## 10-301/601: Introduction to Machine Learning Lecture 17 - Naïve Bayes

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10/31/22

- Announcements:
- HW6 released 10/27, due 11/4 at 11:59 PM
- Only two late days allowed on HW6
- HW6 recitation on Wednesday 11/2; next lecture is on Friday, 11/4
- Exam 2 on 11/10
- All topics between Lecture 8 and Lecture 17 (today's lecture) are in-scope
- Exam 1 content may be referenced but will not be the primary focus of any question
- Fill out the mid-semester survey, due 11/2
- As of 9 AM this morning, only $228 / 405 \approx 56 \%$
- A Bernoulli random variable takes value 1 (or heads) with probability $\phi$ and value 0 (or tails) with probability $1-\phi$
- The pmf of the Bernoulli distribution is

$$
p(x \mid \phi)=\phi^{x}(1-\phi)^{1-x}
$$

- The partial derivative of the log-likelihood is

$$
\begin{aligned}
& \frac{N_{1}}{\hat{\phi}}-\frac{N_{0}}{1-\hat{\phi}}=0 \rightarrow \frac{N_{1}}{\hat{\phi}}=\frac{N_{0}}{1-\hat{\phi}} \\
\rightarrow & N_{1}(1-\hat{\phi})=N_{0} \hat{\phi} \rightarrow N_{1}=\hat{\phi}\left(N_{0}+N_{1}\right) \\
\rightarrow & \hat{\phi}=\frac{N_{1}}{N_{0}+N_{1}}
\end{aligned}
$$

- where $N_{1}$ is the number of 1 's in $\left\{x^{(1)}, \ldots, x^{(N)}\right\}$ and $N_{0}$ is the number of 0 's
- A Bernoulli random variable takes value 1 (or heads) with probability $\phi$ and value 0 (or tails) with probability $1-\phi$


## Poll Question 1:

After flipping your coin 5 times, what is the MLE of your coin?
A. $0 / 5$
B. $1 / 5$
C. $2 / 5$
D. $3 / 5$
E. $\pi / 5$ (TOXIC)
F. $4 / 5$
G. 5/5

- The pmf of the Bernoulli distribution is

$$
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\rightarrow & \hat{\phi}=\frac{N_{1}}{N_{0}+N_{1}}
\end{aligned}
$$

- where $N_{1}$ is the number of 1 's in $\left\{x^{(1)}, \ldots, x^{(N)}\right\}$ and $N_{0}$ is the number of 0 's


## Maximum a <br> Posteriori <br> (MAP)

Estimation

- Insight: sometimes we have prior information we want to incorporate into parameter estimation
- Idea: use Bayes rule to reason about the posterior distribution over the parameters
- MLE finds $\hat{\theta}=\underset{\theta}{\operatorname{argmax}} p(\mathcal{D} \mid \theta)$
- MAP finds $\hat{\theta}=\underset{\theta}{\operatorname{argmax}} p(\theta \mid \mathcal{D})$

$$
\begin{aligned}
& =\underset{\theta}{\operatorname{argmax}} p(\mathcal{D} \mid \theta) p(\theta) / p(\mathcal{D}) \\
& =\underset{\theta}{\operatorname{argmax}} p(\mathcal{D} \mid \theta) p(\theta)
\end{aligned}
$$

$$
=\underset{\theta}{\operatorname{argmax}} \underbrace{\log p(\mathcal{D} \mid \theta)+\log p(\theta)}_{\text {log-posterior }}
$$

# Maximum a Posteriori (MAP) <br> Estimation 

1. Specify the generative story, i.e., the data generating distribution, including a prior distribution

- How on earth do we pick a prior?

