

Great Theoretical Ideas In Computer Science

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CS 15-251 Spring 2005

Lecture 21

March 29, 2005

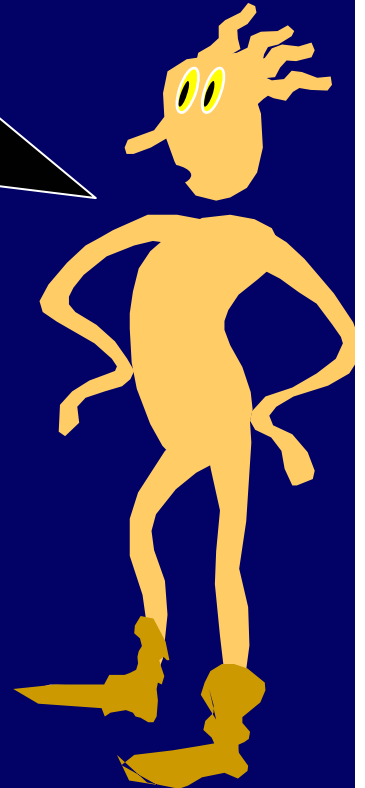
Carnegie Mellon University

15-251
Classics



Today, we will learn
about a formidable tool
in probability that will
allow us to solve
problems that seem
really really messy...

If I randomly put 100 letters
into 100 addressed envelopes,
on average how many letters
will end up in their correct
envelopes?



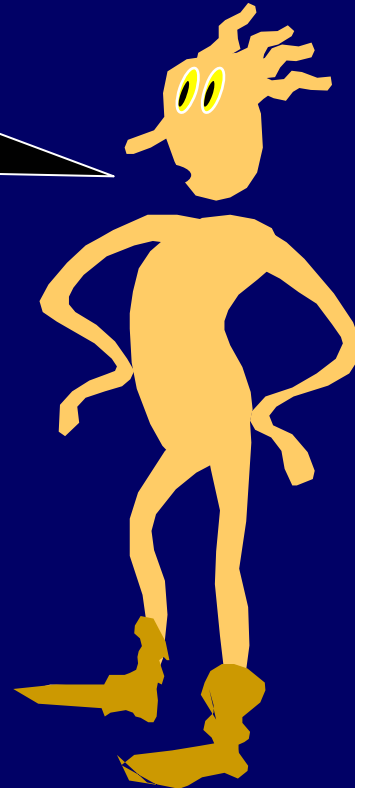
Hmm...

$\sum_k k \cdot \Pr(\text{exactly } k \text{ letters end up in}$
 $\text{correct envelopes})$

$= \sum_k k \cdot (\dots \text{aargh!!} \dots)$



On average, in class of size m , how many pairs of people will have the same birthday?

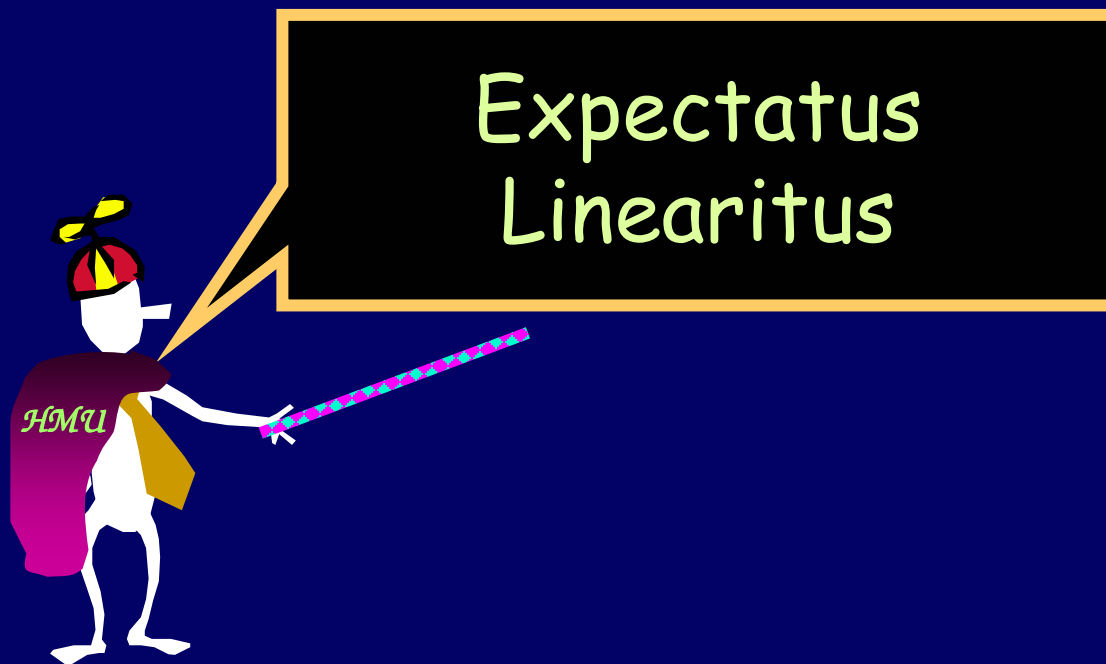


$$\sum_k k \cdot \Pr(\text{exactly } k \text{ collisions})$$

$$= \sum_k k \cdot (\dots \text{aargh!!!!} \dots)$$



The new tool is called
"Linearity of
Expectation"



Random Variable

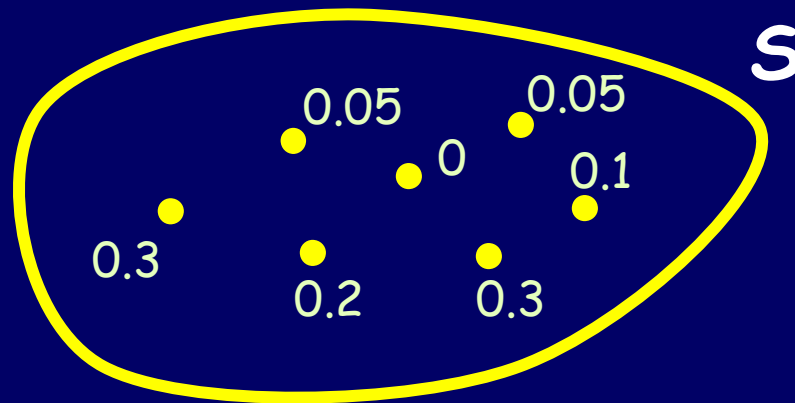
To use this new tool, we will also need to understand the concept of a Random Variable

Today's goal: not too much material, but to understand it well.

Probability Distribution

A (finite) probability distribution D

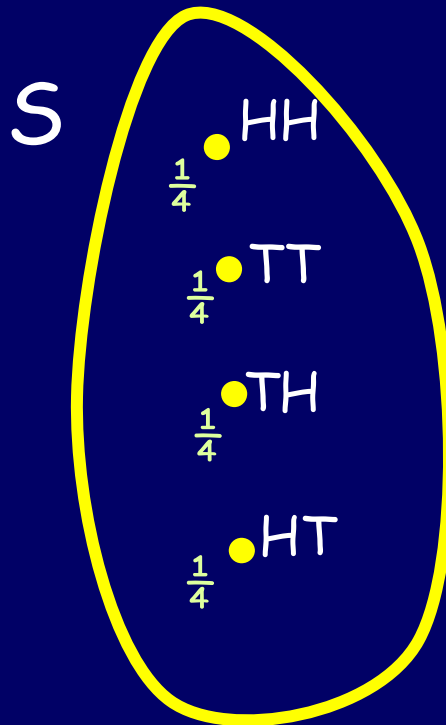
- a finite set S of elements (samples)
- each $x \in S$ has weight or probability $p(x) \in [0,1]$



weights must sum to 1

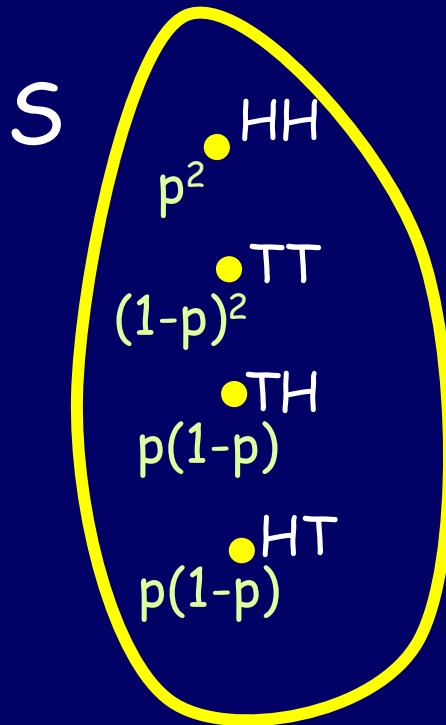
"Sample space"

Flip penny and nickel (unbiased)

