# Mixture Models and the EM Algorithm

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#### **Density Estimation**

Parametric approaches require us to assume a particular distribution function, e.g., gaussian.

The real data may not fit any one simple distribution.

Non-parametric approaches use the data itself to estimate density, e.g., by histogramming (Parzen windows).

Expensive; requires us to store all the data points.

Mixture models offer a compromise approach:

Model the data as a <u>mixture</u> of several parameterized distributions.

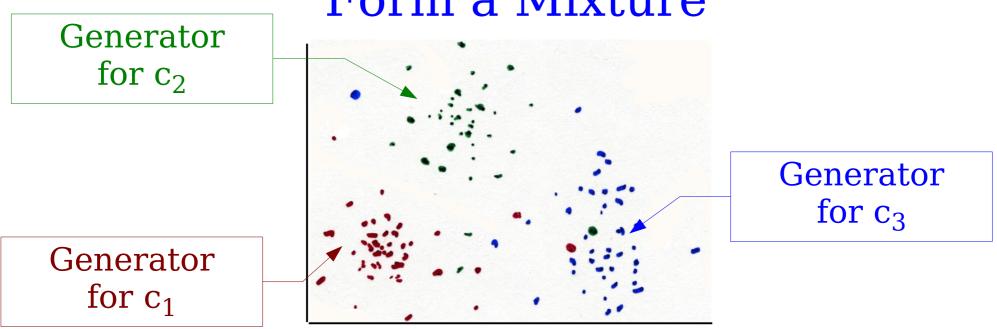
#### A Simple Mixture Model

We can model a dataset as a collection of points generated by a mixture of Gaussians.

- $\mu_i$  mean
- $\sigma_i$  standard deviation
- $\alpha_i$  weight/prevalence/prior probability

$$\left(\sum_{j}\alpha_{j}\right)=1$$
, so  $\alpha_{j}=P(c_{j})$ 

# Gaussian Generators Form a Mixture



Class priors  $P_i = P(c_i)$  (mixture coefficients)

$$\sum_{j} P_{j} = 1$$

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 Density distribution 
$$= \sum_{j} P_{j} \cdot \exp\left[\frac{-(\mathbf{x} - \mathbf{\mu}_{j})^{2}}{\sigma_{j}^{2}}\right]$$

### **Probability Densities**

$$P(j) = prior probability of class c_j$$
  
so  $\sum_{j} P(j) = 1$ 

Probability density of the mixture:

$$p(\mathbf{x}) = \sum_{j=1}^{M} p(\mathbf{x}|j)P(j)$$

Posterior probability:

$$P(j|\mathbf{x}) = \frac{p(\mathbf{x}|j)P(j)}{p(\mathbf{x})}$$

so 
$$\sum_{j} P(j|\mathbf{x}) = 1$$

#### **Conditional Density**

Assume covariance matrix is diagonal with equal elements. Then:

$$p(\mathbf{x}|j) = \frac{1}{(2\pi\sigma_j^2)^{d/2}} \cdot \exp\left\{\frac{-\|\mathbf{x} - \mathbf{\mu_j}\|^2}{2\sigma_j^2}\right\}$$
normalization
term

How can we determine the "most probable" values of  $\mu_j$  and  $\sigma_j$  and P(j), given the dataset  $\{\boldsymbol{x}_i\}$ ?

#### Likelihood of a Dataset

What is the likelihood  $\mathcal{L}$  that a dataset  $\{\mathbf{x}_{i}\}$  was generated by a given mixture model?

$$p(\mathbf{x_i}) = \sum_{j=1}^{M} p(\mathbf{x_i}|j) \cdot P(j)$$

$$p(\{\boldsymbol{x_i}\}) = \prod_{i=1}^n p(\boldsymbol{x_i}) = \mathcal{L}$$

### Log Likelihood

For gradient descent, we want a sum, not a product, because the derivative of a product is messy. So take the negative log.

$$E = -\log \mathcal{L} = -\sum_{i=1}^{n} \log p(\mathbf{x_i})$$
$$= -\sum_{i=1}^{n} \log \left\{ \sum_{j=1}^{M} p(\mathbf{x_i}|j) P(j) \right\}$$

