Scheduling Solutions for Coping with Transient Overload\textsuperscript{1}

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Abstract

For most computer systems, even short periods of overload degrade performance significantly. The number of jobs in the system quickly grows, often exceeding the capacity of the system within just seconds, and response times explode.

In this paper we investigate system behavior under transient overload. We find that the poor behavior of systems under transient overload can at least partly be attributed to the scheduling policy traditionally used in systems. The traditionally-used scheduling policy is Processor-Sharing, or time-sharing, (PS). We derive analytical approximations as well as simulation results for the performance of a single PS queue under transient overload. Simulation and analysis agree. The number of jobs in the system and the system response times grows quite rapidly during overload, and even when the overload period ends, recuperation is very slow.

We propose a new solution for coping with transient overload: SRPT scheduling of jobs (Shortest Remaining Processing Time). We derive analytical approximations for the performance of a single SRPT queue under transient overload, and validate those approximations with simulation.

We evaluate our PS and SRPT approximations under a realistic job size distribution, a Bounded Pareto with a heavy-tailed property. We find that SRPT performs an order of magnitude better with respect to mean response time and mean queue length.

While SRPT might not seem like the best choice for large jobs, particularly under overload, it turns out that under our realistic workload big jobs do not perform worse under SRPT as compared with PS in expectation. We give intuition for this. Finally we pose some interesting open questions on the topic of starvation of large jobs.

Keywords: Overload, HIGH/LOW, scheduling, processor sharing, shortest processing remaining time, time-sharing, shortest job first, SRPT, priority, unfairness, starvation, heavy-tailed behavior, high variance
1 Introduction

Overload is a situation where the rate of work arriving at a system is greater than the system can handle. In this paper we investigate systems under transient overload, where load alternates between a finite period of overload (called HIGH period) followed by a finite period of low load or zero load (LOW period). The mean system load is always below 1. This HIGH/LOW model is described in Section 2.

For most servers (e.g. web servers), even short periods of overload cause the number of jobs in the system to quickly exceed the system capacity, often within a matter of seconds. A large number of jobs implies large allocations of system resources (buffers), as well as many context switches. These effects in turn cause the server to either stop accepting new jobs or crash. Either way, client response times explode.

Recently there has been much attention paid to the problem of overload, particularly in web servers. This research can be divided into 3 primary areas. One solution is to increase the server resources, for example, adding more hardware, bandwidth, or CPU. This is often done by replacing the server by a server farm (e.g. [8, 6, 16]). Another solution is to enhance the OS to better support server software: ([20, 9, 1, 15, 19]). Finally, people try to limit the load at a server, either via installing a proxy cache (at the client or server end) (e.g. [10, 4]) or by admission control ([13]).

Our approach to coping with transient overload is different from the above approaches. Our approach does not require buying more hardware or limiting the number of system users. We simply propose scheduling the jobs in a different order from that traditionally used.

In computer systems today, when multiple jobs contend for a single resource (e.g. CPU or bandwidth), the policy used for scheduling the jobs most closely resembles Processor-Sharing (PS). That is, the desired resource is time-shared among the contending jobs, with each job in turn receiving a small quantum of service.

Processor-Sharing has many provably desirable properties when system load remains below 1, such as low mean response times and fairness (all jobs experience same mean slowdown) [17, 25]. However, in real systems the load fluctuates, sometimes exceeding 1, and it is not clear what the performance of PS is in this case.

The first result in this paper is an approximate analysis of the performance of a PS queue under transient overload. Our analysis is based on combining ideas from Jean-Marie and Robert [14] and Chen, Kella, and Weiss [5]. We validate the assumptions made in our analysis via simulation, which agrees with our analysis. We evaluate our analytic formulas in the case where the job size distribution is a Bounded Pareto distribution, with a heavy-tailed property. This distribution has been shown to be characteristic of computer workloads [18, 12, 7, 21]. Our PS results indicate that performance of a PS server under transient overload can be quite poor, even when mean system load is very low. The performance is dominated by the load during the overload period. The number
of jobs in system increases very rapidly during the overload period and then is slow to recover during the low load period, due to the time-sharing nature of PS scheduling.

In an attempt to improve performance of servers under transient overload, we propose scheduling requests under the well-known Shortest Remaining Processing Time First (SRPT) policy. The motivation is the fact that SRPT minimizes the number of jobs in the system at any time.

Applications have shied away from using SRPT for fear that SRPT "starves" big jobs, particularly under overload [3, 23, 24, 22], however the performance of SRPT under transient overload has never been analyzed. In a recent paper [2] we consider only the case with constant load below 1 (no overload). In [2] we find that, provided the load is not too close to 1 (so as to allow large jobs a turn to run), for many workloads with a heavy-tailed property, the fear of starvation is unsubstantiated. That is, all jobs, including the very largest job have lower queueing time under SRPT scheduling than under FAIR scheduling.

The second result in this paper is an approximate analysis of SRPT scheduling under transient overload. We validate the assumptions made in our analysis via simulation, which agrees with our analysis. We evaluate the analytic formulas in the case where the job size distribution is a Bounded Pareto distribution, with a heavy-tailed property.

The comparison of the PS results with the SRPT results under the Bounded Pareto workload is quite interesting. We find that:

- The improvement of SRPT over PS both in terms of the mean response time and in terms of mean number of jobs in system is substantial (close to an order of magnitude).
- The mean improvements of SRPT over PS are greatest when the overload is close to 1.
- Unfairness to big jobs under SRPT as compared with PS is virtually non-existent under our Bounded Pareto workload.
- The higher the load during the low load period, the better SRPT looks with respect to all of the above metrics.
- For distributions with a lighter tail, e.g. the exponential distribution, the results are not much worse. In fact we can prove that for an exponential distribution, if the load during the overload period exceeds 2, then every single job performs better under SRPT than under PS, in expectation.

The last observation above leads us to speculate on the class of distributions for which SRPT improves upon PS under transient overload. This discussion is covered in Section 7, where we provide analyses, observations, and open problems on this topic.