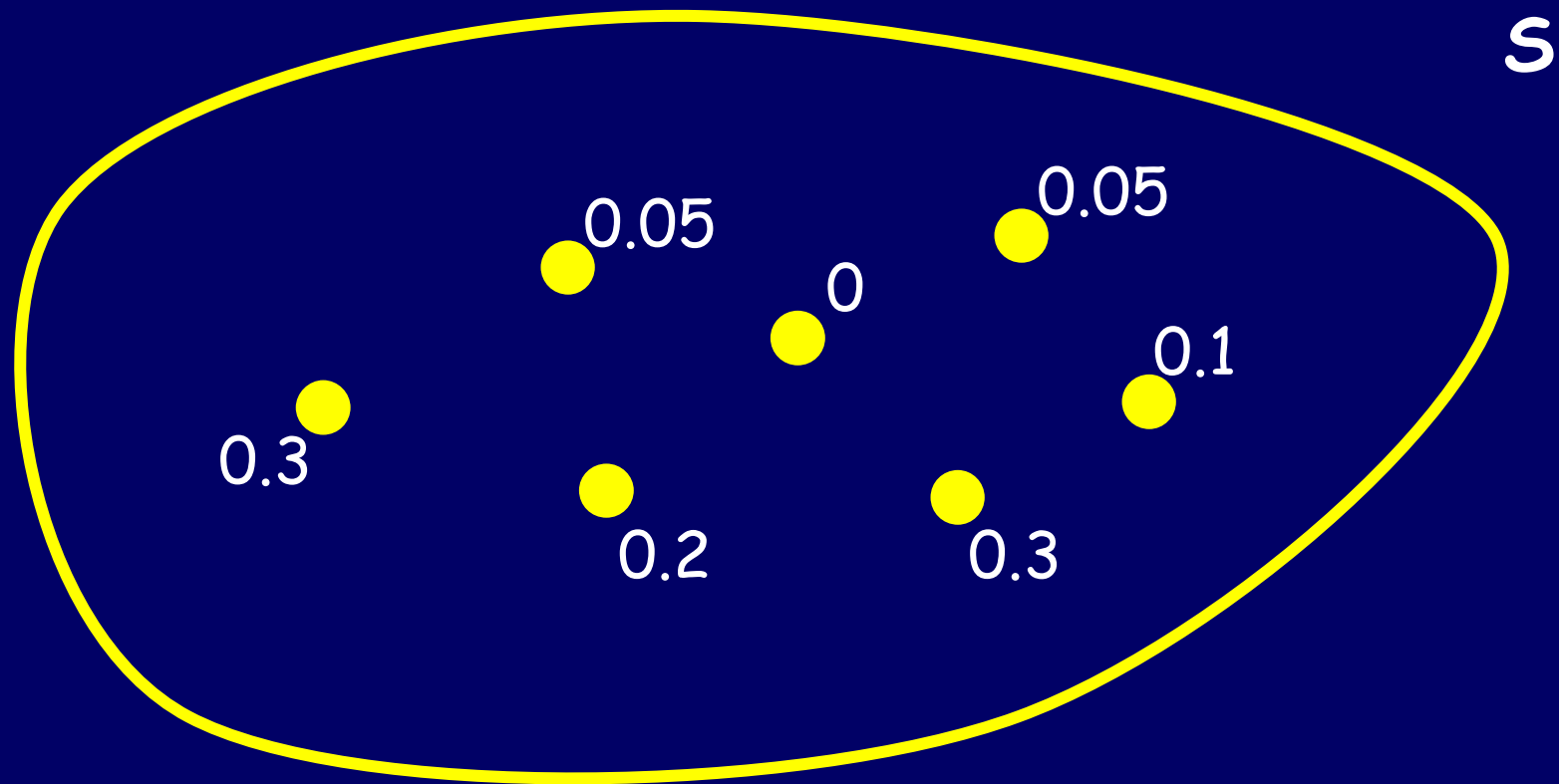


Flip penny and nickel (biased)

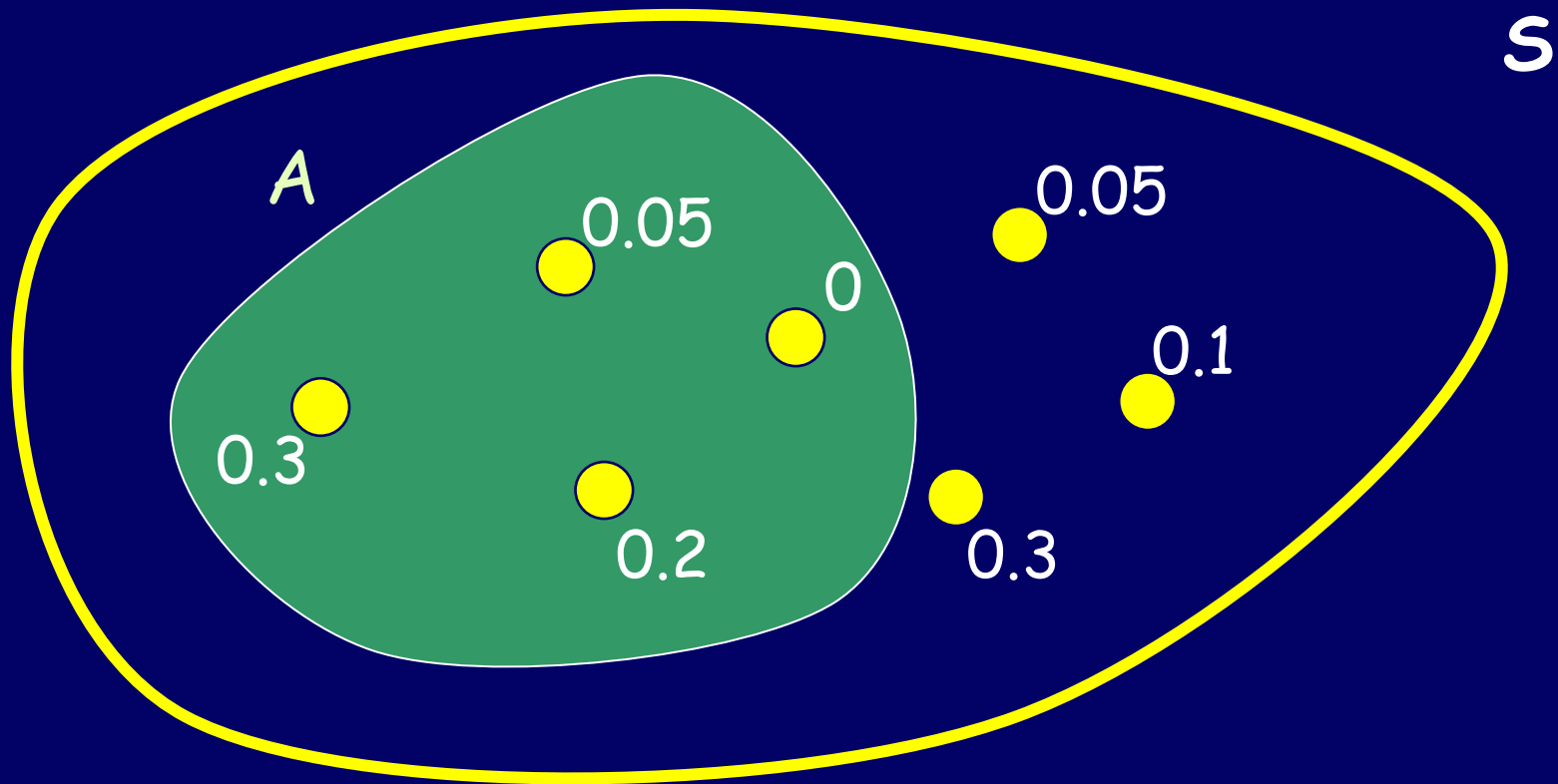
heads probability = p



Probability Distribution



An event is a subset



$$\Pr[A] = \sum_{x \in A} p(x) = 0.55$$

Running Example

I throw a white die and a black die.

Sample space $S =$

(1,1),	(1,2),	(1,3),	(1,4),	(1,5),	(1,6),
(2,1),	(2,2),	(2,3),	(2,4),	(2,5),	(2,6),
(3,1),	(3,2),	(3,3),	(3,4),	(3,5),	(3,6),
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(6,1),	(6,2),	(6,3),	(6,4),	(6,5),	(6,6)

$$\Pr(x) = 1/36 \\ \forall x \in S$$

E = event that $\text{sum} \leq 3$

$$\Pr[E] = |E|/|S| = \text{proportion of } E \text{ in } S = 3/36$$

New concept:
Random Variables

Random Variables

Random Variable: a (real-valued) function on S

Examples:

X = value of white die.

$$X(3,4) = 3, \quad X(1,6) = 1 \text{ etc.}$$

Y = sum of values of the two dice.

$$Y(3,4) = 7, \quad Y(1,6) = 7 \text{ etc.}$$

W = (value of white die)^{value of black die}

$$W(3,4) = 3^4 \quad Y(1,6) = 1^6$$

Z = (1 if two dice are equal, 0 otherwise)

$$Z(4,4) = 1, \quad Z(1,6) = 0 \text{ etc.}$$

Toss a white die and a black die.

Sample space S =

{ (1,1), (1,2), (1,3), (1,4), (1,5), (1,6),
(2,1), (2,2), (2,3), (2,4), (2,5), (2,6),
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(6,1), (6,2), (6,3), (6,4), (6,5), (6,6) }

E.g., tossing a fair coin n times

S = all sequences of $\{H, T\}^n$

D = uniform distribution on S

$$\Rightarrow D(x) = \left(\frac{1}{2}\right)^n \quad \text{for all } x \in S$$

Random Variables (say $n = 10$)

X = # of heads

$$X(\text{HHHTTHTHTT}) = 5$$

Y = (1 if #heads = #tails, 0 otherwise)

$$Y(\text{HHHTTHTHTT}) = 1, Y(\text{THHHHTTTTT}) = 0$$

Notational conventions

Use letters like A, B, E for events.

Use letters like X, Y, f, g for R.V.'s.

