

### **Generative vs. Discriminative Classifiers**



- Goal: Wish to learn f:  $X \rightarrow Y$ , e.g., P(Y|X)
- Generative classifiers (e.g., Naïve Bayes):
  - Assume some functional form for P(X|Y), P(Y)
     This is a 'generative' model of the data!



- Estimate parameters of P(X|Y), P(Y) directly from training data
- Use Bayes rule to calculate P(Y|X=x)
- Discriminative classifiers (e.g., logistic regression)
  - Directly assume some functional form for P(Y|X)
    This is a 'discriminative' model of the data!
  - Estimate parameters of P(Y|X) directly from training data

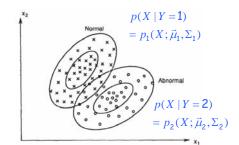


© Eric Xing @ CMU, 2006-2009

Suppose you know the following

. . .

• Class-specific Dist.: P(X|Y)



Bayes classifier:

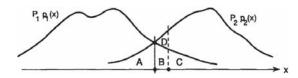
- Class prior (i.e., "weight"): P(Y)
- $P(Y|X) = \frac{P(X|Y)P(Y)}{P(X)}$
- This is a generative model of the data!

### **Optimal classification**



- Theorem: Bayes classifier is optimal!
  - That is

$$error_{true}(h_{Bayes})) \le error_{true}(h), \ \forall h(\mathbf{x})$$



- How to learn a Bayes classifier?
  - Recall density estimation. We need to estimate P(X|y=k), and P(y=k) for all k

© Eric Xing @ CMU, 2006-2009

5

#### **Gaussian Discriminative Analysis**



- learning f:  $X \rightarrow Y$ , where
  - X is a vector of real-valued features,  $\mathbf{X}_n = \langle X_{n,1}...X_{n,m} \rangle$
  - Y is an indicator vector
- What does that imply about the form of P(Y|X)?



• The joint probability of a datum and its label is:

$$p(\mathbf{x}_{n}, y_{n}^{k} = 1 \mid \mu, \sigma) = p(y_{n}^{k} = 1) \times p(\mathbf{x}_{n} \mid y_{n}^{k} = 1, \mu, \sigma)$$
$$= \pi_{k} \frac{1}{(2\pi\sigma^{2})^{1/2}} \exp\left\{\frac{1}{2\sigma^{2}} (\mathbf{x}_{n} - \mu_{k})^{2}\right\}$$

• Given a datum  $\mathbf{x}_n$ , we predict its label using the conditional probability of the label given the datum:

$$p(y_n^k = 1 | \mathbf{x}_n, \mu, \sigma) = \frac{\pi_k \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2} (\mathbf{x}_n - \mu_k)^2\right\}}{\sum_{k'} \pi_{k'} \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2} (\mathbf{x}_n - \mu_k)^2\right\}}$$

### **Conditional Independence**



 X is conditionally independent of Y given Z, if the probability distribution governing X is independent of the value of Y, given the value of Z

$$(\forall i, j, k) P(X = i | Y = j, Z = k) = P(X = i | Z = k)$$

Which we often write

$$P(X \mid Y, Z) = P(X \mid Z)$$

e.g.,

P(Thunder|Rain, Lightning) = P(Thunder|Lightning)

• Equivalent to:

$$P(X, Y \mid Z) = P(X \mid Z)P(Y \mid Z)$$

© Eric Xing @ CMU, 2006-2009

7

### Naïve Bayes Classifier



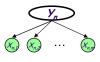
- When X is multivariate-Gaussian vector:
  - The joint probability of a datum and it label is:

$$p(\mathbf{x}_n, y_n^k = 1 \mid \vec{\mu}, \Sigma) = p(y_n^k = 1) \times p(\mathbf{x}_n \mid y_n^k = 1, \vec{\mu}, \Sigma)$$
$$= \pi_k \frac{1}{(2\pi |\Sigma|)^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x}_n - \vec{\mu}_k)^T \Sigma^{-1} (\mathbf{x}_n - \vec{\mu}_k)\right\}$$



