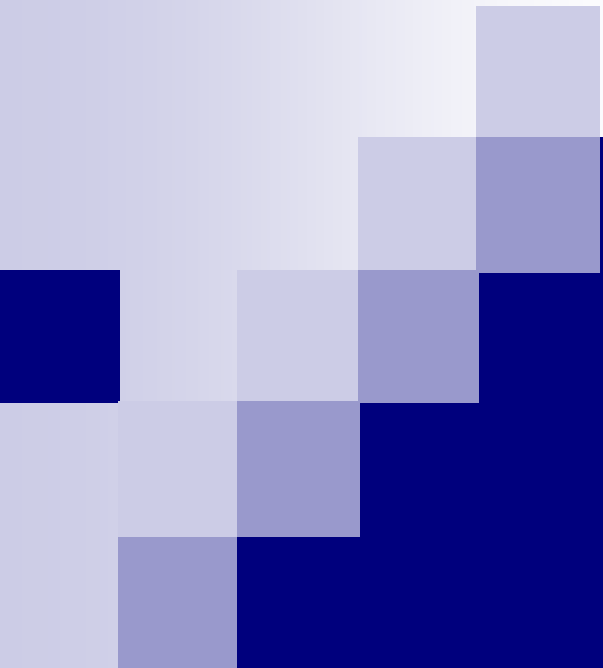


Koller & Friedman: Chapter 16
Jordan: Chapters 13, 15
Uri Lerner's Thesis: Chapters 3,9



Dynamic models 1

Kalman filters, linearization,
Switching KFs, Assumed density filters

Probabilistic Graphical Models – 10708

Carlos Guestrin

Carnegie Mellon University

November 16th, 2005

Announcement

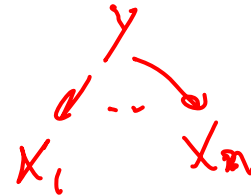


- Special recitation lectures
 - Pradeep will give two special lectures
 - Nov. 22 & Dec. 1: 5-6pm, during recitation
 - Covering: variational methods, loopy BP and their relationship
 - Don't miss them!!!

Adventures of our BN hero

- Compact representation for probability distributions
- Fast inference
- Fast learning
- Approximate inference
- But... Who are the most popular kids?

1. Naïve Bayes



2 and 3.

Hidden Markov models (HMMs)
Kalman Filters

The Kalman Filter



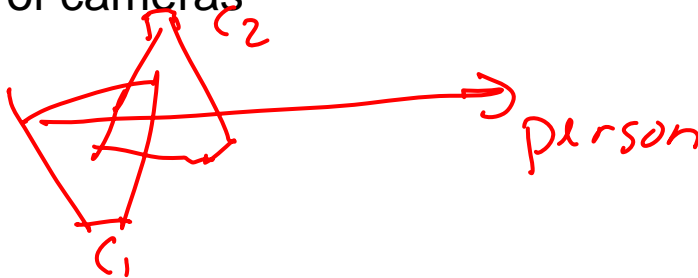
- An HMM with Gaussian distributions
- Has been around for at least 50 years
- Possibly the most used graphical model ever
- It's what
 - does your cruise control
 - tracks missiles
 - controls robots
 - ...
- And it's so simple...
 - Possibly explaining why it's so used
- Many interesting models build on it...
 - Review and extensions today

Example of KF – SLAT

Simultaneous Localization and Tracking

[Funiak, Guestrin, Paskin, Sukthankar '05]

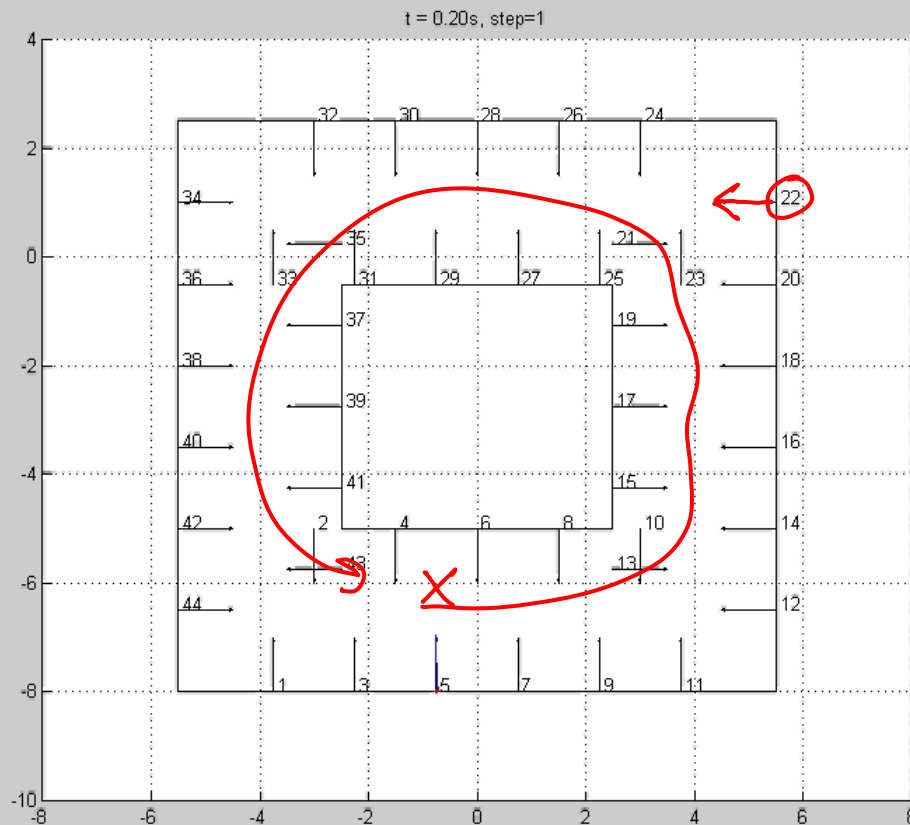
- Place some cameras around an environment, don't know where they are
- Could measure all locations, but requires lots of grad. student (Stano) time
- Intuition:
 - A person walks around
 - If camera 1 sees person, then camera 2 sees person, learn about relative positions of cameras



Example of KF – SLAT

Simultaneous Localization and Tracking

[Funiak, Guestrin, Paskin, Sukthankar '05]



$C_i \leftarrow \text{each camera } i$

$L_t \leftarrow \text{position at time } t$

$p(C, L_t)$

$p(C) \leftarrow \text{"uniform"}$

$p(L_0)$

$p(O_i | C_j, L_i)$

$p(C, L_t | O_{1:t})$

Multivariate Gaussian

$$p(X_1, \dots, X_n) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \right\}$$