#### Gradient Descent on E

$$E = -\sum_{i=1}^{n} \log p(\mathbf{x}_i) = -\sum_{i=1}^{n} \log \left( \sum_{j=1}^{M} p(\mathbf{x}_i|j) P(j) \right)$$

$$\frac{\partial E}{\partial \mu_j} = -\sum_{i=1}^n \left( \frac{1}{p(\mathbf{x_i})} \cdot \sum_{k=1}^M P(k) \frac{\partial}{\partial \mu_j} p(\mathbf{x_i}|k) \right)$$

$$= -\sum_{i=1}^{n} \left( \frac{1}{p(\mathbf{x}_i)} \cdot p(\mathbf{x}_i|j) P(j) \cdot \frac{\|\mathbf{x}_i - \mathbf{\mu}_j\|}{\sigma} \right)$$

$$= -\sum_{i=1}^{n} P(j|\mathbf{x_i}) \cdot \frac{\|\mathbf{x_i} - \mathbf{\mu_j}\|}{\sigma}$$

#### Gradient Descent (cont.)

 $E = -\log \mathcal{L}$  is our error function.

*Do gradient descent on E:* 

$$\frac{\partial E}{\partial \mu_j} = -\sum_{i=1}^n P(j|\mathbf{x_i}) \cdot \frac{\|\mathbf{x_i} - \mu_j\|}{\sigma_j^2}$$

$$\frac{\partial E}{\partial \sigma_j} = \sum_{i=1}^n \left| P(j|\mathbf{x_i}) \cdot \left\{ \frac{d}{\sigma_j} - \frac{\|\mathbf{x_i} - \mathbf{\mu_j}\|^2}{\sigma_j^3} \right\} \right|$$

#### Constrained Gradient Descent

#### There's a problem:

We can't just nudge  $\mu_i$  and  $\sigma_j$  around by small amounts, since their values affect P(j).

#### Must satisfy these constraints:

$$\sum_{j=1}^{M} P(j) = 1$$

$$0 \leq P(j) \leq 1$$

# The EM Algorithm (Expectation-Maximization)

Iterative algorithm for optimizing  $\mu_j$  and  $\sigma_j$ , and  $P_i$  to minimize E.

Definitions:

$$P_{ji} = P(j|\mathbf{x_i}) = \frac{P(\mathbf{x_i}|j) \cdot P(j)}{p(\mathbf{x_i})}$$

Class priors:

$$P(j) = P_j = \frac{1}{N} \sum_{i=1}^{N} P_{ji}$$

$$p(\mathbf{x}) = \sum_{j=1}^{M} p(\mathbf{x}|j) \cdot P(j)$$

#### Expectation (E) Step

Calculate  $P_{ji}$  for all points i and gaussians j,

using current parameter values  $\mu_i$ ,  $\sigma_i^2$ , and  $P_i$ .

#### Maximization (M) Step

$$P_j \leftarrow \frac{1}{N} \sum_{i=1}^{N} P_{ji}$$

$$\mu_{j} \leftarrow \frac{\sum_{i} P_{ji} \cdot \mathbf{x}_{i}}{\sum_{i} P_{ji}} = \frac{\sum_{i} P_{ji} \cdot \mathbf{x}_{i}}{P_{j}}$$

$$\sigma_j^2 \leftarrow \frac{\frac{1}{d} \sum_{i} P_{ji} \cdot ||\mathbf{x}_i - \mathbf{\mu}_j||^2}{P_j}$$

Repeat E and M steps until convergence.