The above results are encouraging with respect to the potential real-world applicability of SRPT scheduling.
2 Model, Relevant Previous Work, and Simplifying Assumptions

2.1 Model

Throughout this paper we will assume an M/G/1 queue. The job sizes will be assumed to be independent and identically distributed with c.d.f. $F(x)$ and p.d.f. $f(x)$. The arrival process will consist of alternating periods of high and low loads. During the high load period (also called the “HIGH” period), jobs arrive with mean arrival rate $\lambda_h$ and create a load of $\rho_h > 1$. The high load period has fixed duration $t_h$. During the low load period (also called the “LOW” period) jobs arrive with mean arrival rate $\lambda_l$ and create a load of $\rho_l < 1$. The low load period has fixed duration $t_l$.

Let $\rho$ denote the average system load. Thus

$$\rho = \frac{t_h \rho_h + t_l \rho_l}{t_h + t_l}$$

We will always assume that the average load $\rho < 1$.

Our system behavior is as follows: During the HIGH period, jobs build up in the system as a function of time. We will derive an expression for this buildup. At the start of the LOW period, there is an accumulation of jobs which we refer to as “the bag of jobs.” Since the load during the LOW period is less than 1, the number in system will start decaying and will converge to a low value, related to $\rho_l$, before the end of the of the LOW period. This is expected since the average load is less than 1.

Some additional notation: The number of jobs in the system $t$ time after the onset of the HIGH period will be denoted by $N_h(t)$. Likewise the number of jobs in the system $t$ time after the onset of the LOW period will be denoted by $N_l(t)$.

2.2 Relevant Previous Work

To the best of our knowledge, the HIGH/LOW workload model with general job size distribution has not been studied under either PS or SRPT scheduling.

For SRPT scheduling, there are no relevant queueing results in the literature dealing with overload.

For PS scheduling, there are two relevant results: one by Chen et. al. [5] and the other by Jean-Marie and Robert [14]. While these results don’t cover the HIGH/LOW model, they are still very relevant in its analysis.

Jean-Marie and Robert’s work [14] can be viewed as analyzing only the HIGH period of the HIGH/LOW model under PS scheduling. They prove that $N_h(t)/t$ converges to
a particular constant in the limit as $t \to \infty$. They also derive the distribution on the residual job sizes in the limit as $t \to \infty$.

Chen et. al.'s work [5] can be viewed as analyzing only the LOW period of the HIGH/LOW model under PS scheduling. Specifically, they assume a PS server running under load $\rho_l < 1$ which begins with a set of initial jobs distributed according to an arbitrary distribution. They obtain a fluid approximation on $N_l$, however it is not expressed as a function of $t$, but rather as a function of $v(t)$, where $v(t)$ is the cumulative amount of processing time per customer allocated by the server up to time $t$.

2.3 Simplifying Assumptions

It seems natural to try to combine the results in [14] and [5] to analyze the HIGH/LOW model under PS. Our goal is to understand the mean response time for a job of size $x$ under the HIGH/LOW model. Unfortunately, we could not see how to directly combine these results to derive mean response time. Part of the problem is that the Chen result doesn’t actually compute $N_l(t)$. Another part of the problem is that the Jean-Marie and Robert result requires $t \to \infty$, whereas we want to consider finite $t_h$ and $t_l$. In our analysis it was therefore necessary for us to resort to two simplifying assumptions under PS and under SRPT:

1. $\frac{N_l(t)}{t} = \text{constant } \forall t$, for all sample paths, and
2. $N_l(t)$ is the same for all sample paths.

These simplifying assumptions will serve several purposes: First, they will allow us to obtain an approximation for expected response time for a job of size $x$ for a HIGH/LOW process under PS scheduling (see Section 3). Second, they will allow us to obtain approximations for all performance metrics of interest for a HIGH/LOW process under SRPT scheduling (expected response time for a job of size $x$, $N_h(t)$, $N_l(t)$) (See Section 4). Third, they will allow us to compare SRPT and PS performance in a HIGH/LOW workload (see Section 5).

Our simplifying assumptions may seem quite strong. However, in Section 6 we will compare the performance numbers derived using the simplifying assumptions with our simulation results which make no assumptions. We will find that provided $t_h$ and $t_l$ are not too small (at least 100 times the mean job size), the simulation numbers agree with the analytically-derived numbers.

3 Analytical Results for PS under HIGH/LOW model

The goal of this section is to derive the expected response time for a job of size $x$ under the HIGH/LOW model with PS scheduling. Throughout we will apply our simplifying
assumptions from Section 2.

Before we begin we will need to review the results of Jean-Marie and Roberts [14] and Chen et al. [5]. In reviewing these results, we will also use our simplifying assumptions to provide intuitive derivations of these results for the benefit of the reader.

3.1 Analysis of the HIGH load period only

**Lemma 0.1** (due to [14]) Consider an M/G/1/PS queue with load $\rho > 1$ and average arrival rate $\lambda_h$. Let $N_h(t)$ denote the number of jobs in the system at time $t$. Then for almost all sample paths,

$$\lim_{t \to \infty} \frac{N_h(t)}{t} = a$$

where $a$ is the solution to $\lambda_h(1 - \int_0^\infty f(x)e^{-ax}dx) = a$.

**Proof via Simplifying Assumptions:** The following is a rough proof based on our approximations made in Section 2. Since the work in the system by time $t$ increases as $(\rho_h - 1)t + o(t)$, it is not difficult to see that for almost all sample paths, $N_h(t) = at + o(t)$, for some $a$. To determine $a$, let us ignore the $o(t)$ term and approximate $N_h(t)$ by $at$, and assume that this holds for all $t$ and for all sample paths.

Now, consider a job of size $x$, arriving at time $t_a$. Then the service received by the job by time $t_0$ is

$$S(t_a, t_0) = \int_{t_a}^{t_0} \frac{dt}{N_h(t)}$$

Clearly the job departs at time $t_d$ such that $S(t_a, t_d) = x$. So, $t_d = t_a e^{ax}$. Thus, at time $t$, if a job of size $x$ arriving at time $t_a$ is present in the system, then $t_a \geq t e^{-ax}$.

