# Spatial Join Selectivity Using Power Laws 

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#### Abstract

We discovered a surprising law governing the spatial join selectivity across two sets of points. An example of such a spatial join is "find the libraries that are within 10 miles of schools". Our law dictates that the number of such qualifying pairs follows a power law, whose exponent we call "pair-count exponent" (PC). We show that this law also holds for self-spatial-joins ("find schools within 5 miles of other schools") in addition to the general case that the two point-sets are distinct. Our law holds for many real datasets, including diverse environments (geographic datasets, feature vectors from biology data, galaxy data from astronomy).

In addition, we introduce the concept of the Box-Occupancy-Product-Sum (BOPS) plot, and we show that it can compute the pair-count exponent in a timely manner, reducing the run time by orders of magnitude, from quadratic to linear. Due to the pair-count exponent and our analysis (Law 1), we can achieve accurate selectivity estimates in constant time $(\mathrm{O}(1))$ without the need for sampling or other expensive operations. The relative error in selectivity is about $30 \%$ with our fast BOPS method, and even better (about $10 \%$ ), if we use the slower, quadratic method.


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## 1-INTRODUCTION

Multi-dimensional and spatial database management systems (DBMS) have attracted a lot of interest. One of the most important operations in a spatial DBMS [Güt 94] is the spatial join, which is the counterpart to the equi-join in a relational DBMS.

The typical query is also called the 'all pairs' query or 'spatial distance join', as in the example, 'Estimate the number of schools that are within 5 miles from libraries'. Spatial distance joins are considered to be among the most exxential joins in application areas, like data mining [CMN99] [NH94]. They are useful in multiple settings, such as the following.

- In geographic information systems (GIS) under the name of overlay queries: for example, 'Find all houses within 2 miles of a river'.
- In urban planning, business planning, commercial intelligence: 'How many households are within 1 mile of our branches and from our competition's branches'.
- In spatial data mining to detect correlations and test hypotheses: for example, 'Find 4bedroom houses that are within 5 miles of a school', or 'How many luxury apartments are within 2 miles of a lake' [NH 94].
- In temporal data mining: 'Find economic embargos that were followed by war within a year', or 'Find network-switch failures that were within 5 seconds of a power surge' [MTV 95] [ $\mathrm{HKM}+96$ ].
- In multimedia and traditional databases: 'Find pairs of stock price changes that are within $\$ 10$ of each other' [FRM 94].

The spatial distance join is defined using two spatial data sets, $A$ and $B$, and a distance function $L$. For a given radius $r$, the spatial distance join computes $\{\langle a, b\rangle \mid a \in A$ and $b \in B, L(a, b)$ $\leq r\}$. A special case arises when the two datasets, $A$ and $B$ are identical. Such joins will be qualified as 'self spatial joins'. We will use the term 'cross spatial joins', when we need to emphasize that the two point sets are distinct. Otherwise, we will simply use the term 'spatial join' to denote a spatial distance join between two distinct datasets.

The goal of this work is to estimate the selectivity of spatial joins among two datasets as opposed to only one. The join selectivity represents the size of the resultant set of the spatial distance join divided by the size of the Cartesian product of the whole data. Estimation of the join selectivity
is important for the following two reasons.

- An accurate estimation is necessary to optimize complex queries. Though there has been quite a lot of work done on how to estimate the selectivity of equi-joins, the problem of estimating the size of spatial joins has received only minimal attention up to now.
- In application areas like the ones mentioned earlier, the size of the spatial distance join (as a function of the radius) is important for evaluating the correlation between datasets. Note that it is generally too costly to obtain the size of the spatial join by simply computing the spatial distance join itself. Therefore, an accurate and inexpensive method is required to estimate the size of spatial distance joins.

Our main contribution is that we observe a 'power law', which holds for many pairs of real datasets. We show how to use this power law to accurately estimate the spatial join selectivities efficiently (in constant time, $\mathrm{O}(1)$ ).

The rest of the paper is organized as follows. Section 2 presents the related work. Section 3 describes our main contribution, the pair-count exponent $\mathscr{P}$ and the fast way to estimate it, through the proposed box-occupancy-product-sum (BOPS). Section 4 discusses implementation and speed issues of the proposed methods. Section 5 gives experimental results, and Section 6 discusses issues for practitioners. Section 7 presents the conclusions.

