

# **Generative vs. Discriminative classifiers**



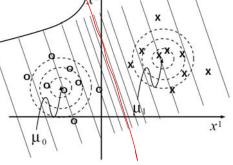
• Goal: Wish to learn f:  $X \rightarrow Y$ , e.g., P(Y|X)

Generative:

 Modeling the joint distribution of all data



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### **Learning Generative and Discriminative Classifiers**



- Goal: Wish to learn f:  $X \rightarrow Y$ , e.g., P(Y|X)
- Generative classifiers (e.g., Naïve Bayes):
  - Assume some functional form for P(X|Y), P(Y)
     This is a 'generative' model of the data!
  - Estimate parameters of P(X|Y) P(Y) directly from training data
  - Use Bayes rule to calculate P(Y|X= x)



- Discriminative classifiers (e.g., logistic regression)
  - Directly assume some functional form for P(Y|X)
     This is a 'discriminative' model of the data!
  - Estimate parameters of P(Y|X) directly from training data



h (x)

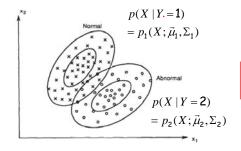
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### Suppose you know the following



. . .

Class-specific Dist.: P(X|Y)



າະໃ ກະໃ Bayes classifier:

$$P(Y|X) = \frac{P(X|Y)P(Y)}{P(X)}$$

$$P(Y=k|X=X) = \frac{P(X=X)P(Y=k)P(Y=k)}{\sum_{i=1}^{n} P(Y=k)P(Y=k)}$$

- Class prior (i.e., "weight"): P(Y)
- This is a generative model of the data!

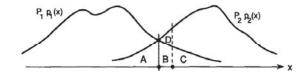
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### **Optimal classification**

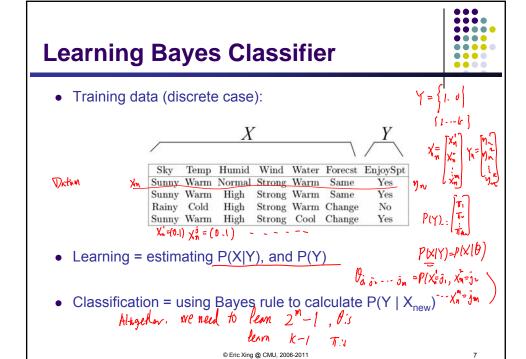


- Theorem: Bayes classifier is optimal!
  - $error_{true}(h_{Bayes})) \leq error_{true}(h), \ \forall h(\mathbf{x})$



- How to learn a Bayes classifier?
  - Recall density estimation. We need to estimate P(X|y=k), and P(y=k) for all k

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### Parameter learning from *iid* data: The Maximum Likelihood Est.



Goal: estimate distribution parameters θ from a dataset of N independent, identically distributed (iid), fully observed, training cases

$$D = \{x_1, \ldots, x_N\}$$

- Maximum likelihood estimation (MLE)
  - 1. One of the most common estimators
  - 2. With iid and full-observability assumption, write  $L(\theta)$  as the likelihood of the data:

$$L(\theta) = P(x_1, x_2, ..., x_N; \theta)$$

$$= P(x; \theta) P(x_2; \theta), ..., P(x_N; \theta)$$

$$= \prod_{n=1}^{N} P(x_n; \theta)$$

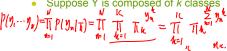
3. pick the setting of parameters most likely to have generated the data we saw:

$$\theta^* = \arg\max_{\alpha} L(\theta) = \arg\max_{\alpha} \log L(\theta)$$

#### How hard is it to learn the optimal classifier?



- How do we represent these? How many parameters?



Likelihood, P(X|Y):

Suppose X is composed of n binary features  $P(X_{n} \cdots X_{n} | y_{n}) = \prod_{n=1}^{n} P(X_{n} | y_{n} \cdot k) = \prod_{n=1}^{n} \prod_{\substack{j \in J \\ j \neq j \\ j \neq j}} \prod_{\substack{j \in J \\ j \neq j \\ j \neq j}} P(X_{n} \cdots X_{n} | y_{n} \cdot k) = \prod_{n=1}^{n} \prod_{\substack{j \in J \\ j \neq j \\ j \neq j}} \prod_{\substack{j \in J \\ j \neq j \\ j \neq j}} P(X_{n} \cdots X_{n} | y_{n} \cdot k) = \prod_{n=1}^{n} \prod_{\substack{j \in J \\ j \neq j \\ j \neq j}} P(X_{n} \cdots X_{n} | y_{n} \cdot k) = \prod_{n=1}^{n} \prod_{\substack{j \in J \\ j \neq j}} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1}^{n} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1}^{n} \prod_{\substack{j \in J \\ j \neq j}} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1}^{n} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1}^{n} \prod_{\substack{j \in J \\ j \neq j}} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1}^{n} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1}^{n} \prod_{\substack{j \in J \\ j \neq j}} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1}^{n} P(X_{n} \mid y_{n} \cdot k) = \prod_{n=1$ 

Complex model → High variance with limited data!!!

$$U(2^m \times k)$$

U(2m xk) when give a lift D.