2. Maximize the log-posterior of $\mathcal{D}=\left\{x^{(1)}, \ldots, x^{(N)}\right\}$

$$
\ell_{M A P}(\theta)=\log p(\theta)+\sum_{i=1}^{N} \log p\left(x^{(i)} \mid \theta\right)
$$

3. Solve in closed form: take partial derivatives, set to 0 and solve

- A Bernoulli random variable takes value 1 (or heads) with probability $\phi$ and value 0 (or tails) with probability $1-\phi$
- The pmf of the Bernoulli distribution is

$$
p(x \mid \phi)=\phi^{x}(1-\phi)^{1-x}
$$

- Assume a Beta prior over the parameter $\phi$, which has pdf

$$
f(\phi \mid \alpha, \beta)=\frac{\phi^{\alpha-1}(1-\phi)^{\beta-1}}{\mathrm{~B}(\alpha, \beta)}
$$

where $\mathrm{B}(\alpha, \beta)=\int_{0}^{1} \phi^{\alpha-1}(1-\phi)^{\beta-1} d \phi$ is a normalizing constant to ensure the distribution integrates to 1

## Beta <br> Distribution

Beta Distribution w/ $\alpha=1$ and $\beta=1$


Beta Distribution $w / \alpha=2$ and $\beta=2$


Beta Distribution w/ $\alpha=2$ and $\beta=5$


Beta Distribution w/ $\alpha=10$ and $\beta=10$



## Why use this strange looking Beta prior?

The Beta distribution is the conjugate prior for the Bernoulli distribution!

- A Bernoulli random variable takes value 1 (or heads) with probability $\phi$ and value 0 (or tails) with probability $1-\phi$
- The pmf of the Bernoulli distribution is

$$
p(x \mid \phi)=\phi^{x}(1-\phi)^{1-x}
$$

- Assume a Beta prior over the parameter $\phi$, which has pdf

$$
f(\phi \mid \alpha, \beta)=\frac{\phi^{\alpha-1}(1-\phi)^{\beta-1}}{\mathrm{~B}(\alpha, \beta)}
$$

where $\mathrm{B}(\alpha, \beta)=\int_{0}^{1} \phi^{\alpha-1}(1-\phi)^{\beta-1} d \phi$ is a normalizing constant to ensure the distribution integrates to 1

- Given $N$ iid samples $\left\{x^{(1)}, \ldots, x^{(N)}\right\}$, the log-posterior is

$$
\begin{aligned}
\ell(\phi)= & \log f(\phi \mid \alpha, \beta)+\sum_{n=1}^{N} \log p\left(x^{(n)} \mid \phi\right) \\
= & \log \frac{\phi^{\alpha-1}(1-\phi)^{\beta-1}}{\mathrm{~B}(\alpha, \beta)}+\sum_{n=1}^{N} \log \phi^{x^{(n)}}(1-\phi)^{1-x^{(n)}} \\
= & (\alpha-1) \log \phi+(\beta-1) \log (1-\phi)-\log \mathrm{B}(\alpha, \beta) \\
& +\sum_{n=1}^{N} x^{(n)} \log \phi+\left(1-x^{(n)}\right) \log (1-\phi) \\
= & \left(\alpha-1+N_{1}\right) \log \phi+\left(\beta-1+N_{0}\right) \log (1-\phi) \\
& -\log \mathrm{B}(\alpha, \beta)
\end{aligned}
$$

Coin
Flipping
MAP

- Given $N$ iid samples $\left\{x^{(1)}, \ldots, x^{(N)}\right\}$, the partial derivative of the log-posterior is

$$
\frac{\partial \ell}{\partial \phi}=\frac{\left(\alpha-1+N_{1}\right)}{\phi}-\frac{\left(\beta-1+N_{0}\right)}{1-\phi}
$$

$$
\rightarrow \hat{\phi}_{M A P}=\frac{\left(\alpha-1+N_{1}\right)}{\left(\beta-1+N_{0}\right)+\left(\alpha-1+N_{1}\right)}
$$