## 2 - RELATED WORK

There has been quite a lot of work on spatial joins recently. See, for example [ORE 86], [BKS 93], [LR 94], [PD 96], [KS 97], [APR+98] and [MP99]. Most of the mentioned work has dealt with developing efficient methods to process spatial intersection joins for two-dimensional data sets [BSW 99] [DNS 91] [SK 96] with little emphasis on the estimation of selectivity. Recently, methods have also been examined and developed for processing spatial distance joins on multidimensional point sets [SSA 97], [KS 98]. The term "similarity join" has frequently also been used for spatial distance joins in the literature. For one-dimensional data, the spatial distance join corresponds to the 'band-join' [DNS 91].

Although not directly related to our spatial join selectivity, we mention earlier attempts to estimate the selectivity of range queries. Typical methods include the milestone 'uniformity and
independence' assumptions [SAC+79]. Although simple to use in a query optimizer, these assumptions are pessimistic and unrealistic [ChR84]. Modern methods include histograms [Poo97], kernel estimators [BKS 99], wavelets [VW 99], and hybrid methods using query feedback [KW 99]. Methods for selectivity estimation of range queries in spatial datasets use multi-dimensional histograms [TS 96], or arguments from the theory of fractals [BF 95]. It should be noted that most of these methods are susceptible to the 'dimensionality curse' [SIL 96] [SCO 92].

Analytical estimates of spatial distance join selectivities are few. We are only aware of the very recent work of [PMT99] and [TSS 98]. In these papers, the points of the two datasets are assumed to be uniformly distributed in the address space. As mentioned earlier, the uniformity assumption was discredited long ago [CHR 84], [FK94] as unrealistic and unfeasible. Our experiments in Section 5 indeed show that it is unrealistic.

In the next sections we proceed with our proposed solution. The major observation is that the selectivity of spatial distance joins follows a power law surprisingly well.

## 3 - PROPOSED METHOD

Our main contribution and its corollaries are discussed below. The problem to be solved is the following.

Given: two point-sets $A$ and $B$ and a radius $r$
Find: the distribution of the count of pairs, as a function of the radius $r$.
That is, is this distribution Gaussian? Is it Poisson? Is it Weibul? It turns out that real datasets do not follow any of the traditional statistical distributions. Instead, we show that the distribution of the pair-wise distances follows a power law. Table 1 lists symbols used in this document. Next, we describe our power law, as well as several useful properties of its exponent.

| Symbol | Definition |
| :--- | :--- |
| $N$ | population of the first point-set |
| $M$ | the population of the second point-set |
| $E$ | embedded dimensionality of the point-set (\# of <br> attributes/axis) |
| $\mathscr{P}$ | pair-count exponent |
| $r$ | radius in the queries |
| $r_{\min }$ | smallest distance between the two point-sets |
| $r_{\max }$ | largest distance |
| $S$ | side of grid cells |

Table 1: Symbols and definitions

## 3.1 - Pair-count function and the PC exponent

We propose to study the probability distribution function of the number of pairs as a function of the distance between those pairs. Specifically, we define and study the pair-count function $P C_{A, B}(r)$, or simply $P C(r)$, of two point-sets A and B used in a spatial join query. It is defined as follows.

Definition 1: For two point-sets A and B , we define $P C_{A, B}(r)$ as the pair-count function, that is, the count of pairs within distance $r$ or less. The first member of the pair should belong to point set $A$, and the second member to point set $B$.

$$
P C_{A, B}(r)=\operatorname{count}(\text { of A-B pairs, within distance } \leq r)
$$

Some observations are helpful:

- Our $P C(r)$ function roughly corresponds to the 'cumulative probability density function' from statistics.
- We typically omit the subscripts A, B for simplicity.
- The implied distance function can be any $L_{p}$ norm. We use the $L_{\text {infinity }}$ norm unless otherwise specified. The reason is that all the upcoming results hold for any $L_{p}$ norm, but the formulas are simpler for the $L_{\text {inffinity }}$ norm.
- For a self spatial join (i.e., $\mathrm{A}==\mathrm{B}$ ) we omit the self-pairs, and we count each pair only once.

That is, if there are $N$ points in the set, we consider $N^{*}(N-1) / 2$ pairs. Again, the upcoming results can be easily adapted to handle any of the omitted cases.

For reasons that will soon be obvious, we define the concept of the pair-count plot:

Definition 2: The pair-count plot, or simply PC-plot, for two point sets A and B is the plot of $P C_{A, B}$ $(r)$ versus $r$, in $\log$-log scales.