• The naïve Bayes simplification

$$\begin{split} p(\mathbf{x}_{n}, y_{n}^{k} = 1 \mid \mu, \sigma) &= p(y_{n}^{k} = 1) \times \prod_{j} p(x_{n,j} \mid y_{n}^{k} = 1, \mu_{k,j}, \sigma_{k,j}) \\ &= \pi_{k} \prod_{j} \frac{1}{(2\pi\sigma_{k,j}^{2})^{1/2}} \exp\left\{ \frac{1}{2\sigma_{k,j}^{2}} (x_{n,j} - \mu_{k,j})^{2} \right\} \end{split}$$



- More generally:  $p(\mathbf{x}_n, y_n \mid \eta, \pi) = p(y_n \mid \pi) \times \prod_{i=1}^m p(x_{n,j} \mid y_n, \eta)$ 
  - Where p(. | .) is an arbitrary conditional (discrete or continuous) 1-D density

### The predictive distribution



• Understanding the predictive distribution

$$p(y_n^k = \mathbf{1} | x_n, \bar{\mu}, \Sigma, \pi) = \frac{p(y_n^k = \mathbf{1}, x_n | \bar{\mu}, \Sigma, \pi)}{p(x_n | \bar{\mu}, \Sigma)} = \frac{\pi_k N(x_n, | \mu_k, \Sigma_k)}{\sum_{k'} \pi_k N(x_n, | \mu_{k'}, \Sigma_{k'})} *$$

• Under naïve Bayes assumption:

$$p(y_n^k = 1 \mid x_n, \bar{\mu}, \Sigma, \pi) = \frac{\pi_k \exp\left\{-\sum_j \left(\frac{1}{2\sigma_{k,j}^2} (x_n^j - \mu_k^j)^2 - \log \sigma_{k,j} - C\right)\right\}}{\sum_{k'} \pi_{k'} \exp\left\{-\sum_j \left(\frac{1}{2\sigma_{k',j}^2} (x_n^j - \mu_k^j)^2 - \log \sigma_{k',j} - C\right)\right\}} **$$

• For two class (i.e., *K*=2), and when the two classes haves the same variance, \*\* turns out to be a logistic function

$$p(y_n^1 = 1 \mid x_n) = \frac{1}{1 + \frac{z_1 \exp\left[-\sum_{i} \frac{1}{z_p^2} (z_i^j \cdot \mu_i^j)^2 - \log \sigma_i - C\right]}{z_1 \exp\left[-\sum_{i} \frac{1}{z_p^2} (z_i^j \cdot \mu_i^j)^2 - \log \sigma_i - C\right]}} = \frac{1}{1 + \exp\left[-\sum_{j} \left(x_n^j \frac{1}{\sigma_j^2} (\mu_1^j - \mu_2^j) + \frac{1}{\sigma_j^2} ([\mu_1^j]^2 - [\mu_2^j]^2)\right) + \log \frac{(1-z_1)}{z_1}\right]}$$

© Eric Xing @ CMU, 2006-2009

a

#### The decision boundary



• The predictive distribution

$$p(y_n^1 = 1 \mid x_n) = \frac{1}{1 + \exp\left\{-\sum_{i=1}^M \theta_i x_n^j - \theta_0\right\}} = \frac{1}{1 + e^{-\theta^T x_n}}$$

• The Bayes decision rule:

$$\ln \frac{p(y_n^1 = 1 \mid x_n)}{p(y_n^2 = 1 \mid x_n)} = \ln \left( \frac{\frac{1}{1 + e^{-\theta^T x_n}}}{\frac{e^{-\theta^T x_n}}{1 + e^{-\theta^T x_n}}} \right) = \theta^T x_n$$



• For multiple class (i.e., K>2), \* correspond to a softmax function

$$p(y_n^k = 1 | x_n) = \frac{e^{-\theta_k^T x_n}}{\sum_{i} e^{-\theta_j^T x_n}}$$

## **Generative vs. Discriminative Classifiers**



- Goal: Wish to learn f:  $X \rightarrow Y$ , e.g., P(Y|X)
- Generative classifiers (e.g., Naïve Bayes):
  - Assume some functional form for P(X|Y), P(Y)
     This is a 'generative' model of the data!