- $\alpha-1$ is a "pseudocount" of the number of 1 's (or heads) you've "observed"
- $\beta-1$ is a "pseudocount" of the number of 0 's (or tails) you've "observed"
- Suppose $\mathcal{D}$ consists of ten 1 's or heads $\left(N_{1}=10\right)$ and two 0 's or tails ( $N_{0}=2$ ):

Coin
Flipping
MAP:
Example

$$
\phi_{M L E}=\frac{10}{10+2}=\frac{10}{12}
$$

- Using a Beta prior with $\alpha=2$ and $\beta=5$, then

$$
\phi_{M A P}=\frac{(2-1+10)}{(2-1+10)+(5-1+2)}=\frac{11}{17}<\frac{10}{12}
$$

- Suppose $\mathcal{D}$ consists of ten 1 's or heads $\left(N_{1}=10\right)$ and two 0 's or tails ( $N_{0}=2$ ):
Coin
Flipping
MAP:
Example

$$
\phi_{M L E}=\frac{10}{10+2}=\frac{10}{12}
$$

- Using a Beta prior with $\alpha=101$ and $\beta=101$, then

$$
\phi_{M A P}=\frac{(101-1+10)}{(101-1+10)+(101-1+2)}=\frac{110}{212} \approx \frac{1}{2}
$$

- Suppose $\mathcal{D}$ consists of ten 1 's or heads $\left(N_{1}=10\right)$ and two 0 's or tails ( $N_{0}=2$ ):

Coin
Flipping
MAP:
Example

$$
\phi_{M L E}=\frac{10}{10+2}=\frac{10}{12}
$$

- Using a Beta prior with $\alpha=1$ and $\beta=1$, then

$$
\phi_{M A P}=\frac{(1-1+10)}{(1-1+10)+(1-1+2)}=\frac{10}{12}=\phi_{M L E}
$$

You should be able to...

- Recall probability basics, including but not limited to: discrete and continuous random variables, probability mass functions, probability density functions, events vs. random variables, expectation and variance, joint


## MLE/MAP Learning Objectives

 probability distributions, marginal probabilities, conditional probabilities, independence, conditional independence- State the principle of maximum likelihood estimation and explain what it tries to accomplish
- State the principle of maximum a posteriori estimation and explain why we use it
- Derive the MLE or MAP parameters of a simple model in closed form


## Text Data

- https://www.nytimes.com/20 22/10/13/movies/halloween-ends-review.html
- https://www.nytimes.com/20 22/10/20/business/the-spirit-of-halloween.html
- https://www.theonion.com/b iden-issues-urgent-warning-for-americans-to-decide-wha1849597566
'Halloween Ends' Review: It Probably Doesn't
David Gordon Green wraps up his reboot trilogy for a horror
franchise that never stays dead for long.



Biden Issues Urgent Warning For Americans To Decide What To Be For Halloween Now
| 9/30/22 5:30AM | Alerts
The Spirit c


## Text Data



Bag-of-Words Model

| $x_{1}$ <br> ("hat") | $\begin{gathered} x_{2} \\ \text { ("cat") } \end{gathered}$ | $\begin{gathered} x_{3} \\ \left(" \operatorname{cog}^{\prime}\right) \end{gathered}$ | $\begin{gathered} x_{4} \\ (\text { "fish") } \end{gathered}$ | $\begin{gathered} x_{5} \\ \text { ("mom") } \end{gathered}$ | $\begin{gathered} x_{6} \\ (\text { "dad" }) \end{gathered}$ | $\begin{gathered} y \\ \text { (Dr. Seuss) } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 0 | 0 | 0 | 0 | 1 |

Bag-of-Words Model


## Bag-of-Words Model

| $x_{1}$ <br> ("hat") | $x_{2}$ <br> ("cat") | $x_{3}$ <br> ("dog") | $x_{4}$ <br> ("fish") | $x_{5}$ <br> ("mom") | $x_{6}$ <br> $(" d a d ")$ | $y$ <br> (Dr. Seuss) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 0 | 0 | 0 | 0 | 1 |
| 0 | 0 | 1 | 0 | 0 | 0 | 0 |