R.V. = random variable

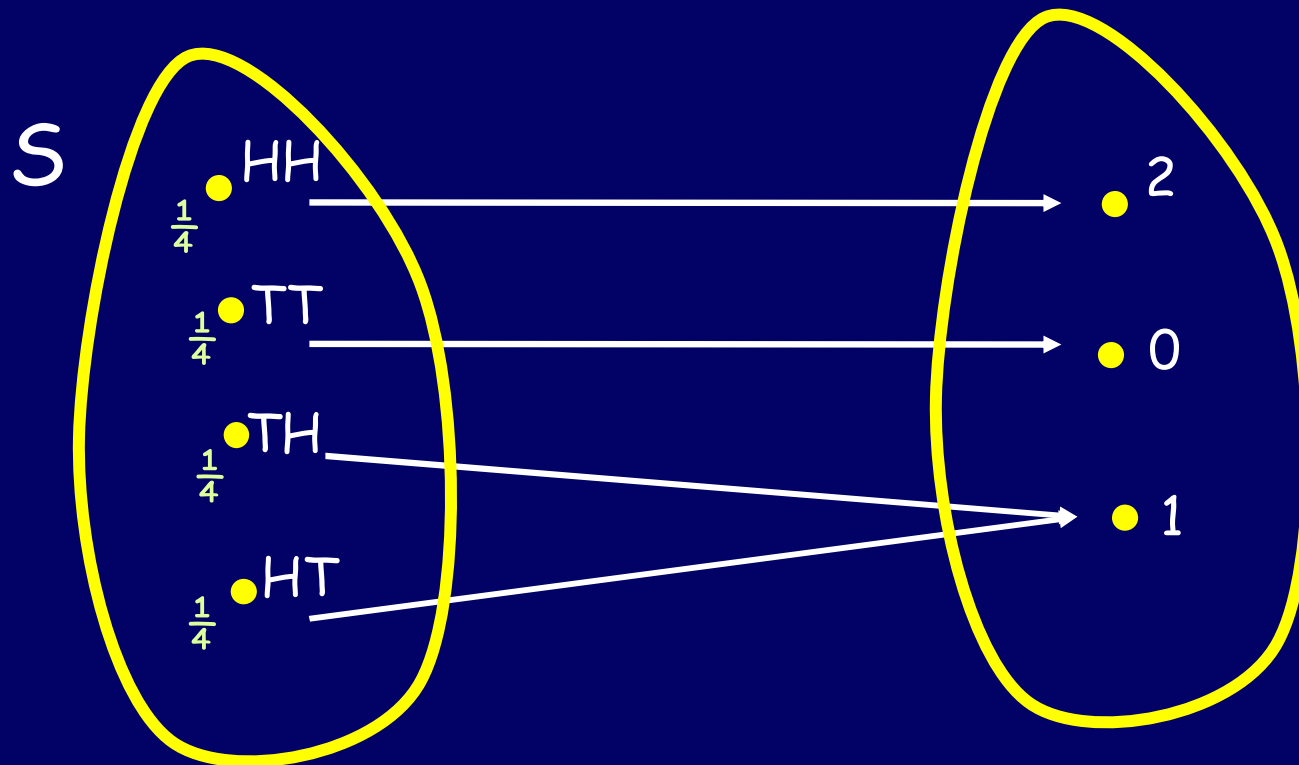
Two views of random variables

Think of a R.V. as

- a function from S to the reals P
- or think of the induced distribution on P

Two coins tossed

$X: \{TT, TH, HT, HH\} \rightarrow \{0, 1, 2\}$
counts the number of heads



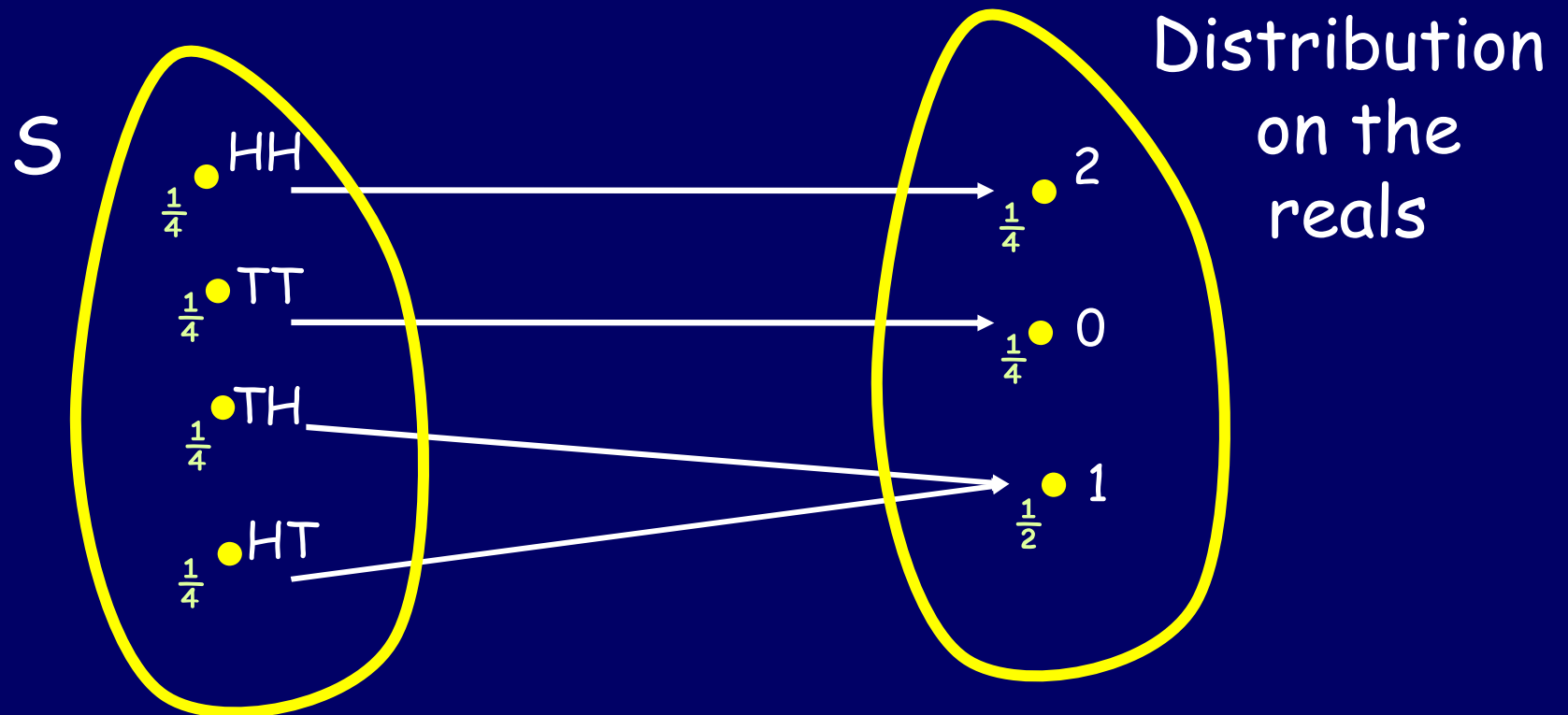
Two views of random variables

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Two views of random variables

Think of a R.V. as

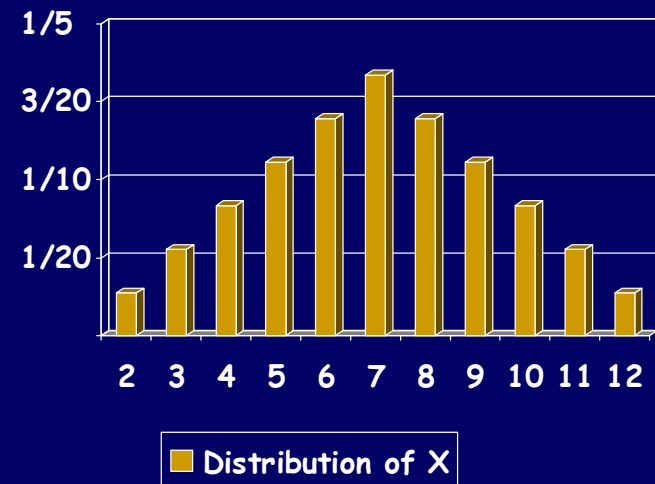
- a function from S to the reals P
- or think of the induced distribution on P

Two dice

I throw a white die and a black die.

Sample space $S =$

{ (1,1), (1,2), (1,3), (1,4), (1,5), (1,6),
(2,1), (2,2), (2,3), (2,4), (2,5), (2,6),
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(6,1), (6,2), (6,3), (6,4), (6,5), (6,6) }



X = sum of both dice

function with $X(1,1) = 2$, $X(1,2)=X(2,1)=3$, ..., $X(6,6)=12$

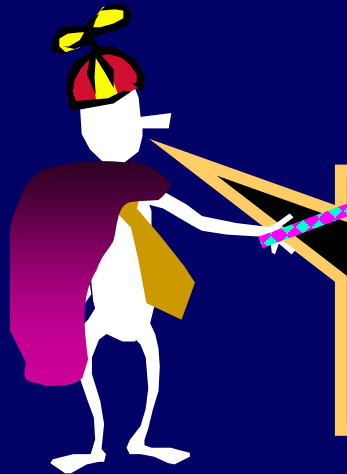
It's a floor wax *and* a dessert topping



It's a function on
the sample space S .



It's a variable with a
probability distribution
on its values.



You should be comfortable
with both views.

From Random Variables to Events

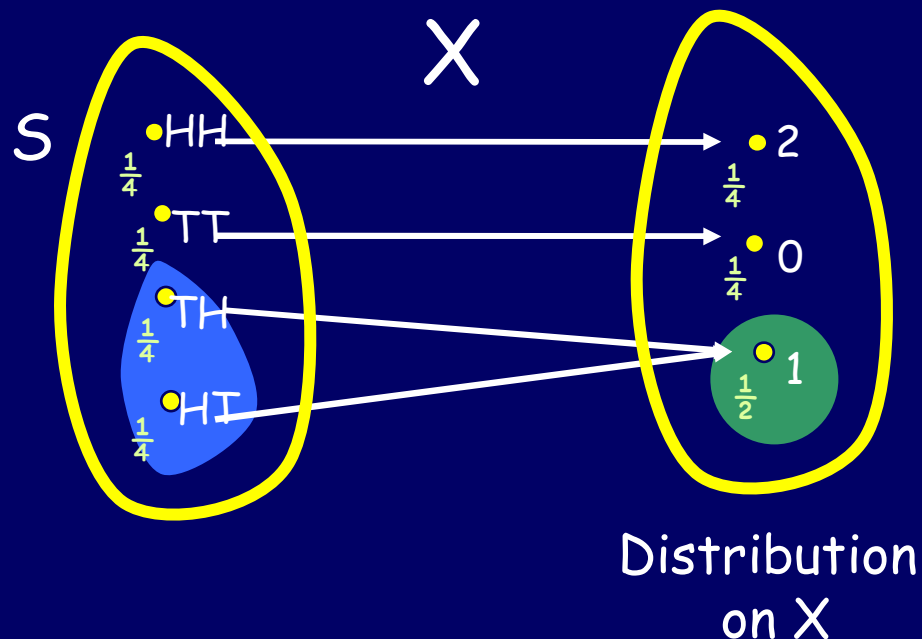
For any random variable X and value a ,
we can define the event A that $X=a$.

$$\Pr(A) = \Pr(X=a) = \Pr(\{x \in S \mid X(x)=a\}).$$

Two coins tossed

$X: \{TT, TH, HT, HH\} \rightarrow \{0, 1, 2\}$
counts the number of heads

$$\Pr(X = a) = \Pr(\{x \in S \mid X(x) = a\})$$



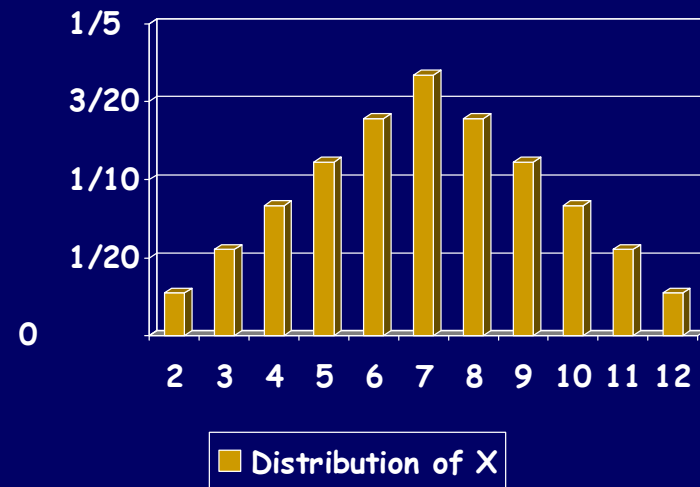
$$\begin{aligned}\Pr(X = 1) \\ &= \Pr(\{x \in S \mid X(x) = 1\}) \\ &= \Pr(\{TH, HT\}) = \frac{1}{2}.\end{aligned}$$