- Estimate parameters of P(X|Y), P(Y) directly from training data
- Use Bayes rule to calculate P(Y|X= x)
- Discriminative classifiers:
  - Directly assume some functional form for P(Y|X)
     This is a 'discriminative' model of the data!
  - Estimate parameters of P(Y|X) directly from training data



© Eric Xing @ CMU, 2006-2009

11

### **Linear Regression**

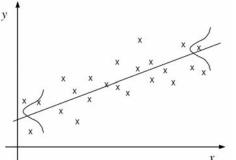


• The data:

$$\{(x_1, y_1), (x_2, y_2), (x_3, y_3), \dots, (x_N, y_N)\}$$



- Both nodes are observed:
  - X is an input vector
  - Y is a response vector
     (we first consider y as a generic continuous response vector, then we consider the special case of classification where y is a discrete indicator)
- A regression scheme can be used to model p(y|x) directly, rather than p(x,y)



© Eric Xing & CIVIU, 2000-2009

### **Linear Regression**



- Assume that Y (target) is a linear function of X (features):
  - e.g.

$$\hat{y} = \theta_0 + \theta_1 x_1 + \theta_2 x_2$$

 let's assume a vacuous "feature" X<sup>0</sup>=1 (this is the intercept term, why?), and define the feature vector to be:

$$\mathbf{x} = [1, x_1, x_2]$$

• then we have the following general representation of the linear function:

$$\hat{y} = \mathbf{x}^T \theta$$

- Our goal is to pick the optimal  $\theta$  . How!
  - We seek  $\theta$  that minimize the following **cost function**:

$$J(\theta) = \frac{1}{2} \sum_{i=1}^{n} (\hat{y}_{i}(\vec{x}_{i}) - y_{i})^{2}$$

© Eric Xing @ CMU, 2006-2009

. .

## The Least-Mean-Square (LMS) method



• Consider a gradient descent algorithm:

$$\theta_j^{t+1} = \theta_j^t - \alpha \frac{\partial}{\partial \theta_j} J(\theta)$$

• Now we have the following descent rule:

$$\theta_j^{t+1} = \theta_j^t + \alpha \sum_{i=1}^n (y_i - \bar{\mathbf{x}}_i^T \theta^t) x_i^j$$

• For a single training point, we have:

$$\theta_i^{t+1} = \theta_i^t + \alpha (y_i - \bar{\mathbf{x}}_i^T \theta^t) x_i^j$$

- This is known as the LMS update rule, or the Widrow-Hoff learning rule
- This is actually a "stochastic", "coordinate" descent algorithm
- This can be used as a on-line algorithm

© Eric Xing @ CMU, 2006-2009

## **Probabilistic Interpretation of LMS**



 Let us assume that the target variable and the inputs are related by the equation:

$$y_i = \boldsymbol{\theta}^T \mathbf{x}_i + \boldsymbol{\varepsilon}_i$$

where  $\pmb{\varepsilon}$  is an error term of unmodeled effects or random noise

Now assume that ε follows a Gaussian N(0,σ), then we have:

$$p(y_i \mid x_i; \theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y_i - \theta^T \mathbf{x}_i)^2}{2\sigma^2}\right)$$

• By independence assumption:

$$L(\theta) = \prod_{i=1}^{n} p(y_i \mid x_i; \theta) = \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left(-\frac{\sum_{i=1}^{n} (y_i - \theta^T \mathbf{x}_i)^2}{2\sigma^2}\right)$$

© Eric Xing @ CMU, 2006-2009

15

# Probabilistic Interpretation of LMS, cont.



• Hence the log-likelihood is:

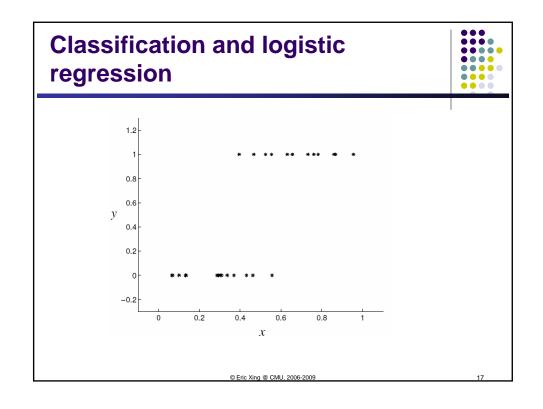
$$l(\theta) = n \log \frac{1}{\sqrt{2\pi\sigma}} - \frac{1}{\sigma^2} \frac{1}{2} \sum_{i=1}^{n} (y_i - \theta^T \mathbf{x}_i)^2$$

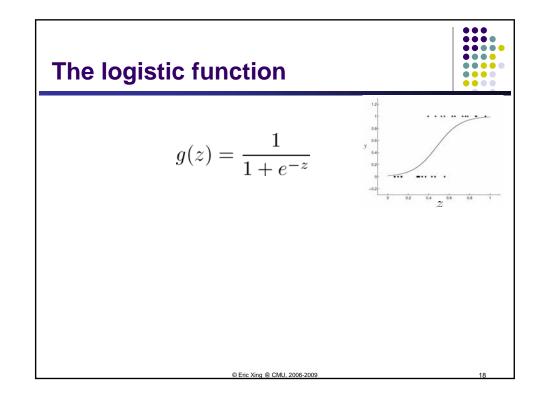
• Do you recognize the last term?

Yes it is: 
$$J(\theta) = \frac{1}{2} \sum_{i=1}^{n} (\mathbf{x}_i^T \theta - y_i)^2$$

 Thus under independence assumption, LMS is equivalent to MLE of θ!

© Eric Xing @ CMU, 2006-2009





# Logistic regression (sigmoid classifier)

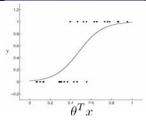


• The condition distribution: a Bernoulli

$$p(y | x) = \mu(x)^{y} (1 - \mu(x))^{1-y}$$

where  $\mu$  is a logistic function

function 
$$\mu(x) = \frac{1}{1 + e^{-\theta^T x}}$$



- We can used the brute-force gradient method as in LR
- But we can also apply generic laws by observing the p(y|x) is an exponential family function, more specifically, a generalized linear model (see future lectures ...)

© Eric Xing @ CMU, 2006-2009

10

### **Training Logistic Regression: MCLE**



- Estimate parameters  $\theta = <\theta_0, \ \theta_1, \dots \ \theta_m>$  to maximize the **conditional likelihood** of training data
- Training data  $\mathcal{D} = \{(x_1, y_1), \dots, (x_N, y_N)\}$
- Data likelihood =  $\prod_{i=1}^{N} P(x_i, y_i; \theta)$
- $\bullet \quad \text{Data conditional likelihood} = \prod_{i=1}^N P(x_i|y_i;\theta)$

$$\theta = \arg\max_{\theta} \ln \prod_{i} P(y_i|x_i;\theta)$$

© Eric Xing @ CMU, 2006-200

### **Expressing Conditional Log Likelihood**



$$l(\theta) \equiv \ln \prod_{i} P(y_i|x_i;\theta) = \sum_{i} \ln P(y_i|x_i;\theta)$$

