

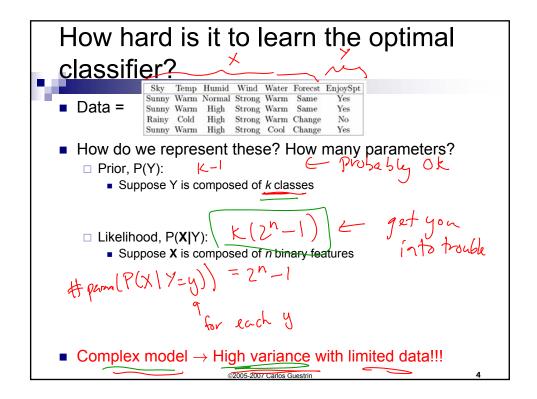
Optimal classification
$$P(Y|X)$$

Theorem: Bayes classifier h_{Bayes} is optimal! $f_{Y|X}$
 $f_{Bayes}(X) = g_{Y(X)} + argus P(Y=g|X=X)$

That is $error_{true}(h_{Bayes})) \leq error_{true}(h)$, $\forall h(X)$

Proof: $p(error_{X}) = \int_{x} p(error_{X}|x)p(x)dx$

Theorem: Bayes classifier h_{Bayes} is optimal! $f_{Y|X}$
 $f_{Y|X}$



Conditional Independence

- X is conditionally independent of Y given Z, if the probability distribution governing X is independent of the value of Y, given the value of Z $(\forall i, j, k) P(X = i | Y = j, Z = k) = P(X = i | Z = k)$
- e.g., P(Thunder|Rain, Lightning) = P(Thunder|Lightning)TIRIL: Thunder independent of raining given lightning: XIYIZ
- Equivalent to:

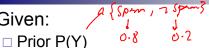
$$P(X,Y \mid Z) = P(X \mid Z)P(Y \mid Z)$$

The Naïve Bayes assumption XIIXZIY Naïve Bayes assumption: □ Features are independent given class: $P(X_1, X_2|Y) = P(X_1|X_2, Y)P(X_2|Y)$ $= P(X_1|Y)P(X_2|Y)$ $P(X_1...X_n|Y) = \prod P(X_i|Y)$ How many parameters now? Suppose X is composed of n binary features X.(2n-1) parameters

The Naïve Bayes Classifier









- □ *n* conditionally independent features **X** given the class Y
- \square For each X_i , we have likelihood $P(X_i|Y)$

■ Decision rule:

Decision rule:
$$y^* = h_{NB}(\mathbf{x}) = \arg\max_{y} P(y) P(x_1, \dots, x_n \mid y)$$
$$= \arg\max_{y} P(y) \prod_{i} P(x_i \mid y)$$

■ If assumption holds, NB is optimal classifier!

MLE for the parameters of NB



- Given dataset 🞾
 - \square Count(A=a,B=b) \leftarrow number of examples where A=a and B=b

MLE for NB, simply:

Prior:
$$P(Y=y) = \frac{Count(Y=y)}{|D|}$$

□ Likelihood: $P(X_i=x_i|Y_i=y_i) = \frac{Count(X_i=x_i, Y_i=y_i)}{Count(Y_i=y_i)}$

Subtleties of NB classifier 1 – Violating the NB assumption

Usually, features are not conditionally independent:

$$P(X_1...X_n|Y) \neq \prod_i P(X_i|Y)$$

- Actual probabilities P(Y|X) often biased towards 0 or 1
- Nonetheless, NB is the single most used classifier out there
 - □ NB often performs well, even when assumption is violated
 - □ [Domingos & Pazzani '96] discuss some conditions for good performance

©2005-2007 Carlos Guestrin

ç

Subtleties of NB classifier 2 – Insufficient training data

- What if you never see a training instance where X_1 =a when Y=b? $P(X_1 = 1) = \emptyset$
 - □ e.g., Y={SpamEmail}, X₁={'Enlargement'}
 - $\Box P(X_1=a \mid Y=b) = 0$
- Thus, no matter what the values $X_2,...,X_n$ take:

$$P(Y=b \mid X_1=a, X_2, ..., X_n) = 0 \ 0$$

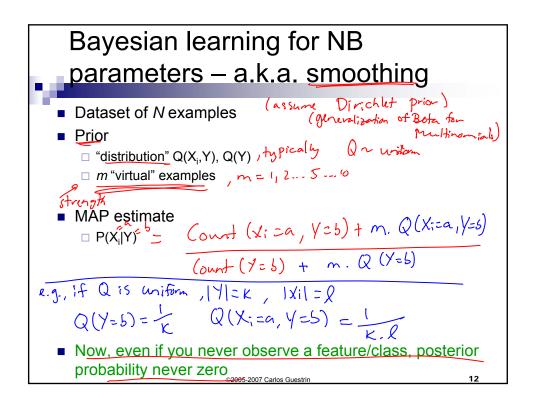
$$P(Y=b \mid X_1=a, X_2, ..., X_n) = 0 \ 0$$

$$P(X_1=a, X_2, ..., X_n) = 0 \ 0$$

What now???

©2005-2007 Carlos Guestrir

MAP for Beta distribution $P(\theta \mid \mathcal{D}) = \frac{\theta^{\beta_H + \alpha_H - 1}(1 - \theta)^{\beta_T + \alpha_T - 1}}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$ $\theta = \arg\max_{\theta} P(\theta \mid \mathcal{D}) = \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim \frac{\beta_H + \alpha_H - 1}{B(\beta_H + \alpha_H,$



Text classification



- Classify e-mails
 - □ Y = {Spam,NotSpam}
- Classify news articles
 - ☐ Y = {what is the topic of the article?}
- Classify webpages
 - ☐ Y = {Student, professor, project, ...}
- What about the features X?
 - The text!

©2005-2007 Carlos Guestrin

Features X are entire document – X; for ith word in article



Article from rec.sport.hockey

Path: cantaloupe.srv.cs.cmu.edu!das-news.harvard.e

From: xxx@yyy.zzz.edu (John Doe)
Subject: Re: This year's biggest and worst (opinic

Ann 03 09:53:39 GMT I can only comment on the Kings, but the most

obvious candidate for pleasant surprise is Alex Zhitnik. He came highly touted as a defensive defenseman, but he's clearly much more than that. Great skater and hard shot (though wish he were more accurate). In fact, he pretty much allowed the Kings to trade away that huge defensive liability Paul Coffey. Kelly Hrudey is only the biggest disappointment if you thought he was any good to begin with. But, at best, he's only a mediocre goaltender. A better choice would be Tomas Sandstrom, though not through any fault of his own, but because some thugs in Toronto decided

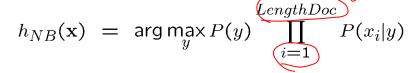
NB for Text classification



- P(X|Y) is huge!!!
 - □ Article at least 1000 words, $\mathbf{X} = \{X_1, ..., X_{1000}\}$
 - □ X_i represents ith word in document, i.e., the domain of X_i is entire vocabulary, e.g., Webster Dictionary (or more), (10,000 words), etc.

explicitly # (P(XIX)) -

- NB assumption helps a lot!!!
 - \square P(X_i=x_i|Y=y) is just the probability of observing word x_i in a document on topic y



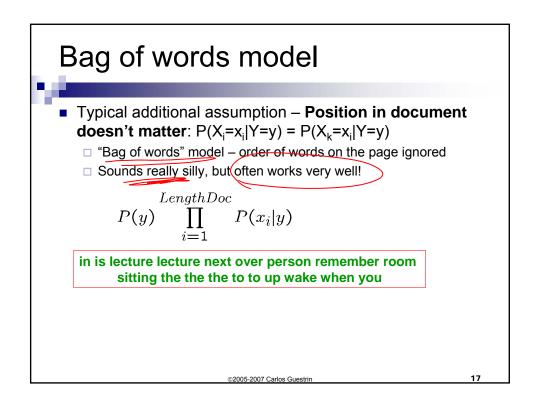
Bag of words model

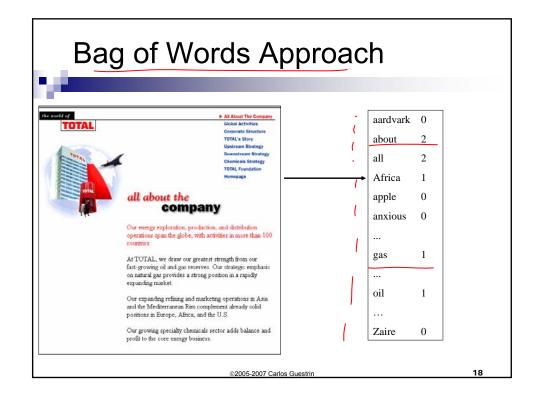


- Typical additional assumption Position in document doesn't matter: $P(X_i=x_i|Y=y) = P(X_k=x_i|Y=y)$
 - □ "Bag of words" model order of words on the page ignored
 - □ Sounds really silly, but often works very well!