Figure 1 - The Pair-count plot of California datasets (CA-str cross joined with CA-wat) (a) linear scales, and (b) log-log scales

Figure 1 presents (a) a pair-count plot for real datasets in linear scales, and (b) the same paircount plot in $\log -\log$ scales (b). The datasets are explained in Section 5. The question is whether functions obey any rules? It turns out that many of them indeed follow a law, specifically a power law, as we discuss next. The experiments we have done with many real datasets show that many of them result in a PC-plot that is almost linear (within 1.5\% MLS error and typically less) for a suitable range of distances $r$ (radius from $r_{1}$ to $r_{2}$ ). Considering this, we present our major result.

Law 1 (PAIR-COUNT): For several real datasets and for a usable range of scales, the pair-count $P C(r)$ of pairs within distance $r$ or less follows a power law:

$$
\begin{equation*}
P C(r)=K \cdot r^{\mathscr{Q}} \tag{1}
\end{equation*}
$$

where $K$ is a proportionality constant. Equivalently Definition 3 follows.

Definition 3: The exponent of the law is defined as the pair-count exponent $\mathscr{P}$ as

$$
\begin{equation*}
\mathscr{P}=\frac{\bar{o}(\log (P C(r)))}{\partial(\log (r))} \tag{2}
\end{equation*}
$$

Figure 1 b shows the pair-count plot for the same pair of datasets as Figure 1a in log-log scales. The plots are clearly linear, for a significant range of scales. This range is usually most sought after for queries; we are not interested in radii much smaller or larger than the typical distances involved in the dataset.

Figure 2 shows PC-Plots and fitting lines for two cross-joins of California datasets, a streets cross joined with railroads and $\mathbf{b}$ streets cross joined with water. The description of these datasets and additional $P C(r)$ plots are shown later in Section 5, which deals with our experiments.


Figure 2 - PC-Plots and slopes of the fitting lines and the pair-count exponent $\mathscr{P}$ for two pairs of California datasets: (a) streets cross joined with railroads; (b) streets cross joined with water.

## 3.2-Properties of the pair-count exponent $\mathscr{P}$

The following observations show some of the interesting properties of the pair-count exponent $\mathscr{P}$.
Observation 1: The pair-count exponent $\mathscr{P}$ includes the "correlation fractal dimension" $D_{2}$ as a special case.
Justification: When the second dataset is identical to the first, the PC exponent is, by definition, equal to the "correlation fractal dimension" [Belussi_95]. Intuitively, this is the 'intrinsic' dimensionality of the dataset.

Observation 2: The pair-count exponent $\mathscr{P}$ is invariant to affine transformations, namely to translation, rotation, and uniform scaling.

Justification: By 'uniform scaling' we mean that all the axes are scaled by the same amount. Translation and rotation do not affect the distances ${ }^{5}$ and thus leave the plots unchanged. Uniform scaling scales all the distances, and thus shifts the plot to the left or the right. Its slope, however, remains the same.

Observation 3: The pair-count exponent $\mathscr{P}$ is invariant to sampling.
Justification: Sampling is useful when we deal with large datasets, although our upcoming BOPS algorithm can handle huge datasets even better. It is useful that our power law holds for subsets of our data. The intuitive argument is as follows. Consider a dataset A with $N$ points and a sampling rate $p_{a}\left(0 \leq \mathrm{p}_{\mathrm{a}} \leq 1\right)$, that is the sample has $N^{*} p_{a}$ points. Similarly, let $M$ be the number of points in dataset B , and let $p_{b}$ be its sampling rate. Consider a point $a_{l}$ from the dataset A and let $a_{l}(r)$ be the


Figure 3 - Illustration of the effects of sampling on the pair-count exponent $\mathscr{G}$. The PC-plots for the full datasets and for $20 \%, 10 \%$ and $5 \%$ samples. (a) California pol X wat and (b) Galaxy dev and $\mathrm{X} \exp$.

5
This is clearly true for the Euclidean distance. As we mention later (Observation 4), the pair-count exponent is the same, no matter what $L p$ metric we use. Combining Observation 2 and Observation 4, we see that the pair-count exponent remains the same, for any affine transformation, for any $L p$ metric.
number of its B-type neighbors within distance $r$. After sampling, it will have $p_{a}(r) * p_{b}$ neighbors on the average. Thus, the total number of pairs in the two samples within distance $r$ will be the original $P C(r)$ times $p_{a} * p_{b}$ on the average. This will not change the slope of the PC-plot: it will only lower the position of the plot, by $\log \left(p_{a}{ }^{*} p_{b}\right)$.

Figure 3 shows the $P C(r)$ plots for two pairs of datasets. In (a) it shows California political cross joined with California water and in (b) it shows Galaxydev cross-joined with Galaxy-exp, as well as their 20\%, 10\% and 5\% samples. Notice that the plots are linear, and those corresponding


Figure 4 - Effects of the distance functions to obtain PC-plots. to samples are parallel to the full dataset.