Now the expected number of jobs of size between $x$ and $x + dx$ which arrive during time $[0, t]$ is $\lambda_h(t f(x)dx).$ Since the arrival process is Poisson these jobs can be assumed to arrive uniformly over $[0, t]$. Thus, we expect that about $\lambda_h t(1 - e^{-ax})f(x)dx$ of these jobs will still be present in the system at time $t$. Averaging over the possible job sizes $x$, we get that $N_h(t) = \int_0^\infty \lambda_h(t- te^{-ax})f(x)dx.$ Equating $N_h(t)$ to $at$ gives us $a = \int_0^\infty \lambda_h(1-e^{-ax})f(x)dx.$

Since we will be interested in evaluating various metrics for PS, we will make the results more compact. Let $L_g(s)$ denote the Laplace transform of a function $g$, i.e. $L_g(s) = \int_0^\infty g(x)e^{-sx}dx.$ Thus we observe that $a$ satisfies $a = \lambda_h(1 - L_f(a))$. Moreover observing that $L_f(s) = (1 - L_f(s))/s$, we can write $a$ in the following form which will be useful later.

$$L_f(a) = \frac{1}{\lambda_h}$$
We now obtain an approximation for the number of jobs with remaining size $> y$ at the end of the HIGH period, using our simplifying assumptions.\(^1\)

Consider a job of (original) size $z$. Arguing as above, this job will have remaining size $> y$ at time $t_h$ iff it’s arrival time, $t_a$, is such that $t_a > t_h e^{-a(z-y)}$

Thus the total number of jobs of size $(z, z + dz)$ which have remaining size $> y$ at time $t_h$ will be

$$\lambda_h t_h f(z) (1 - e^{-a(z-y)})dz$$

Integrating 4 over all possible job sizes greater than $y$ gives us the total number of jobs which have remaining size $> y$ at time $t_h$, which is

$$\int_y^\infty \lambda_h t_h f(z) (1 - e^{-a(z-y)})dz$$

Let $F_r$ denote the c.d.f of the remaining sizes of the jobs at the end of the HIGH period. Since the total number of jobs at time $t_h$ is $\rho(t_h)$, using 5 gives,

$$\overline{F_r}(y) = \frac{\int_y^\infty \lambda_h f(z) (1 - e^{-a(z-y)})dz}{\lambda_h}$$

After some manipulation, we observe that the Laplace Transform of $\overline{F_r}$ can be written simply as

$$L_{\overline{F_r}}(s) = \frac{1 - \lambda_h \overline{F}(s)}{s - \lambda_h}$$

3.2 Analysis of the LOW load period only

We now consider just the low load period ($\rho_l < 1$), starting with $N_l(0)$ jobs with remaining sizes distributed according to c.d.f. $F_r$. New jobs arrive into the system at rate $\lambda_l$ where the new job sizes have c.d.f. $F$. We now use our simplifying assumptions to derive an expression for the number of jobs in the system. Chen, Kella and Weiss [5] have analyzed the low-load only system and obtained a fluid limit expression for the number of jobs in the system. The [5] result is similar to our result in Equation 11.

Denote the number of jobs in this system at time $t$ by $N_l(t)$, and let us assume that $N_l(t)$ is constant for all sample paths. $N_l(t)$ will consist of two types of jobs. Jobs which were present at time 0 and those which arrived at time greater than 0. We will call these jobs of Type 1 and Type 2 respectively.

To determine $N_l(t)$, consider a job of size $x$ which arrives at time $t_a$. This job will complete at time $t_d$ such that the service received by the job during time $t_a$ to $t_d$ is $x$. Now, the service received by the job during time $t_a$ to $t_d$ under PS, will be $\int_{t_a}^{t_d} \frac{dt}{N_l(t)}$.\(^2\)

\(^1\)[14] also study the problem of residual job sizes.
Let us define $v(t) = \int_0^t \frac{dy}{N_i(y)}$, where $v(t)$ is the cumulative service per customer allocated by the server up to time $t$. Thus a job of size $x$ arriving at time $t_a$ is present in the system at time $t$ iff $v(t) - v(t_a) < x$. Thus the number of Type 1 jobs still present at time $t$ will be $N_i(0)\overline{F}_r(v(t))$. To obtain the number of Type 2 jobs, consider the jobs which arrive during time $y$ and $y + dy$. There will be approximately $\lambda_i dy$ such jobs. Out of these the number still present at time $t$ will be $\lambda_i dy \overline{F}_r(v(t) - v(y))$. Integrating over $y$ from 0 to $t$ will give us the total number of Type 2 jobs still present at time $t$. Adding the number of Type 1 and Type 2 jobs we thus obtain,

$$N_i(t) = N_i(0)\overline{F}_r(v(t)) + \int_0^t \lambda_i \overline{F}_r(v(t) - v(y))dy$$  \hspace{1cm} (8)

We do not know of a way to solve Equation 8 directly to obtain the number in system as a function of time. However, we can obtain this indirectly by solving for the number in system and solving for $t$ as a function of $v$. We will denote these by $N_i^v$ and $t^v$ respectively.

Observing that $v(0) = 0$ and $\frac{dv}{dt} = \frac{1}{n(t)}$, Equation 8 can be written as

$$N_i^v(v) = N_i(0)\overline{F}_r(v) + \int_0^v \lambda_i \overline{F}_r(v - z)N_i^v(z)dz$$  \hspace{1cm} (9)

We can now solve Equation 9. Using Laplace Transforms this gives us

$$L_{N_i^v}(s) = N_i(0)L_{\overline{F}_r}(s) + \lambda L_{\overline{F}_r}(s)L_{N_i^v}(s)$$  \hspace{1cm} (10)

which yields

$$L_{N_i^v}(s) = \frac{N_i(0)L_{\overline{F}_r}(s)}{1 - \lambda L_{\overline{F}_r}(s)}$$  \hspace{1cm} (11)

Using Equation 7 we get,

$$L_{N_i^v}(s) = \frac{N_i(0)}{s - a} \cdot \frac{(1 - \lambda L_{\overline{F}_r}(s))}{(1 - \lambda L_{\overline{F}_r}(s))}$$  \hspace{1cm} (12)

To obtain $t^v$ observe that $\frac{dv}{dt^v} = N_i^v$, thus we get,

$$L_{t^v}(s) = \frac{1}{s} \cdot \frac{N_i(0)}{s - a} \cdot \frac{(1 - \lambda L_{\overline{F}_r}(s))}{(1 - \lambda L_{\overline{F}_r}(s))}$$  \hspace{1cm} (13)

### 3.3 Analysis of Response times for the HIGH/LOW model

Using the results above we can now obtain the expressions for the response times as a function of job size under transient overload.