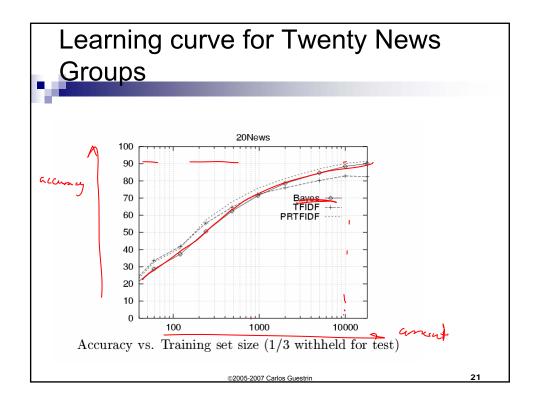
$$P(y) \prod_{i=1}^{LengthDoc} P(x_i|y)$$

When the lecture is over, remember to wake up the person sitting next to you in the lecture room.

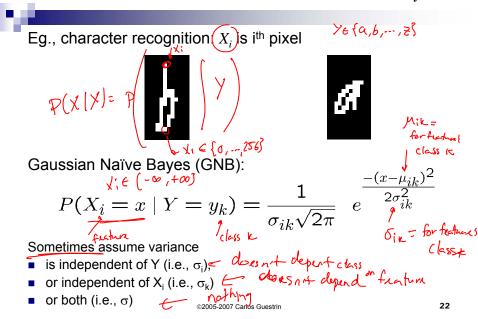


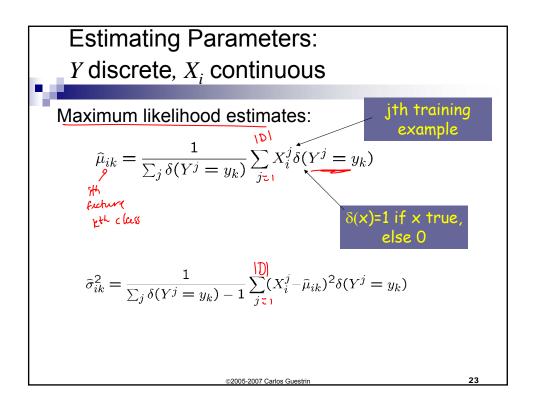


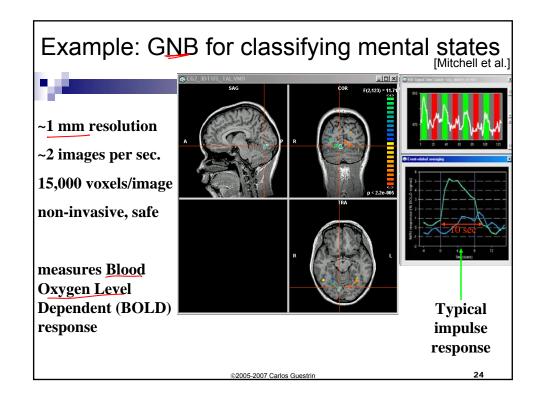


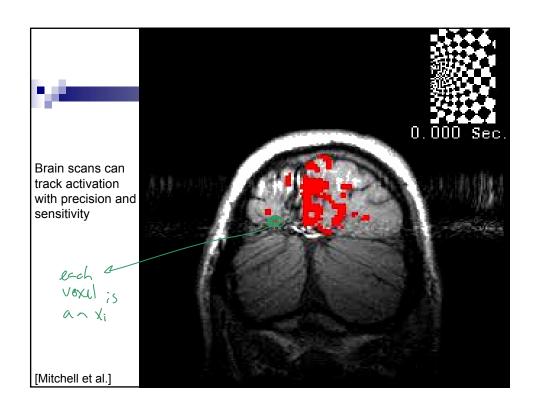


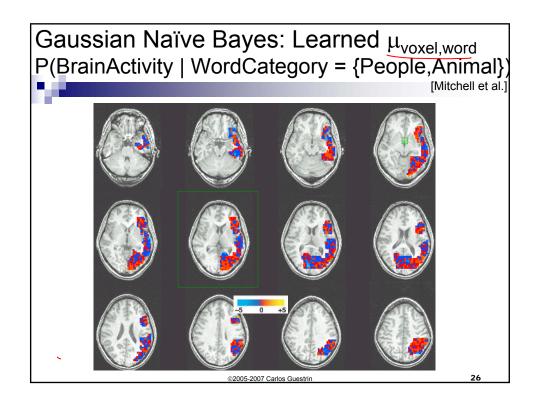
What if we have continuous X_i ?