Tables 3 and 4 summarize their $\mathscr{P}$ values.

Observation 4: The pair-count exponent $\mathscr{P}$ is invariant to the $L_{p}$ distance used.
Justification: Consider the 'sphere' that each $L_{p}$ metric defines (see Figure 4). Let $\operatorname{vol}(p, r)$ be the volume of an n-dimensional $L_{p}$ '‘sphere' of radius $r$. For $p=2$, this is indeed a sphere; for $p=$ infinity this is an $n$ dimensional cube, etc. Our power law states that the number of type-B neighbors of a type-A point grows as $r^{\mathscr{D}}$ or, equivalently it grows as volume ${ }^{\Phi / E}$. Then, if $P C_{p}(r)$ denotes the number of neighbors within $L_{p}$ distance $r$, we have:


Figure 5 - The shapes of $L_{\infty} L_{l}$ and $L_{2}$ norms in 2-d.

$$
\begin{equation*}
P C_{A B}\left(r, L_{P}\right)=P C\left(r, L_{\infty}\right) \cdot(\operatorname{vol}(p, r) / \operatorname{vol}(p, r))^{\Phi / E} \tag{3}
\end{equation*}
$$

therefore, the number of pairs will only differ by a multiplicative constant for different values of $p$ in the $L_{p}$ metric. Figure 5 shows the effect of norm invariance on the cross join of two California
datasets (political and water). It is clear that the three $L_{p}$ metrics chosen result in parallel lines. Therefore, for the rest of this work, we will only focus on the $L_{i n f}$ metric. We can conclude that the pair-count exponent shows an intrinsic property of the two point-sets, and it is independent of the particular $L_{p}$ distance function used to build the PC plot.

## 4 - IMPLEMENTATION AND SPEED ISSUES

By the definition of the 'pair-count exponent', we need to estimate the pair-counts for several distances $r$. Each of them requires $O\left(N^{*} M\right)$ operations, which are quadratic on the size of the input datasets. This is prohibitive for large datasets. The question becomes: how we can accelerate the computation of $\mathscr{P}$. This is precisely the topic of this section.


Figure 6 - A grid superimposed over a point-set to count $C_{A i}$ and $C_{B i}$

## 4.1-A faster way to compute the 'pair-count exponent' $\mathscr{P}$

Here we give a Lemma, which computes of the pair-count exponent $O(N+M)$ and thus performs dramatically faster for huge datasets. A crucial concept that we introduce is the Box-Occupancy-Product-Sum (BOPS), which is defined as follows. Consider the address space of two point-sets in a $n$-dimensional space, and impose an $n$-grid with grid-cells of side $s$ (or, equivalently, radius $r=s / 2$ ). Focusing on the $i$-th cell, let $C_{A, i}, C_{B, i}$ be the counts ('occupancies') of points from the first and from the second point-set, respectively, as illustrated in Figure 6.

Definition 2: The 'Box-Occupancy-Product-Sum" (BOPS) of a grid with cell side $s$ is defined as the sum of products of occupancies as

$$
\begin{equation*}
B O P S(s)=\sum_{i} C_{A, i} * C_{B, i} \tag{4}
\end{equation*}
$$

and the BOPS plot is the plot of $\operatorname{BOPS}(s)$ as a function of the grid side $s$, in log-log scales.

Lemma 1 (BOPS): The pair-count exponent $\mathscr{P}$ for a given radius is equal to the box-occupancy-product-sum (BOPS) for the doubled radius; that is

$$
\begin{equation*}
P C(s / 2) \approx B O P S(s) \tag{5}
\end{equation*}
$$

Proof: The fundamental assumption is that the densities of points are smooth functions. Thus, if a point $p_{1}$ of set A has $x$ neighbors from the set B within radius $r$, so does a close-by neighbor $p_{2}$ that also belongs to set A .

Thus, for a given cell side $s$ and another given cell (say, the $i$-th one), consider one of the points of the set A. This point has a number of neighbors proportional to $C_{B, i}$ neighbors from the set B within radius $s / 2$. Thus, the $i$-th cell contributes with

$$
\begin{equation*}
C_{A, i} * C_{B, i} \tag{6}
\end{equation*}
$$

pairs. Adding up the contributions of all the cells, we have

$$
\begin{equation*}
P C(s / 2)=\sum_{i} C_{A, i} \cdot C_{B, i} \tag{7}
\end{equation*}
$$

which completes the proof.
QED
Corollary: The BOPS follows a power law with its exponent equal to the "pair-count exponent".

$$
\begin{equation*}
B O P S(s)=s^{p} \tag{8}
\end{equation*}
$$

Proof: Trivial, from Lemma 1 and Law 1.
QED

We are going to use the estimation $\mathrm{PC}(\mathrm{r})=\operatorname{BOPS}(2 r)$ for the rest of this work. The 'BOPS' Lemma has important efficiency implications which are vital for large datasets. Next we show how to use this Lemma for fast selectivity estimations.