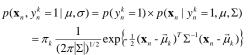
you will have a diff est.

### **Gaussian Discriminative Analysis**



- learning f:  $X \rightarrow Y$ , where
  - X is a vector of real-valued features,  $\mathbf{X}_n = \langle X_n^1, ... X_n^m \rangle$
  - Y is an indicator vector
- What does that imply about the form of P(Y|X)?
  - The joint probability of a datum and its label is:





Given a datum  $\mathbf{x}_{n}$ , we predict its label using the conditional probability of the label given the datum:

$$p(y_n^k = 1 | \mathbf{x}_n, \mu, \Sigma) = \frac{\pi_k \frac{1}{(2\pi|\Sigma|)^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x}_n - \vec{\mu}_k)^T \Sigma^{-1}(\mathbf{x}_n - \vec{\mu}_k)\right\}}{\sum_{k'} \pi_{k'} \frac{1}{(2\pi|\Sigma|)^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x}_n - \vec{\mu}_k)^T \Sigma^{-1}(\mathbf{x}_n - \vec{\mu}_{k'})\right\}}$$

### **Conditional Independence**



• X is **conditionally independent** of Y given Z, if the probability distribution governing X is independent of the value of Y, given the value of Z

$$(\forall i, j, k) P(X = i | Y = j, Z = k) = P(X = i | Z = k)$$

Which we often write

$$P(X \mid Y, Z) = P(X \mid Z)$$

e.g.,

P(Thunder|Rain, Lightning) = P(Thunder|Lightning)

• Equivalent to:

$$P(X, Y \mid Z) = P(X \mid Z)P(Y \mid Z)$$

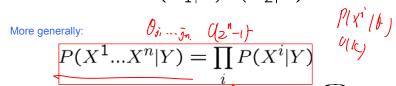
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#### The Naïve Bayes assumption



- Naïve Bayes assumption:
  - Features are conditionally independent given class:

$$P(X_1, X_2|Y) = P(X_1|X_2, Y)P(X_2|Y)$$
  
=  $P(X_1|Y)P(X_2|Y)$   $\rightarrow O(n)$ 



- How many parameters now?
  - Suppose X is composed of m binary features

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### The Naïve Bayes Classifier



- Given:
  - Prior P(Y)
  - m conditionally independent features X given the class Y
  - For each  $X_n$ , we have likelihood  $P(X_n|Y)$
- Decision rule:

$$\begin{aligned} \boldsymbol{y}^* &= h_{NB}(\mathbf{x}) &= \arg\max_{\boldsymbol{y}} P(\boldsymbol{y}) P(\boldsymbol{x}^1, \dots, \boldsymbol{x^m} | \ \boldsymbol{y}) \\ &= \arg\max_{\boldsymbol{y}} P(\boldsymbol{y}) \prod_{i} P(\boldsymbol{x}^i | \boldsymbol{y}) \end{aligned}$$

• If assumption holds, NB is optimal classifier!

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## The A Gaussian Discriminative Naïve Bayes Classifier



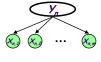
- When X is multivariate-Gaussian vector:
  - The joint probability of a datum and it label is:

$$p(\mathbf{x}_n, y_n^k = 1 | \vec{\mu}, \Sigma) = p(y_n^k = 1) \times p(\mathbf{x}_n | y_n^k = 1, \vec{\mu}, \Sigma)$$
$$= \pi_k \frac{1}{(2\pi |\Sigma|)^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x}_n - \vec{\mu}_k)^T \Sigma^{-1}(\mathbf{x}_n - \vec{\mu}_k)\right\}$$



• The naïve Bayes simplification

$$p(\mathbf{x}_n, \mathbf{y}_n^k = 1 \mid \mu, \sigma) = p(\mathbf{y}_n^k = 1) \times \prod_j p(\mathbf{x}_n^j \mid \mathbf{y}_n^k = 1, \mu_k^j, \sigma_k^j)$$
$$= \pi_k \prod_j \frac{1}{\sqrt{2\pi\sigma_k^j}} \exp\left\{ -\frac{1}{2} \left( \frac{\mathbf{x}_n^j - \mu_k^j}{\sigma_k^j} \right)^2 \right\}$$