Go, Dog. Go!
(by P. D. Eastman)
by P.D.Dastman

| $x_{1}$ <br> ("hat") | $x_{2}$ <br> ("cat") | $x_{3}$ <br> ("dog") | $x_{4}$ <br> ("fish") | $x_{5}$ <br> ("mom") | $x_{6}$ <br> ("dad") | $y$ <br> (Dr. Seuss) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 0 | 0 | 0 | 0 | 1 |
| 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 0 | 0 | 0 | 1 | 0 | 0 | 1 |

Bag-of-Words Model


## Bag-of-Words Model

| $x_{1}$ <br> ("hat") | $x_{2}$ <br> ("cat") | $x_{3}$ <br> ("dog") | $x_{4}$ <br> ("fish") | $x_{5}$ <br> $($ ("mom") | $x_{6}$ <br> $(" d a d ")$ | $y$ <br> (Dr. Seuss) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 0 | 0 | 0 | 0 | 1 |
| 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 0 | 0 | 0 | 1 | 0 | 0 | 1 |
| 0 | 0 | 0 | 0 | 1 | 0 | 0 |

## Are You <br> My <br> 

Mother?
Are You My Mother? (by P. D. Eastman)


- Define a decision rule
- Given a test data point $x^{\prime}$, predict its label $\hat{y}$ using the posterior distribution $P\left(Y=y \mid X=x^{\prime}\right)$


## Recall: <br> Building a <br> Probabilistic Classifier

- Common choice: $\hat{y}=\operatorname{argmax} P\left(Y=y \mid X=x^{\prime}\right)$ $y$
- Model the posterior distribution
- Option 1 - Model $P(Y \mid X)$ directly as some function of $X$ (recall: logistic regression)
- Option 2 - Use Bayes' rule (today!):

$$
P(Y \mid X)=\frac{P(X \mid Y) P(Y)}{P(X)} \propto P(X \mid Y) P(Y)
$$

- Define a decision rule
- Given a test data point $x^{\prime}$, predict its label $\hat{y}$ using the posterior distribution $P\left(Y=y \mid X=x^{\prime}\right)$


## How hard is modelling $P(X \mid Y)$ ?

- Common choice: $\hat{y}=\operatorname{argmax} P\left(Y=y \mid X=x^{\prime}\right)$ $y$
- Model the posterior distribution
- Option 1 - Model $P(Y \mid X)$ directly as some function of $X$ (recall: logistic regression)
- Option 2 - Use Bayes' rule (today!):

$$
P(Y \mid X)=\frac{P(X \mid Y) P(Y)}{P(X)} \propto P(X \mid Y) P(Y)
$$

How hard is modelling $P(X \mid Y)$ ?

| $x_{1}$ <br> ("hat") | $x_{2}$ <br> ("cat") | $x_{3}$ <br> ("dog") | $x_{4}$ <br> ("fish") | $x_{5}$ <br> ("mom") | $x_{6}$ <br> ("dad") | $P(X \mid Y=1)$ | $P(X \mid Y=0)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 0 | 0 | 0 | 0 | 0 | $\theta_{1}$ | $\theta_{64}$ |
| 1 | 0 | 0 | 0 | 0 | 0 | $\theta_{2}$ | $\theta_{65}$ |
| 1 | 1 | 0 | 0 | 0 | 0 | $\theta_{3}$ | $\theta_{66}$ |
| 1 | 0 | 1 | 0 | 0 | 0 | $\theta_{4}$ | $\theta_{67}$ |
|  |  |  |  |  |  |  |  |
| $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ |
| 1 | 1 | 1 | 1 | 1 | 1 | $1-\sum_{i=1}^{63} \theta_{i}$ | $1-\sum_{i=64}^{126} \theta_{i}$ |