## 4.2-Algorithms

The problem is defined as follows.
Given two point-sets A and B in $n$-dimensional space,
Estimate their pair-count exponent $\mathscr{P}$ and the proportionality constant $K$.
We developed a single-pass algorithm to obtain the BOPS plot. Specifically, the algorithm is linear $O(N+M)$ over the total number of points in both datasets. If $l$ is the number of points that we want in the BOPS plot (ie., number of grid-sizes), then the complexity of our algorithm is $O((N+M) * l * n)$, where $n$ is the dimensionality of the input point-sets. Below is a brief algorithm to generate the BOPS-plot and the estimate of the pair-count exponent.

Without loss of generality, due to Observation 2,
Normalize the address space of the datasets to the unit hyper-cube;
For each desirable grid-size $s=1 / 2^{\mathrm{j}}, \mathrm{j}=1,2, \ldots, l$;
For each point $a$ of dataset A
Decide which grid cell it falls in (say, the i-th cell);
Increment the count $C_{A} i$;
For each point $b$ of dataset B
Decide which grid cell it falls in (say, the i-th cell);
Increment the count $C_{B, i}$;
Compute the sum of product occupancies $\operatorname{BOPS}(s)=\sum C_{A, i} * C_{B, i}$;
Print the values of $\log (s / 2)$ and $\log (\operatorname{BOPS}(s))$ as the BOPS-plot;
Perform a linear interpolation and report the slope $\mathscr{P}$ and the;
proportionality constant $K$.
Figure 7 - Algorithm for calculating BOPS plots.

## 4.3-Estimation of selectivity

Here we describe exactly how to estimate the spatial join selectivities, exploiting our two major observations, the pair-count law and the BOPS lemma. More specifically, the problem is as follows.

Given two point-sets A and B , and a radius $r$,
Estimate the count of pairs $P C(r)$.

We distinguish the following methods, depending on what else we are given:

- PC plot estimation: Through previously kept statistics on the PC plot, suppose that we already know the pair-count exponent $\mathscr{P}$ and the proportionality constant $K$. Then we estimate immediately the PC plot as $P C(r)=K * r^{g}$
- BOPS plot estimation: We assume that we are given only the dataset, without any statistics about the data. Then, we generate the BOPS plot for several values of grid-side $s$, and we estimate the slope $\mathscr{P}$ and the constant $K$, as explained in the algorithm in Figure 7. Notice that we not only obtain our estimate, but we also provide $\mathscr{P}$ and $K$ for future upcoming queries.

An obvious trick to approximate the BOPS plot is to do sampling first. We discuss its relative merit in Section 5.

## 5 - EXPERIMENTS

We implemented our method and checked whether the power law holds for different data sets. For the sake of clarity we named the datasets used in the experiments. Point-sets come in groups; thus, each dataset is characterized by its group name, a dash '-' and the dataset name. Their characteristics are as follows.


Figure 7 - Real data used in the experiments. (a) California: CA-pol and CA-wat, (2-dimensional point-sets), (b) Iris: setosa, versicolor and virginica (4-dimensional point-sets) and (c) Galaxy: class dev and $\exp$ (2-dimensional point-sets).

California - Two-dimensional sets of points, they refer to geographical coordinates in California. The four files contain data features from streets (CA-str with 62,933 points), railways (CA-rai with 31,059 points), political borders (CA-pol with 46,850 points), and natural water systems (CA-wat, with 72,066 points).
Iris - This set contains three files, each of which describes a few properties of a specific flower type of Iris. The points are 4-dimensional (sepal length, sepal width, petal length, petal width); the species are 'virginica', 'versicolor' and 'setosa', and there are 50 points from each species. This is a well-known dataset in the literature of machine learning and statistics, which we obtained from the UC-Irvine Repository.

Galaxy - Galaxies come from the SLOAN telescope: ( $\mathrm{x}, \mathrm{y}$ ) coordinates, plus class label. There are 82,277 in the 'dev' class (deVaucouleurs), and 70,405 in the 'exp' class (exponential).