- More generally:  $p(\mathbf{x}_n, y_n \mid \eta, \pi) = p(y_n \mid \pi) \times \prod_{j=1}^m p(x_n^j \mid y_n, \eta)$ 
  - Where p(. | .) is an arbitrary conditional (discrete or continuous) 1-D density

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#### The predictive distribution

• Understanding the predictive distribution

$$p(y_n^k = \mathbf{1} \mid x_n, \bar{\mu}, \Sigma, \pi) = \frac{p(y_n^k = \mathbf{1}, x_n \mid \bar{\mu}, \Sigma, \pi)}{p(x_n \mid \bar{\mu}, \Sigma)} = \frac{\pi_k N(x_n, |\mu_k, \Sigma_k)}{\sum_{k'} \pi_k N(x_n, |\mu_{k'}, \Sigma_{k'})} *$$

• Under naïve Bayes assumption:

$$p(y_n^k = 1 \mid x_n, \bar{\mu}, \Sigma, \pi) = \frac{\pi_k \exp\left\{-\sum_j \left(\frac{1}{2} \left(\frac{x_n^j - \mu_k^j}{\sigma_k^j}\right)^2 - \log \sigma_k^j - C\right)\right\}}{\sum_{k} \pi_{k'} \exp\left\{-\sum_j \left(\frac{1}{2} \left(\frac{x_n^j - \mu_{k'}^j}{\sigma_{k'}^j}\right)^2 - \log \sigma_{k'}^j - C\right)\right\}} \quad **$$

• For two class (i.e., *K*=2), and when the two classes haves the same variance, \*\* turns out to be a logistic function

$$\begin{aligned} & p(y_n^1 = \mathbf{1} \mid x_n) &= \frac{1}{1 + \frac{\pi_i \exp\left[-\sum_{n=1}^{j} (x_n^j - \mu_n^j)^2 - \log \sigma_j - c\right]}{\pi_i \exp\left[-\sum_{n=1}^{j} (x_n^j - \mu_n^j)^2 - \log \sigma_j - c\right]}} = \frac{1}{1 + \exp\left\{-\sum_{n=1}^{j} (x_n^j - \mu_n^j)^2 + \frac{1}{\sigma_j^2} (\mu_n^j)^2 - [\mu_n^j]^2 - [\mu_n^j]^2\right\} + \log\frac{(1-\pi_i)^2}{\pi_i}}{\pi_i \exp\left[-\sum_{n=1}^{j} (x_n^j - \mu_n^j)^2 - (\mu_n^j)^2 - [\mu_n^j]^2\right] + \log\frac{(1-\pi_i)^2}{\pi_i}}{\pi_i \exp\left[-\sum_{n=1}^{j} (x_n^j - \mu_n^j)^2 - [\mu_n^j]^2\right]}} \end{aligned}$$

$$= \frac{1}{1 + e^{-\theta^T x_n}}$$

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### The decision boundary

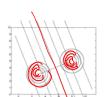


• The predictive distribution

$$p(y_n^1 = 1 \mid x_n) = \frac{1}{1 + \exp\left\{-\sum_{j=1}^M \theta_j x_n^j - \theta_0\right\}} = \frac{1}{1 + e^{-\theta^T x_n}}$$

• The Bayes decision rule:

$$\ln \frac{p(y_n^1 = 1 \mid x_n)}{p(y_n^2 = 1 \mid x_n)} = \ln \left( \frac{1}{1 + e^{-\theta^T x_n}} / \frac{e^{-\theta^T x_n}}{1 + e^{-\theta^T x_n}} \right) = \theta^T x_n = \emptyset$$



• For multiple class (i.e., K>2), \* correspond to a softmax function

$$p(y_n^k = 1 | x_n) = \frac{e^{-\theta_k^T x_n}}{\sum_i e^{-\theta_j^T x_n}}$$

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### **Summary:** The Naïve Bayes Algorithm



- Train Naïve Bayes (examples)
  - for each\* value y

- estimate  $\pi_k = P(Y = y_k)$
- for each\* value x<sub>ii</sub> of each attribute X<sub>i</sub>



$$\begin{split} Y^{new} \leftarrow \arg\max_{y_k} P(Y = y_k) \prod_i P(X^i = x_{ij} | Y = y_k) \\ Y^{new} \leftarrow \arg\max_{y_k} \pi_k \prod_i \theta_{ijk} \end{split}$$

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### **Generative vs. Discriminative Classifiers**



- Goal: Wish to learn f:  $X \rightarrow Y$ , e.g., P(Y|X)
- Generative classifiers (e.g., Naïve Bayes):
  - Assume some functional form for P(X|Y), P(Y) This is a 'generative' model of the data!



- Estimate parameters of P(X|Y), P(Y) directly from training data
- Use Bayes rule to calculate P(Y|X= x)



 Directly assume some functional form for P(Y|X) This is a 'discriminative' model of the data!