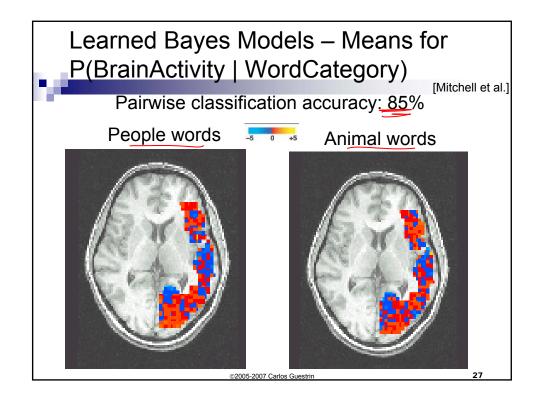


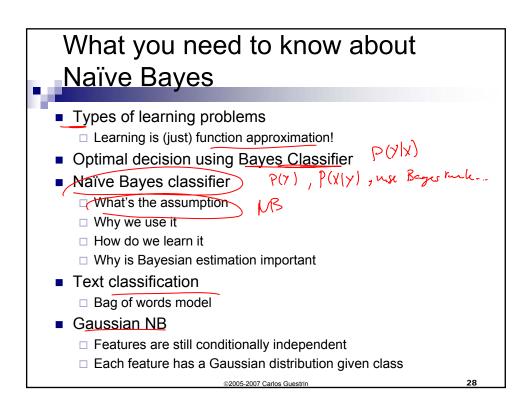




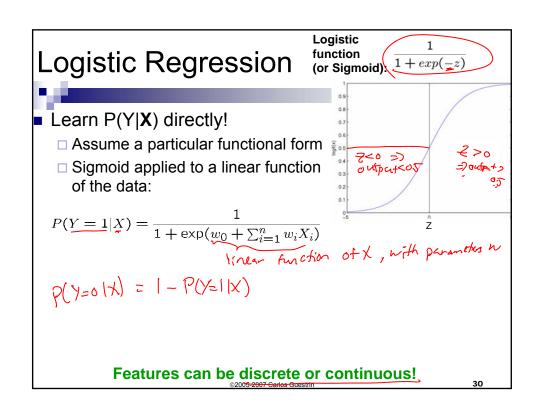








Generative v. Discriminative classifiers - Intuition Want to Learn: h: $X \mapsto Y$ □ X – features ityon have League (x) = argmax P(y) xxx □ Y – target classes Bayes optimal classifier – P(Y|X)Generative classifier, e.g., Naïve Bayes: NB while \square Assume some functional form for P(X|Y), P(Y)☐ Estimate parameters of P(X|Y), P(Y) directly from training data □ Use Bayes rule to calculate $P(Y|X=x) \sim P(Y)$, P(X=x|Y)This is a 'generative' model vityon went to classify Indirect computation of P(Y|X) through Bayes rule ■ But, can generate a sample of the data, $P(X) = \sum_{y} P(y) P(X|y)$ Discriminative classifiers, e.g., Logistic Regression: Assume some functional form for P(Y|X) directly can del P(Y|X) □ Estimate parameters of P(Y|X) directly from training data □ This is the 'discriminative' model MScriminate classes, l.g., Directly learn P(Y|X) But cannot obtain a sample of the data, because P(X) is not available



