STAT 383C: Statistical Modeling I

Lecture 12 — September 3

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Disclaimer: These scribe notes have been slightly proofread and may have typos etc.

12.1 How collinearity effect ridge and lasso

Let us us assume that we have m copies of the same feature (similar to add m-1 copies).

12.1.1 Ridge

$$\min\left[(y - X\beta)^T (y - X\beta) + \lambda\beta^T \beta \right]$$
$$\hat{\beta} = \left(X^T X + \lambda \mathbf{I} \right)$$
$$a = \frac{X^T y}{X^T X + \lambda}$$
(12.1)

The same for m copies.

$$\hat{\beta}' = \left(x^T x \cdot J + \lambda \mathbf{I}\right)^{-1} x^T y \tag{12.2}$$

Using

$$X^{T}X = (x^{T}x)\mathbf{J}$$

$$X^{T}y = (x^{T}y)\mathbf{1}_{m}$$
(12.3)

Eq. ?? becomes

$$\hat{\beta}' = (x^T y)(c\mathbf{J} + \lambda \mathbf{I})_{1_m}^{-1}$$
(12.4)

Now, using the following expressions

$$\mathbf{J} = \mathbf{1}_m \cdot \mathbf{1}_m^T = (m, \mathbf{1}_m)$$

$$c\mathbf{J} = (cm, \mathbf{1}_m)$$

$$c\mathbf{J} + \lambda \mathbf{I} = (cm + \lambda, \mathbf{1}_m)$$

(12.5)

Eq. ?? becomes

$$\hat{\beta}' = (x^T y) \frac{1_m}{cm + \lambda} \tag{12.6}$$

Therefore

$$\hat{\beta}' = ax \frac{c+\lambda}{cm+\lambda} = ax \frac{1+\frac{\lambda}{X^T X}}{m+\frac{\lambda}{X^T X}} \approx \frac{a}{m}$$
(12.7)

The last equality holds for small λ .

Now let's solve the HW question.

$$x_{i1} = x_{i2} = x_{i3} = \dots = x_{im} = x \tag{12.8}$$

Then

$$\sum_{i} \left(y_{i} - x_{i}\beta_{1} - \sum_{j \neq i} X_{ij}\beta_{j} \right)^{2} + \lambda \sum_{j} \beta_{j}^{2}$$

$$= \sum_{i} \left(y_{i} - \sum_{j} X_{i1} \left(\beta_{11}' + \beta_{12}' + \dots + \beta_{1m}' \right) - \sum_{j} X_{ij}\beta_{1} \right)^{2} + \lambda \sum_{j=1}^{m} \left(\beta_{1j}' \right)^{2} + \sum_{j=2}^{p} \left(\beta_{j}' \right)^{2}$$

$$= \left\{ \sum_{i} \left(y_{i} - \sum_{j} X_{i1} \left(\beta_{11}' + \beta_{12}' + \dots + \beta_{1m}' \right) - \sum_{j} X_{ij}\beta_{1} \right)^{2} + \lambda \sum_{j=1}^{p} \left(\beta_{j}' \right)^{2} \right\} + \sum_{j=1}^{m} \left(\beta_{j1}' \right)^{2} - \beta_{1}'^{2}$$
(12.9)

When n is large, the first part of the optimization dominates, and so we can make the following approximate argument. Say we keep the first part pinned at its original optimal value,

$$\beta_{11}' + \beta_{12}' + \dots + \beta_{1m}' = a \tag{12.10}$$

Minimizing

$$\min\sum_{k=1}^{m} \left(\beta_{ij}^{\prime}\right)^2 \tag{12.11}$$

Consequently, all β'_{ij} should be the same, such that

$$\sum_{j} \beta'_{ij} = a \tag{12.12}$$

12.2 Logistic regression

12.2.1 Multiclass logistic regression

b for every class.

$$P\left(y=k|x_{j}\right) \propto e^{\beta_{m}^{T}x_{i}} = \frac{e^{\beta_{k}^{T}x}}{\sum_{i}e^{\beta_{i}^{T}x_{i}}}$$
(12.13)

Assuming

$$\beta_k' = 0_p \tag{12.14}$$

and applying the transformation

$$\beta'_j = \beta_j - \beta_k, j \neq k \tag{12.15}$$

Eq. ?? becomes

$$P(y = k|x_j) = \frac{1}{\sum_{c} e^{(\beta_c - \beta_k)^T x_i}}$$
(12.16)

Discriminative, because does not care about x distribution (generated). It cares only about how (Y|x) was generated.

12.2.2 Generative analogued (Linear discriminant analysis)

$$X_i | y_i = K \sim N(\mu_k, \Sigma_k), \tag{12.17}$$

where μ and Σ are known. how to decide which class involves any particular points (Point classification).

$$P\left(y_i = k\right) = \pi_k \tag{12.18}$$

$$P(Y = 1|x_{\xi}) \propto P(x|Y = 1) P(Y = 1)$$

= $\frac{1}{(2\pi)^{p/2} |\Sigma|} \exp\left(-\frac{(x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1)}{2}\right) \pi_1$ (12.19)

Decision rule. Classify as 1 if

$$P\left(y_i = 1|x\right) \ge P\left(Y = 0|x\right) \tag{12.20}$$

Similarly,

$$\log P\left(y_i = 1|x\right) \ge \log P\left(Y = 0|x\right) \tag{12.21}$$

$$-\frac{1}{2}\log|\Sigma_{1}| - \frac{(x-\mu_{1})^{T}\Sigma_{1}^{-1}(x-\mu_{1})}{2} + \log\pi_{1} \ge 1$$

$$1 + \log\pi_{1} = \frac{(x-\mu_{2})^{T}\Sigma_{2}^{-1}(x-\mu_{2})}{2} + \log\pi_{1} \ge 1$$
(12.22)

$$-\frac{1}{2}\log|\Sigma_2| - \frac{(x-\mu_2)^T \Sigma_2^{-1} (x-\mu_2)}{2} + \log \pi_2$$
(12.22)

If

$$\Sigma_1 = \Sigma_2, \tag{12.23}$$

then we get a linear decision boundary.

$$-\frac{(x-\mu_1)^T \Sigma^{-1} (x-\mu_1) + (x-\mu_2)^T \Sigma^{-1} (x-\mu_2)}{2} \ge \log \frac{\pi_2}{\pi_1}$$
(12.24)

$$2\mu_1^T \Sigma^{-1}(x-\mu_1) - 2\mu_2^T \Sigma^{-1}(x-\mu_2) + \frac{\mu_1^T \Sigma^{-1} \mu_1 - \mu_2^T \Sigma^{-1} \mu_2}{2} \ge \log \frac{\pi_2}{\pi_1}$$
(12.25)

$$-\left(x - \frac{\mu_1 + \mu_2}{2}\right)^T \Sigma^{-1}(\mu_2 - \mu_1) \ge \log \frac{\pi_2}{\pi_1}$$
(12.26)