# SDS 321: Introduction to Probability and Statistics <br> Lecture 13: Expectation and Variance and joint distributions 

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## Multiple random variables

So far we have been talking about single random variables and associated PMF's. However, often we are interested in multiple random variables.

- Consider two discrete random variables $X$, and $Y$ associated with the same experiment.
- The joint PMF of $X$ and $Y$ are defined as $p_{X, Y}(x, y)=P(X=x, Y=y)$ for all pairs of values $x, y X$ and $Y$ can take.
- This is none other than $P(\{X=x\} \cap\{Y=y\})$.
- Of course the order does not matter.


## Properties of the joint PMF

- Recall that if $A_{1}, A_{2}, \ldots, A_{K}$ is a partition of $\Omega$,

$$
P(B)=P\left(\bigcup_{k}\left(B \cap A_{k}\right)\right)=\sum_{k} P\left(B \cap A_{k}\right) .
$$

- $\{X=x\}$ is the disjoint union of $\{X=x\} \cap\{Y=y\}$ for all $y$ values $Y$ can take.
- $\{X=x\} \cap\{Y=y\}$ is none other than $\{X=x, Y=y\}$
- We can extend this to PMFs: $\sum_{y} P(X=x, Y=y)=$

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## Contingency tables

Alice says that there are more left handed women than left handed men. Bob gives her some numbers to count probabilities.

|  | Right Handed (L=0) | Left handed (L=1) |  |
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- $P(X=1)=1 / 2 \leftarrow$ Marginal probability!
- $P(L=1)=10 / 100 \leftarrow$ Marginal probability!


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- $P(X=1)=1 / 2 \leftarrow$ Marginal probability!
- $P(L=1)=10 / 100 \leftarrow$ Marginal probability!
- Remember! These really are estimated numbers, and hence approximations. I am estimating the fraction of left handed men in a population via my sample!


## Functions of multiple random variables

- $E(g(X, Y))=\sum_{x, y} g(x, y) P(X=x, Y=y)$.
- Let $g(X, Y)=a X+b Y$.
- $E(g(X, Y))=\sum_{x, y}(a x+b y) P(X=x, Y=y)=a E[X]+b E[Y]$.
- What if $g(X, Y)=a X^{2}+b Y^{2}+c$ ?
- $E[g(X, Y)]=a E\left[X^{2}\right]+b E\left[Y^{2}\right]+c$
- Common Mistake: $E[g(X, Y)] \neq g(E[X], E[Y])$ ! unless $g$ is linear in $X$ and $Y$ !


## Multiple random variables

How about three random variables?

- We will write $p_{X, Y, Z}(x, y, z)=P(X=x, Y=y, Z=z)$
- The rules are the same:
- $P(X=x, Y=y)=\sum_{z} P(X=x, Y=y, Z=z)$.
- $P(X=x)=\sum_{y, z} P(X=x, Y=y, Z=z)$.
- $P(Y=y)=\sum_{x, z} P(X=x, Y=y, Z=z)$.
- $P(Z=z)=\sum_{x, y} P(X=x, Y=y, Z=z)$.
- $\sum_{x, y, z} P(X=x, Y=y, Z=z)=1$.
- Generalizes easily to more than 3 random variables.


## Linearity of expectation

Perhaps one of the most useful and powerful results!

- $E[a X+b Y+c Z+d]=a E[X]+b E[Y]+c E[Z]+d$
- More generally,

$$
E\left[a_{1} X_{1}+a_{2} X_{2}+\cdots+a_{n} X_{n}\right]=a_{1} E\left[X_{1}\right]+a_{2} E\left[X_{2}\right]+\ldots a_{n} E\left[X_{n}\right]
$$

- This is extremely general! $x_{1}, \ldots, X_{n}$ do not have to be mutually independent for this to hold!
- This generalizes to $E\left[\sum_{i} a_{i} f\left(X_{i}\right)\right]=\sum_{i} a_{i} E\left[f\left(X_{i}\right)\right]$, as long as the expectations are well defined.