# EM Finds a Local Minimum in E (or Local Maximum in \( \alpha \)

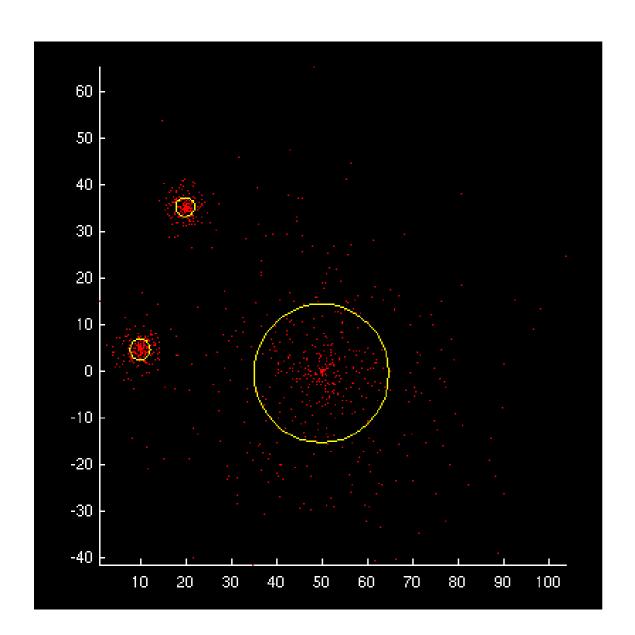
$$\frac{\partial E}{\partial \mu_{j}} = 0 = -\sum_{i=1}^{n} P(j|\mathbf{x}_{i}) \cdot \frac{\|\mathbf{x}_{i} - \mu_{j}\|}{\sigma_{j}^{2}}$$

$$\mu_{j} \sum_{i=1}^{n} \frac{P(j|\mathbf{x}_{i})}{\sigma_{j}^{2}} = \frac{\sum_{i=1}^{n} P(j|\mathbf{x}_{i}) \mathbf{x}_{i}}{\sigma_{j}^{2}}$$

$$\mu_{j} = \frac{\sum_{i=1}^{n} P(j|\mathbf{x}_{i}) \mathbf{x}_{i}}{\sum_{i=1}^{n} P(j|\mathbf{x}_{i})}$$

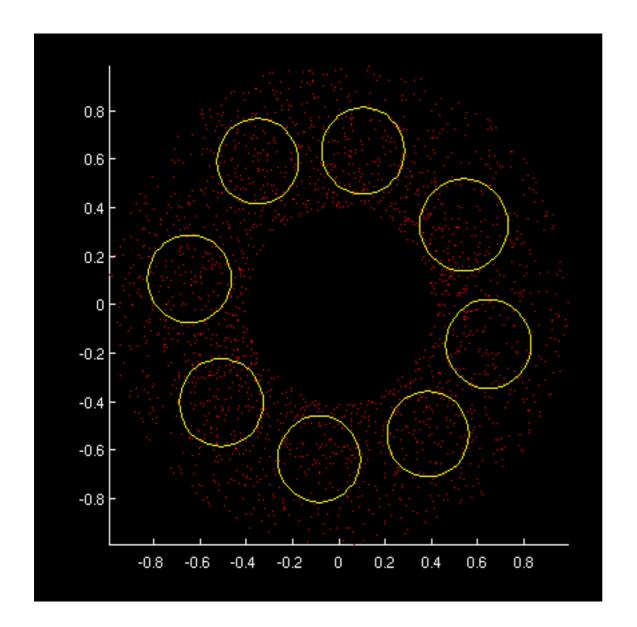
 $\mu_j$  is the weighted mean of the  $\mathbf{x_i}$ 's credited to the jth mixture component.

#### EM Demo



EmDataSet =4 emdemo

#### EM Demo



EmDataSet = 3 emdemo

## Cheapo Heuristic

Not as good as split/merge (defined later), but easy to program:

If a component captures fewer than 1/(2M) points, reset its  $\mu$  to a random  $\mathbf{x}_i$  and recalculate its  $\sigma^2$ .

Assumes the P(j) values are roughly equal.

EmHeuristic = 1 emdemo

#### Characteristics of EM

Learns in a small number of iterations.

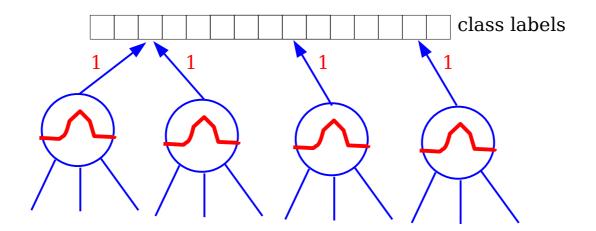
Can get stuck in local minima.

But you can add heuristics to help unstick the algorithm.

Must decide in advance how many Gaussians.

#### Williamson: Gaussian ARTMAP

1. Use an RBF network to do pattern classification:



Each unit votes for one class. Tally votes from all active units. The class with the most votes wins. No LMS training.

2. Use a variant of EM to train the gaussians.

# "Match Tracking" in ARTMAP

Establish a match threshold  $\rho$ .