We classify the jobs into 3 types. Jobs which arrive during the HIGH period and finish during the HIGH period itself are type 1 jobs. Jobs which arrive during the HIGH period but finish in the LOW period are type 2 jobs. Jobs which arrive during the LOW period and finish during the same LOW period are type 3 jobs.
Note, that since the average load is less than 1, very few jobs which arrive during the LOW period will continue during the next HIGH period. Thus we consider only the three types of jobs mentioned above.

Suppose $J$ is a job of size $x$. If it is a type 1 job, then we know that it must have arrived during time $0$ to $t_h e^{-ax}$. Assuming it arrives uniformly during this interval, we obtain

$$E[T(x)|\text{Job is of type 1}] = \frac{1}{t_h e^{-ax}} \int_0^{t_h e^{-ax}} y(e^{ax} - 1)dy = \frac{1}{2} t_h (1 - e^{-ax})$$

(14)

Suppose $J$ is of type 2, then it must have arrived during time $t_h e^{-ax}$ to $t_h$. If it arrives at time $y$, then it spends $t_h - y$ time during the HIGH period and its remaining size at the beginning of the LOW period is $r(y) = x - \frac{1}{a} \log \frac{t_h}{y}$. Observe that by Equation 13 this job finishes at time $t_u(r(y))$. Thus we obtain,

$$E[T(x)|\text{Job is of type 2}] = \frac{1}{k(x)} \int_{t_h e^{-ax}}^{t_h} [(t_h - y) + t_u(x - \frac{1}{a} \log \frac{t_h}{y})]dy = \frac{1}{2} k(x) + \frac{1}{k(x)} \int_{t_h e^{-ax}}^{t_h} t_u(x - \frac{1}{a} \log \frac{t_h}{y})dy = \frac{1}{2} k(x) + \frac{1}{1 - e^{-ax}} \int_0^x at'v(x - z)e^{-az}dz$$

(15)

where $k(x) = t_h (1 - e^{-ax})$.

Finally we consider the case when $J$ is a job of type 3. To do this, we need to first digress and define quantity $t_w$.

Consider the amount of work accumulated during the HIGH period. This is approximately $(\rho_h - 1)t_h$. Thus the time until this work is removed from the system during the LOW period will be around $t_w = (\rho_h - 1)t_h / (1 - \rho_l)$.

Returning to the case of a job of type 3, we can separate jobs of type 3 into 2 cases: those which arrive before time $t_w$ and those which arrive after time $t_w$.

Let us first consider the jobs which arrive during time $t_w$ from the start of the LOW period.

Suppose a job arrives at time $t'_{a}$ from the start of the LOW period, then it will finish at time $t_u(x + v(t'_{a}))$. Given that a job arrives during the time $0$ to $t_w$, it will arrive uniformly during this interval (since the arrival process is Poisson). Thus,

$$E[T(x)|\text{Job is of type 3, and arrives in (0, t_w)]}$$
\[
\begin{align*}
E[T(x)] &= \frac{1}{t_w} \int_{0}^{t_w} (t^v(x + v(y)) - y)dy \\
&= \frac{1}{t_w} \int_{0}^{\infty} \left( t^v(z + x) - t^v(z) \right) N^v_t(z)dz 
\end{align*}
\] (16)

Observe that for each of the Type 1, 2 and 3 jobs considered above, the response time of a job is proportional to \( t_h \).

For type 3 jobs which arrive after time \( t_w \), their response time will be independent of \( t_h \). Thus if \( t_h \) is quite large (compared to the mean job size) we can assume that the response time for these jobs is negligible compared to jobs of type 1, type 2 and jobs of type 3 which arrive during time 0 to \( t_w \). Hence we will approximate the response time of type 3 jobs arriving during \( t_w \) to \( t_1 \) by 0.

So, finally we obtain

\[
E[T(x)|\text{Job is of type 3}] = \frac{1}{t_1} \int_{0}^{\infty} \left( t^v(z + x) - t^v(z) \right) N^v_t(z)dz 
\] (17)

Finally to obtain the approximate expected response time for a job of size \( x \), we find the unconditional response time. Thus we obtain that

\[
E[T(x)]_{PS} = p_1 E(T_1(x)) + p_2 E(T_2(x)) + p_3 E(T_3(x)) 
\] (18)

where \( E[T_i(x)] \) is the expected response time for a job of type \( i \), \( i = 1, 2, 3 \), as obtained in Equations 14, 15 and 17. The probability \( p_1 \) that a job of size \( x \) is of type 1 is \( \lambda_h t_h e^{-ax}/(\lambda_h t_h + \lambda t_l) \). The probability \( p_2 \) that a job of size \( x \) is of type 2 is \( (\lambda_h t_h (1 - e^{-ax}))/ (\lambda_h t_h + \lambda t_l) \). Finally the probability \( p_3 \) that a job of size \( x \) is of type 3 is \( \lambda t_l/(\lambda_h t_h + \lambda t_l) \).

Substituting the values in Equation 18 and doing some manipulation gives

\[
E[T(x)]_{PS} = \frac{\lambda_h t_h}{\lambda_h t_h + \lambda t_l} \left[ \frac{1}{2} \left( 1 - e^{-ax} \right) t_h + \int_{0}^{x} at^v(x - z)e^{-a z}dz \right] \\
+ \frac{\lambda t_l}{\lambda_h t_h + \lambda t_l} \left[ \int_{0}^{\infty} \left( t^v(z + x) - t^v(z) \right) N^v_t(z)dz \right] 
\] (19)

Observe that if \( \rho_l = 0 \),

\[
E[T(x)]_{PS} = \frac{1}{2} \left( 1 - e^{-ax} \right) t_h + \int_{0}^{x} at^v(x - z)e^{-a z}dz 
\] (20)
4 Analytical Results for SRPT under HIGH/LOW model

In the previous section we derived an approximation for the response time for a job of size $x$ under PS scheduling. The goal of this section is to derive an approximation for the response time for a job of size $x$ under the HIGH/LOW model with SRPT scheduling.

4.1 Number of jobs during HIGH period

Let $x_o$ be defined such that $\rho_h(x_o) = 1$.

Consider a job of size $x > x_o$. This job will never run as long as there are jobs of size less than $x$ in the system. Since $\rho_h(x) > 1$, it follows that for almost all sample paths $\omega$ there will be a time $t(\omega)$ such that the work made up by jobs of size less than $x$ will be non-zero for all $t > t(\omega)$. Hence, if the job $x$ arrives after time $t(\omega)$ it will remain in the system at least until the end of the HIGH period. Assuming that $t_h$ is large compared to $t(\omega)$, we will use the approximation that all jobs of size $x > x_o$ are held back during the HIGH period.