Eigenfaces - Two datasets ('lyf' with 11,900 points; and 'tyf' with 3,456 points) come from the Informedia project [WKS+96] at Carnegie Mellon University. Each face was processed with the eigenfaces method [TP 91], resulting in 16-dimensional points.

Our experiments are designed to answer the following questions.

- How often do real datasets follow the proposed power law? How good is the linear fit?
- How accurate is our 'box-occupancy-product-sum' Lemma?
- What are the effects on sampling and affine transformations on them?
- How fast is the BOPS method, compared to other estimations of $P C(r)$ ?

Due to lack of space, we present here mainly the graphs of some of the datasets used in the experiments.

## 5.1-Accuracy of 'PC’ Law

We present our experiments in two groups, two-dimensional geographical datasets (California and Galaxy data), and higher-dimensionality ones (Iris, Eigenfaces).


Figure 8 - PC plots and the pair-count exponents $\mathscr{P}$ of geographical data. First row: Galaxy datasets (a) cross join of 'dev' and 'exp', (b) self join of 'dev', (c) self join of 'exp'. Second row California datasets (d) cross join of CA-pol and CA-wat, (e) self join of CA-pol, (f) self join of CA-wat.

### 5.1.1-Geographical datasets

The immediate application for the pair-count exponent is to estimate the selectivities for cross spatial joins. Thus, the natural candidates to show that this method works are geographical datasets. Figure 8 shows the pair-count exponent for California and Galaxy datasets, and it can be seen that the PC plots are linear for a suitable range of $r$. The slopes of the fitting lines are also shown, and these give us the proportionality constant that will be used to estimate the selectivities in cross or self joins.

### 5.1.2 - Higher Dimensional datasets

Figure 9 presents the PC-plots, the fitting lines and the pair-count exponent $\mathscr{P}$ for the Eigenfaces datasets which are 16-dimensional data. It can be seen that our power law remains quite accurate for high-dimensional datasets. Recurring conclusions from all the above experiments are

1. The linear fit implied by our 'pair-count' law is extremely precise, for a wide variety of diverse datasets.
2. For self-joins, as well as for cross-joins, the correlation coefficient of the fit is at least 0.995 (where ' 1 ' is the value of perfect linear correlation).
3. Especially for the high-dimensional datasets, the self-join exponent is significantly lower than the embedding dimensionality of the data. For example, in Eigenfaces, the intrinsic dimensionality is between 4.5 to 6.7 (values of $\mathcal{P}$ varies from 4.49 for self-join of 'lyf' to 6.73 for the cross-join of 'tyf' and 'lyf'), while the embedding dimensionality $E$ was 16 . This implies that these $n$-dimensional points are not even close to being uniformly distributed (if they were, then $\mathscr{P}=16$ ). Thus, any analysis making the uniform assumption will be very inaccurate, since the dimensionality of the data ( $\mathscr{P}$ or $E$ ) is in the exponent!


Figure 9-PC Plots and the pair-count exponent $\mathscr{P}$ of the Eigenfaces datasets, (a) self join of 'lyf' dataset, (b) self join of 'tyf' dataset, (c) cross join of 'lyf' and 'tyf' datasets.

## 5.2 - Sampling

We present further experiments in order to illustrate Observation 3, which states that PC plots are invariant to sampling. Figure 10 presents the pair-count exponents obtained from PC plots (points) and BOPS plots (lines). All plots are clearly parallel. Table 2 shows the results for the Galaxy and California datasets when the pair-count exponent was calculated for self-joins. Sampling clearly has negligible effects on the PC exponent. Table 3 shows the results for the same datasets using the paircount exponent obtained from PC plots and from BOPS plots.



Figure 10 - PC-plots and corresponding BOPS plots for (a) California datasets and (b) Galaxy datasets. Both plots are shown for the full datasets and three levels of sampling.

|  | Galaxy |  | California |  |  |  |
| :--- | ---: | ---: | :--- | :--- | ---: | ---: |
| Sampling rate | dev | dexp | pol | wat | str |  |
| $100 \%$ | 1.876 | 1.928 | 1.650 | 1.529 | 1.838 |  |
| $20 \%$ | 1.875 | 1.932 | 1.643 | 1.562 | 1.701 |  |
| $10 \%$ | 1.873 | 1.952 | 1.631 | 1.694 | 1.661 |  |
| $5 \%$ | 1.880 | 2.146 | 1.515 | 1.711 | 1.623 |  |

Table 2: The pair-count exponents $\mathscr{P}$ for samples of Galaxy ('dev' and 'exp') and California (CA_pol, CA_wat and CA_str) datasets for self-joins.