Two dice

I throw a white die and a black die. X = sum

Sample space S =

(1,1),	(1,2),	(1,3),	(1,4),	(1,5),	(1,6),
(2,1),	(2,2),	(2,3),	(2,4),	(2,5),	(2,6),
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(6,1),	(6,2),	(6,3),	(6,4),	(6,5),	(6,6)



$$\Pr(X = 6)$$

$$= \Pr(\{x \in S \mid X(x) = 6\})$$

$$= 5/36.$$

$$\Pr(X = a) = \Pr(\{x \in S \mid X(x) = a\})$$

From Random Variables to Events

For any random variable X and value a , we can define the event A that $X=a$.

$$\Pr(A) = \Pr(X=a) = \Pr(\{x \in S \mid X(x)=a\}).$$



X has a
distribution on
its values

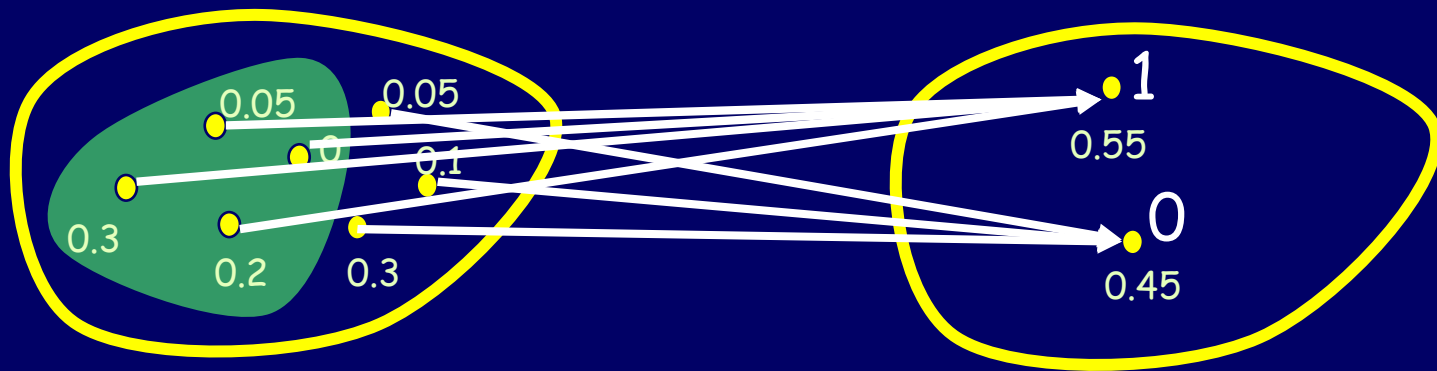


X is a function
on the sample space S

From Events to Random Variables

For any event A , can define the indicator random variable for A :

$$\begin{aligned} X_A(x) &= 1 && \text{if } x \in A \\ &= 0 && \text{if } x \notin A \end{aligned}$$



Definition: expectation

The expectation, or expected value of a random variable X is written as $E[X]$, and is

$$E[X] = \sum_{x \in S} \Pr(x) X(x) = \sum_k k \Pr(X = k)$$

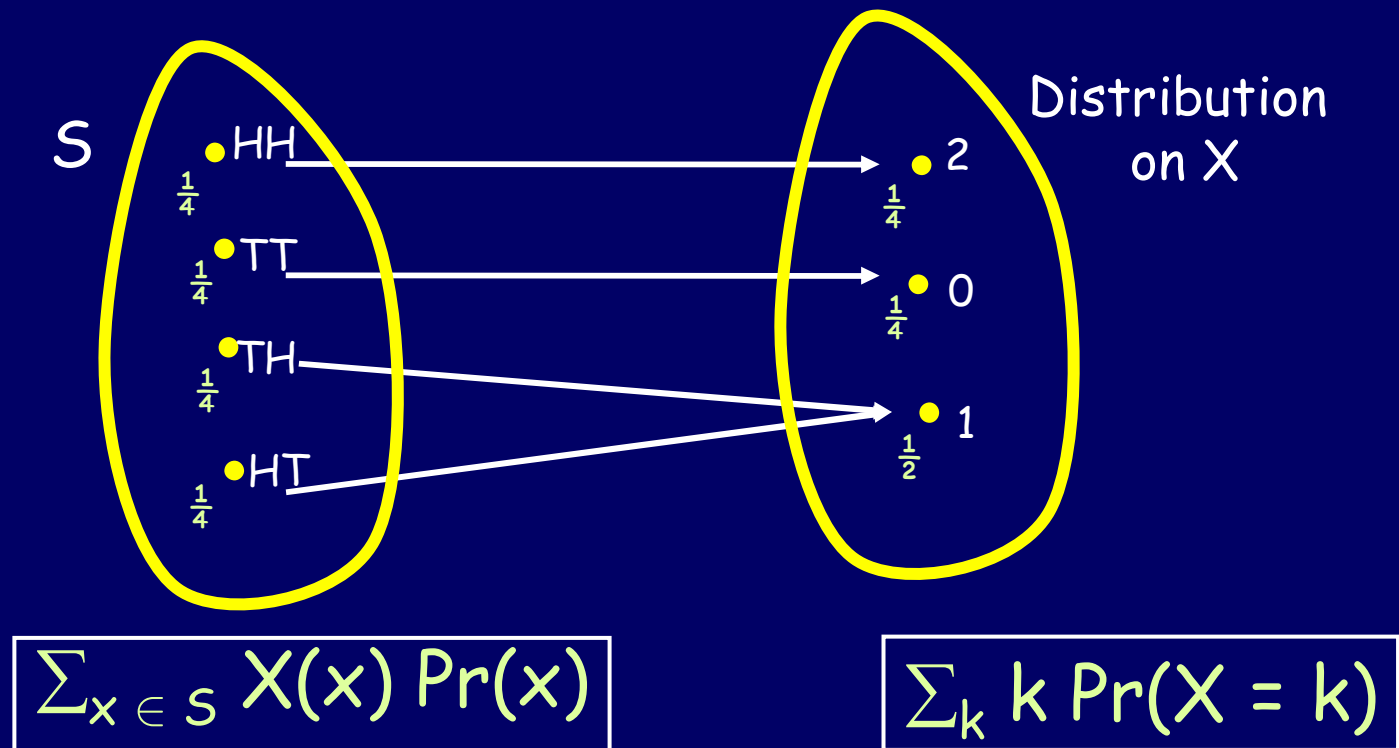


X is a function
on the sample space S



X has a
distribution on
its values

Thinking about expectation



$$E[X] = \frac{1}{4} * 0 + \frac{1}{4} * 1 + \frac{1}{4} * 1 + \frac{1}{4} * 2 = 1.$$

$$E[X] = \frac{1}{4} * 0 + \frac{1}{2} * 1 + \frac{1}{4} * 2 = 1.$$

A quick calculation...

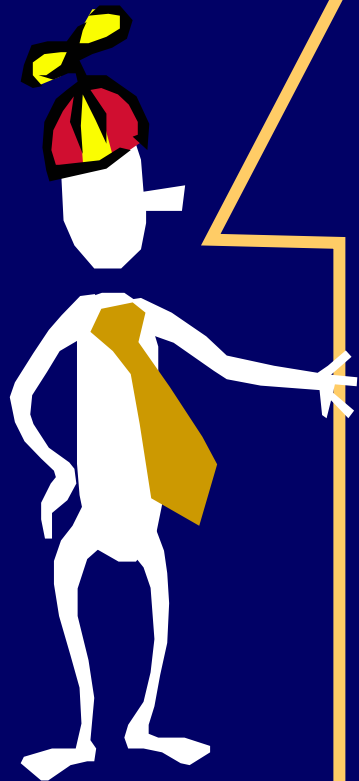
What if I flip a coin 3 times? Now what is the expected number of heads?

$$E[X] = (1/8) \times 0 + (3/8) \times 1 + (3/8) \times 2 + (1/8) \times 3 = 1.5$$

But $\Pr[X = 1.5] = 0...$

Moral: don't always expect the expected.
 $\Pr[X = E[X]]$ may be 0!

Type checking



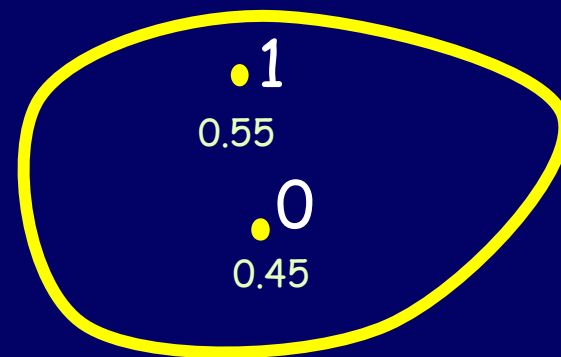
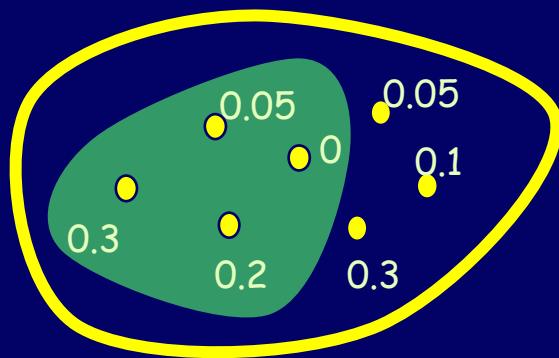
A Random Variable is the type of thing you might want to know an expected value of.

If you are computing an expectation, the thing whose expectation you are computing is a random variable.