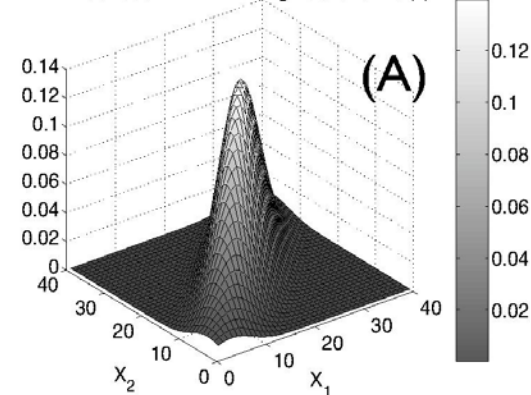
Mean vector:

$$\mu = \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_d \end{pmatrix}$$

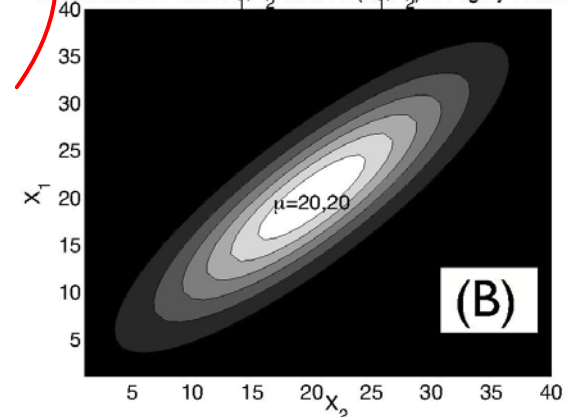
Covariance matrix:

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \sigma_{12} & \dots \\ \sigma_{12} & \sigma_2^2 & \dots \\ \vdots & \vdots & \ddots \end{pmatrix}$$

2D Gaussian PDF With High Covariance (Σ)



Gaussian PDF over X_1, X_2 where $\Sigma(X_1, X_2)$ is Highly Positive



Conditioning a Gaussian

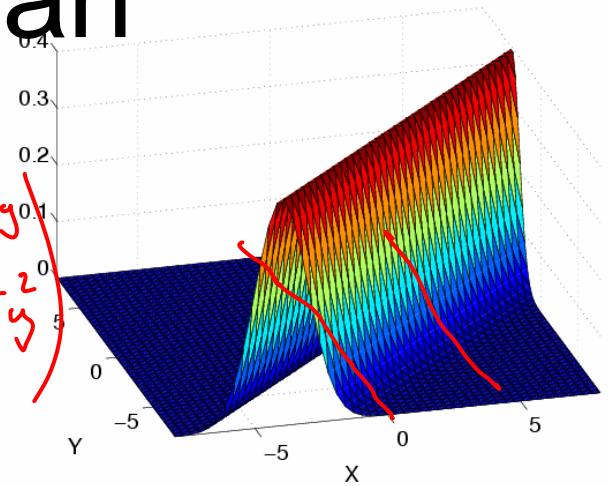
Joint Gaussian:

$$\square p(X, Y) \sim N(\mu; \Sigma)$$

Conditional linear Gaussian:

$$\square p(Y|X) \sim N(\mu_{Y|X}; \sigma^2)$$

$$\mu = \begin{pmatrix} \mu_x \\ \mu_y \end{pmatrix} \quad \Sigma = \begin{pmatrix} \sigma_x^2 & \sigma_{xy} \\ \sigma_{xy} & \sigma_y^2 \end{pmatrix}$$

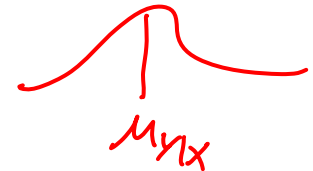


$$\frac{p(Y, X)}{p(X)}$$

$$\mu_{Y|X} = \mu_Y + \frac{\sigma_{YX}}{\sigma_X^2} (x - \mu_x)$$

$$\sigma_{Y|X}^2 = \sigma_Y^2 - \frac{\sigma_{YX}^2}{\sigma_X^2}$$

← doesn't depend on x



Gaussian is a “Linear Model”

$$\mu_{Y|X} = \mu_Y + \frac{\sigma_{YX}}{\sigma_X^2}(x - \mu_x)$$

■ Conditional linear Gaussian:

□ $p(Y|X) \sim N(\beta_0 + \beta X; \sigma_{Y|X}^2)$

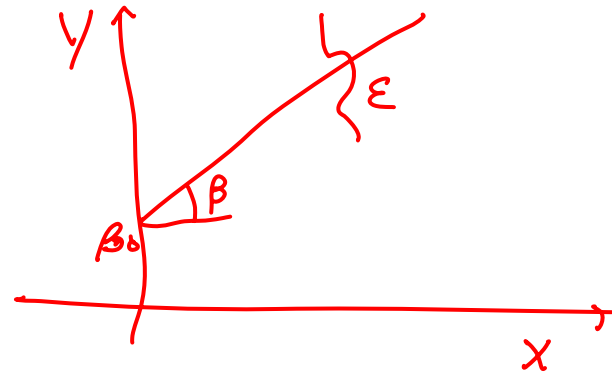
$$\sigma_{Y|X}^2 = \sigma_Y^2 - \frac{\sigma_{YX}^2}{\sigma_X^2}$$

$$Y = \beta_0 + \beta X + \varepsilon$$

$\varepsilon \sim N(0, \sigma_{Y|X}^2)$

$$\beta_0 = \mu_Y - \frac{\sigma_{YX}}{\sigma_X^2} \mu_X$$

$$\beta = \frac{\sigma_{YX}}{\sigma_X^2}$$



Conditioning a Gaussian

Joint Gaussian:

□ $p(X,Y) \sim N(\mu; \Sigma)$

$$\Sigma = \begin{pmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{pmatrix}$$

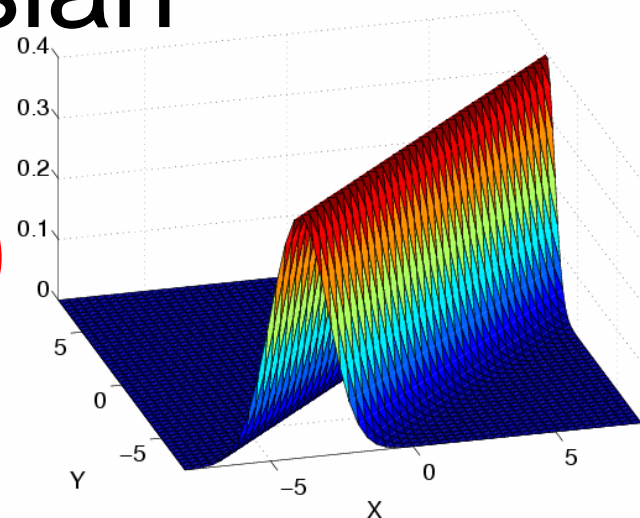
Conditional linear Gaussian:

□ $p(Y|X) \sim N(\mu_{Y|X}; \Sigma_{YY|X})$

$$\mu_{Y|X} = \mu_Y + \Sigma_{YX} \Sigma_{XX}^{-1} (x - \mu_x)$$

$$\Sigma_{YY|X} = \Sigma_{YY} - \Sigma_{YX} \Sigma_{XX}^{-1} \Sigma_{XY}$$

doesn't depend on x



Conditional Linear Gaussian (CLG) – general case

■ Conditional linear Gaussian:

□ $p(Y|X) \sim N(\beta_0 + BX; \Sigma_{YY|X})$

$$\mu_{Y|X} = \mu_Y + \Sigma_{YX} \Sigma_{XX}^{-1} (x - \mu_x)$$

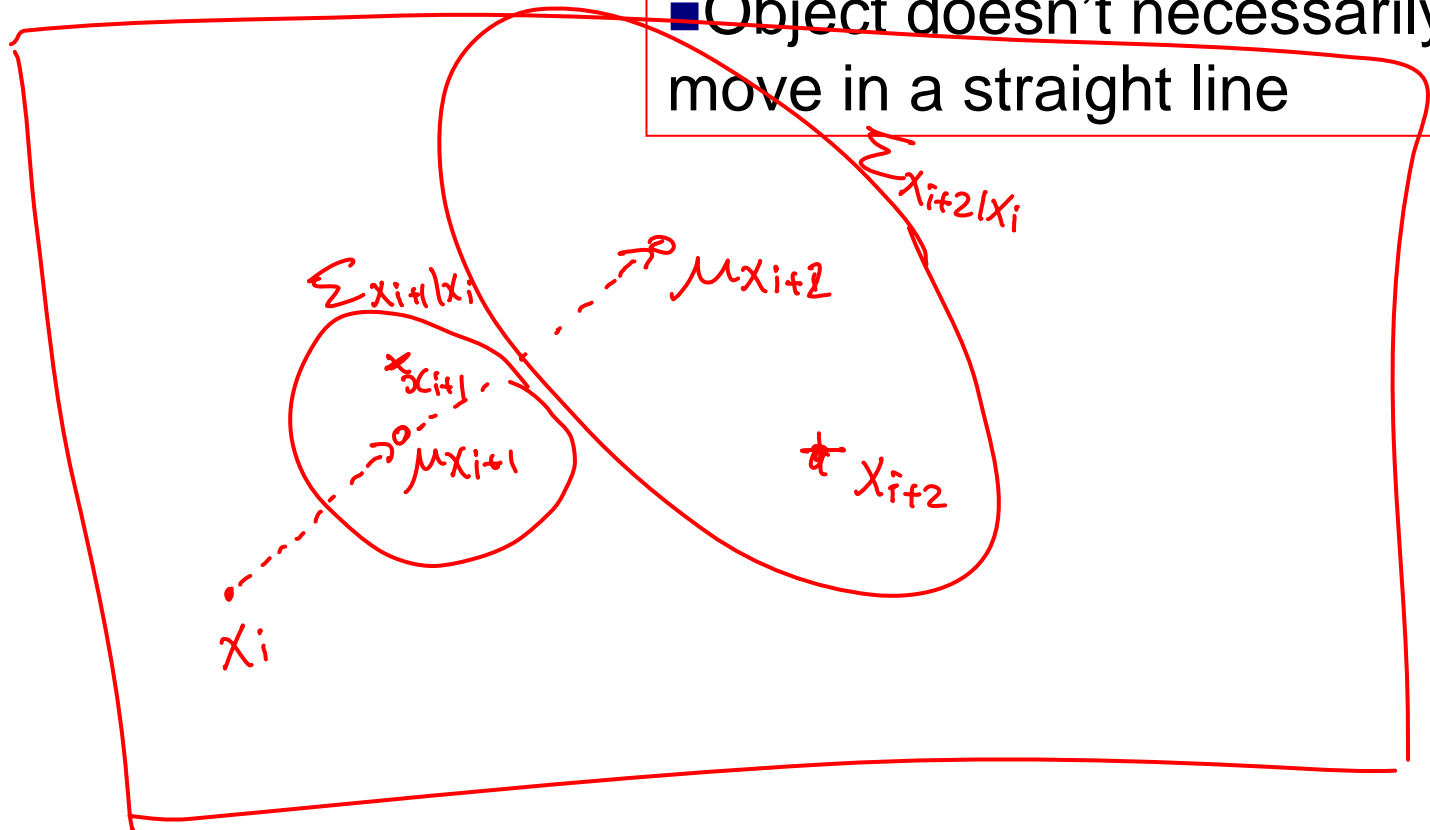
$$\Sigma_{YY|X} = \Sigma_{YY} - \Sigma_{YX} \Sigma_{XX}^{-1} \Sigma_{XY}$$

$$\beta_0 = \mu_Y + \Sigma_{YX} \Sigma_{XX}^{-1} \mu_X$$

$$B = -\Sigma_{YX} \Sigma_{XX}^{-1}$$

Understanding a linear Gaussian – the 2d case

- Variance increases over time (motion noise adds up)
- Object doesn't necessarily move in a straight line

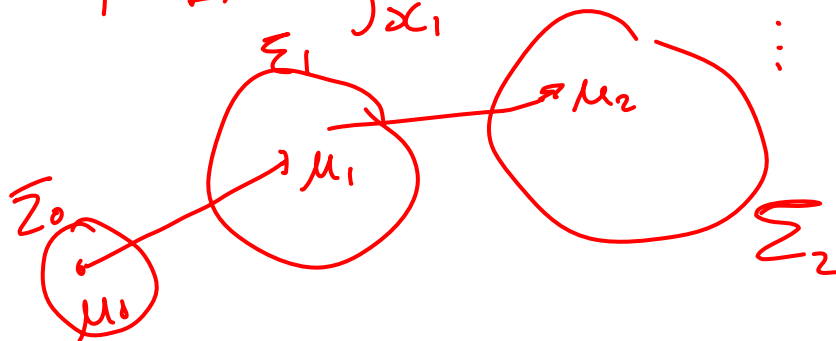


Tracking with a Gaussian 1

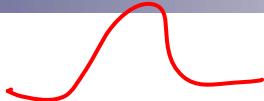
- $p(X_0) \sim N(\mu_0, \Sigma_0)$
- $p(X_{i+1}|X_i) \sim N(B X_i + \beta; \Sigma_{X_{i+1}|X_i})$