- Assume features are conditionally independent given the label:

$$
P(X \mid Y)=\prod_{d=1}^{D} P\left(X_{d} \mid Y\right)
$$

- Pros:
- Significantly reduces computational complexity
- Also reduces model complexity, combats overfitting
- Cons:
- Is a strong, often illogical assumption
- We'll see a relaxed version of this later in the semester when we discuss Bayesian networks
- Define a model and model parameters
- Make the naïve Bayes assumption
- Assume independent, identically distributed (iid) data
- Parameters: $\pi=P(Y=1), \theta_{d, y}=P\left(X_{d}=1 \mid Y=y\right)$


## Recipe for Naïve Bayes

- Write down an objective function
- Maximize the log-likelihood
- Optimize the objective w.r.t. the model parameters
- Solve in closed form: take partial derivatives, set to 0 and solve

$$
\begin{aligned}
\ell_{\mathcal{D}}(\pi, \boldsymbol{\theta}) & =\log P\left(\mathcal{D}=\left\{\boldsymbol{x}^{(1)}, y^{(1)}, \ldots, \boldsymbol{x}^{(N)}, y^{(N)}\right\} \mid \pi, \boldsymbol{\theta}\right) \\
& =\log \prod_{n=1}^{N} P\left(\boldsymbol{x}^{(n)}, y^{(n)} \mid \pi, \boldsymbol{\theta}\right)=\log \prod_{n=1}^{N} P\left(\boldsymbol{x}^{(n)} \mid y^{(n)}, \boldsymbol{\theta}\right) P\left(y^{(n)} \mid \pi\right) \\
& =\log \prod_{n=1}^{N}\left(\prod_{d=1}^{D} P\left(x_{d}^{(n)} \mid y^{(n)}, \theta_{d, 1}, \theta_{d, 0}\right)\right) P\left(y^{(n)} \mid \pi\right) \\
& =\sum_{n=1}^{N}\left(\sum_{d=1}^{D} \log P\left(x_{d}^{(n)} \mid y^{(n)}, \theta_{d, 1}, \theta_{d, 0}\right)\right)+\log P\left(y^{(n)} \mid \pi\right) \\
& =\sum_{n: y^{(n)}=1}\left(\sum_{d=1}^{D} \log P\left(x_{d}^{(n)} \mid \theta_{d, 1}\right)\right) \\
& +\sum_{n: y^{(n)}=0}\left(\sum_{d=1}^{D} \log P\left(x_{d}^{(n)} \mid \theta_{d, 0}\right)\right)+\sum_{n=1}^{N} \log P\left(y^{(n)} \mid \pi\right)
\end{aligned}
$$

- Binary label
- $Y \sim \operatorname{Bernoulli}(\pi)$
- $\hat{\pi}=N_{Y=1} / N_{N}$
- $N=\#$ of data points
- $N_{Y=1}=\#$ of data points with label 1
- Binary features
- $X_{d} \mid Y=y \sim \operatorname{Bernoulli}\left(\theta_{d, y}\right)$
- $\hat{\theta}_{d, y}={ }^{N_{Y=y, X_{d}=1} / N_{Y=y}}$
- $N_{Y=y}=\#$ of data points with label $y$
- $N_{Y=y, X_{d}=1}=\#$ of data points with label $y$ and feature $X_{d}=1$


## Setting the Parameters via MLE

Poll Question 2:
Given this
dataset, what is the MLE of $\pi$ ?