Units count as "active" only if  $P(x_i|j) > \rho$ .

All other units are reset to zero; they do not vote.

If the network guesses the wrong class, increase  $\rho$  slightly and try again.

If  $\rho$  gets too high and all units are reset, then add a new unit to handle this data point.

#### Performance of Gaussian ARTMAP

Note: EM is a batch (offline) learning algorithm. Guassian ARTMAP uses an online variant.

Did well on several tasks:

Letter image classification

Landsat satellite image segmentation

Speaker-independent vowel recognition

Match tracking helps Gaussian ARTMAP outperform EM by "backpropagating" the effects of erroneous classifications.

## Offline Calculation of $\mu$ and $\sigma^2$

$$\mu = \langle x \rangle$$

$$\sigma^{2} = \langle (x - \mu)^{2} \rangle$$

$$= \langle x^{2} - 2x\mu + \mu^{2} \rangle$$

$$= \langle x^{2} \rangle - 2\langle x \rangle \mu + \langle \mu^{2} \rangle$$

$$= \langle x^{2} \rangle - 2\mu^{2} + \mu^{2}$$

$$= \langle x^{2} \rangle - \langle x \rangle^{2}$$

### On-line Calculation of $\mu$

$$\mu_n = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\mu_{n+1} = \frac{n \cdot \mu_n + x_{n+1}}{n+1}$$

$$= \frac{n}{n+1} \mu_n + \frac{x_n}{n+1}$$

$$= \left(1 - \frac{1}{n+1}\right)\mu_n + \frac{1}{n+1}x_{n+1}$$

#### On-line Calculation of $\sigma^2$

$$\sigma_n^2 = \frac{1}{n} \sum_{i=1}^n \left( x_i - \mu_n \right)^2$$

$$\sigma_{n+1}^2 = \left(1 - \frac{1}{n+1}\right)\sigma_n^2 + \frac{1}{n+1}(x_{n+1} - \mu_{n+1})^2$$

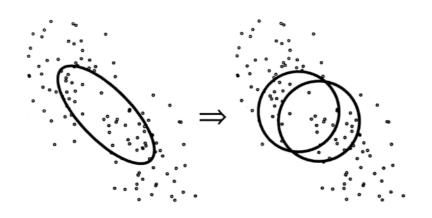
 $\sigma_n^2$  is slightly biased because  $\mu_n$  changes, but the effect is not significant.

# Split and Merge EM (Ueda, Nakano, Gharamani, and Hinton)

Split a component if it does a poor job estimating the local density.

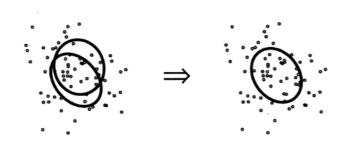
Example: a component stuck between two clusters will have low density near its mean and high density near the true cluster centers.

To split, make two copies, and perturb each one away from the mean by a small amount.



## Split and Merge EM (cont.)

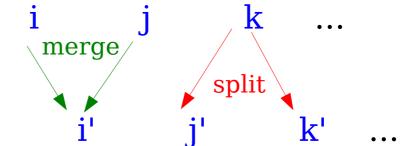
Merge two components if their parameters are close. Set the merged component's parameters to the weighted average.



Combine one merge step with one split step, so the number of mixture components M stays the same.

### Combined Split and Merge Steps

Old components:



New components:

Run a mini-EM step to adapt elements i', j', and k'. Then run full EM to asymptote.

If overall likelihood is not improved, undo the split/merge and try a different set of candidates.

Candidates are ranked heuristically; only need to look at about 5.

### When To Merge

Define 
$$P_i(\Theta^*) = \langle P(i|x_1; \Theta^*), \dots, P(i|x_N; \Theta^*) \rangle^T$$

Merge criterion:

$$J_{merge}(i,j;\Theta^*) = \frac{P_i(\Theta^*)^T P_j(\Theta^*)}{\|P_i(\Theta^*)\| \ \|P_j(\Theta^*)\|}$$

 $J_{merge}$  will be high if the vectors  $P_i(\Theta^*)$  and  $P_j(\Theta^*)$  point in roughly the same direction, meaning their gaussians capture roughly the same set of points.