Secondly, for jobs of size $x < x_o$ the expected response time for a job of size $x$ is $o(t_h)$, as shown in [2]. The reason for this is that if we consider the busy periods during which jobs of size $\leq x$ are executed, then at most one job of size $x > x$ can affect a particular busy period. Hence, the response time for a job of size $x$ is not affected significantly by jobs of size $x > x$, and hence is $o(t_h)$, if $\rho(x) < 1$.

Using the observations above we approximate the number of jobs at time $t$ as

$$N_h(t) = \lambda_h t F(x_o)$$

and the distribution of the job sizes of the accumulated jobs as

$$f_r(x) = \begin{cases} \frac{f(x)}{F(x_o)}, & x > x_o \\ 0, & \text{otherwise} \end{cases}$$

4.2 Number of jobs during LOW period

During the LOW period, $\rho_l < 1$, and thus the jobs accumulated during the HIGH period will start receiving some share of the processor. Let us consider what the accumulated jobs look like during the LOW period. At time $t = 0$, the jobs in the bag are of size $x_o$ and greater. As time progresses, jobs will be cleared from the bag starting from the jobs of size $x_o$. Let $x(t)$ be the size of the smallest job remaining in the bag at time $t$. We will first approximate $x(t)$ as a function of $t$.

Clearly $x(0) = x_o$. Consider the scenario at time $t$. Consider how much time it takes to advance $x(t)$ by $dx$ amount. This will simply be the amount of work of size between
\( x(t) \) and \( x(t) + dx \) present in the system at time \( t \). Suppose it takes \( dt \) time to do this. Let us estimate \( dt \) in terms of \( dx \). The expected amount of work that needs to be done during the \( dt \) time can be divided into 3 parts:

1. The expected work made up of jobs in the original bag with sizes between \( x(t) \) and \( x(t) + dx \). This work totals to \( \lambda_h t_h x(t) f(x(t))dx \).

2. Pending work due to the new arrivals (during the LOW period) with sizes between \( x(t) \) and \( x(t) + dx \). Note that these jobs were not worked upon, until now. This new work will be \( \lambda_t x(t) f(x(t))dx \).

3. The fresh work which arrives during this \( dt \) amount of time. We need only consider work made up by jobs of size less than \( x(t) \), since \( x(t) \) will be affected by other jobs. This work will be \( \rho_t(x(t))dt \).

Using these observations we can write,

\[
dt = \lambda_h t_h x(t) f(x(t))dx + \lambda_t x(t) f(x(t))dx + \rho_t(x(t))dt
\]

Hence,

\[
\frac{dt}{\lambda_h t_h + \lambda_t} = \frac{x(t) f(x(t))dx}{1 - \rho_t(x(t))}
\]

Integrating the l.h.s. of Equation 23 from 0 to \( t \) and the r.h.s. from \( x_o \) to \( x(t) \) and observing that \( \frac{d\rho_t(x)}{dx} = \lambda_t x f(x) \) and \( \rho_t(x_o) = 1 \), we get,

\[
\frac{1}{\lambda_t} \log \frac{\lambda_h t_h + \lambda_t t}{\lambda_h t_h} = \frac{1}{\lambda_t} \log \frac{1 - \rho_t(x_o)}{1 - \rho_t(x(t))}
\]

Thus we obtain\(^2\)

\[
\rho_t(x(t)) = \frac{\lambda_t (t + t_h)}{\lambda_h t_h + \lambda_t t}
\]

Having obtained \( x(t) \), \( N_t(t) \) can be readily determined, since by the above arguments the number of jobs at time \( t \) will be approximately \((\lambda_h t_h + \lambda_t t) \bar{F}(x(t))\). Thus,

\[
N_t(t) = (\lambda_h t_h + \lambda_t t) \bar{F}(x(t))
\]

Finally, we obtain the inverse of \( x(t) \) which we denote by \( \bar{x}(t) \). Thus \( \bar{x}(x(t)) = t \). This will be useful in obtaining the approximation for the expected response time. Clearly \( \bar{x}(x) \) is only defined for \( x > x_o \). Then Equation 24 gives\(^3\)

\[
\bar{x}(x) = t_h \frac{\rho_h(x) - 1}{1 - \rho_t(x)}
\]

\(^2\)For the case when \( \lambda_t = 0 \), we get \( \rho_t(x(t)) = 1 + \frac{1}{t_h} \)

\(^3\)For the case when \( \lambda_t = 0 \), we get \( \bar{x}(x) = t_h (\rho_h(x) - 1) \)
4.3 Response times as a function of job size

To derive the expected response time under SRPT, we first consider a job of size less than \( x_o \). For such a job we know that the expected response time is a constant, hence \( o(t_h) \). This holds whether the job arrives during the HIGH period or during the LOW period.

Consider a job of size \( x \), \( x > x_o \). Suppose the job arrives during the HIGH period. Then it will wait throughout the HIGH period, and approximately \( \hat{t}(x) \) time during the LOW period. Given that the job arrives during the HIGH period, it will arrive uniformly in the interval since the arrival process is Poisson. Thus

\[
E[T(x)\mid \text{job arrives during HIGH period}]_{SRPT} = \frac{1}{2} t_h + \hat{t}(x)
\]  

(27)

Now consider the case when the job of size \( x \), \( x > x_o \) arrives during the LOW period. Again, if the job arrives after time \( \hat{t}(x) \), then its response time will be \( o(t_h) \).

Finally, given that the job arrives during the low period between time 0 to \( \hat{t}(x) \), its response time will be approximately \( \frac{1}{2} \hat{t}(x) \).

The probability that a job arrives during the HIGH period is \( \lambda_h t_h / (\lambda_h t_h + \lambda_l t_l) \). The probability that a job arrives during the interval \([0, \hat{t}(x)]\) of the LOW period is \( \lambda_l \hat{t}(x) / (\lambda_h t_h + \lambda_l t_l) \).