Understanding the sigmoid
$$g(w_0 + \sum_{i} w_i x_i) = \frac{1}{1 + e^{w_0 + \sum_{i} w_i x_i}}$$

$$v_0 = -2, w_1 = -1$$

$$v_0 = 0, w_1 = -0.5$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_2 = 0$$

$$v_3 = 0$$

$$v_4 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_2 = 0$$

$$v_3 = 0$$

$$v_4 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_2 = 0$$

$$v_3 = 0$$

$$v_4 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_2 = 0$$

$$v_3 = 0$$

$$v_4 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_2 = 0$$

$$v_3 = 0$$

$$v_4 = 0$$

$$v_1 = 0$$

$$v_1 = 0$$

$$v_2 = 0$$

$$v_3 = 0$$

$$v_4 = 0$$

$$v_1 = 0$$

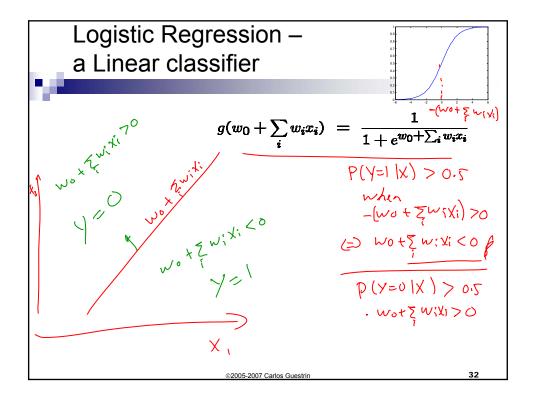
$$v_1 = 0$$

$$v_2 = 0$$

$$v_3 = 0$$

$$v_4 = 0$$

$$v_4$$



Very convenient!

$$P(Y = 1 | X = \langle X_1, ... X_n \rangle) = \frac{1}{1 + exp(w_0 + \sum_i w_i X_i)}$$

implies

$$P(Y = 0|X = < X_1, ... X_n >) = \frac{exp(w_0 + \sum_i w_i X_i)}{1 + exp(w_0 + \sum_i w_i X_i)}$$

implies

$$\frac{P(Y=0|X)}{P(Y=1|X)} = exp(w_0 + \sum_i w_i X_i)$$

/ linear classification rule!

implies

$$\ln \frac{P(Y=0|X)}{P(Y=1|X)} = w_0 + \sum_i w_i X_i$$

©2005-2007 Carlos Guestrin

33

Logistic regression v. Naïve Bayes



- Consider learning f: $X \rightarrow Y$, where
 - □ X is a vector of real-valued features, < X1 ... Xn >
 - ☐ Y is boolean
- Could use a Gaussian Naïve Bayes classifier
 - $\hfill \square$ assume all X_i are conditionally independent given Y
 - □ model $P(X_i | Y = y_k)$ as Gaussian $N(\mu_{ik}, \sigma_i)$
 - □ model P(Y) as Bernoulli(θ ,1- θ)
- What does that imply about the form of P(Y|X)?

©2005-2007 Carlos Guestrin

Logistic regression v. Naïve Bayes



- Consider learning f: X → Y, where
 - □ X is a vector of real-valued features, < X1 ... Xn >
 - ☐ Y is boolean
- Could use a Gaussian Naïve Bayes classifier
 - □ assume all X_i are conditionally independent given Y
 - \square model P(X_i | Y = y_k) as Gaussian N(μ_{ik} , σ_i)
 - \square model P(Y) as Bernoulli(θ ,1- θ)
- What does that imply about the form of P(Y|X)?

$$P(Y = 1|X = < X_1, ...X_n >) = \frac{1}{1 + exp(w_0 + \sum_i w_i X_i)}$$

Cool!!!!

©2005-2007 Carlos Guestrin

35

Derive form for P(Y|X) for continuous X_i



$$\begin{split} P(Y=1|X) &= \frac{P(Y=1)P(X|Y=1)}{P(Y=1)P(X|Y=1) + P(Y=0)P(X|Y=0)} \\ &= \frac{1}{1 + \frac{P(Y=0)P(X|Y=0)}{P(Y=1)P(X|Y=1)}} \\ &= \frac{1}{1 + \exp(\ln\frac{P(Y=0)P(X|Y=0)}{P(Y=1)P(X|Y=1)})} \\ &= \frac{1}{1 + \exp(\ln\frac{1-\theta}{\theta}) + \left[\sum_{i} \ln\frac{P(X_{i}|Y=0)}{P(X_{i}|Y=1)}\right)} \end{split}$$

©2005-2007 Carlos Guestrin

Ratio of class-conditional probabilities



$$\ln \frac{P(X_i|Y=0)}{P(X_i|Y=1)}$$

$$P(X_i = x \mid Y = y_k) = \frac{1}{\sigma_i \sqrt{2\pi}} e^{\frac{-(x - \mu_{ik})^2}{2\sigma_i^2}}$$