• Recall the logistic function:  $\mu = \frac{1}{1 + e^{-\theta^T x}}$ 

and conditional likelihood:  $P(y|x) = \mu(x)^y (1 - \mu(x))^{1-y}$ 

$$\begin{split} l(\theta) &= \sum_{i} \ln P(y_{i}|x_{i};\theta) &= \sum_{i} y_{i} \ln u(x_{i}) + (1-y_{i}) \ln(1-\mu(x_{i})) \\ &= \sum_{i} y_{i} \ln \frac{u(x_{i})}{1-\mu(x_{i})} + \ln(1-\mu(x_{i})) \\ &= \sum_{i} y_{i} \theta^{T} x_{i} - \theta^{T} x_{i} + \ln(1+e^{-\theta^{T} x_{i}}) \\ &= \sum_{i} (y_{i}-1) \theta^{T} x_{i} + \ln(1+e^{-\theta^{T} x_{i}}) \end{split}$$

21

## **Maximizing Conditional Log Likelihood**



• The objective:

$$l(\theta) = \ln \prod_{i} P(y_i|x_i;\theta)$$
$$= \sum_{i} (y_i - 1)\theta^t x_i + \ln(1 + e^{-\theta^T x_i})$$

- Good news:  $I(\theta)$  is concave function of  $\theta$
- Bad news: no closed-form solution to maximize I(θ)

© Eric Xing @ CMU, 2006-2009

#### The Newton's method



• Finding a zero of a function

$$\theta^{t+1} := \theta^t - \frac{f(\theta^t)}{f'(\theta^t)}$$

© Eric Xing @ CMU, 2006-2009

23

### The Newton's method (con'd)



• To maximize the conditional likelihood  $l(\theta)$ :

$$l(\theta) = \sum_{i} (y_i - 1)\theta^T x_i + \ln(1 + e^{-\theta^T x_i})$$

since l is convex, we need to find  $\theta^*$  where  $l'(\theta^*)=0$ !

• So we can perform the following iteration:

$$\theta^{t+1} := \theta^t + \frac{l'(\theta^t)}{l''(\theta^t)}$$

© Eric Xing @ CMU, 2006-2009

### The Newton-Raphson method



• In LR the  $\theta$  is vector-valued, thus we need the following generalization:

$$\theta^{t+1} := \theta^t + H^{-1} \nabla_{\theta^t} l(\theta^t)$$

- $\nabla$  is the gradient operator over the function
- H is known as the Hessian of the function

© Eric Xina @ CMU, 2006-2009

25

#### The Newton-Raphson method



• In LR the  $\theta$  is vector-valued, thus we need the following generalization:

$$\theta^{t+1} := \theta^t + H^{-1} \nabla_{\theta^t} l(\theta^t)$$

•  $\nabla$  is the gradient operator over the function

$$\nabla_{\theta} l(\theta) = \sum_{i} (y_i - u_i) x_i = \mathbf{X}^T (\mathbf{y} - \mathbf{u})$$

• H is known as the Hessian of the function

$$H = \nabla_{\theta} \nabla_{\theta} l(\theta) = \sum_{i} u_{i} (1 - u_{i}) x_{i} x_{i}^{T} = \mathbf{X}^{T} \mathbf{R} \mathbf{X}$$
 where  $R_{ii} = u_{i} (1 - u_{i})$ 

© Eric Xing @ CMU, 2006-2009

# **Iterative reweighed least squares** (IRLS)



• Recall in the least square est. in linear regression, we have:

$$\theta = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

which can also derived from Newton-Raphson

• Now for logistic regression:

$$\theta^{t+1} = \theta^t + H^{-1} \nabla_{\theta^t} l(\theta^t)$$

$$= \theta^t - (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{u} - \mathbf{y})$$

$$= (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \{ \mathbf{X}^T \mathbf{R} \mathbf{X} \theta^t - \mathbf{X}^T (\mathbf{u} - \mathbf{y}) \}$$

$$= (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \mathbf{X}^T \mathbf{R} \mathbf{z}$$

© Eric Xing @ CMU, 2006-2009

27

## Logistic regression: practical issues



- NR (IRLS) takes  $O(N+d^3)$  per iteration, where N= number of training cases and d= dimension of input x, but converge in fewer iterations
- Quasi-Newton methods, that approximate the Hessian, work faster.
- Conjugate gradient takes O(Nd) per iteration, and usually works best in practice.
- Stochastic gradient descent can also be used if N is large c.f. perceptron rule:

© Eric Xing @ CMU, 2006-2009

### **Generative vs. Discriminative Classifiers**



- Goal: Wish to learn f:  $X \rightarrow Y$ , e.g., P(Y|X)
- Generative classifiers (e.g., Naïve Bayes):
  - Assume some functional form for P(X|Y), P(Y)
     This is a 'generative' model of the data!