Poll Question 3:
Given this dataset, what is the MLE of $\theta_{3,1}$ ?

| $x_{1}$ | $x_{2}$ | $x_{3}$ | $y$ |
| :---: | :---: | :---: | :---: |
| 1 | 0 | 1 | 0 |
| 0 | 1 | 0 | 1 |
| 0 | 1 | 1 | 1 |
| 0 | 0 | 1 | 0 |
| 1 | 0 | 1 | 0 |
| 1 | 0 | 1 | 1 |

A. 0/6
B. $1 / 6$
C. $2 / 6$
D. $3 / 6$
E. $4 / 6$
F. 5/6
G. 6/6
H. 7/6 (TOXIC)

- Binary label
- $Y \sim \operatorname{Bernoulli}(\pi)$
- $\hat{\pi}=N_{Y=1} / N_{N}$
- $N=\#$ of data points
- $N_{Y=1}=\#$ of data points with label 1
- Binary features
- $X_{d} \mid Y=y \sim \operatorname{Bernoulli}\left(\theta_{d, y}\right)$
- $\hat{\theta}_{d, y}={ }^{N_{Y=y, X_{d}=1} / N_{Y=y}}$
- $N_{Y=y}=\#$ of data points with label $y$
- $N_{Y=y, X_{d}=1}=\#$ of data points with label $y$ and feature $X_{d}=1$
- Binary label
- $Y \sim \operatorname{Bernoulli}(\pi)$
- $\hat{\pi}=N_{Y=1} / N$
- $N=\#$ of data points
- $N_{Y=1}=\#$ of data points with label 1
- Discrete features ( $X_{d}$ can take on one of $K$ possible values)
- $X_{d} \mid Y=y \sim$ Categorical $\left(\theta_{d, 1, y}, \ldots, \theta_{d, K-1, y}\right)$
- $\hat{\theta}_{d, k, y}=N_{Y=y, X_{d}=k} / N_{Y=y}$
- $N_{Y=y}=\#$ of data points with label $y$
- $N_{Y=y, X_{d}=k}=\#$ of data points with label $y$ and feature $X_{d}=k$
- Binary label
- $Y \sim \operatorname{Bernoulli}(\pi)$
- $\hat{\pi}=N_{Y=1} / N_{N}$
- $N=\#$ of data points
- $N_{Y=1}=\#$ of data points with label 1
- Real-valued features

$$
\begin{aligned}
& \cdot X_{d} \mid Y=y \sim \operatorname{Gaussian}\left(\mu_{d, y}, \sigma_{d, y}^{2}\right) \\
& \text { - } \hat{\mu}_{d, y}=\frac{1}{N_{Y=y}} \sum_{n: y^{(n)}=y} x_{d}^{(n)} \\
& \text { - } \hat{\sigma}_{d, y}^{2}=\frac{1}{N_{Y=y}} \sum_{n: y^{(n)}=y}\left(x_{d}^{(n)}-\hat{\mu}_{d, y}\right)^{2}
\end{aligned}
$$

- $N_{Y=y}=\#$ of data points with label $y$
- Discrete label ( $Y$ can take on one of $M$ possible values)
- $Y \sim$ Categorical $\left(\pi_{1}, \ldots, \pi_{M}\right)$
- $\hat{\pi}_{m}=N_{Y=m} /{ }_{N}$
- $N=\#$ of data points
- $N_{Y=m}=\#$ of data points with label $m$
- Real-valued features
- $X_{d} \mid Y=y \sim \operatorname{Gaussian}\left(\mu_{d, y}, \sigma_{d, y}^{2}\right)$
- $\hat{\mu}_{d, y}=\frac{1}{N_{Y=y}} \sum_{n: y^{(n)}=y} x_{d}^{(n)}$
- $\hat{\sigma}_{d, y}^{2}=\frac{1}{N_{Y=y}} \sum_{n: y^{(n)}=y}\left(x_{d}^{(n)}-\hat{\mu}_{d, y}\right)^{2}$
- $N_{Y=y}=\#$ of data points with label $y$