©2005-2007 Carlos Guestrin

3

Derive form for P(Y|X) for continuous X_i

$$P(Y = 1|X) = \frac{P(Y = 1)P(X|Y = 1)}{P(Y = 1)P(X|Y = 1) + P(Y = 0)P(X|Y = 0)}$$

$$= \frac{1}{1 + \exp((\ln \frac{1-\theta}{\theta}) + \sum_{i} \ln \frac{P(X_{i}|Y=0)}{P(X_{i}|Y=1)})}$$

$$\sum_{i} \left(\frac{\mu_{i0} - \mu_{i1}}{\sigma_{i}^{2}} X_{i} + \frac{\mu_{i1}^{2} - \mu_{i0}^{2}}{2\sigma_{i}^{2}} \right)$$

$$P(Y = 1|X) = \frac{1}{1 + \exp(w_0 + \sum_{i=1}^{n} w_i X_i)}$$

©2005-2007 Carlos Guestrin

Gaussian Naïve Bayes v. Logistic Regression



Set of Gaussian
Naïve Bayes parameters
(feature variance
independent of class label)

Set of Logistic Regression parameters

- Representation equivalence
 - ☐ But only in a special case!!! (GNB with class-independent variances)
- But what's the difference???
- LR makes no assumptions about P(X|Y) in learning!!!
- Loss function!!!
 - $\hfill\Box$ Optimize different functions \to Obtain different solutions

©2005-2007 Carlos Guestrin

39

Logistic regression for more than 2 classes



■ Logistic regression in more general case, where $Y \in \{Y_1 ... Y_R\}$: learn R-I sets of weights

©2005-2007 Carlos Guestrin

Logistic regression more generally



■ Logistic regression in more general case, where $Y \in \{Y_1 ... Y_R\}$: learn R-I sets of weights

for k < R

$$P(Y = y_k | X) = \frac{\exp(w_{k0} + \sum_{i=1}^{n} w_{ki} X_i)}{1 + \sum_{j=1}^{R-1} \exp(w_{j0} + \sum_{i=1}^{n} w_{ji} X_i)}$$

for k=R (normalization, so no weights for this class)

$$P(Y = y_R | X) = \frac{1}{1 + \sum_{j=1}^{R-1} \exp(w_{j0} + \sum_{i=1}^{n} w_{ji} X_i)}$$

Features can be discrete or continuous!

©2005-2007 Carlos Guestrin

4 1

Loss functions: Likelihood v. Conditional Likelihood



Generative (Naïve Bayes) Loss function:

Data likelihood

$$\begin{aligned} \ln P(\mathcal{D} \mid \mathbf{w}) &= \sum_{j=1}^{N} \ln P(\mathbf{x}^{j}, y^{j} \mid \mathbf{w}) \\ &= \sum_{j=1}^{N} \ln P(y^{j} \mid \mathbf{x}^{j}, \mathbf{w}) + \sum_{j=1}^{N} \ln P(\mathbf{x}^{j} \mid \mathbf{w}) \end{aligned}$$

- Discriminative models cannot compute P(xi|w)!
- But, discriminative (logistic regression) loss function:

Conditional Data Likelihood

$$\ln P(\mathcal{D}_Y \mid \mathcal{D}_X, \mathbf{w}) = \sum_{i=1}^{N} \ln P(y^j \mid \mathbf{x}^j, \mathbf{w})$$

□ Doesn't waste effort learning P(X) – focuses on P(Y|X) all that matters for classification

©2005-2007 Carlos Guestrin

Expressing Conditional Log Likelihood



$$l(\mathbf{w}) = \sum_{j} y^{j} \ln P(y^{j} = 1 | \mathbf{x}^{j}, \mathbf{w}) + (1 - y^{j}) \ln P(y^{j} = 0 | \mathbf{x}^{j}, \mathbf{w})$$

©2005-2007 Carlos Guestrii

43

Maximizing Conditional Log Likelihood

$$l(\mathbf{w}) \equiv \ln \prod_{j} P(y^{j} | \mathbf{x}^{j}, \mathbf{w})$$

$$= \sum_{i} y^{j} (w_{0} + \sum_{i} w_{i}x_{i}^{j}) - \ln(1 + exp(w_{0} + \sum_{i} w_{i}x_{i}^{j}))$$

Good news: $l(\mathbf{w})$ is concave function of $\mathbf{w} \to \mathsf{no}$ locally optimal solutions