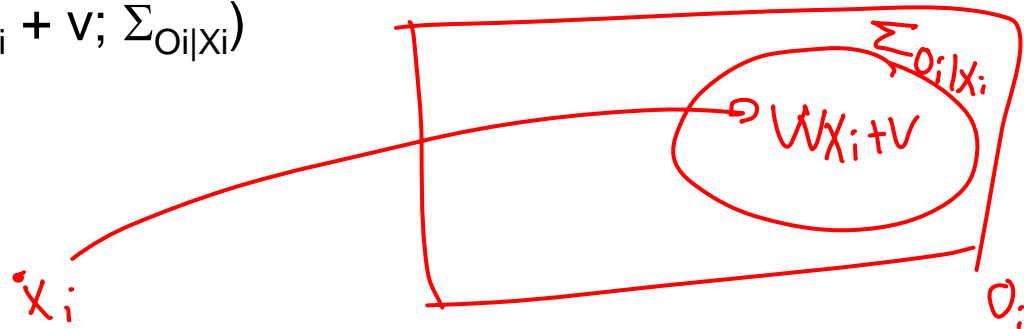
$$p(X_1) = \int_{x_0} p(x_0) \cdot p(X_1|x_0) dx_0$$

$$p(X_2) = \int_{x_1} p(x_1) p(X_2|x_1) dx_1$$

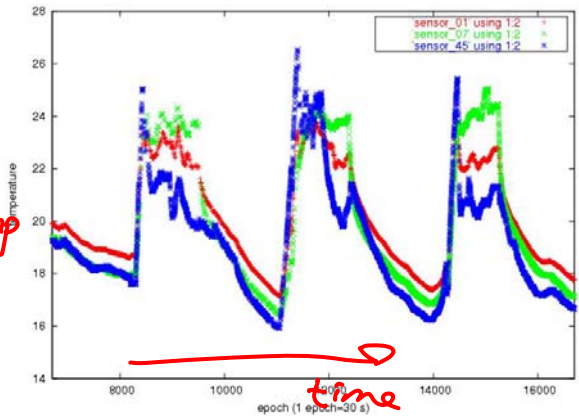
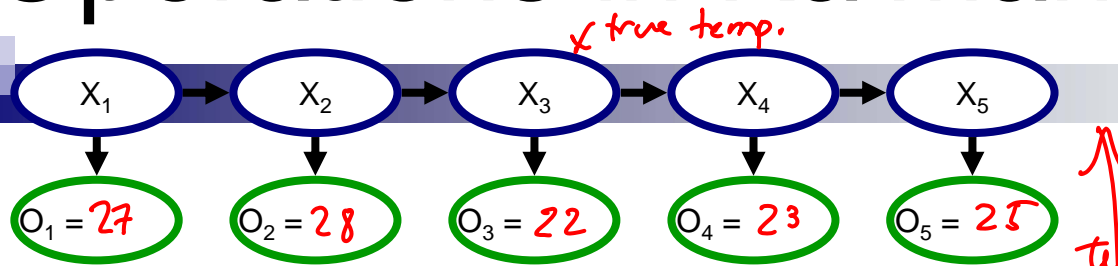


Tracking with Gaussians 2 – Making observations

- We have $p(X_i)$ 
- Detector observes $O_i = o_i$
- Want to compute $p(X_i | O_i = o_i)$
- Use Bayes rule:
$$p(X_i | o_i) = \frac{p(X_i) \cdot p(o_i | X_i)}{p(o_i)}$$
- Require a CLG observation model
 - $p(O_i | X_i) \sim N(W X_i + v; \Sigma_{O_i | X_i})$



Operations in Kalman filter



- Compute $p(X_t \mid O_{1:t} = o_{1:t})$
- Start with $p(X_0)$
- At each time step t :
 - **Condition** on observation

$$p(X_t \mid o_{1:t}) \propto p(X_t \mid o_{1:t-1})p(o_t \mid X_t)$$
 - **Prediction** (Multiply transition model)

$$p(X_{t+1}, X_t \mid o_{1:t}) = p(X_{t+1} \mid X_t)p(X_t \mid o_{1:t})$$
 - **Roll-up** (marginalize previous time step)

$$p(X_{t+1} \mid o_{1:t}) = \int_{x_t} p(X_{t+1}, x_t \mid o_{1:t}) dx_t$$
- I'll describe one implementation of KF, there are others
 - Information filter

Canonical form

$$\begin{aligned} p(X_1, \dots, X_n) &= \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \right\} \\ &= K \exp \left\{ \eta^T \mathbf{x} - \frac{1}{2} \mathbf{x}^T \Lambda^{-1} \mathbf{x} \right\} \end{aligned}$$

- Standard form and canonical forms are related:

$$\begin{aligned} \mu &= \Lambda^{-1} \underline{\eta} \\ \Sigma &= \underline{\Lambda}^{-1} \end{aligned}$$

- Conditioning is easy in canonical form
- Marginalization easy in standard form

Conditioning in canonical form

$$p(X_t | o_{1:t}) \propto p(X_t | o_{1:t-1}) p(\underline{o_t} | X_t)$$

■ First multiply: $p(A, B) = p(A) p(B | A)$

$$p(A) : \eta_1, \overset{\text{K}}{\Lambda_1} \quad \overset{\text{K}}{\Lambda_1} \quad \overset{\text{K}}{\Lambda_1}$$

$$p(B | A) : \eta_2, \overset{2K}{\Lambda_2} \quad - \quad \begin{pmatrix} \Lambda_1 & 0 \\ 0 & 0 \end{pmatrix} \quad \begin{pmatrix} \Lambda_2 \end{pmatrix}$$

$$p(A, B) : \underbrace{\eta_3 = \eta_1 + \eta_2}, \underbrace{\Lambda_3 = \Lambda_1 + \Lambda_2}$$

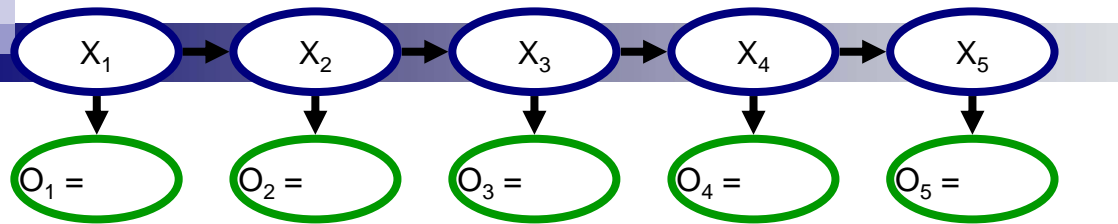
■ Then, condition on value $B = y$ $p(A | B = y)$

$$\eta_{A|B=y} = \underline{\eta_A - \Lambda_{AB} \cdot y}$$

$$\Lambda_{AA|B=y} = \Lambda_{AA}$$

$$\Lambda_3 = \begin{pmatrix} \Lambda_{AA} & \Lambda_{AB} \\ \Lambda_{BA} & \Lambda_{BB} \end{pmatrix}$$

Operations in Kalman filter



- Compute $p(X_t \mid O_{1:t} = o_{1:t})$
- Start with $p(X_0)$
- At each time step t :