Conclusions from the above experiments are as follows.
1). The pair-count exponent $\mathscr{P}$ is practically unaffected by sampling, for reasonable sample sizes (e.g., equal or higher than $10 \%$ ).
2). Whatever the sampling rate, the corresponding BOPS plot on the samples is very close to the pair- count plot of the samples. This means that whatever the time that sampling can save, BOPS applied on the samples will outperform, with practically the same accuracy.

The estimation of $\mathscr{P}$ obtained from BOPS results on relative error practically always less than $5 \%$. Only when the sampled size of a dataset is very small, the BOPS plot results in a $9 \%$ error; indeed, $9 \%$ of error is also a reasonable value.

| $\begin{aligned} & \text { Sampling } \\ & \text { rate } \end{aligned}$ | Galaxy dev x exp |  | California pol x wat |  | California pol x str |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\mathscr{P}$ from PC | $\mathscr{P}$ from BOPS | $\mathscr{P}$ from PC | $\mathscr{P}$ from BOPS | $\mathscr{P}$ from PC | $\mathcal{P}$ from BOPS |
| 100\% | 1.915 | 1.963 | 1.835 | 1.819 | 1.783 | 1.743 |
| 20\% | 1.915 | 1.963 | 1.833 | 1.825 | 1.776 | 1.759 |
| 10\% | 1.902 | 1.965 | 1.839 | 1.816 | 1.783 | 1.715 |
| 5\% | 1.918 | 1.736 | 1.856 | 1.786 | 1.752 | 1.725 |

Table 3: The pair-count exponent $\mathscr{P}$ values (PC and BOPS) for joins on sampled data from Galaxy ('dev' and 'exp') datasets and also on California_pol, California_wat, California_str datasets.

## 5.3-Accuracy of Selectivity Estimations

We see that the pair-count Law is obeyed (Figures 8 and 9). We also have just seen (Figure 10 and Table 3) that our BOPS Lemma leads to very close approximations for the pair-count exponent. The question now becomes how precise the selectivity estimation $P C(r)$ can be by using,
(a) our Law 1 and
(b) our estimates from BOPS.

|  | Galaxy |  |  | California |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | dev x exp | dev x dev | exp x exp | pol x wat | pol x pol | wat x wat |
| PC plot <br> estimation | 0.027 | 0.0164 | 0.023 | 0.028 | 0.029 | 0.067 |
| BOPS plot <br> estimation | 0.138 | 0.247 | 0.259 | 0.164 | 0.305 | 0.348 |

Table 4 - Geometric average of the relative error of selectivity estimation.
Table 4 shows the relative error for the selectivities calculated by $\frac{P C(r)-\hat{P C}(r)}{P C(r)}$, and we report the geometric average values for several values of $r$. The top row estimates $\hat{P} C(r)$ as follows.

Step (a): Compute the PC plot.
Step (b): Fit the line to obtain the estimation.
In order to measure the relative error in estimating the selectivities of queries, we compared paircount exponent methods to the real prediction given by Law 1. Table 4 presents the geometric average of the relative error of the PC plot by the pair-count exponents $\mathscr{P}$ when we compare the values obtained from PC and BOPS plots with the actual figures given by Law 1.

## 5.4-Timing results

The question now becomes: (a) how long it takes to estimate the PC exponent with the PC plot and (b) how long it takes to obtain the estimation from the BOPS plot. Table 5 reports the wall clock times for each plot on an Intel Pentium II 450 MHz , running Windows NT. Both methods were implemented in C++ language.

We can see in Table 5 that there is a huge difference in the CPU time when calculating the PC plots and BOPS plots. Calculating the pair-count exponent using BOPS method save orders of magnitude. Moreover, BOPS plots give a fast and accurate approximation of $\mathscr{P}$. Sampling also gives a close approximation of $\mathscr{P}$, but is much more time-consuming because all the dataset must be scanned in order to generate the sample before to apply the PC plot. When we compare the time needed to obtain the pair-count exponent for a dataset sampled to $10 \%$ of the data (a limit to preserve the
accuracy of the estimation), BOPS still remains much faster than sampling technique, from 5.27 seconds for the whole dataset for BOPS to 2.11 minutes for a $10 \%$ sampling for PC plot.