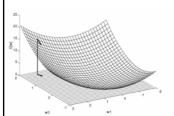
Bad news: no closed-form solution to maximize $l(\mathbf{w})$

Good news: concave functions easy to optimize

©2005-2007 Carlos Guestrin

Optimizing concave function – Gradient ascent





Gradient:
$$\nabla_{\mathbf{w}} l(\mathbf{w}) = [\frac{\partial l(\mathbf{w})}{\partial w_0}, \dots, \frac{\partial l(\mathbf{w})}{\partial w_n}]'$$

Update rule:
$$\Delta \mathbf{w} = \mathring{\eta } \nabla_{\mathbf{w}} l(\mathbf{w})$$

$$w_i \leftarrow w_i + \eta \frac{\partial l(\mathbf{w})}{\partial w_i}$$

Gradient ascent is simplest of optimization approaches
 e.g., Conjugate gradient ascent much better (see reading)

@2005-2007 Carlos Guestrin

45

Maximize Conditional Log Likelihood: Gradient ascent



$$l(\mathbf{w}) = \sum_{j} y^{j} (w_{0} + \sum_{i}^{n} w_{i} x_{i}^{j}) - \ln(1 + exp(w_{0} + \sum_{i}^{n} w_{i} x_{i}^{j}))$$

Gradient ascent algorithm: iterate until change < ε

For all
$$i$$
, $w_i \leftarrow w_i + \eta \sum_j x_i^j [y^j - \hat{P}(Y^j = 1 \mid \mathbf{x}^j, \mathbf{w})]$

repeat

©2005-2007 Carlos Guestrin

That's all M(C)LE. How about MAP?



- One common approach is to define priors on w
 - □ Normal distribution, zero mean, identity covariance
 - □ "Pushes" parameters towards zero
- Corresponds to Regularization
 - □ Helps avoid very large weights and overfitting
 - □ Explore this in your homework
 - □ More on this later in the semester
- MAP estimate

$$\mathbf{w}^* = \arg\max_{\mathbf{w}} \ln \left[p(\mathbf{w}) \prod_{j=1}^{N} P(y^j \mid \mathbf{x}^j, \mathbf{w}) \right]$$

©2005-2007 Carlos Guestrin

47

Gradient of M(C)AP



$$rac{\partial}{\partial w_i}$$
 In $\left[p(\mathbf{w})\prod_{j=1}^N P(y^j\mid \mathbf{x}^j,\mathbf{w})
ight]$

$$p(\mathbf{w}) = \prod_{i} \frac{1}{\kappa \sqrt{2\pi}} e^{\frac{-w_i^2}{2\kappa^2}}$$

©2005-2007 Carlos Guestrin

MLE vs MAP



Maximum conditional likelihood estimate

$$\begin{split} \mathbf{w}^* &= \arg \max_{\mathbf{w}} \ln \left[\prod_{j=1}^N P(y^j \mid \mathbf{x}^j, \mathbf{w}) \right] \\ w_i &\leftarrow w_i + \eta \sum_j x_i^j [y^j - \hat{P}(Y^j = 1 \mid \mathbf{x}^j, \mathbf{w})] \end{split}$$

Maximum conditional a posteriori estimate

$$\mathbf{w}^* = \arg\max_{\mathbf{w}} \ln \left[p(\mathbf{w}) \prod_{j=1}^N P(y^j \mid \mathbf{x}^j, \mathbf{w}) \right]$$

$$w_i \leftarrow w_i + \eta \left\{ -\lambda w_i + \sum_j x_i^j [y^j - \hat{P}(Y^j = 1 \mid \mathbf{x}^j, \mathbf{w})] \right\}$$

©2005-2007 Carlos Guestrin

49

What you should know about Logistic Regression (LR)



- Gaussian Naïve Bayes with class-independent variances representationally equivalent to LR
 - □ Solution differs because of objective (loss) function
- In general, NB and LR make different assumptions
 - \square NB: Features independent given class \rightarrow assumption on P(X|Y)
 - \square LR: Functional form of P(Y|X), no assumption on P(X|Y)
- LR is a linear classifier
 - □ decision rule is a hyperplane
- LR optimized by conditional likelihood
 - □ no closed-form solution
 - □ concave → global optimum with gradient ascent
 - □ Maximum conditional a posteriori corresponds to regularization

©2005-2007 Carlos Guestrin

Acknowledgements



 Some of the material is the presentation is courtesy of Tom Mitchell

©2005-2007 Carlos Guestrin