• Estimate parameters of P(Y|X) directly from training data

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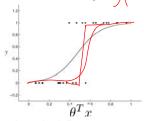
### Logistic regression (sigmoid classifier)



• The condition distribution: a Bernoulli

$$p(y | x) = \mu(x)^{y} (1 - \mu(x))^{1-y}$$

where 
$$\mu$$
 is a logistic function 
$$\mu(x) = \frac{1}{1 + e^{-\theta^T x}}$$



- We can used the brute-force gradient method as in LR
- But we can also apply generic laws by observing the p(y|x) is an exponential family function, more specifically, a generalized linear model (see future lectures ...)

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### **Training Logistic Regression: MCLE**



- Estimate parameters  $\theta = <\theta_0, \theta_1, \dots \theta_m>$  to maximize the conditional likelihood of training data
- Training data  $\mathcal{D} = \{(x_1, y_1), \dots, (x_N, y_N)\}$
- Data likelihood =  $\prod_{i=1}^{N} P(x_i, y_i; \theta)$
- Data conditional likelihood =  $\prod^{N} P(\mathbf{Y}_{i}|\mathbf{Y}_{i};\theta)$

$$\theta = \arg \max_{\theta} \ln \prod_{i} P(y_i|x_i;\theta)$$

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### **Expressing Conditional Log Likelihood**



$$l(\theta) \equiv \ln \prod_{i} P(y_i|x_i;\theta) = \sum_{i} \ln P(y_i|x_i;\theta)$$

• Recall the logistic function:  $\mu = \frac{1}{1 + e^{-\theta^T x}}$ 

and conditional likelihood:  $P(y|x) = \mu(x)^y (1 - \mu(x))^{1-y}$ 

$$\begin{split} l(\theta) &= \sum_{i} \ln P(y_{i}|x_{i};\theta) &= \sum_{i} y_{i} \ln u(x_{i}) + (1-y_{i}) \ln (1-\mu(x_{i})) \\ &= \sum_{i} y_{i} \ln \frac{u(x_{i})}{1-\mu(x_{i})} + \ln (1-\mu(x_{i})) \\ &= \sum_{i} y_{i} \theta^{T} x_{i} - \theta^{T} x_{i} + \ln (1+e^{-\theta^{T} x_{i}}) \\ &= \sum_{i} (y_{i} - 1) \theta^{T} x_{i} + \ln (1+e^{-\theta^{T} x_{i}}) \end{split}$$

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### **Maximizing Conditional Log Likelihood**



• The objective:

 $l(\theta) = \ln \prod_{i} P(y_i|x_i;\theta)$  $= \sum_{i} (y_i - 1)\theta^t x_i + \ln(1 + e^{-\theta^T x_i})$ 



- Good news:  $I(\theta)$  is concave function of  $\theta$
- 1/ 0)
- Bad news: no closed-form solution to maximize  $l(\theta)$

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#### **Gradient Ascent**



$$l(\theta) = \ln \prod_{i} P(y_{i}|x_{i}; \theta)$$

$$= \sum_{i} (y_{i} - 1)\theta^{T}x_{i} + \ln(1 + e^{-\theta^{T}x_{i}}) = \sum_{i} (y_{i} - 1)\theta^{T}x_{i} - \ln \mu(\theta^{T}x_{i})$$

• Property of sigmoid function:

$$\mu = \frac{1}{1+e^{-t}}$$

$$\frac{d\mu}{dt} = \mu(1 - \mu)$$

• The gradient:

$$\frac{\partial l(\theta)}{\partial \theta_{N}} = (\text{N-I}) \ \, \text{$\chi_{1}$} - \frac{1}{\text{N}} \ \, \frac{1}{\text{N}} \ \, = (\text{N-I}) \, \text{$\chi_{n}$} - (1-\text{N}) \, \text{$\chi_{n}$}.$$

The gradient ascent algorithm iterate until change  $< \epsilon$ 

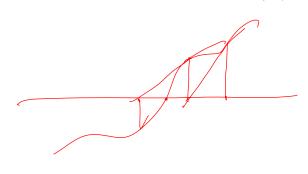
For all i, 
$$\theta_j \leftarrow \theta_j + \eta \sum_i (y_i - P(y_i = 0 | x_i; \theta)) x_i^j$$
 repeat

#### The Newton's method



• Finding a zero of a function

$$\theta^{t+1} := \theta^t - \underbrace{f(\theta^t)}_{f'(\theta^t)}$$



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### The Newton's method (con'd)



• To maximize the conditional likelihood  $l(\theta)$ :

$$l(\theta) = \sum_{i} (y_i - 1)\theta^T x_i + \ln(1 + e^{-\theta^T x_i})$$

since l is convex, we need to find  $\theta^*$  where  $l'(\theta^*)=0$ !