| Datasets |  | PC-Plot | BOPS |
| :---: | :---: | :---: | :---: |
| California | pol x wat ( $100 \%$ of data) | 7,752.50 | 3.44 |
|  | pol $x$ wat ( $10 \%$ of data) | 73.36 | 0.5 |
|  | str x rai <br> ( $100 \%$ of data) | 4,434.27 | 2.55 |
|  | str x rai <br> ( $10 \%$ of data) | 42.64 | 0.47 |
|  | pol x str <br> ( $100 \%$ of data) | 7,664.28 | 3.44 |
|  | pol $x$ str ( $10 \%$ of data) | 66.58 | 0.53 |
| Galaxy | $\operatorname{dev} \mathrm{x} \exp$ <br> ( $100 \%$ of data) | 13,078.38 | 5.27 |
|  | $\operatorname{dev} x \exp$ <br> ( $10 \%$ of data) | 126.98 | 0.72 |
| Iris | setosa x virginica | 5.32 | 0.01 |
|  | virginica x versicolor | 4.98 | 0.01 |

Table 5 - Clock time in seconds to obtain the pair-count exponent by PC-plots and BOPS-plots.

Table 5 reports the times needed to build each plot for several pairs of datasets. It also shows the times, when only samples are fed into the two algorithms. The sampling rate is reported on each row, and it is the same for both datasets. The observations are the following:
1). Our BOPS method is up to four order of magnitude faster.
2). In fact, BOPS on the full sets is still faster than the PC plots on the samples ( $10 \%$ sampling rate), up to 20 times! Thus, we conclude that the BOPS plot is a fast and accurate
tool for selectivity estimation of spatial joins.

## 6 - DISCUSSION

Our discussion addresses two questions, which are
a) How often should we expect the 'pair-count' law to hold?
b) How can we use it to do other extrapolations?

## 6.1 - How often?

We mention that power laws regularly occur in real datasets. In fact, our 'pair- count' law is obeyed by the self-join of any self-similar dataset, in which case the 'pair-count' exponent is exactly the correlation fractal dimension $\mathrm{D}_{2}$ of that dataset. It is well-known that vast majority of real datasets are self-similar [BF 95], coastlines, with fractal dimension 1.1-1.3, stock prices (fractal dimension $=$
 As we have just seen, the same is true for the self-joins of our real datasets ( 1.9 for the GALAXY datasets, 1.5-1.8 for the CA datasets, 1.9-2.9 for the 4-dimensional IRIS datasets, and 4.5-5.4 for the 16-dimensional Eigenfaces datasets).

## 6.2-Other extrapolations

There is a wealth of estimations that we can perform whenever a pair of real datasets obeys the paircount law, and the invariant properties of the pair-count exponent $\mathscr{P}$. One extrapolation is to estimate the minimum distance $r_{\text {min }}$ between the closest pair of points. The formula is

$$
\begin{align*}
& P C\left(r_{\min }\right)=1=K r_{\min }^{\mathscr{P}} \\
& r_{\min }=K^{-1 / g} \tag{11}
\end{align*}
$$

The justification comes straightforward from Law 1. We can also estimate the distance $r_{c}$ of the $c$-th closest pair and the formula is

$$
\begin{equation*}
P C\left(r_{c}\right)=K r_{c}{ }^{\mathscr{P}} \tag{12}
\end{equation*}
$$

Additional extrapolations can be performed for subsets and supersets of the two original datasets since the pair-count exponent $\mathscr{P}$ is not affected by sampling.

## 7 - CONCLUSIONS

The main contribution of this work is the identification of a power law, namely the 'pair-count' law. This is the first and only published law that governs the distribution of pair-wise distances between two real, $n$-dimensional point-sets. This law leads to the estimation of spatial join selectivities through a simple formula, which is extremely accurate, less than $9 \%$ of error. Given the pair-count exponent $\mathscr{P}$, the selectivity estimations can be performed in constant time $(O(1))$ without the need for sampling or any other costly operations. Additional contributions include the following:

- The identification of several invariant properties of the pair-count exponent $\mathscr{P}$. It is invariant to rotation, translation, scaling, sampling. Moreover, this holds for any $L_{p}$ norm.
- Efficiency issues: the introduction of the BOPS concept (box-occupancy-product-sum). It allows a fast estimation of the pair-count exponent $\mathscr{P}$. Its response time is orders of magnitude better than the straightforward estimation using the pair-count function $P C(r)$. Thanks to the BOPS plot, the whole concept of the pair-count exponent becomes practical. In fact, our method used on the full sets, is still significantly faster than the PC plots on samples.
- Experiments on many, diverse datasets. The experiments show that (a) the pair-count law holds for a surprisingly large number of real datasets and (b) that our BOPS approximation is highly accurate. The error is less than $9 \%$ for the pair-count exponent $\mathcal{P}$ and less than $35 \%$ for the selectivity estimation.

Future research could focus on the discovery of additional power laws in real, spatial datasets, as well as on explaining the reasons why these laws hold.

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