Indicator R.V.s: $E[X_A] = \Pr(A)$

For event A , the indicator random variable for A :

$$\begin{aligned} X_A(x) &= 1 && \text{if } x \in A \\ &= 0 && \text{if } x \notin A \end{aligned}$$



$$E[X_A] = 1 \times \Pr(X_A = 1) = \Pr(A)$$

Adding Random Variables

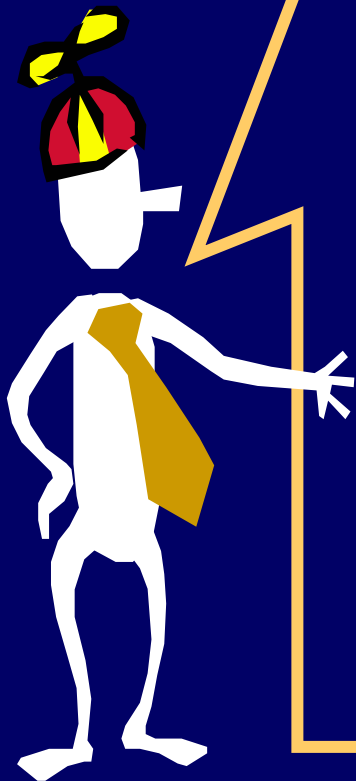
If X and Y are random variables
(on the same set S), then
 $Z = X + Y$ is also a random variable.

$$Z(x) \equiv X(x) + Y(x)$$

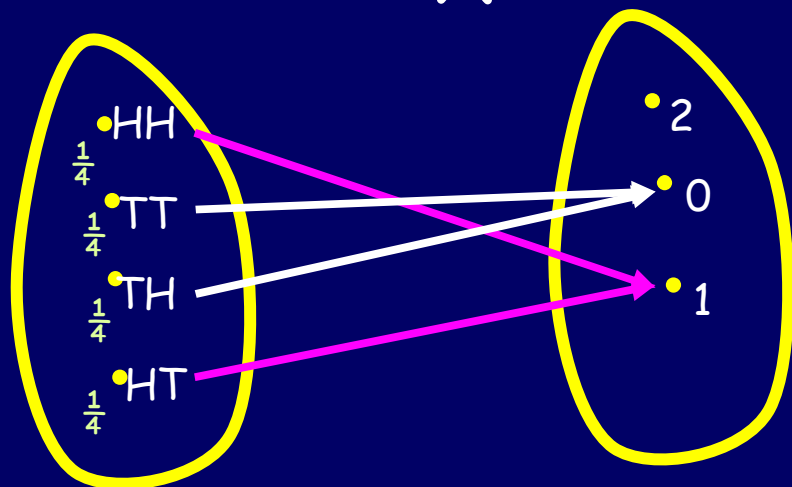
E.g., rolling two dice.

$X = 1^{\text{st}}$ die, $Y = 2^{\text{nd}}$ die,

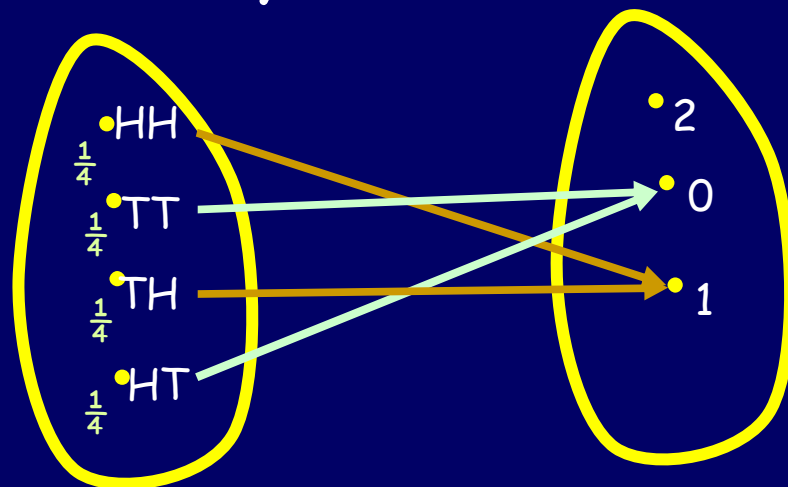
$Z = \text{sum of two dice.}$



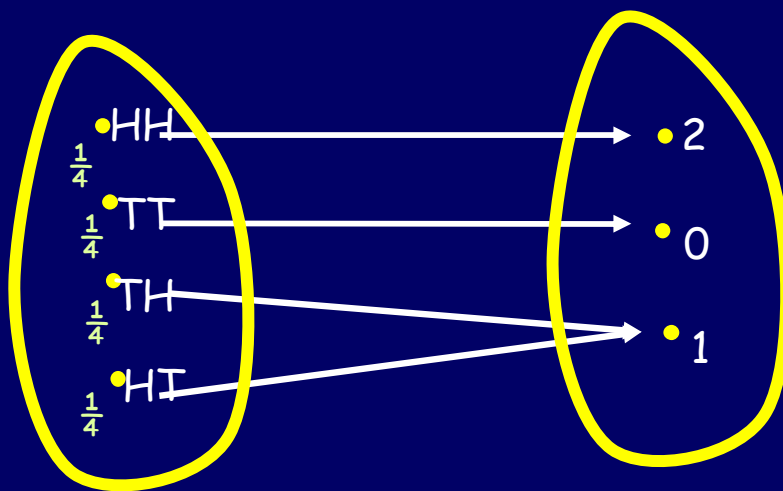
X



Y



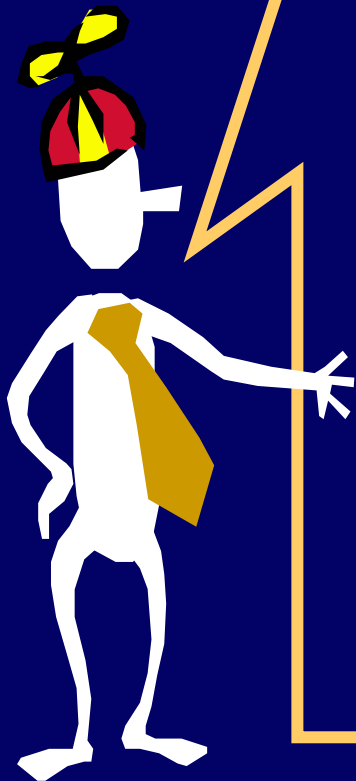
$$Z = X + Y$$



Adding Random Variables

Example: Consider picking a random person in the world. Let X = length of the person's left arm in inches. Y = length of the person's right arm in inches. Let $Z = X + Y$. Z measures the combined arm lengths.

Formally, $S = \{\text{people in world}\}$,
 $D = \text{uniform distribution on } S$.



Independence

Two random variables X and Y are independent if for every a, b , the events $X=a$ and $Y=b$ are independent.

How about the case of
 $X=1^{\text{st}}$ die, $Y=2^{\text{nd}}$ die?
 $X = \text{left arm}$, $Y = \text{right arm}$?

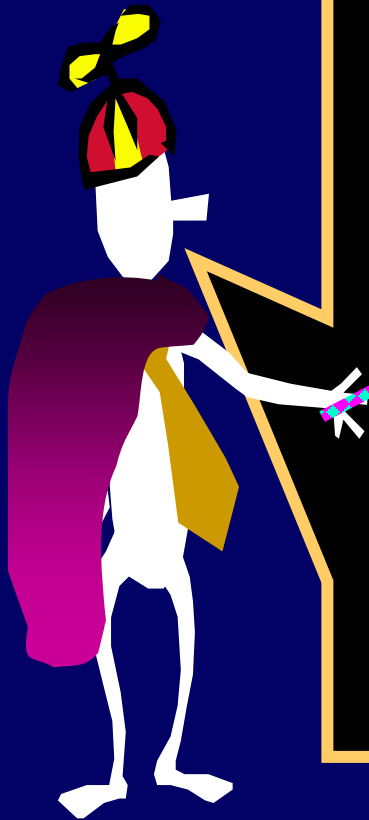


Linearity of Expectation

If $Z = X + Y$, then

$$E[Z] = E[X] + E[Y]$$

Even if X and Y are not independent.



Linearity of Expectation

If $Z = X + Y$, then

$$E[Z] = E[X] + E[Y]$$

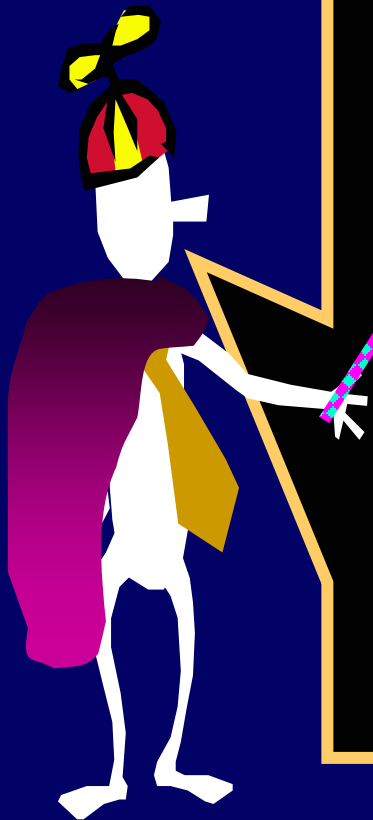
Proof:

$$E[X] = \sum_{x \in S} \text{Pr}(x) X(x)$$

$$E[Y] = \sum_{x \in S} \text{Pr}(x) Y(x)$$

$$E[Z] = \sum_{x \in S} \text{Pr}(x) Z(x)$$

$$\text{but } Z(x) = X(x) + Y(x)$$



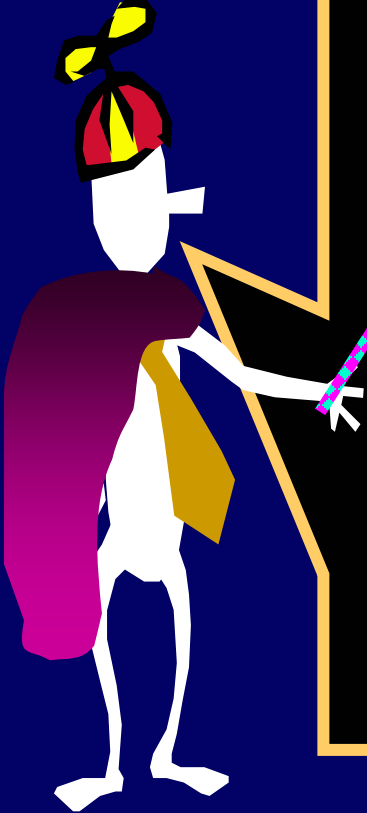
Linearity of Expectation

E.g., 2 fair flips:

X = 1st coin, Y = 2nd coin.

$Z = X + Y$ = total # heads.

What is $E[X]$? $E[Y]$? $E[Z]$?