## Expectation of $Y \sim \operatorname{Binomial}(n, p)$

Remember that a Binomial $(n, p)$ random variable is nothing other than the sum of $n$ independent Bernoulli's!

- $Y=\sum_{i=1}^{n} X_{i}$, where $X_{i} \sim \operatorname{Bernoulli}(n, p)$.
- We know that $E\left[X_{i}\right]=p$.
- Using our newfound tool, we have:

$$
E[Y]=E\left[\sum_{i} X_{i}\right]=\sum_{i} E\left[X_{i}\right]=n p .
$$

- We do not need the mutual independence of the Bernoullis to get this result!


## Balls and bins

I am throwing $m$ distinguishable balls into $n$ distinguishable bins. What is the expected number of empty bins (call this $Y$ )?

- Let $X_{i}= \begin{cases}1 & \text { The } i^{\text {th }} \text { bin is empty } \\ 0 & \text { Otherwise }\end{cases}$
- We want $E[Y]$.
- $E[Y]=$
- $E\left[X_{i}\right]=$


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- $E\left[X_{i}\right]=P($ No ball falls in bin $i)=$


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- $E[Y]=E\left[\sum_{i} X_{i}\right]=\sum_{i} E\left[X_{i}\right]$
- $E\left[X_{i}\right]=P($ No ball falls in bin $i)=(1-1 / n)^{m}$
- $E[Y]=n(1-1 / n)^{m}$


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- $E[Y]=n(1-1 / n)^{m}$
- When $m=n$, for large $n, E[Y]=n(1-1 / n)^{n} \approx n / e$.


## Conditional PMF

So we have started thinking about how knowing about one random variable alters out belief about another random variable. This brings us to conditional PMFs!

- The conditional PMF of a random variable $X$, conditioned on a particular event $A$ with $P(A)>0$, is defined by:

$$
p_{X \mid A}(x)=P(X=x \mid A)=\frac{P(\{X=x\} \cap A)}{P(A)}
$$

- So we have

$$
\sum_{x} P(X=x \mid A)=\sum_{x} \frac{P(\{X=x\} \cap A)}{P(A)}=\frac{\sum_{x} P(\{X=x\} \cap A)}{P(A)}
$$

- But $A$ can be written as a disjoint union of the events $\{X=x\} \cap A$ for all numerical values $X$ takes.
- Total probability rule gives: $P(A)=\sum_{x} P(\{X=x\} \cap A)$, and so $\sum_{x} P(X=x \mid A)=1$.


## Conditioning one random variable on another

Let $X$ and $Y$ be two random variables associated with the same experiment. Now the knowledge of $Y=y$ for some particular value $y$ provides us with partial knowledge about what value $X$ may take.

- The conditional PMF of $X$ given $Y$ is given by

$$
p_{X \mid Y}(x, y)=P(X=x \mid\{Y=y\})
$$

- Using the same set of rules as before we can write:

$$
P(X=x \mid Y=y)=\frac{P(X=x, Y=y)}{P(Y=y)}=\frac{p_{X, Y}(x, y)}{p_{Y}(y)}
$$

- For any fixed $y$ such that $P(Y=y)>0$, we also have:
$\sum_{x} P(X=x \mid Y=y)=1$.
- So, a conditional PMF satisfies the properties of a PMF.


## Conditional PMF

Bob and Alice are interested in finding out the conditional probability of being left handed given a person is a man. Bob finds his data again.

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- $P(L=1 \mid X=0)$ is just the fraction of all men who are left handed


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- $P(L=1 \mid X=0)=7 / 50$.


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- Let us plug in the formula. $P(L=1, X=0)=$


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- Let us plug in the formula. $P(L=1, X=0)=7 / 100$.
- $P(X=0)=50 / 100$. So $\frac{P(L=1, X=0)}{P(X=0)}=7 / 50$.


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- Let us plug in the formula. $P(L=1, X=0)=7 / 100$.
- $P(X=0)=50 / 100$. So $\frac{P(L=1, X=0)}{P(X=0)}=7 / 50$.
- What is $P(L=0 \mid X=0)$ ? Its just the fraction of all men who are right handed! So 43/50.