- Estimate parameters of P(X|Y), P(Y) directly from training data
- Use Bayes rule to calculate P(Y|X=x)
- Discriminative classifiers:
  - Directly assume some functional form for P(Y|X)
     This is a 'discriminative' model of the data!



• Estimate parameters of P(Y|X) directly from training data

© Eric Xing @ CMU, 2006-2009

20

# Naïve Bayes vs Logistic Regression



- Consider Y boolean, X continuous, X=<X1 ... Xm>
- Number of parameters to estimate:

NB: 
$$p(y | \mathbf{x}) = \frac{\pi_k \exp\left\{-\sum_{j} \left(\frac{1}{2\sigma_{k,j}^2} (x_j - \mu_{k,j})^2 - \log \sigma_{k,j} - C\right)\right\}}{\sum_{k'} \pi_{k'} \exp\left\{-\sum_{j} \left(\frac{1}{2\sigma_{k',j}^2} (x_j - \mu_{k',j})^2 - \log \sigma_{k',j} - C\right)\right\}} **$$

LR: 
$$\mu(x) = \frac{1}{1 + e^{-\theta^T x}}$$

- Estimation method:
  - NB parameter estimates are uncoupled
  - LR parameter estimates are coupled

© Eric Xing @ CMU, 2006-2009

# Naïve Bayes vs Logistic Regression



- Asymptotic comparison (# training examples → infinity)
- when model assumptions correct
  - NB, LR produce identical classifiers
- when model assumptions incorrect
  - LR is less biased does not assume conditional independence
  - therefore expected to outperform NB

© Eric Xing @ CMU, 2006-2009

21

# Naïve Bayes vs Logistic Regression



- Non-asymptotic analysis (see [Ng & Jordan, 2002])
- convergence rate of parameter estimates how many training examples needed to assure good estimates?

NB order log m (where m = # of attributes in X) LR order m

 NB converges more quickly to its (perhaps less helpful) asymptotic estimates

© Eric Xing @ CMU, 2006-2009

# Rate of convergence: logistic regression



• Let  $h_{Dis,m}$  be logistic regression trained on n examples in m dimensions. Then with high probability:

$$\epsilon(h_{Dis,n}) \le \epsilon(h_{Dis,\infty}) + O\left(\sqrt{\frac{m}{n}\log \frac{n}{m}}\right)$$

- Implication: if we want  $\epsilon(h_{Dis,m}) \leq \epsilon(h_{Dis,\infty}) + \epsilon_0$  for some small constant  $\varepsilon_0$ , it suffices to pick order m examples
  - $\rightarrow$  Convergences to its asymptotic classifier, in order m examples
  - result follows from Vapnik's structural risk bound, plus fact that the "VC Dimension" of an m-dimensional linear separators is m