Thus the expected response time for a job of size \( x \geq x_o \) under SRPT will be

\[
E[T(x)]_{SRPT} = \frac{\lambda_h t_h}{\lambda_h t_h + \lambda_l t_l} \left( \frac{1}{2} t_h + \hat{t}(x) \right) + \frac{\lambda_l \hat{t}^2(x)}{\lambda_h 2t_h}
\]

(28)

Observe that when \( \rho_h \) and \( \rho_l \) are fixed \( E[T(x)] \) varies linearly with \( t_h \). However if the job size is less than \( x_o \), we know that \( E[T(x)] \) is \( o(t_h) \). Hence, for large \( t_h \) the ratio of response times for a job of size \( < x_o \) to that of size \( > x_o \) will tend to 0. Thus we will approximate \( E[T(x)] \) by 0 for \( x < x_o \). We will show later in Section 6 that this approximation is quite good.

5 Comparison of PS and SRPT based on analytical results

In Sections 3 and 4 we derived the approximation for the expected response time of a job of size \( x \) under SRPT and PS for the HIGH/LOW model. These results are summarized in Equations 28 and 19 respectively. It is difficult to compare these two equations analytically for generally-distributed workloads, therefore in this section we compare Equations 28 and 19 evaluated on a particular real-world workload. Although we only show results for this particular workload, we will make observations which apply to more general workloads as well, and we will discuss intuitions for why these results hold more generally. In Section 7 we return to the discussion of other workloads.
Our workload assumes job sizes have a Bounded Pareto distribution. This distribution is defined as follows:

Recall a *Pareto* distribution with parameter $\alpha$, is defined such that

$$P r[X > x] \sim x^{-\alpha}, \text{ where } 0 < \alpha < 2$$

The *Bounded-Pareto* distribution [11] is characterized by three parameters: $\alpha$, the exponent of the power law; $k$, the smallest possible job; and $p$, the largest possible job. The probability density function for the Bounded Pareto $B(k, p, \alpha)$ is defined as:

$$f(x) = \frac{\alpha k^\alpha}{1 - (k/p)^\alpha} x^{-\alpha - 1} \quad k \leq x \leq p.$$ 

In this paper, we consider the distribution $B(k, p, \alpha)$ obtained by keeping the mean fixed (at 3000) and the maximum value fixed (at $p = 10^{10}$ and $\alpha = 1.5$). These parameters correspond to typical values for Web workloads taken from [7]. Throughout we normalize the distribution by scaling down the job sizes by a factor of 3000, leaving the mean as 1. We refer to this normalized distribution as $B(\alpha = 1.5)$.

Pareto and Bounded-Pareto distributions have been shown to be characteristic of the job size distributions in many computer workloads [18, 12, 7, 21]. These distributions have 3 important properties:

1. Decreasing failure rate (Pareto) or mostly-decreasing (Bounded Pareto).
2. Infinite variance (Pareto) or very high variance (Bounded Pareto).
3. The *heavy-tailed* property, which we define as: “A very small fraction of the largest jobs (e.g., 1%) comprise more than half the total load.”

We now compare Equations 28 and 19 evaluated on our $B(\alpha = 1.5)$ distribution with respect to two metrics:

1. Number of jobs in system as a function of time (Section 5.1).
2. Response time for a jobs of size $x$ (Section 5.2).

### 5.1 Number of jobs in system

We now consider the number of jobs in the system as a function of time under PS and SRPT. The number in system is an interesting practical metric. Consider as an example a Web server which services its requests in SRPT order, as opposed to the traditional PS service order. The number of requests in the system corresponds to the number of simultaneously open connections in the Web server. The greater this number the more
overhead is required by the Web server. Furthermore, if this number gets too high, the Web server may crash or simply stop accepting requests.

Figure 1 shows the number of jobs as a function of time for various values of $\rho_h$ with $\rho_l = 0$, under SRPT and PS scheduling for the $B(\alpha = 1.5)$ distribution. Here $t_h = 1000$ and $t_l$ is chosen such that the average load is 0.7. Observe that a value of 1000 on the x-axis indicates the end of an HIGH period. Observe that the area under the curves is proportional to the mean number of jobs in the system (hence to the mean response time).

We now state a few general observations based on Figure 1. For each observation, we provide intuition, and discuss why it should hold for more general workloads as well.

**Observation 1** The mean number of jobs in the system under SRPT is significantly less than that under PS.

This is due to the fact that SRPT is known to always minimize the number of jobs in the system (under all conditions) since it always works on that job which can be completed most quickly.

**Observation 2** Though still significant, the relative advantage of SRPT over PS (with respect to number of jobs in the system) decreases at higher values of $\rho_h$. (See Figures 1a and 1b).

This occurs since as $\rho_h$ grows very high, both PS and SRPT retain almost all jobs during
the HIGH period and thus the difference in their performance depends only on their behavior during the low period.

**Observation 3** SRPT recuperates from overload faster than PS. Specifically, the curve for the number of jobs under SRPT during the LOW period appears convex, while this may not be true for PS.

This follows since SRPT works on jobs with the smallest remaining size first, thus the rate of clearance of jobs is maximum in the beginning of the LOW period and then decreases. Also, since PS timeshares among all the jobs, it somewhat delays getting jobs out at the beginning of the LOW period. This can be observed in Figure 1b (right). Thus SRPT not only accumulates fewer jobs, but it also gets them out as quickly as possible.

Finally, a subtle, but important observation:

**Observation 4** Given a fixed $\rho_l$, the number in system in Figure 1 does not depend on the average load (provided $\rho < 1$). In general, the average load $\rho$ can be made arbitrarily low or arbitrarily close to 1 by choosing $t_1$ accordingly. However, the the number of jobs during the HIGH period or the LOW period will not change.

This is an important departure from the usual M/G/1 queueing model in the sense that the performance metric does not depend on the average load, but only on the load during the overload period.

Finally, we find that the trends look similar for the case when $\rho_l > 0$. Notice that the number of jobs during the HIGH period is independent of $\rho_l$. Hence $N_h(t)$ looks identical for all values of $\rho_l$. The only noticeable difference is that that it takes a longer time (stretched by $1/(1 - \rho_l)$) during the LOW period for the number of jobs to go down to zero.

### 5.2 Expected Response Times as a function of job size

The point of this section is to determine whether the large jobs “starve” under SRPT scheduling as compared with PS scheduling. We do this by observing the expected response time for large jobs.