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- What is $P(L=0 \mid X=0)$ ? Its just the fraction of all men who are right handed! So 43/50.
- $P(L=0 \mid X=0)+P(L=1 \mid X=0)=1$ !


## Conditional PMF

- Remember that a conditional PMF is a valid PMF.
- Since $P(X=x \mid Y=y)=\frac{P(X=x, Y=y)}{P(Y=y)}$, we also have the multiplication rule:
- $P(X=x, Y=y)=P(X=x \mid Y=y) P(Y=y)$
- But $P(X=x, Y=y)=P(Y=y, X=x)$, and so we also have: $P(X=x, Y=y)=P(Y=y \mid X=x) P(X=x)$.
- Same as multiplication rule from before!
- We can also draw trees to get conditional probabilities!


## Independence of random variables

- Lets first consider two events $\{X=x\}$ and $A$. We know that these two events are independent if $P(\{X=x\}, A)=P(\{X=x\}) P(A)$
- In other words if $P(A)>0$, then $P(X=x \mid A)=P(X=x)$, i.e. knowing the occurrence of $A$ does not change our belief about $\{X=x\}$.
- We will call the random variable $X$ and event $A$ to be independent if

$$
P(X=x, A)=P(X=x) P(A) \quad \text { For all } x
$$

- Two random variables are said to be independent if

$$
P(X=x, Y=y)=P(X=x) P(Y=y) \quad \text { For all } x \text { and } y
$$

- To put it a bit differently,

$$
P(X=x \mid Y=y)=P(X=x) \quad \text { For all } x \text { and } y \text { such that } P(Y=y)>0
$$

## A super important implication

We saw that $E[X+Y]=E[X]+E[Y]$ no matter whether $X$ and $Y$ are independent or not.

- If $X$ and $Y$ are independent, $E[X Y]=E[X] E[Y]$
- $E[X Y]=\sum_{x, y} x y P(X=x, Y=y)=\sum_{x, y} x y P(X=x) P(Y=y)$

$$
=\left(\sum_{x} x P(X=x)\right)\left(\sum_{y} y P(Y=y)\right)=E[X] E[Y]
$$

- In fact, $E[g(X) h(Y)]=E[g(X)] E[h(Y)]$


## Variance of sum of independent random variables

Let $X$ and $Y$ be two independent random variables. What is $\operatorname{var}(X+Y)$ ?

- Remember! $\operatorname{var}(X+Y)=E\left[(X+Y)^{2}\right]-(E[X+Y])^{2}$
- $E\left[(X+Y)^{2}\right]=E\left[X^{2}+Y^{2}+2 X Y\right]=E\left[X^{2}\right]+E\left[Y^{2}\right]+2 E[X Y]$

$$
=E\left[X^{2}\right]+E\left[Y^{2}\right]+2 E[X] E[Y]
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$E[X+Y]^{2}=(E[X]+E[Y])^{2}=E[X]^{2}+E[Y]^{2}+2 E[X] E[Y]$
$\operatorname{var}(X+Y)=E\left[(X+Y)^{2}\right]-(E[X+Y])^{2}$


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$$
=E\left[X^{2}\right]+E\left[Y^{2}\right]+2 E[X] E[Y]
$$

$$
E[X+Y]^{2}=(E[X]+E[Y])^{2}=E[X]^{2}+E[Y]^{2}+2 E[X] E[Y]
$$

$$
\operatorname{var}(X+Y)=E\left[(X+Y)^{2}\right]-(E[X+Y])^{2}
$$

$$
=\underbrace{E\left[X^{2}\right]-E[X]^{2}}_{\operatorname{var}(X)}+\underbrace{E\left[Y^{2}\right]-E[Y]^{2}}_{\operatorname{var}(Y)}=\operatorname{var}(X)+\operatorname{var}(Y)
$$

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Let $X$ and $Y$ be two independent random variables. What is $\operatorname{var}(X+Y)$ ?