	1,1,2	
1,0,1	HH	0,1,1
HT	0,0,0	TH
	TT	

Linearity of Expectation

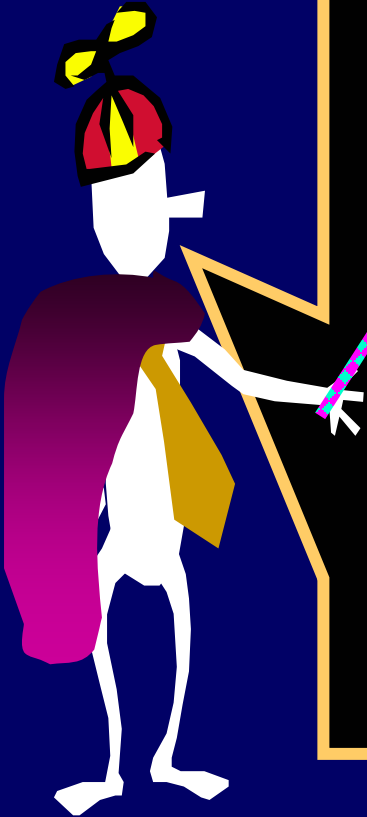
E.g., 2 fair flips:

X = at least one coin heads,

Y = both coins are heads, $Z = X + Y$

Are X and Y independent?

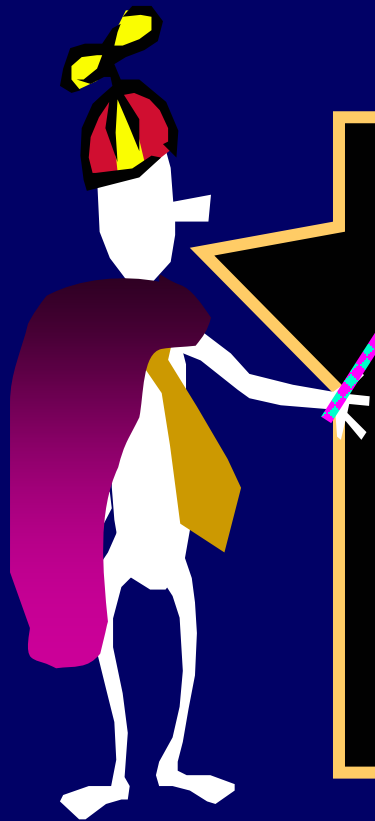
What is $E[X]$? $E[Y]$? $E[Z]$?



	1,1,2	
1,0,1	HH	1,0,1
HT	0,0,0	TH
	TT	

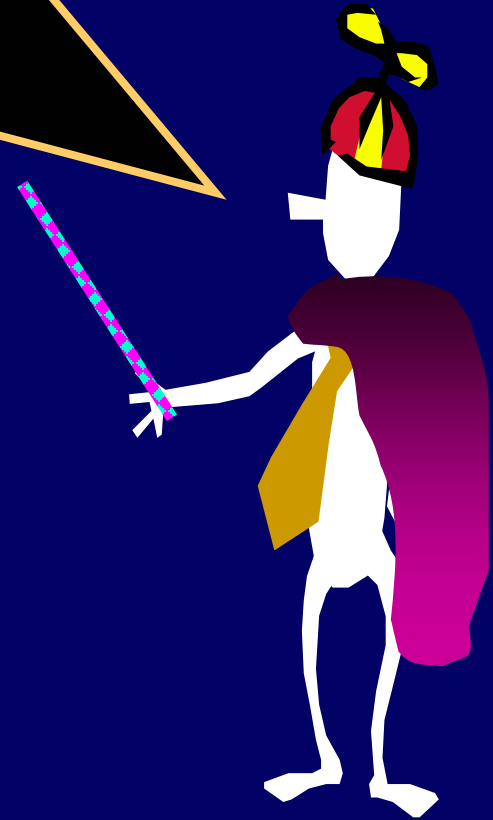
By induction

$$E[X_1 + X_2 + \dots + X_n] = E[X_1] + E[X_2] + \dots + E[X_n]$$

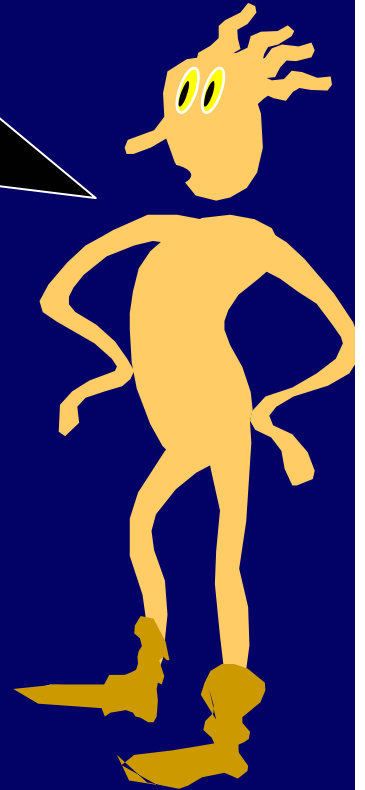


the expectation
of the sum
=
the sum of the
expectations

It is finally time
to show off our
probability
prowess...



If I randomly put 100 letters
into 100 addressed envelopes,
on average how many letters
will end up in their correct
envelopes?



Hmm...

$$\sum_k k \cdot \Pr(\text{exactly } k \text{ letters end up in correct envelopes}) \\ = \sum_k k \cdot (\dots \text{aargh!!} \dots)$$



Use Linearity of Expectation

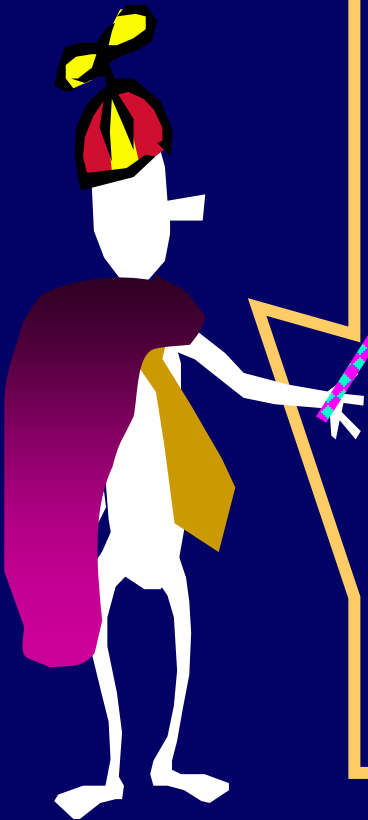
Let A_i be the event the i^{th} letter ends up in its correct envelope.

Let X_i be the indicator R.V. for A_i .

$$X_i = \begin{cases} 1 & \text{if } A_i \text{ occurs} \\ 0 & \text{otherwise} \end{cases}$$

Let $Z = X_1 + \dots + X_{100}$.

We are asking for $E[Z]$.



Use Linearity of Expectation

Let A_i be the event the i^{th} letter ends up in the correct envelope.

Let X_i be the indicator R.V. for A_i .

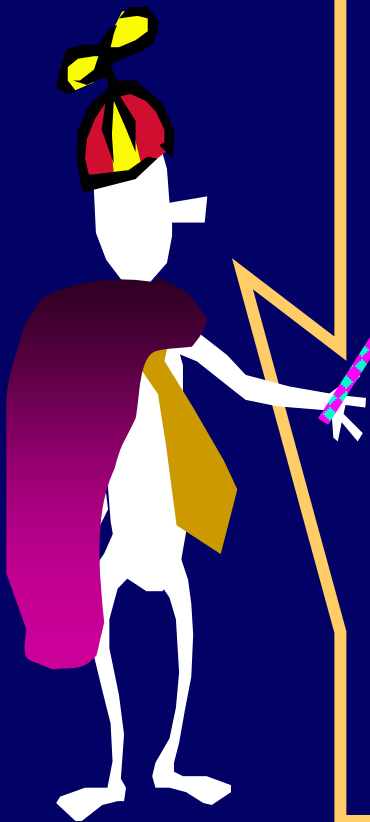
Let $Z = X_1 + \dots + X_n$. We are asking for $E[Z]$.

What is $E[X_i]$?

$$E[X_i] = \Pr(A_i) = 1/100.$$

What is $E[Z]$?

$$\begin{aligned} E[Z] &= E[X_1 + \dots + X_{100}] \\ &= E[X_1] + \dots + E[X_{100}] \\ &= 1/100 + \dots + 1/100 = 1. \end{aligned}$$



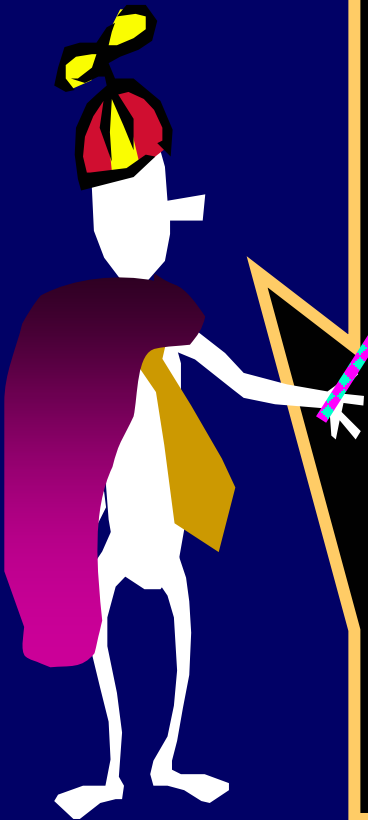
Use Linearity of Expectation

So, in expectation, 1 card will be in the same position as it started.

Pretty neat: it doesn't depend on how many cards!

Question: were the X_i independent?

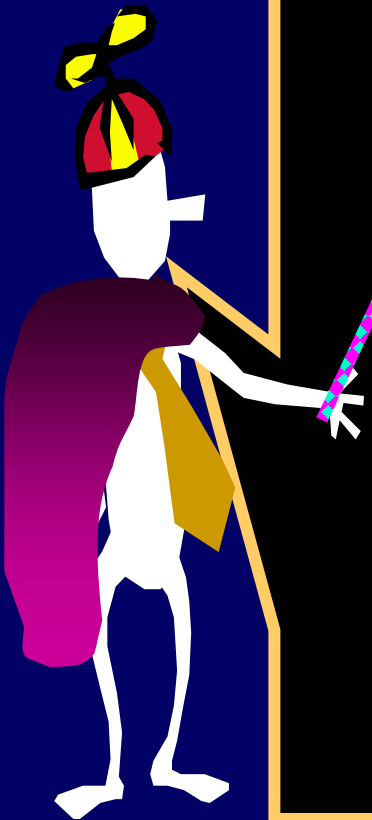
No! E.g., think of $n=2$.