- **Condition** on observation

$$p(X_t \mid o_{1:t}) \propto p(X_t \mid o_{1:t-1})p(o_t \mid X_t)$$

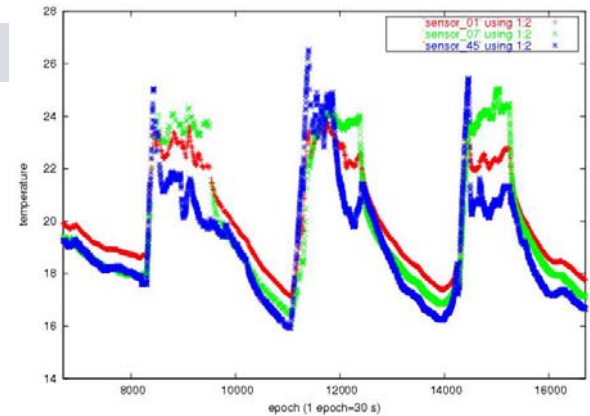
- **Prediction** (Multiply transition model)

$$p(X_{t+1}, X_t \mid o_{1:t}) = p(X_{t+1} \mid X_t)p(X_t \mid o_{1:t})$$

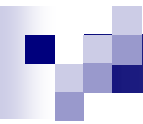
- **Roll-up** (marginalize previous time step)

$$p(X_{t+1} \mid o_{1:t}) = \int_{X_t} p(X_{t+1}, x_t \mid o_{1:t}) dx_t$$

} matrix ops



Prediction & roll-up in canonical form


$$p(X_{t+1} | o_{1:t}) = \int_{X_t} p(X_{t+1} | x_t) p(x_t | o_{1:t}) dx_t$$

- First multiply: $p(A, B) = p(A)p(B | A)$

add matrices

- Then, marginalize X_t : $p(A) = \int_B p(A, b) db$

$$\eta_A^m = \eta_A - \Lambda_{AB} \Lambda_{BB}^{-1} \eta_B$$

$$\Lambda_{AA}^m = \Lambda_{AA} - \Lambda_{AB} \Lambda_{BB}^{-1} \Lambda_{BA}$$

Handwritten: $\Lambda_3 = \begin{pmatrix} \Lambda_{AA} & \Lambda_{AB} \\ \Lambda_{BA} & \Lambda_{BB} \end{pmatrix}$

What if observations are not CLG?

- Often observations are not CLG

- CLG if $O_i = B X_i + \beta_o + \varepsilon$ $\leftarrow N(0, \Sigma_\varepsilon)$

- Consider a motion detector

- $O_i = 1$ if person is likely to be in the region

$$p(O_i=1 | X_i) = \begin{cases} 0 & \text{if outside the box} \\ 1 & \text{if inside the box} \end{cases}$$

- Posterior is not Gaussian

$$p(X_i | O_i=1)$$

$$p(X_i):$$



Linearization: incorporating non-linear evidence

- $p(O_i|X_i)$ not CLG, but...
- Find a Gaussian approximation of $p(X_i, O_i) = p(X_i) p(O_i|X_i)$
- Instantiate evidence $O_i=o_i$ and obtain a Gaussian for $p(X_i|O_i=o_i)$
- Why do we hope this would be any good?
 - Locally, Gaussian may be OK

Linearization as integration

$$\hat{\Sigma} = \begin{pmatrix} \hat{\sigma}_{o_i}^2 & \hat{\sigma}_{o_i x_i} \\ \hat{\sigma}_{x_i o_i} & \hat{\sigma}_{x_i}^2 \end{pmatrix}$$

- Gaussian approximation of $p(X_i, O_i) = p(X_i) p(O_i | X_i)$

$$\hat{\mu} = \begin{pmatrix} \hat{\mu}_{o_i} \\ \mu_{x_i} \end{pmatrix}$$

- Need to compute moments

$$\square E[O_i] = \int_{x_i, o_i} o_i \cdot p(x_i) \cdot p(o_i | x_i) dx_i do_i = \hat{\mu}_{o_i}$$

$$\square E[O_i^2] = \int_{x_i, o_i} o_i^2 \cdot \dots \Rightarrow \hat{\sigma}_{o_i}^2 = \hat{\mu}_{o_i}^2$$

$$\square E[O_i X_i] = \int_{x_i, o_i} o_i \cdot x_i \cdot \dots \Rightarrow \hat{\sigma}_{o_i x_i} = \hat{\mu}_{o_i} \mu_{x_i}$$

- Note: Integral is product of a Gaussian with an arbitrary function

Linearization as numerical integration

gaussian

■ Product of a Gaussian with arbitrary function

$$\int_x w(x) f(x) dx$$

■ Effective numerical integration with Gaussian quadrature method

- Approximate integral as **weighted sum over integration points** $\rightarrow \langle w_j, x_j \rangle$
- Gaussian quadrature defines location of points and weights

$$\int_x w(x) f(x) dx \approx \sum_{j=1}^N w_j f(x_j)$$

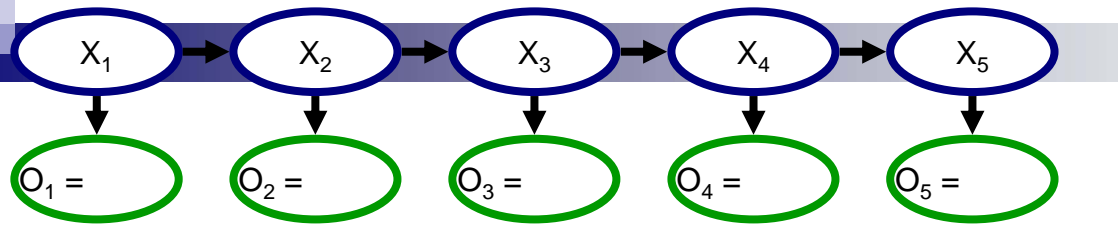
■ Exact if arbitrary function is **polynomial of bounded degree**

■ **Number of integration points exponential** in number of dimensions d

■ **Exact monomials** requires exponentially fewer points

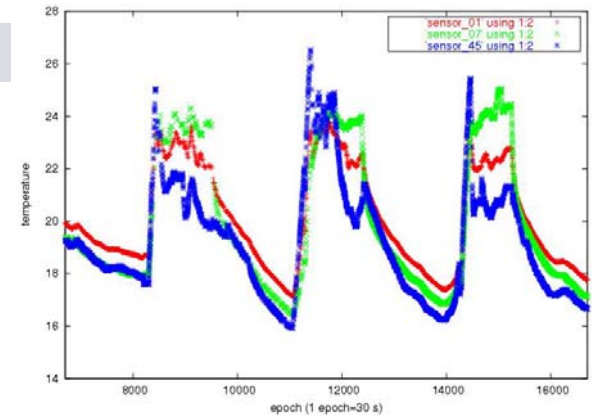
- For **$2d+1$ points**, this method is equivalent to effective **Unscented Kalman filter**
- **Generalizes to many more points**

Operations in non-linear Kalman filter



- Compute $p(X_t \mid O_{1:t} = o_{1:t})$
- Start with $p(X_0)$
- At each time step t :
 - **Condition** on observation (use **numerical integration**)
$$p(X_t \mid o_{1:t}) \propto p(X_t \mid o_{1:t-1})p(o_t \mid X_t)$$
 - **Prediction** (Multiply transition model, use **numerical integration**)
$$p(X_{t+1}, X_t \mid o_{1:t}) = p(X_{t+1} \mid X_t)p(X_t \mid o_{1:t})$$
 - **Roll-up** (marginalize previous time step)

$$p(X_{t+1} \mid o_{1:t}) = \int_{X_t} p(X_{t+1}, x_t \mid o_{1:t}) dx_t$$



What if the person chooses different motion models?