© Eric Xing @ CMU, 2006-2009

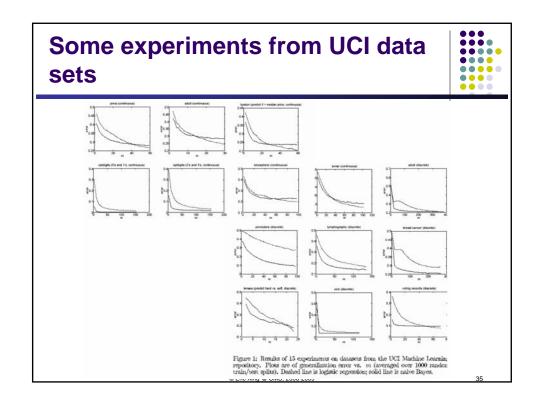
33

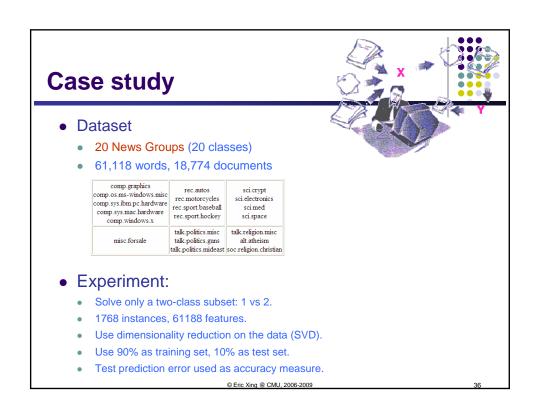
# Rate of convergence: naïve Bayes parameters

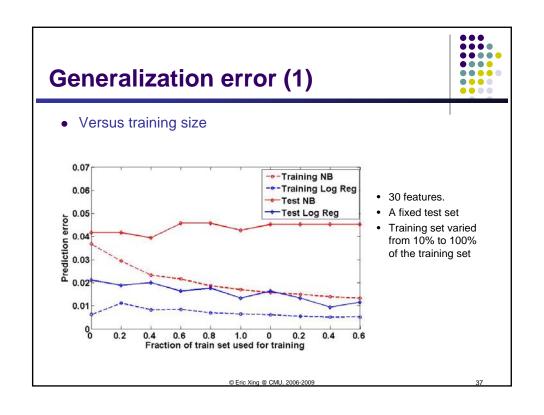


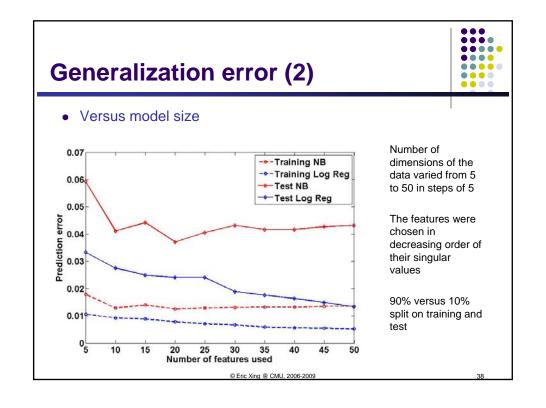
- Let any  $\varepsilon_1$ ,  $\delta$ >0, and any  $n \ge 0$  be fixed. Assume that for some fixed  $\rho_0$  > 0, we have that  $\rho_0 \le p(y=T) \le 1 - \rho_0$
- Let  $n = O((1/\epsilon_1^2)\log(m/\delta))$
- Then with probability at least 1- $\delta$ , after n examples:
  - 1. For discrete input,  $\begin{aligned} |\hat{p}(x_i|y=b) p(x_i|y=b)| &\leq \epsilon_1 \\ |\hat{p}(y=b) p(y=b)| &\leq \epsilon_1 \end{aligned} \qquad \text{for all $i$ and $b$}$
  - 2. For continuous inputs,  $\begin{aligned} |\hat{\mu}_{i|y=b} \mu_{i|y=b}| &\leq \epsilon_1 \\ |\hat{\sigma}_{i|y=b}^2 \sigma_{i|y=b}^2| &\leq \epsilon_1 \end{aligned} \qquad \text{for all $i$ and $b$}$

© Eric Xing @ CMU, 2006-2009









### **Summary**



- Naïve Bayes classifier
  - What's the assumption
  - Why we use it
  - How do we learn it
- Logistic regression
  - Functional form follows from Naïve Bayes assumptions
  - For Gaussian Naïve Bayes assuming variance
  - For discrete-valued Naïve Bayes too
  - But training procedure picks parameters without the conditional independence assumption
- Gradient ascent/descent
  - - General approach when closed-form solutions unavailable
- · Generative vs. Discriminative classifiers
  - Bias vs. variance tradeoff

© Eric Xing @ CMU, 2006-2009