Use Linearity of Expectation

General approach:

- View thing you care about as expected value of some RV.
- Write this RV as sum of simpler RVs (typically indicator RVs).
- Solve for their expectations and add them up!



Example

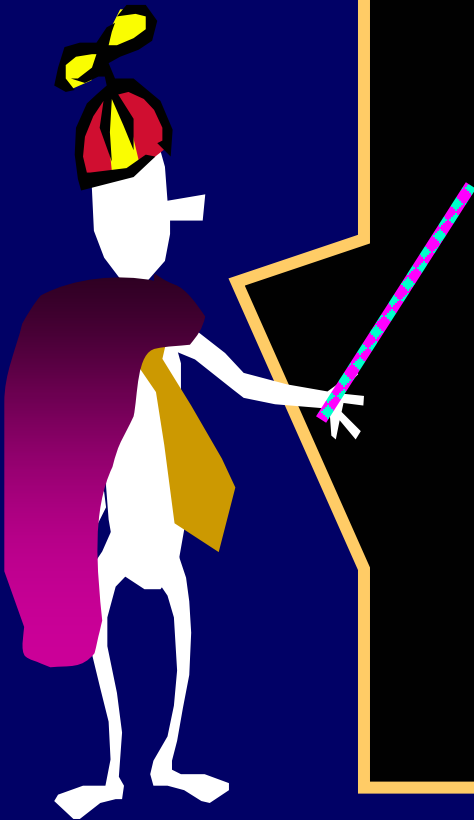
We flip n coins of bias p .
What is the expected
number of heads?

We could do this by summing

$$\sum_k k \Pr(X = k)$$

$$= \sum_k k \binom{n}{k} p^k (1-p)^{n-k}$$

But now we know a better way

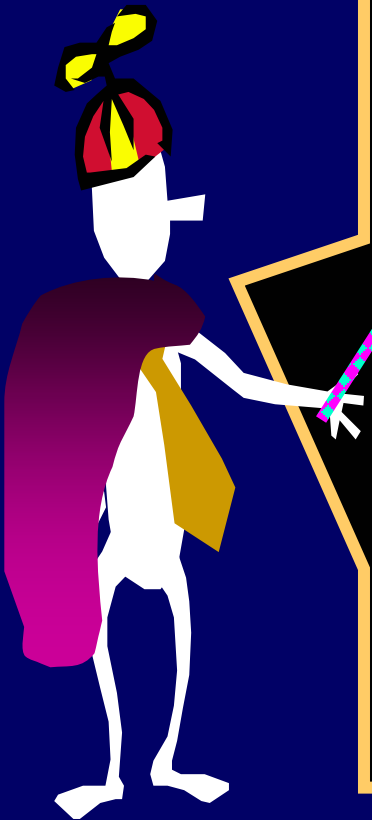


Let X = number of heads
when n independent coins of
bias p are flipped.

Break X into n simpler RVs,

$$X_i = \begin{cases} 0, & \text{if the } i^{\text{th}} \text{ coin is tails} \\ 1, & \text{if the } i^{\text{th}} \text{ coin is heads} \end{cases}$$

$$E[X] = E[\sum_i X_i] = ?$$

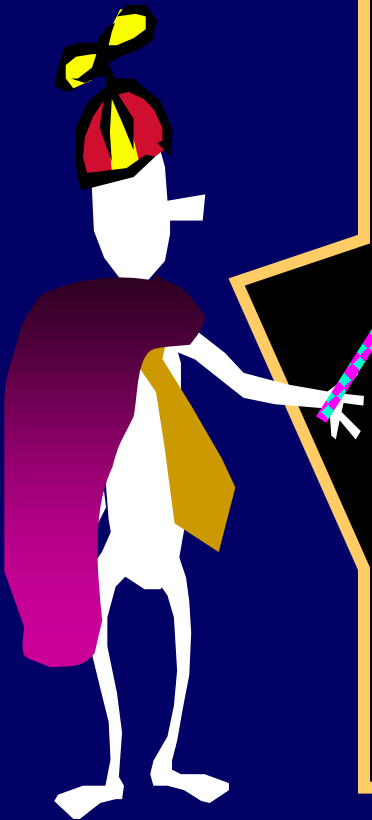


Let X = number of heads
when n independent coins of
bias p are flipped.

Break X into n simpler RVs,

$$X_i = \begin{cases} 0, & \text{if the } i^{\text{th}} \text{ coin is tails} \\ 1, & \text{if the } i^{\text{th}} \text{ coin is heads} \end{cases}$$

$$E[X] = E[\sum_i X_i] = np$$



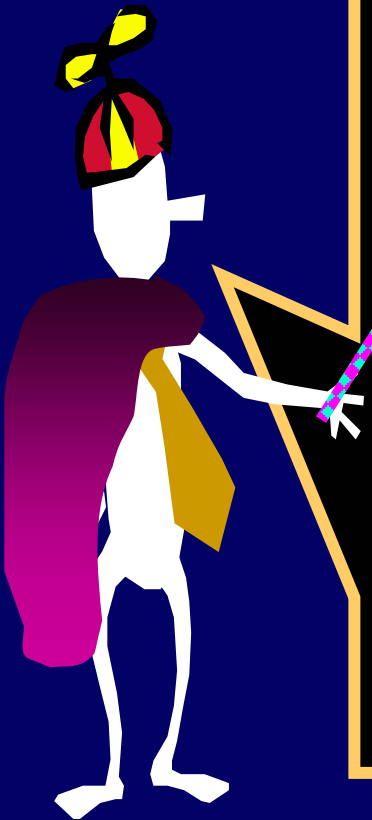
What about Products?

If $Z = XY$, then
 $E[Z] = E[X] \times E[Y]$?

No!

X =indicator for "1st flip is heads"
 Y =indicator for "1st flip is tails".

$E[XY]=0$.



But it is true if RVs are independent

Proof:

$$E[X] = \sum_a a \times \Pr(X=a)$$

$$E[Y] = \sum_b b \times \Pr(Y=b)$$

$$\begin{aligned} E[XY] &= \sum_c c \times \Pr(XY = c) \\ &= \sum_c \sum_{a,b:ab=c} c \times \Pr(X=a \cap Y=b) \\ &= \sum_{a,b} ab \times \Pr(X=a \cap Y=b) \\ &= \sum_{a,b} ab \times \Pr(X=a) \Pr(Y=b) \\ &= E[X] E[Y] \end{aligned}$$



E.g., 2 fair flips.

X = indicator for 1st coin being heads,

Y = indicator for 2nd coin being heads.

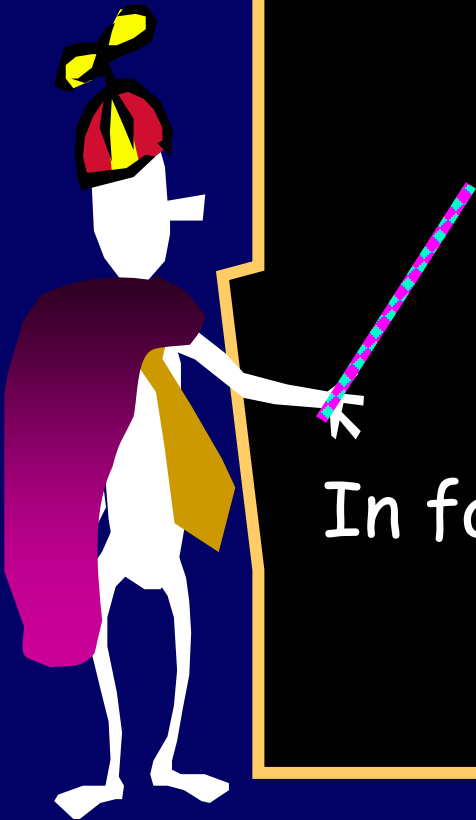
XY = indicator for "both are heads".

$$E[X] = \frac{1}{2}, E[Y] = \frac{1}{2}, E[XY] = \frac{1}{4}.$$

$$E[X * X] = E[X]^2?$$

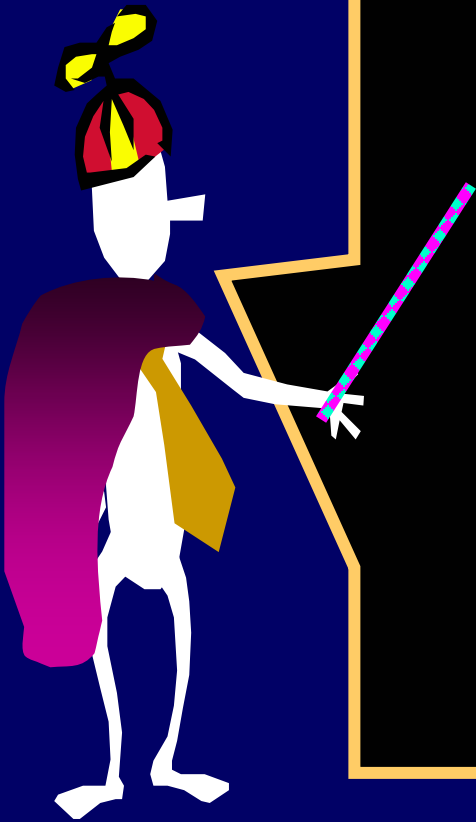
No: $E[X^2] = \frac{1}{2}, E[X]^2 = \frac{1}{4}.$

In fact, $E[X^2] - E[X]^2$ is called the *variance* of X .



Most of the time, though,
power will come from using
sums.

Mostly because
Linearity of Expectations
holds even if RVs are
not independent.



Another problem

On average, in class of size m , how many pairs of people will have the same birthday?



$$\sum_k k \cdot \Pr(\text{exactly } k \text{ collisions})$$

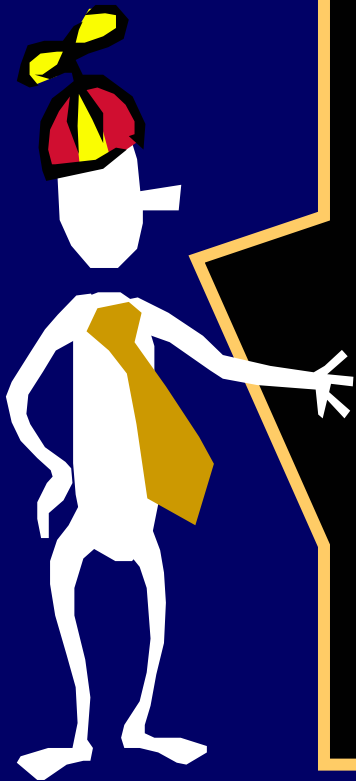
$$= \sum_k k \cdot (\dots \text{aargh!!!!} \dots)$$

Use linearity of expectation.