## Visualizing Gaussian Naive Bayes

- Fisher (1936) used 150 measurements of flowers from 3 different species: Iris setosa (0), Iris virginica (1), Iris versicolor (2) collected by Anderson (1936)

| Species | Sepal Length | Sepal Width | Deleted two of the <br> four features, so that <br> input space is 2D |
| :--- | :--- | :--- | :--- |
| 0 | 4.3 | 3.0 |  |
| 0 | 4.9 | 3.6 |  |
| 0 | 5.3 | 3.7 |  |
| 1 | 4.9 | 2.4 |  |
| 1 | 5.7 | 2.8 |  |
| 1 | 6.3 | 3.3 |  |
| 1 | 6.7 | 3.0 |  |

# Visualizing <br> Gaussian <br> Nailve <br> Bayes <br> (2 classes) 



# Visualizing <br> Gaussian <br> Nailve <br> Bayes <br> (2 classes) 



Classification with Naive Bayes

## Visualizing Gaussian Nailve <br> Bayes <br> (2 classes, <br> equal <br> variances)



Classification with Naive Bayes

Visualizing Gaussian Nailve<br>Bayes<br>(2 classes, learned variances)



Classification with Naive Bayes

Visualizing Gaussian Nailve<br>Bayes<br>(3 classes, equal variances)



Classification with Naive Bayes

Visualizing Gaussian Nailve<br>Bayes<br>(3 classes, learned variances)



Visualizing<br>Gaussian<br>Naïve<br>Bayes<br>(2 classes,<br>learned<br>variances)



# Visualizing Gaussian Nailve <br> Bayes <br> (2 classes, <br> learned <br> variances) 



## Bernoulli <br> Naïve <br> Bayes: <br> Making <br> Predictions <br> - Given a test data point $\boldsymbol{x}^{\prime}=\left[x_{1}^{\prime}, \ldots, x_{D}^{\prime}\right]^{T}$ <br> $$
\begin{aligned} P\left(Y=1 \mid x^{\prime}\right) & \propto P(Y=1) P\left(x^{\prime} \mid Y=1\right) \\ & =\hat{\pi} \prod_{d=1}^{D} \hat{\theta}_{d, 1}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 1}\right)^{1-x_{d}^{\prime}} \end{aligned}
$$ <br> $$
P\left(Y=0 \mid x^{\prime}\right) \propto(1-\hat{\pi}) \prod_{d=1}^{D} \hat{\theta}_{d, 0}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 0}\right)^{1-x_{d}^{\prime}}
$$ <br> $$
\hat{y}=\left\{\begin{array}{l} 1 \text { if } \hat{\pi} \prod_{d=1}^{D} \hat{\theta}_{d, 1}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 1}\right)^{1-x_{d}^{\prime}}> \\ \quad(1-\hat{\pi}) \prod_{d=1}^{D=} \hat{\theta}_{d, 0}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 0}\right)^{1-x_{d}^{\prime}} \\ 0 \text { otherwise } \end{array}\right.
$$

What if some Word-Label pair never appears in our training data?

- Given a test data point $\boldsymbol{x}^{\prime}=\left[x_{1}^{\prime}, \ldots, x_{D}^{\prime}\right]^{T}$

$$
\begin{aligned}
P\left(Y=1 \mid x^{\prime}\right) & \propto P(Y=1) P\left(x^{\prime} \mid Y=1\right) \\
& =\hat{\pi} \prod_{d=1}^{D} \hat{\theta}_{d, 1}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 1}\right)^{1-x_{d}^{\prime}}
\end{aligned}
$$

$$
P\left(Y=0 \mid x^{\prime}\right) \propto(1-\hat{\pi}) \prod_{d=1}^{D} \hat{\theta}_{d, 0}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 0}\right)^{1-x_{d}^{\prime}}
$$

$$
\hat{y}=\left\{\begin{array}{l}
1 \text { if } \hat{\pi} \prod_{d=1}^{D} \hat{\theta}_{d, 1}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 1}\right)^{1-x_{d}^{\prime}}> \\
\quad(1-\hat{\pi}) \prod_{d=1}^{a=1} \hat{\theta}_{d, 0}^{x_{d}^{\prime}}\left(1-\hat{\theta}_{d, 0}\right)^{1-x_{d}^{\prime}} \\
0 \text { otherwise }
\end{array}\right.
$$