• So we can perform the following iteration:

$$\theta^{t+1} := \theta^t + \frac{l'(\theta^t)}{l''(\theta^t)}$$

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#### The Newton-Raphson method



• In LR the  $\theta$  is vector-valued, thus we need the following generalization:

$$\theta^{t+1} := \theta^t + H^{-1} \nabla_{\theta^t} l(\theta^t)$$

- ullet  $\nabla$  is the gradient operator over the function
- H is known as the Hessian of the function

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### The Newton-Raphson method



• In LR the  $\theta$  is vector-valued, thus we need the following generalization:

$$\theta^{t+1} := \theta^t + H^{-1} \nabla_{\theta^t} l(\theta^t)$$

•  $\nabla$  is the gradient operator over the function

$$\nabla_{\theta} l(\theta) = \sum_{i} (y_i - u_i) x_i = \mathbf{X}^T (\mathbf{y} - \mathbf{u})$$

• H is known as the Hessian of the function

$$H = \nabla_{\theta} \nabla_{\theta} l(\theta) = \sum_{i} u_{i} (1 - u_{i}) x_{i} x_{i}^{T} = \mathbf{X}^{T} \mathbf{R} \mathbf{X}$$
where  $R_{ii} = u_{i} (1 - u_{i})$ 

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### Iterative reweighed least squares (IRLS)



• Recall in the least square est. in linear regression, we have:

$$\theta = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

which can also derived from Newton-Raphson

Now for logistic regression:

$$\theta^{t+1} = \theta^t + H^{-1} \nabla_{\theta^t} l(\theta^t)$$

$$= \theta^t - (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{u} - \mathbf{y})$$

$$= (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \{ \mathbf{X}^T \mathbf{R} \mathbf{X} \theta^t - \mathbf{X}^T (\mathbf{u} - \mathbf{y}) \}$$

$$= (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \mathbf{X}^T \mathbf{R} \mathbf{z}$$

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#### **IRLS**



• Recall in the least square est. in linear regression, we have:

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which can also derived from Newton-Raphson

• Now for logistic regression:

$$\theta^{t+1} = (\mathbf{X}^T \mathbf{R} \mathbf{X})^{-1} \mathbf{X}^T \mathbf{R} \mathbf{z}$$

where 
$$\mathbf{z} = \mathbf{X}\theta^t - \mathbf{R}^{-1}(\mathbf{u} - \mathbf{y})$$

and 
$$R_{ii} = u_i(1 - u_i)$$

#### **Convergence curves** 0.008 0.006 0.004 comp.windows.x alt.atheism rec.autos VS. vs. vs. comp.graphics rec.sport.baseball rec.motorcycles Legend: - X-axis: Iteration #; Y-axis: error - In each figure, red for IRLS and blue for gradient descent @ Erie Xing @ CMU, 2006-2018

### Logistic regression: practical issues



- NR (IRLS) takes O(N+w<sup>3</sup>) per iteration, where N = number of training cases and d = dimension of input x, but converge in fewer iterations
- Quasi-Newton methods, that approximate the Hessian, work faster.
- Conjugate gradient takes O(Nd) per iteration, and usually works best in practice.
- Stochastic gradient descent can also be used if N is large c.f. perceptron rule:

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#### **Case Study: Text classification**



- Classify e-mails
  - Y = {Spam,NotSpam}
- Classify news articles
  - Y = {what is the topic of the article?}
- Classify webpages
  - Y = {Student, professor, project, ...}
- What about the features X?
  - The text!



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### Features X are entire document – $X^i$ for $i^{th}$ word in article





### Bag of words model



- Typical additional assumption Position in document doesn't matter: P(Xi=xi|Y=y) = P(Xk=xi|Y=y)
  - "Bag of words" model order of words on the page ignored
  - Sounds really silly, but often works very well!

$$P(y) \prod_{i=1}^{LengthDoc} P(x^i|y)$$
 or  $P(y) \prod_{k=1}^{LengthVol} P(w^k|y)$ 

When the lecture is over, remember to wake up the person sitting next to you in the lecture room.