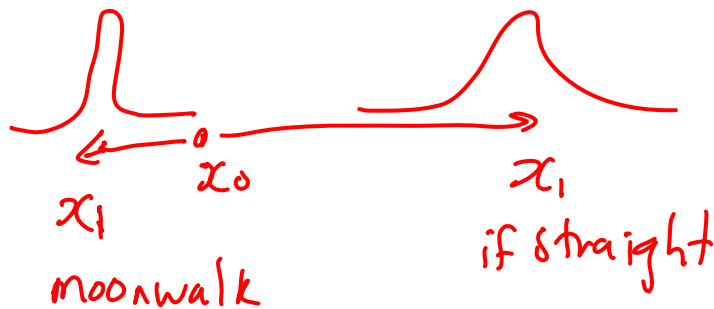
- With probability θ , move more or less straight
- With probability $1-\theta$, do the “moonwalk”

The moonwalk



What if the person chooses different motion models?

- With probability θ , move more or less straight
- With probability $1-\theta$, do the “moonwalk”



Switching Kalman filter

- At each time step, choose one of k motion models:

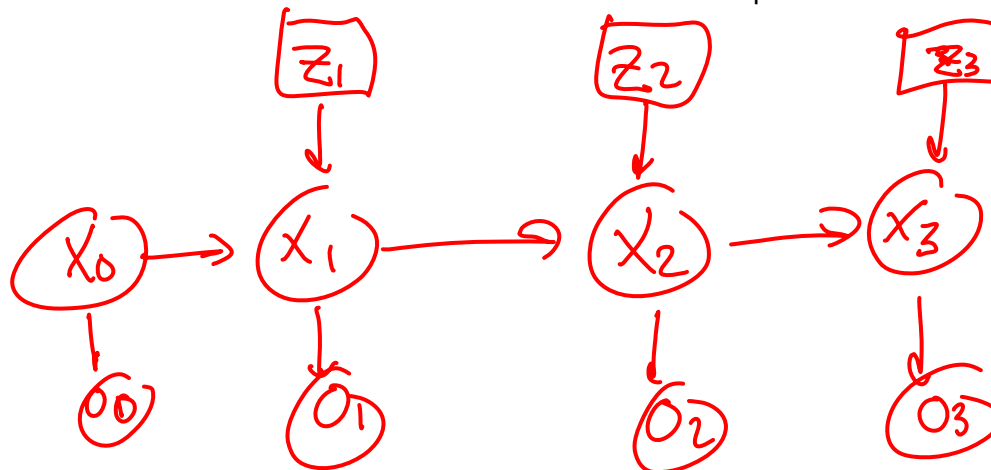
- You never know which one!

- $p(X_{i+1}|X_i, Z_{i+1})$

- CLG indexed by Z_i

- $p(X_{i+1}|X_i, Z_{i+1}=j) \sim N(\beta^j_0 + B^j X_i; \Sigma^j_{X_{i+1}|X_i})$

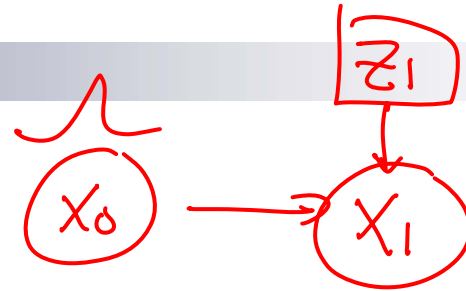
depending on motion model



Inference in switching KF – one step

- Suppose

- $p(X_0)$ is Gaussian
- Z_1 takes one of two values
- $p(X_1|X_0, Z_1)$ is CLG



- Marginalize X_0

$$p(X_1|Z_1) = \int_{x_0} p(x_0) \cdot p(X_1|x_0, Z_1) dx_0$$

- Marginalize Z_1

$$p(X_1) = \sum_j p(X_1|Z_1=j) \cdot P(Z_1=j)$$

- Obtain mixture of two Gaussians!



Multi-step inference

$$x_i \rightarrow x_{i+1}^{z_{i+1}}$$

- Suppose

- $p(x_i)$ is a mixture of m Gaussians
- Z_{i+1} takes one of two values
- $p(x_{i+1}|x_i, Z_{i+1})$ is CLG



$$p(x_i) = \sum_{k=1}^m w_k N(\mu_k, \Sigma_k)$$

- Marginalize x_i

$$p(x_{i+1} | z_{i+1} = j) = \int_{x_i} p(x_{i+1} | x_i, z_i = j) \cdot p(x_i) dx_i$$

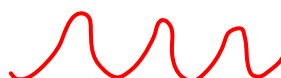
$$= \sum_{k=1}^m w_k \int_{x_i} p(x_{i+1} | x_i, z_i = j) N(\mu_k, \Sigma_k) dx_i$$

- Marginalize Z_i

$$\hookrightarrow p(x_{i+1}) = \sum_j p(z_{i+1} = j) \cdot p(x_{i+1} | z_{i+1} = j)$$