# What if some Word-Label pair never appears in our training data? 

| $x_{1}$ <br> ("hat") | $x_{2}$ <br> ("cat") | $x_{3}$ <br> ("dog") | $x_{4}$ <br> ("fish") | $x_{5}$ <br> ("mom") | $x_{6}$ <br> ("dad") | $y$ <br> (Dr. Seuss) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 0 | 0 | 0 | 0 | 1 |
| 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 0 | 0 | 0 | 1 | 0 | 0 | 1 |
| 0 | 0 | 0 | 0 | 1 | 0 | 0 |

The Cat in the Hat gets a Dog (by ???)

- If some $\hat{\theta}_{d, y}=0$ and that word appears in our test data $\boldsymbol{x}^{\prime}$, then $P\left(Y=y \mid \boldsymbol{x}^{\prime}\right)=0$ even if all the other features in $x^{\prime}$ point to the label being $y$ !
- The model has been overfit to the training data
- We can address this with a prior over the parameters!
- Binary label
- $Y \sim \operatorname{Bernoulli}(\pi)$
- $\hat{\pi}={ }^{N_{Y=1} / N}$
- $N=\#$ of data points
- $N_{Y=1}=\#$ of data points with label 1


## Setting the Parameters via MAP

- Binary features
- $X_{d} \mid Y=y \sim \operatorname{Bernoulli}\left(\theta_{d, y}\right)$ and $\theta_{d, y} \sim \operatorname{Beta}(\alpha, \beta)$
- $\hat{\theta}_{d, y}={ }^{N_{Y=y, X_{d}=1}+(\alpha-1)} /_{N_{Y=y}+(\alpha-1)+(\beta-1)}$
- $N_{Y=y}=\#$ of data points with label $y$
- $N_{Y=y, X_{d}=1}=\#$ of data points with label $y$ and feature $X_{d}=1$
- Common choice: $\alpha=2, \beta=2$
- Naïve Bayes is a generative model


## Logistic Regression vs. Nailve Bayes

- By modelling $P(X \mid Y)$ and $P(Y)$, we can generate new data points:

1. Sample a label $y \sim P(Y)$
2. Sample features $x_{d} \sim P\left(X_{d} \mid Y=y\right)$

- Logistic regression is a discriminative model
- By modelling $P(Y \mid X)$, we can only discriminate (or distinguish) between classes.
- Naïve Bayes and logistic regression form a generativediscriminative model pair
- Recall that under certain conditions, the Gaussian Naïve Bayes (GNB) decision boundary is linear
- If the Naïve Bayes assumption holds, then in the limit of infinite training data, GNB and logistic regression learn the same (linear) decision boundary!
- In general, Naïve Bayes performs well when data is scarce but logistic regression has lower asymptotic error.


## Logistic Regression vs. Naïve Bayes <br> (Ng and <br> Jordan, 2001)




optdigits ( 0 's and 1's, continuous)




- Dotted line: logistic regression
- Solid line: Naïve Bayes


## You should be able to...

- Write the generative story for Naive Bayes
- Create a new naïve Bayes classifier using your favorite probability distribution as the event model
- Apply the principle of maximum likelihood estimation (MLE) to learn the parameters of Bernoulli naïve Bayes


## Naïve Bayes Learning Objectives

- Motivate the need for MAP estimation through the deficiencies of MLE
- Apply the principle of maximum a posteriori (MAP) estimation to learn the parameters of Bernoulli naïve Bayes
- Select a suitable prior for a model parameter
- Describe the tradeoffs of generative vs. discriminative models
- Implement Bernoulli naïve Bayes
- Describe how the variance affects whether a Gaussian naïve Bayes model will have a linear or nonlinear decision boundary