Suppose we have n people
each with a uniformly chosen
birthday from 1 to 366.

X = number of pairs of people
with the same birthday.

$$E[X] = ?$$



X = number of pairs of people
with the same birthday.

$$E[X] = ?$$

Use $m(m-1)/2$ indicator variables,
one for each pair of people.

$X_{jk} = 1$ if person j and person k
have the same birthday; else 0.

$$\begin{aligned} E[X_{jk}] &= (1/366) 1 + (1 - 1/366) 0 \\ &= 1/366 \end{aligned}$$



X = number of pairs of people
with the same birthday.

$$E[X] = E[\sum_{j \leq k \leq m} X_{jk}]$$

There are many dependencies
among the indicator variables.
E.g., X_{12} and X_{13} and X_{23} are
dependent.

But we don't care!

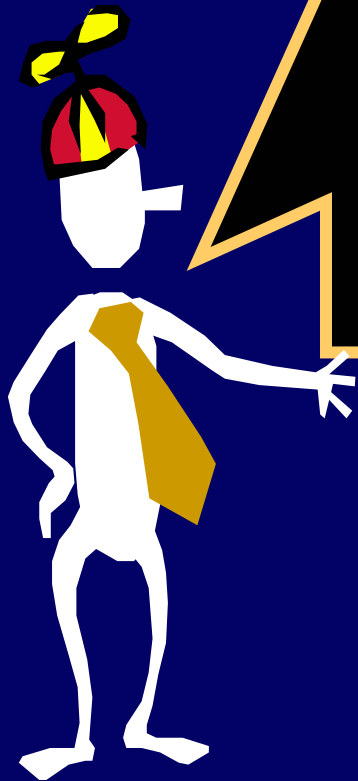


X = number of pairs of people
with the same birthday.

$$E[X] = E[\sum_{j \leq k \leq m} X_{jk}]$$

$$= \sum_{j \leq k \leq m} E[X_{jk}]$$

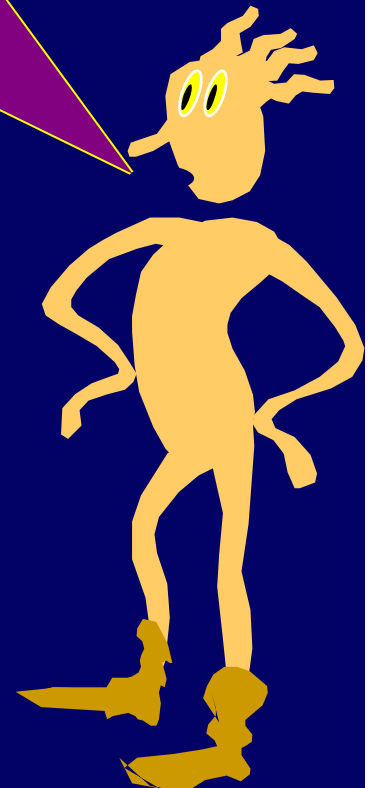
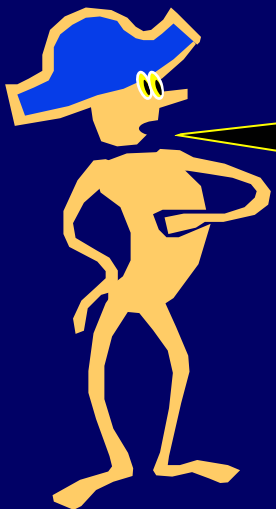
$$= m(m-1)/2 \times 1/366$$



Step right up...

You pick a number $n \in [1..6]$.
You roll 3 dice. If any match
 n , you win \$1. Else you pay me
\$1. Want to play?

Hmm...
let's see



Analysis

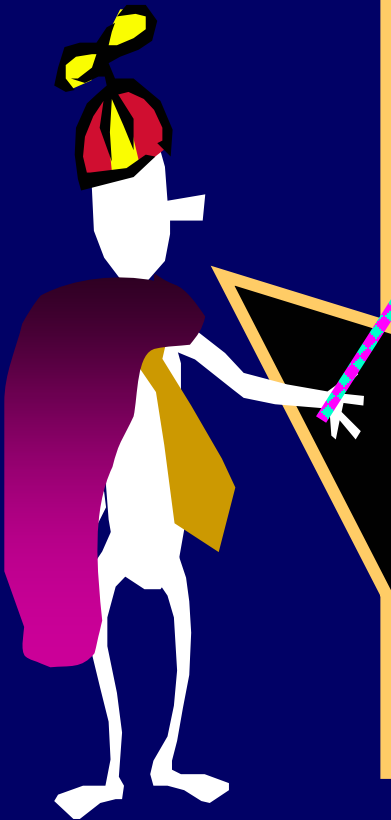
A_i = event that i -th die matches.

X_i = indicator RV for A_i .

Expected number of dice that match:

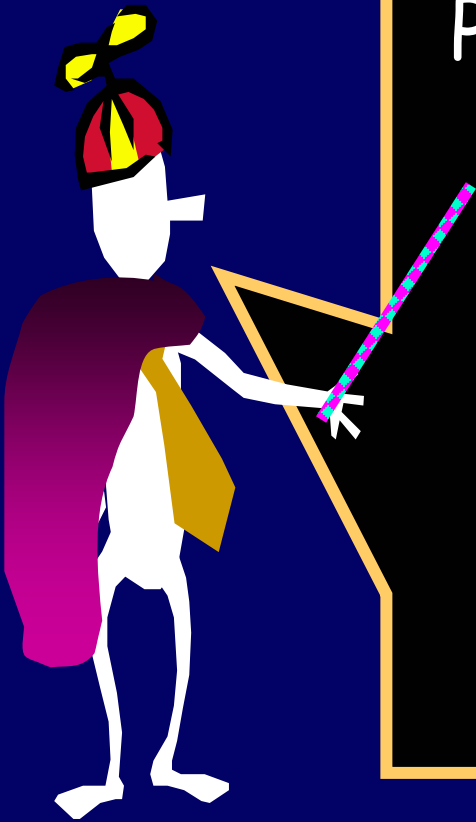
$$E[X_1 + X_2 + X_3] = 1/6 + 1/6 + 1/6 = \frac{1}{2}.$$

But this is not the same as
 $\Pr(\text{at least one die matches})$.



Analysis

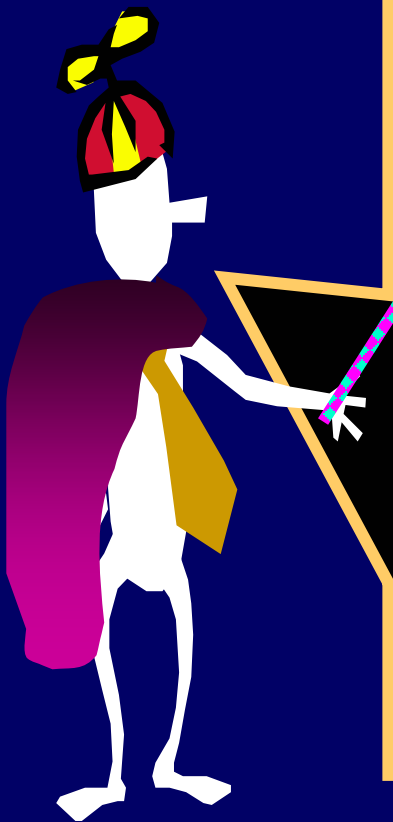
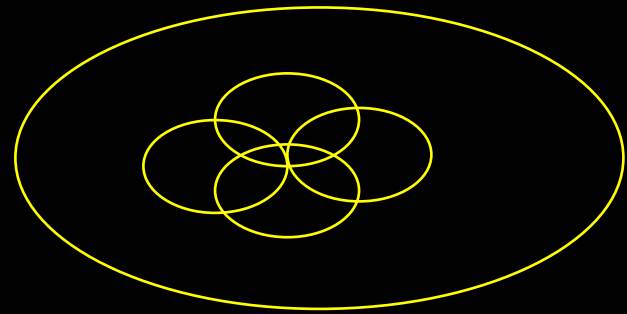
$$\begin{aligned}\Pr(\text{at least one die matches}) \\ &= 1 - \Pr(\text{none match}) \\ &= 1 - (5/6)^3 = 0.416.\end{aligned}$$



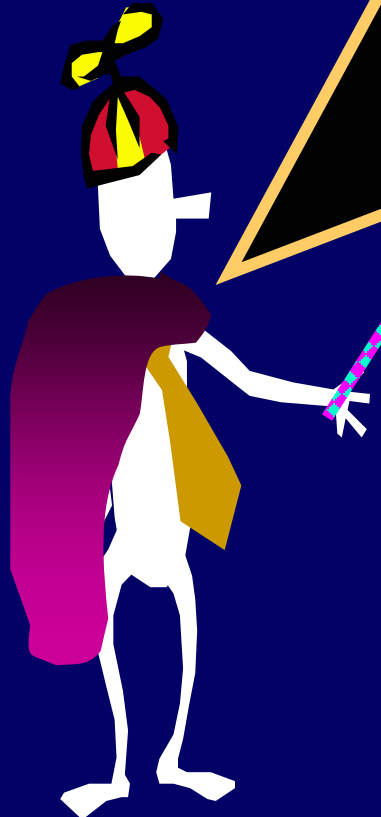
What's going on?

Say we have a collection of events $A_1, A_2, \dots$

How does $E[\# \text{ events that occur}]$ compare to $\Pr(\text{at least one occurs})$?



What's going on?


$$\begin{aligned} E[\# \text{ events that occur}] \\ &= \sum_k \Pr(k \text{ events occur}) \times k \\ &= \sum_{(k > 0)} \Pr(k \text{ events occur}) \times k \end{aligned}$$

$$\begin{aligned} \Pr(\text{at least one event occurs}) \\ &= \sum_{(k > 0)} \Pr(k \text{ events occur}) \end{aligned}$$

What's going on?

Moral #1: be careful you are modeling problem correctly.

Moral #2: watch out for carnival games.