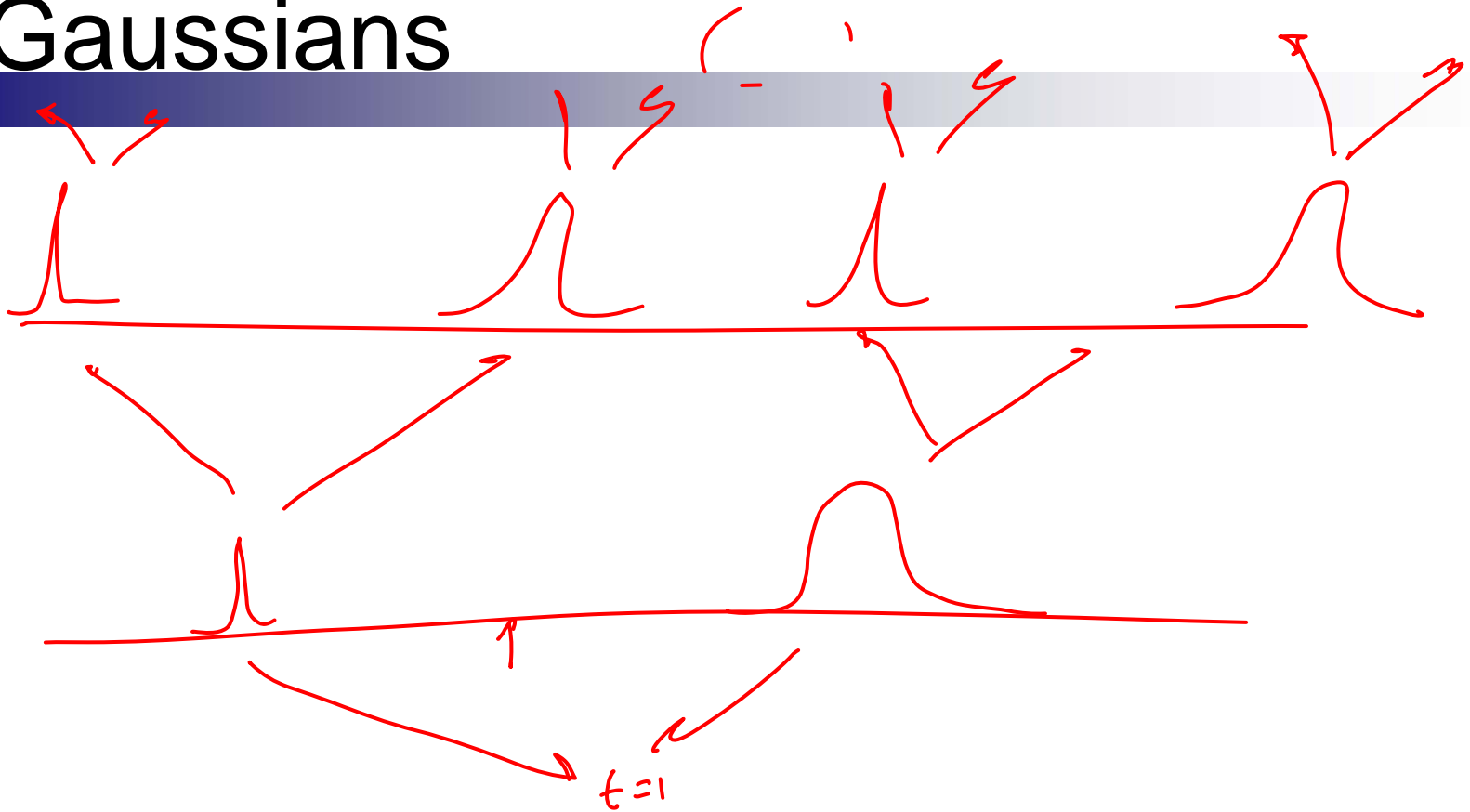
- Obtain mixture of $2m$ Gaussians!

- Number of Gaussians grows exponentially!!!



humps

Gaussians



Computational complexity of inference in switching Kalman filters

- Switching Kalman Filter with (only) 2 motion models
- Query:
- **Problem is NP-hard!!!** [Lerner & Parr `01]
 - Why “!!!”?
 - Graphical model is a tree:
 - Inference efficient if all are discrete
 - Inference efficient if all are Gaussian
 - But not with hybrid model (combination of discrete and continuous)

Bounding number of Gaussians

- $P(X_i)$ has 2^m Gaussians, but...
 - usually, most are bumps have low probability and overlap:
-
- **Intuitive approximate inference:**
 - Generate $k.m$ Gaussians
 - Approximate with m Gaussians

Collapsing Gaussians – Single Gaussian from a mixture

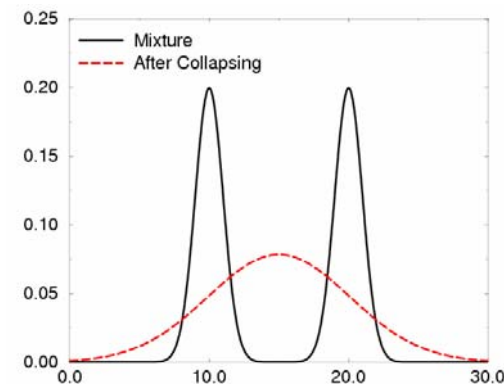
- Given mixture $P < w_i; N(\mu_i, \Sigma_i) >$
- Obtain approximation $Q \sim N(\mu, \Sigma)$ as:

$$\mu = \sum_i w_i \mu_i$$

$$\Sigma = \sum_i w_i \Sigma_i + \sum_i w_i (\mu_i - \mu)(\mu_i - \mu)^T$$

- **Theorem:**

- P and Q have same first and second moments
- **KL projection:** Q is single Gaussian with lowest KL divergence from P

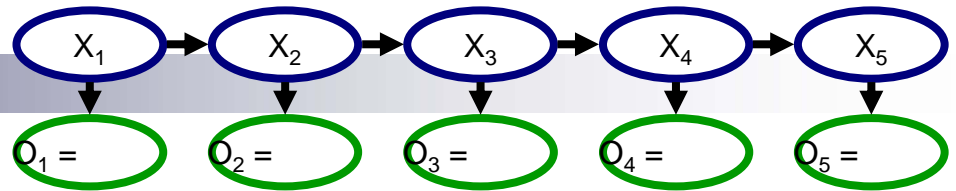


Collapsing mixture of Gaussians into smaller mixture of Gaussians

- Hard problem!
 - Akin to clustering problem...

- Several heuristics exist
 - *c.f.*, Uri Lerner's Ph.D. thesis

Operations in non-linear switching Kalman filter



- Compute mixture of Gaussians for $p(X_t \mid O_{1:t} = o_{1:t})$
- Start with $p(X_0)$
- At each time step t :
 - For each of the m Gaussians in $p(X_i \mid o_{1:i})$:
 - **Condition** on observation (use **numerical integration**)
 - **Prediction** (Multiply transition model, use **numerical integration**)
 - Obtain k Gaussians
 - **Roll-up** (marginalize previous time step)
 - **Project** $k.m$ Gaussians into m' Gaussians $p(X_i \mid o_{1:i+1})$

Assumed density filtering

- Examples of very important **assumed density filtering**:

- Non-linear KF
- Approximate inference in switching KF

- General picture:

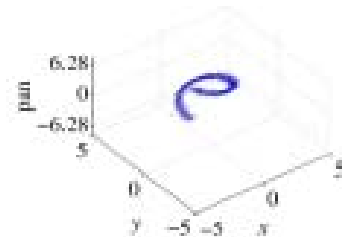
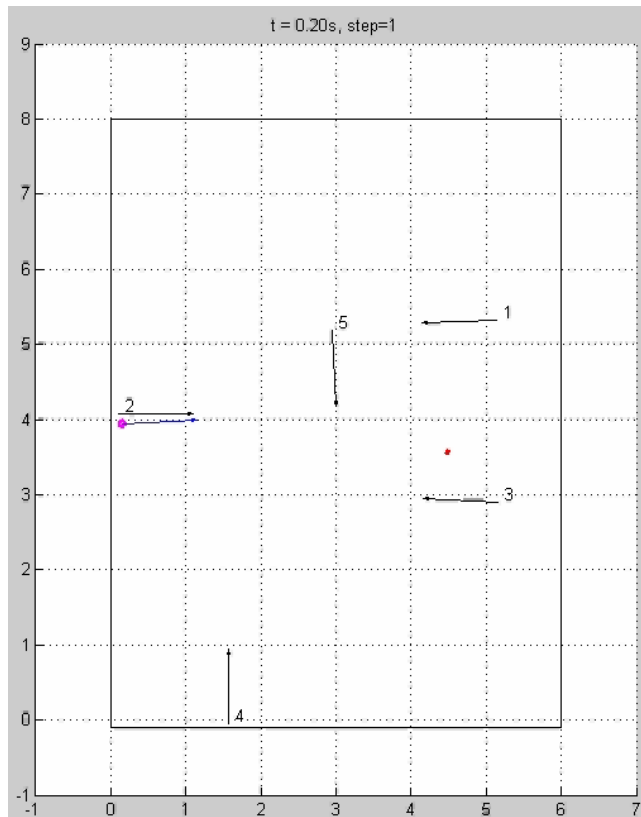
- Select an **assumed density**
 - e.g., single Gaussian, mixture of m Gaussians, ...
- After conditioning, prediction, or roll-up, **distribution no-longer representable with assumed density**
 - e.g., non-linear, mixture of $k.m$ Gaussians,...
- **Project** back into assumed density
 - e.g., numerical integration, collapsing,...

When non-linear KF is not good enough

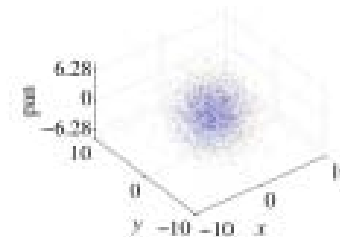
- Sometimes, distribution in non-linear KF is not approximated well as a single Gaussian
 - e.g., a banana-like distribution
- Assumed density filtering:
 - Solution 1: **reparameterize problem** and solve as a **single Gaussian**
 - Solution 2: more typically, **approximate as a mixture of Gaussians**

Reparameterized KF for SLAT

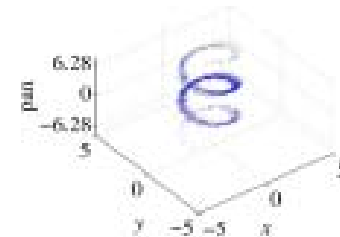
[Funiak, Guestrin, Paskin, Sukthankar '05]



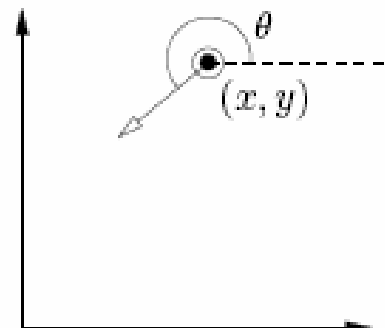
(a) true posterior



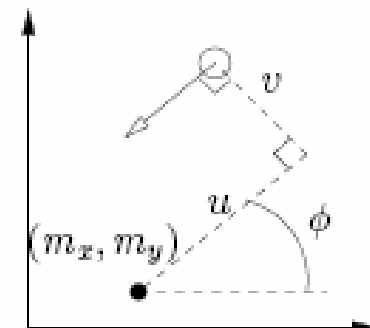
(b) Gaussian in absolute parameters



(c) Gaussian in relative parameters

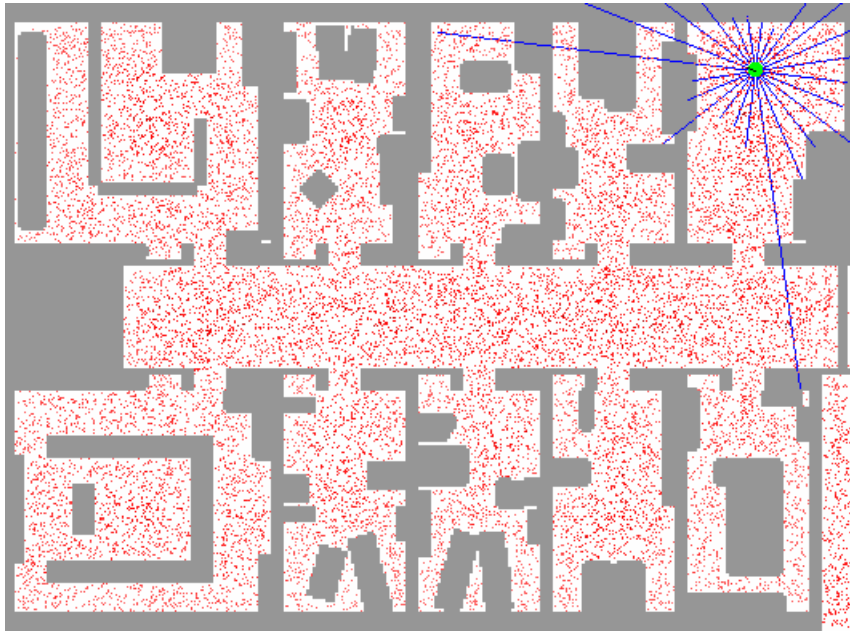


(a) absolute parameters



(b) ROP parameters

When a single Gaussian ain't good enough

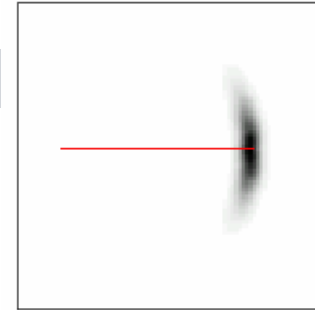


[Fox et al.]

- Sometimes, smart parameterization is not enough
 - Distribution has multiple hypothesis
- Possible solutions
 - Sampling – particle filtering
 - Mixture of Gaussians
 - ...
- Quick overview of one such solution...

Approximating non-linear KF with mixture of Gaussians

- Robot example:



- $P(X_i)$ is a Gaussian, $P(X_{i+1})$ is a banana
- Approximate $P(X_{i+1})$ as a mixture of m Gaussians
 - e.g., using discretization, sampling,...
- Problem:
 - $P(X_{i+1})$ as a mixture of m Gaussians
 - $P(X_{i+2})$ is m bananas
- One solution:
 - Apply collapsing algorithm to project m bananas in m' Gaussians

What you need to know

■ Kalman filter

- Probably most used BN
- Assumes Gaussian distributions
- Equivalent to linear system
- Simple matrix operations for computations

■ Non-linear Kalman filter

- Usually, observation or motion model not CLG
- Use numerical integration to find Gaussian approximation

■ Switching Kalman filter

- Hybrid model – discrete and continuous vars.
- Represent belief as mixture of Gaussians
- Number of mixture components grows exponentially in time
- Approximate each time step with fewer components

■ Assumed density filtering

- Fundamental abstraction of most algorithms for dynamical systems
- Assume representation for density
- Every time density not representable, project into representation