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in is lecture lecture next over person remember room sitting the the to to up wake when you

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### **NB** with Bag of Words for text classification



- Learning phase:
  - Prior P(Y)
    - Count how many documents you have from each topic (+ prior)
  - P(X<sup>i</sup>|Y)
    - For each topic, count how many times you saw word in documents of this topic (+ prior)
- Test phase:
  - For each document x<sub>new</sub>
    - Use naïve Bayes decision rule

$$h_{NB}(\mathbf{x}_{\text{new}}) = \arg \max_{y} P(y) \prod_{i=1}^{LengthDoc} P(x_{\text{new}}^{i}|y)$$

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### Back to our 20 NG Case study

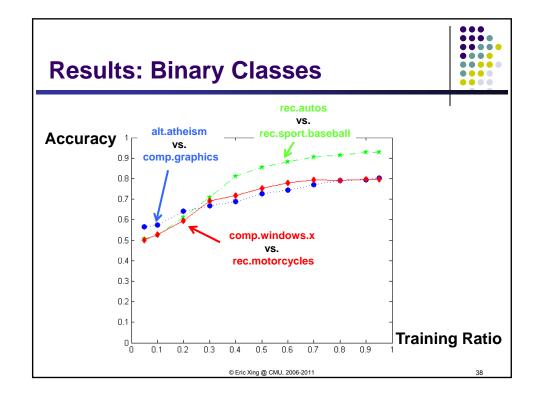


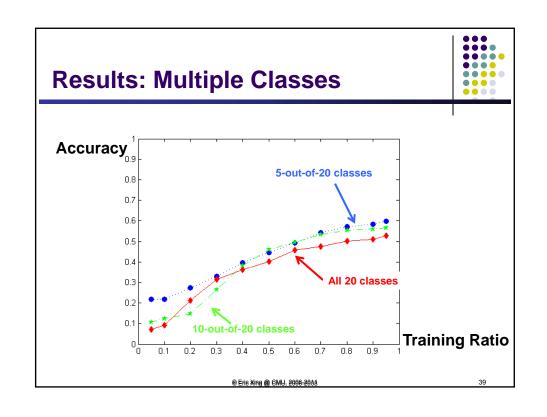
- Dataset
  - 20 News Groups (20 classes)
  - 61,118 words, 18,774 documents

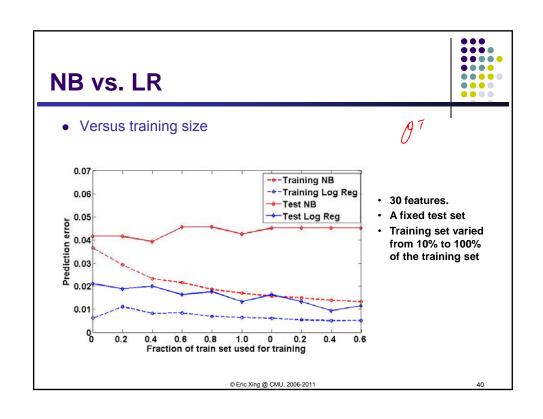
comp.graphics comp.os.ms-windows.misc comp.sys.ibm.pc.hardware comp.sys.mac.hardware comp.windows.x	rec.autos rec.motorcycles rec.sport.baseball rec.sport.hockey	sci.crypt sci.electronics sci.med sci.space
misc forsale	talk.politics.misc talk.politics.guns talk.politics.mideast	talk.religion.misc alt.atheism soc.religion.christia

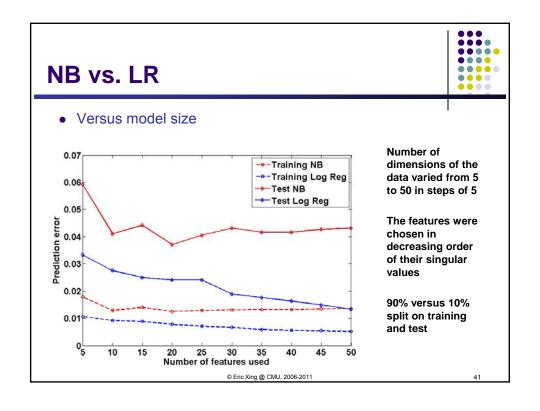
- Experiment:
  - Solve only a two-class subset: 1 vs 2.
  - 1768 instances, 61188 features.
  - Use dimensionality reduction on the data (SVD).
  - Use 90% as training set, 10% as test set.
  - Test prediction error used as accuracy measure.

$$Accuracy = \frac{\displaystyle\sum_{\substack{\text{below set}\\ \text{# of test samples}}} I(\textit{predict}_i == \textit{true label}_i)}{\textit{# of test samples}}$$
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### **Generative vs. Discriminative Classifiers**



- Goal: Wish to learn f:  $X \rightarrow Y$ , e.g., P(Y|X)
- Generative classifiers (e.g., Naïve Bayes):
  - Assume some functional form for P(X|Y), P(Y)
     This is a 'generative' model of the data!
  - Estimate parameters of P(XJY) (P(Y)) directly from training data
  - Use Bayes rule to calculate P(Y|X=x)



- Discriminative classifiers:
  - Directly assume some functional form for P(Y|X)
    This is a 'discriminative' model of the data!
  - Estimate parameters of P(Y|X) directly from training data