- Remember! $\operatorname{var}(X+Y)=E\left[(X+Y)^{2}\right]-(E[X+Y])^{2}$
- $E\left[(X+Y)^{2}\right]=E\left[X^{2}+Y^{2}+2 X Y\right]=E\left[X^{2}\right]+E\left[Y^{2}\right]+2 E[X Y]$ $=E\left[X^{2}\right]+E\left[Y^{2}\right]+2 E[X] E[Y]$

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$$

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$$
=\underbrace{E\left[X^{2}\right]-E[X]^{2}}_{\operatorname{var}(X)}+\underbrace{E\left[Y^{2}\right]-E[Y]^{2}}_{\operatorname{var}(Y)}=\operatorname{var}(X)+\operatorname{var}(Y)
$$

- Variance of sum of independent random variables equals the sum of the variances!


## Independence of several random variables

- Three random variables $X, Y$ and $Z$ are said to be independent if

$$
P(X=x, Y=y, Z=z)=P(X=x) P(Y=y) P(Z=z) \quad \text { For all } x, y, z
$$

- If $X, Y, Z$ are independent, then so are $f(X), g(Y)$ and $h(Z)$.
- Also, any random variable $f(X, Y)$ and $g(Z)$ are independent.
- Are $f(X, Y)$ and $g(Y, Z)$ independent?


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- Not necessarily, both have $Y$ in common.
- For $n$ independent random variables, $X_{1}, X_{2}, \ldots, X_{n}$, we also have:

$$
\operatorname{var}\left(X_{1}+X_{2}+X_{3}+\cdots+X_{n}\right)=\operatorname{var}\left(X_{1}\right)+\operatorname{var}\left(X_{2}\right)+\cdots+\operatorname{var}\left(X_{n}\right)
$$

## Variance of a Binomial

Consider $n$ independent Bernoulli variables $X_{1}, X_{2}, \ldots, X_{n}$, each with probability $p$ of having value " 1 ". The sum $Y=\sum_{i} X_{i}$ is a $\operatorname{Binomial}(n, p)$ random variable.

- We saw last time that $E[Y]=\sum_{i} E\left[X_{i}\right]=n p$. What about the variance?
- Recall that $\operatorname{var}\left(X_{i}\right)=p(1-p)$ for $i \in\{1,2, \ldots, n\}$.
- $\operatorname{var}(Y)=\operatorname{var}\left(X_{1}+X_{2}+\cdots+X_{n}\right)=\sum_{i=1}^{n} \operatorname{var}\left(X_{i}\right)=n p(1-p)$.


## Conditional Independence

- Very similar to conditional independence of events!
- $X$ and $Y$ are conditionally independent, given a positive probability event $A$ if

$$
P(X=x, Y=y \mid A)=P(X=x \mid A) P(Y=y \mid A) \quad \text { For all } x \text { and } y
$$

- Same as saying $P(X=x \mid Y=y, A)=P(X=x \mid A)$, i.e.
- Once you know that $A$ has occurred, knowing $\{Y=y\}$ has occurred does not give you any more information!
- Like we learned before, conditional independence does not imply unconditional independence.


## Example-conditionally independent but not marginally

- I separately phone two students (Alice and Bob) and tell them the midterm grade.
- To each, I report the same grade, $G \in\{A+, A \ldots, C\}$.
- The signal is bad and, Alice and Bob each independently make an educated guess of what I said.
- Let the grades guessed by Alice and Bob be $X$ and $Y$.
- Are $X$ and $Y$ marginally independent?


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- What if I tell you that $G=A-$ ?
- Are $X$ and $Y$ conditionally independent given $\{G=A-\}$.
- YES! Because if we know the grade I actually said, the two variables are no longer dependent.


## Example-marginally independent but not conditionally

- I toss two dice independently and $X$ and $Y$ are the readings on them.
- Are $X$ and $Y$ independent?
- Now I tell you that $X+Y=12$. Are they still independent?