Figure 2 shows the expected response time as a function of job size under SRPT versus PS for the $B(\alpha = 1.5)$ distribution, when $\rho_l = 0$. The job size is expressed as a percentile of the job size distribution (where 100 percentile indicates the very largest job). Note that due to the choice of the x-axis, the area under the PS (respectively, SRPT) curve corresponds to the mean response time under PS (respectively, SRPT). In Figure 3 we consider the case when $\rho_l = 0.5$. However we use a 3 stage hyper-exponential approximation to $B(1.5)$,

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4The plots for this case are not included for lack of space.
Figure 2: Analytic results: Response times as a function of job size for (a) $\rho_h = 1.3, \rho_l = 0$ and (b) $\rho_h = 2, \rho_l = 0$. Both figures assume the $B(1.5)$ distribution. $t_h$ is assumed to be 1000, $\rho_l = 0$ and $t_l$ is chosen such that the average load is 0.7.

Figure 3: Analytic results: Response times as a function of job size under (a) $\rho_h = 1.3, \rho_l = 0.5$ and (b) $\rho_h = 2, \rho_l = 0.5$. Both figures assume the 3 phase approximation to the $B(1.5)$ distribution. $t_h$ is 1000, and $t_l$ is chosen such that the average load is 0.7.
in order to evaluate the response times (Recall that for \( \rho_l > 0 \), evaluating \( E[T(x)] \) for PS requires the use of the Laplace Transform of the job size distribution).

5.2.1 Case when \( \rho_l = 0 \)

We first consider the case when \( \rho_l = 0 \). The following observations hold for all plots in Figure 2 and are easily explained:

**Observation 5** Under SRPT jobs of size less than \( x_o \) have an approximate response time of 0 when compared with \( t_h \).

**Observation 6** Large jobs do not necessarily suffer under SRPT as compared with PS (as is commonly believed). (See Figure 2.)

To see why this is the case, observe that although large jobs do badly under SRPT, they do almost equally badly under PS. The point is that the average amount of service received by a large job under PS during a HIGH period is negligible compared to its size. Thus this job stays in the system throughout the HIGH period (since its arrival). Moreover it is among the last of the jobs to complete during the LOW period, since its remaining size at the beginning of the LOW period is large compared to other remaining jobs.

The observations above make a very strong case for SRPT. Not only is there a significant improvement in the mean response time under SRPT, but this improvement does not come at the cost of starving large jobs.

Finally, we note (as in Observation 4) that the results above only depend on \( \rho_h \). In particular, they are not a consequence of fixing the average load to a value of 0.7.

5.2.2 Case when \( \rho_l > 0 \)

**Observation 7** The extent of starvation when \( \rho_l > 0 \) is much less than that when \( \rho_l = 0 \).

To see why this is the case, first observe that the growth rate of jobs during the HIGH period is the same for both \( \rho_l = 0 \) and \( \rho_l = 0.5 \). Thus the difference in the response times arises due to the behavior of the system during the LOW period. Observe that the number of jobs is non-zero for a longer time (by a factor of \( 1/(1 - \rho_l) \)) during the LOW period as compared to the case when \( \rho_l = 0 \). Thus the poor behavior of PS with respect to clearing jobs out of the system as noted in Observation 3 is accentuated for the case when \( \rho_l > 0 \).
6 Simulation Comparison

The discussion in Section 5 was based on analytical results obtained in Sections 3 and 4. In this section we compare how well our approximations agree with the actual numbers obtained by simulation.

We simulate PS and SRPT under the HIGH/LOW model for various distributions. The goal of the simulation is two-fold.

1. To figure out how large \( t_h \) needs to be so that the analytical results match simulation.
2. To figure out the number of the HIGH/LOW cycles required (denoted by \( n \)), so that the metrics of interest averaged over these cycles converge to their mean.

Figure 4 shows the response times as a function of the percentile of job size under PS and SRPT, both under theory and simulation. The job size distribution is the 3 stage hyper-exponential approximation to \( \text{B}(1.5) \), with mean job size scaled down to 1, which we denote by \( B_h(1.5) \). The parameters \( t_h \) and \( n \) are set to 1000 and 100 respectively. Observe that for PS the approximation matches simulation almost exactly. For SRPT the approximation is close except for the point at \( x_o \). We find that, as \( t_h \) is made higher, the simulation and analysis results grow closer for SRPT.

![Simulation comparison results](image)

Figure 4: Simulation results as compared with analytical approximation results, \( \rho_h = 2, \rho_t = 0.5 \).

In general for all workload distributions we tested\(^5\), we found that our analytical approximation results match almost exactly with simulation when \( t_h \) is large. We found that simulation matched analysis even for \( t_h \) only about 100-1000 times the mean job size, provided \( \rho_h \) was not too close too 1. For low variability distributions \( t_h \) and \( n \) may be lower. The comparison of simulation and analysis is an even closer match when considering the metric number in system as a function of time.

\(^5\) More plots not shown here for lack of space.
7 SRPT vs. PS under more general workloads – Open problems

In Section 5 we compared the performance of SRPT vs. PS on a $B(\alpha = 1.5)$ distribution under transient overload. We found that SRPT had two desirable properties: (1) It improved significantly upon PS with respect to mean response time, and (2) SRPT did not treat large jobs worse than they would be treated under PS.

The point of this section is to explore whether these desirable properties of SRPT also translate to other distributions.

We first show that the desirable properties for low variability distributions as well. We prove that for the exponential distribution ($C^2 = 1$), SRPT improves upon PS for every job.

**Observation 8** For an exponential with $\rho_h > 2$ and $\rho_l = 0$ every job has lower expected response time under SRPT as compared with PS.

**Proof:**

Without loss of generality assume that the mean job size is 1. Then $\lambda_h = \rho_h$, and $a = \rho_h - 1$. Recall the expression for response time under PS,

$$E[T(x)]_{PS} = (1 - e^{-ax})\frac{t_h}{2} + \int_0^x at^v(x - z)e^{-az}dz$$

Now using the expression for $t^v$ we note that the

$$L_{t^v}(s) = \frac{1}{s} \cdot \frac{at_h}{s - a} (1 - \lambda_hL_{F_r}(s))$$

Note that $F_r(x) = e^{-x}$, hence $L_{F_r}(x) = \frac{1}{s+1}$.