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## Naïve Bayes vs Logistic Regression



- Consider Y boolean, X continuous, X=<X1 ... Xm>
- Number of parameters to estimate:

NB: 
$$p(y | \mathbf{x}) = \frac{\pi_k \exp \left\{ -\sum_{j} \left( \frac{1}{2\sigma_{k,j}^2} (x_j - \mu_{k,j})^2 - \log \sigma_{k,j} - C \right) \right\}}{\sum_{k'} \pi_{k'} \exp \left\{ -\sum_{j} \left( \frac{1}{2\sigma_{k',j}^2} (x_j - \mu_{k',j})^2 - \log \sigma_{k',j} - C \right) \right\}} **$$

LR: 
$$\mu(x) = \frac{1}{1 + e^{-\theta^T x}}$$

- Estimation method:
  - NB parameter estimates are uncoupled
  - LR parameter estimates are coupled

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## Naïve Bayes vs Logistic Regression



- Asymptotic comparison (# training examples → infinity)
- when model assumptions correct
  - NB, LR produce identical classifiers
- when model assumptions incorrect
  - LR is less biased does not assume conditional independence
  - therefore expected to outperform NB

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### Naïve Bayes vs Logistic Regression



- Non-asymptotic analysis (see [Ng & Jordan, 2002])
- convergence rate of parameter estimates how many training examples needed to assure good estimates?

NB order log m (where m = # of attributes in X) LR order m

 NB converges more quickly to its (perhaps less helpful) asymptotic estimates

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### Rate of convergence: logistic regression



• Let  $h_{Dis,m}$  be logistic regression trained on n examples in m dimensions. Then with high probability:

$$\epsilon(h_{Dis,n}) \le \epsilon(h_{Dis,\infty}) + O\left(\sqrt{\frac{m}{n}\log\frac{n}{m}}\right)$$

- Implication: if we want  $\epsilon(h_{Dis,m}) \leq \epsilon(h_{Dis,\infty}) + \epsilon_0$  for some small constant  $\varepsilon_0$ , it suffices to pick order m examples
  - $\rightarrow$  Convergences to its asymptotic classifier, in order m examples
  - result follows from Vapnik's structural risk bound, plus fact that the "VC Dimension" of an m-dimensional linear separators is m

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### Rate of convergence: naïve Bayes parameters

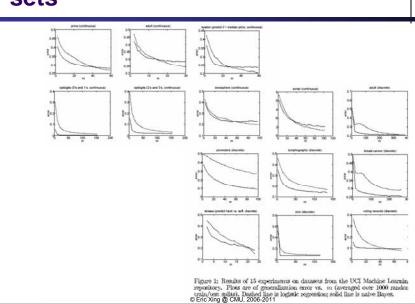


- Let any  $\varepsilon_1$ ,  $\delta$ >0, and any  $n \ge 0$  be fixed. Assume that for some fixed  $\rho_0 > 0$ , we have that  $\rho_0 \le p(y = T) \le 1 - \rho_0$
- Let  $n = O((1/\epsilon_1^2)\log(m/\delta))$
- Then with probability at least 1- $\delta$ , after n examples:
  - 1. For discrete input,  $|\hat{p}(x_i|y=b) p(x_i|y=b)| \leq \epsilon_1 \\ |\hat{p}(y=b) p(y=b)| \leq \epsilon_1$  for all i and b
  - $|\hat{\mu}_{i|y=b} \mu_{i|y=b}| \leq \epsilon_1 \\ |\hat{\sigma}_{i|y=b}^2 \sigma_{i|y=b}^2| \leq \epsilon_1$  for all i and b

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### Some experiments from UCI data sets





### **Summary**



- Naïve Bayes classifier
  - What's the assumption
  - Why we use it
  - How do we learn it
- Logistic regression
  - Functional form follows from Naïve Bayes assumptions
  - For Gaussian Naïve Bayes assuming variance
  - For discrete-valued Naïve Bayes too
  - But training procedure picks parameters without the conditional independence assumption
- Gradient ascent/descent
  - - General approach when closed-form solutions unavailable
- Generative vs. Discriminative classifiers
  - Bias vs. variance tradeoff

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### **Appendix**



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### Subtleties of NB classifier 1 – Violating the NB assumption



- Often the  $X^i$  are not really conditionally independent
- We use Naïve Bayes in many cases anyway, and it often works pretty well
  - Often the right classification, even when not the right probability (see [Domingos&Pazzani, 1996])
  - But the resulting probabilities  $P(Y|X_{new})$  are biased toward 1 or 0 (why?)