Choose candidate pairs (i,j) to give the highest value of  $J_{merae}(i,j;\Theta^*)$ 

#### When To Split

Estimate the local density of kth model around x:

stimate the local density of kth model 
$$f_k(x; \Theta^*) = \frac{\sum_{n=1}^{N} \delta(x - x_n) P(k|x_n; \Theta^*)}{\sum_{n=1}^{N} P(k|x_n; \Theta^*)}$$

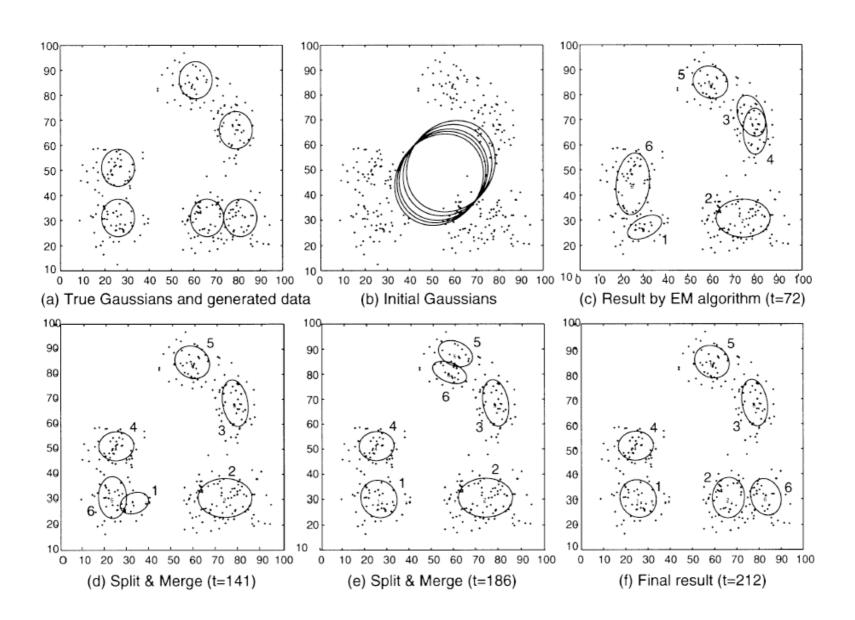
where  $\delta$  is the empirical density distribution.

Check Kullback-Leibler divergence of local data density estimate  $f_k$ with density  $p_k$  of kth model as specified by current parameter estimate  $\Theta^*$ :

$$J_{split}(k; \Theta^*) = \int f_k(x; \Theta^*) \log \frac{f_k(x; \Theta^*)}{p_k(x; \Theta^*)} dx$$

A large value of  $J_{split}$  suggests that model k has a poor estimate of local density, so splitting it might be helpful.

#### Performance of SMEM

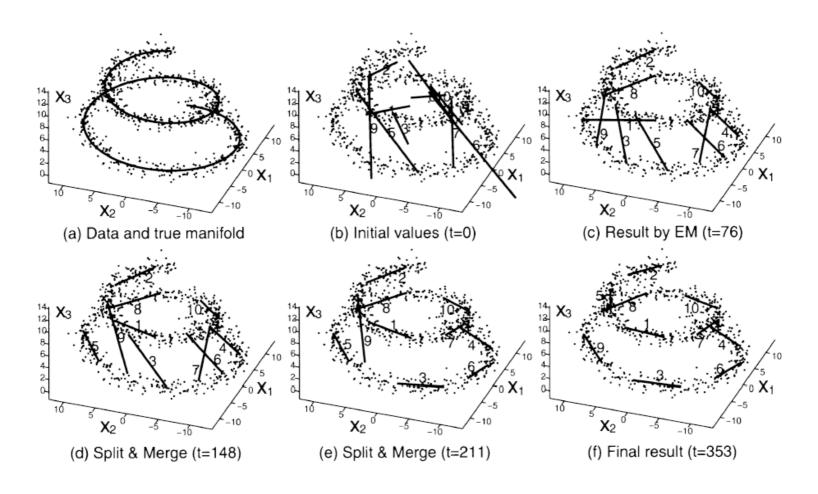


#### Factor Analyzers

- More complex approach than EM.
- Able to handle high-dimensional data with low-dimensional embedded structure.

#### Mixture of Factor Analyzers

Project high-D space down to lower-D space. Compute a mixture of low-D functions.



# **Image Reconstruction**