Thus,

$$L_{t^v}(s) = \frac{at_h}{s(s+1)}$$

which gives $t^v(y) = at_h(1 - e^{-y})$. Thus,

$$E[T(x)]_{PS} = (1 - e^{-ax})\frac{t_h}{2} + a^2 \int_0^x t_h(1 - e^{-x})e^{-az}dz$$

$$E[T(x)]_{PS} = (a + 1)\frac{t_h}{2} + \frac{1}{2a - 1}e^{-ax}t_h - \frac{a^2}{a - 1}e^{-x}t_h$$

(29)

By Equation 28 we get,

$$E[T(x)]_{SRPT} = (a + 1)t_h \int_0^x ye^{-y}dy - \frac{t_h}{2}$$
Figure 5: No starvation for big jobs under Exponential job size distribution with $\rho_h = 2$, $\rho_l = 0$

which gives,

$$E[T(x)]_{SRPT} = (a + \frac{1}{2})t_h - (a + 1)(e^{-x} + xe^{-x})t_h$$  \hspace{1cm} (30)$$

We now show that for $a > 1$, $\forall x$, $E[T(x)]_{PS} > E[T(x)]_{SRPT}$.

To show that $E[T(x)]_{PS} > E[T(x)]_{SRPT}$, it suffices to show that

$$\frac{a^2}{a - 1}e^{-x} - \frac{a + 1}{2(a - 1)}e^{-ax} < (a + 1)(e^{-x} + xe^{-x})$$

Or equivalently, we must show that:

$$\frac{1}{a - 1}e^{-x} < \frac{a + 1}{2(a - 1)}e^{-ax} + (a + 1)xe^{-x}$$

$$1 < (a^2 - 1)x + \frac{a + 1}{2}e^{(-a+1)x}$$

Define

$$g(x) = 1 - (a^2 - 1)x - \frac{a + 1}{2}e^{(-a+1)x}$$

Clearly $g(0) = (1 - a)/2$ which is less than 0. $g'(x) = (1 - a^2)(1 - e^{(-a+1)x}/2)$, which is less than 0, for $x \geq 0$. Thus $g(x) < 0$ for all $x > 0$, and the proof follows.  \hspace{.5cm} \blacksquare$

Figure 5 shows response time as a function of job size when the job size distribution is Exponential when $\rho_h = 2$ and $\rho_l = 0$. The figure shows that every job prefers SRPT to PS in expectation and furthermore that the improvement in mean response time of SRPT over PS is significant (a factor of about 3, judging by the area under the curves). Observe that in the above proof, we required that $\rho_h > 2$. We now show that if $\rho_h$ is close to 1, then big jobs fair worse under SRPT as compared with PS.
Figure 6: Big jobs suffer under SRPT for Exponential job size distribution with $\rho_h = 1.1$, $\rho_l = 0$

**Observation 9** For an exponential job size distribution with $\rho_h \approx 1$ and $\rho_l = 0$, big jobs perform worse under SRPT as compared with PS.

The proof is evident from Figure 6. We now provide some intuition for why the case of $\rho_h \approx 1$ is bad for SRPT as compared with PS with respect to large jobs.

Consider a job of size a little bigger than $x_0$. If $\rho_h \approx 1$, then there is a big difference under SRPT vs. PS with respect to the remaining size on this job at the end of the HIGH period. SRPT has a large remaining size (the original size of the job). In contrast under PS since $\rho_h$ is low, big jobs receive quite a bit of service during the high period, hence their remaining size is small, and therefore response times in this case are lower under PS than SRPT (see Figure 6).

However even in the case of $\rho_h \approx 1$, we find that for the $B(\alpha = 1.5)$ distribution, starvation is largely diminished.

**Observation 10** When $\rho_h = 1.1$, and the job size distribution is $B(\alpha = 1.5)$, only 0.06% of jobs do worse under SRPT as compared with PS, and only at most 20% worse.

To explain the above observation, we have to return to the heavy-tailed property exhibited by the $B(\alpha = 1.5)$ distribution. Recall that the heavy-tailed property says that very few of the largest jobs carry all the weight. For distributions with a heavy-tailed property, the work received by the very largest jobs under PS during the HIGH period will be negligible compared to their very large size. Thus these jobs perform comparably under SRPT and PS. Hence no “starvation.”

We leave it as an open problem to determine the exact conditions under which SRPT has desirable properties as compared with PS. We conjecture that these conditions are
related to \( \rho_h \), and the nature of the job size distribution, in particular the weight of its tail, its overall range, and its variance.

8 Conclusion

In this paper we propose a solution for coping with transient overload in systems. Our proposal is that the traditional scheduling policy used in systems (PS scheduling) be replaced by SRPT scheduling of jobs. We consider a HIGH-LOW workload model which alternates between a period of overload (HIGH) and a period of low load (LOW). We obtain analytical approximations on the response time for a job of size \( x \) under the HIGH-LOW model and other metrics. Our analytical approximation results match closely to simulation.

We find the behavior of both SRPT and PS under the HIGH-LOW model to be quite different from that under a tradition \( M/G/1, \rho < 1 \) queue. For example, the main property of PS – fairness (in the sense of the same mean slowdown for all jobs) – no longer holds under the HIGH-LOW model.

We evaluate our analytical results on a distribution characteristic of today’s computer workloads, and find that, for the HIGH-LOW model, SRPT improves upon PS for every job size (in expectation), and that the improvement in mean response time can be quite significant.

Our analysis sheds light on some general reasons for why SRPT performs well as compared with PS under the HIGH-LOW model. Our analysis shows that PS is particularly ineffective in dealing with periods of temporary overload. Due to its time-sharing nature, it deteriorates the performance of all the jobs, which is actually unfair to small jobs. Moreover, PS is particularly slow at getting the system “back to normal” once the overload has disappeared. By contrast, SRPT accumulates far fewer jobs during the overload period, and is also much more efficient at getting them out once the overload period is over.

Our analysis sheds light on why SRPT performs well as compared with PS even on large jobs. The reason is that while it seems obvious that large jobs won’t receive service under SRPT during the overload period, it turns out that they also don’t receive much service under PS during the overload period. This is particularly true when the distribution has a heavy-tailed property.

Our results have implications for real world systems, where fluctuations in load are common. Under PS scheduling, these fluctuations can result in large buildups of jobs which continue to effect the system for a long time. In SRPT, as we’ve shown, the effect of load fluctuations is much less severe.
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References