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### Subtleties of NB classifier 2 – Insufficient training data



- What if you never see a training instance where w<sup>1000</sup>>0 when Y=b?
  - e.g., Y={SpamEmail or not},  $w^{1:999} \equiv \{\text{'pill'}, \text{'enhancement'}, \text{'Rolex'}, \text{'enlarge'} \dots \}$
  - P( enlargement>0 | Y=T) = 0
- Thus, no matter what the values w<sub>1</sub>,...,w<sub>n</sub>/'enlargement' take:
  - $P(Y=T | w^1, w^2, ..., enlargement, ..., w^k) = 0$
- What now???

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### Learning NB: parameter estimation



Maximum Likelihood Estimate (MLE):
 choose θ that maximizes probability of observed data D

$$\hat{\theta} = \arg\max_{\theta} P(\mathcal{D}|\theta)$$

Maximum a Posteriori (MAP) estimate:
 choose θ that is most probable given prior probability and the data

$$\begin{split} \hat{\theta} &= & \arg \max_{\theta} p(\theta | \mathcal{D}) \\ &= & \arg \max_{\theta} \frac{P(\mathcal{D} | \theta) p(\theta)}{P(\mathcal{D})} \end{split}$$

Bayesian estimate:

$$\hat{\theta} = \int \theta p(\theta|\mathcal{D}) d\theta$$

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#### MAP for the parameters of NB



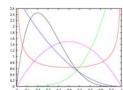
#### Discrete features:

• Maximum a Posteriori (MAP) estimate: (MAP's):

$$\hat{\theta} = \arg \max_{\theta} \frac{P(\mathcal{D}|\theta)\mathcal{P}(\theta)}{P(\mathcal{D})}$$

- Given prior:
  - Consider binary feature
  - $\theta$  is a Bernoulli rate

$$P(\theta; \alpha_T, \alpha_F) = \frac{\Gamma(\alpha_T + \alpha_F)}{\Gamma(\alpha_T)\Gamma(\alpha_F)} \theta^{\alpha_T - 1} (1 - \theta)^{\alpha_F - 1} = \frac{\theta^{\alpha_T - 1} (1 - \theta)^{\alpha_F - 1}}{B(\alpha_T, \alpha_F)}$$



• Let  $\beta_a$ =Count(X=a)  $\leftarrow$  number of examples where X=a

$$P(\theta \mid \mathcal{D}) = \frac{\theta^{\beta_T + \alpha_T - 1} (1 - \theta)^{\beta_F + \alpha_F - 1}}{B(\beta_T + \alpha_T, \beta_F + \alpha_F)} \sim Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$$

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### Bayesian learning for NB parameters – a.k.a. smoothing



- Posterior distribution of  $\theta$ 
  - Bernoulli:  $P(\theta \mid \mathcal{D}) = \frac{\theta^{\beta_T + \alpha_T 1} (1 \theta)^{\beta_F + \alpha_F 1}}{B(\beta_T + \alpha_T, \beta_F + \alpha_F)} \sim Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$
  - $\qquad \text{Multinomial} \qquad P(\theta \mid \mathcal{D}) = \frac{\prod_{j=1}^K \theta_j^{\beta_j + \alpha_j 1}}{B(\beta_1 + \alpha_1, \dots, \beta_K + \alpha_K)} \sim Dirichlet(\beta_1 + \alpha_1, \dots, \beta_K + \alpha_K)$
- MAP estimate:

$$\hat{\theta} = \arg \max_{\theta} P(\theta|\mathcal{D}) =$$

- Beta prior equivalent to extra thumbtack flips
- As  $N \to \infty$ , prior is "forgotten"
- But, for small sample size, prior is important!

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#### MAP for the parameters of NB



- Dataset of N examples
  - Let  $\beta_{iab}$ =Count(X<sup>i</sup>=a,Y=b)  $\leftarrow$  number of examples where X<sub>i</sub>=a and Y=b
  - Let  $\gamma_b$ =Count(Y=b)
- Prior

 $Q(X^i|Y) \propto Multinomial(\alpha_{i1}, ..., \alpha_{iK})$  or  $Multinomial(\alpha/K)$  $Q(Y) \propto Multinomial(\tau_{i1}, ..., \tau_{iM})$  or  $Multinomial(\tau/M)$ 

m "virtual" examples

MAP estimate

$$\hat{\pi}_k = \arg \max_{\pi_k} \prod_k P(Y = y_k; \pi_k) P(\pi_k | \vec{\tau}) = ?$$

$$\hat{\theta}_{ijk} = \arg \max_{\theta_{ijk}} \prod_j P(X^i = x_{ij} | Y = y_k; \theta_{ijk}) P(\theta_{ijk} | \vec{\alpha}_{ik}) = ?$$

Now, even if you never observe a feature/class, posterior probability never zero
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