

Homework 1: Solution

*Lecturer: Aarti Singh***Acknowledgement:** *The TA graciously thanks Rafael Stern for providing most of the solutions.***1.1 Problem 1****1.1.1 1.1**

Weighing 6 balls vs 6 balls yields only two possible outcomes: left heavier or right heavier. Hence, any weighing scheme that at any point weighs 6 vs 6 can yield, in 3 weighings, at most $2 * 3 * 3 = 18$ possible outcomes. However, to distinguish the special ball and determine whether it is heavier or lighter, we need to our method to differentiate between 24 possible cases. $18 < 24$, and hence, we need at least 4 weighings. Since the optimal method uses 3 weighings, this is suboptimal.

1.1.2 1.2

We can use similar reasoning as 1.1.

Suppose that the outcome of the first weighing is that the two sets of 3 balls are equally heavy, then the odd ball must be in the remaining 6 balls. We must then be able to differentiate between 12 cases. Therefore, we need at least 3 additional weighings to find the odd ball in the remaining 6 balls.

We can then conclude that any method that weighs 3 balls versus 3 balls in its first weighing needs at least 4 weighings.

1.1.3 1.3

See figure 1.1.3

1.2 Problem 2**1.2.1 2.1**

To prove that $H(X|Y) \leq H(X)$, recall that $H(X) = H(X|Y) + I(X;Y)$. Hence, it is enough to show that $I(X;Y) \geq 0$. Observe that $I(X;Y) = D_{KL}(p(x,y);p(x)p(y))$. Thus, by Gibb's Inequality, $I(X;Y) \geq 0$ and the desired result is established.

Next, say that a variable has distribution $S(p)$ if $P(X = 1) = p$ and $P(X = -1) = 1 - p$. Define $X \sim S(0.5)$ and $Z \sim S(0.5)$ independent random variables. Consider $Y = XZ$. Observe that Y is independent of X and, thus, $I(X;Y) = 0$. Observe that, given no value of Z , X and Y are independent, that is, $I(X;Y|Z = z) > 0$. Thus, $I(X;Y|Z) > 0 = I(X;Y)$.

1.2.2 2.2

Lets prove this result by induction. The case $n = 2$ was shown in class, hence it is only necessary to prove the inductive step. Lets assume that the equality works for n and show that it works for $n + 1$.

$$\begin{aligned}
 H(X_1, \dots, X_{n+1}) &= - \sum_{x_1, \dots, x_{n+1}} p(x_1, \dots, x_{n+1}) \log(p(x_1, \dots, x_{n+1})) = \\
 &= - \sum_{x_1, \dots, x_{n+1}} p(x_1, \dots, x_{n+1}) \log(p(x_{n+1})) - \sum_{x_1, \dots, x_{n+1}} p(x_1, \dots, x_{n+1}) \log(p(x_1, \dots, x_n | x_{n+1})) = \\
 &= - \sum_{x_{n+1}} p(x_{n+1}) \log(p(x_{n+1})) - \sum_{x_{n+1}} p(x_{n+1}) \sum_{x_1, \dots, x_n} p(x_1, \dots, x_n | x_{n+1}) \log(p(x_1, \dots, x_n | x_{n+1})) = \\
 &= H(X_{n+1}) + \sum_{x_{n+1}} p(x_{n+1}) H(p(x_1, \dots, x_n | x_{n+1})) =
 \end{aligned}$$

Calling $p(x_{a_1}, \dots, x_{a_n} | x_{b_1}, \dots, x_{b_n}, x_{n+1}) = q_{x_{n+1}}(x_{a_1}, \dots, x_{a_n} | x_{b_1}, \dots, x_{b_n})$

$$= H(X_{n+1}) + \sum_{x_{n+1}} p(x_{n+1}) H(q_{x_{n+1}}(x_1, \dots, x_n)) =$$

By the induction hypothesis:

$$\begin{aligned}
 \sum_{x_{n+1}} p(x_{n+1}) H(q_{x_{n+1}}(x_1, \dots, x_n)) &= \sum_{x_{n+1}} p(x_{n+1}) \sum_i H(q_{x_{n+1}}(x_{n-i} | x_n, \dots, x_{n-i+1})) = \\
 &= \sum_{i \geq 1} H(X_{n+1-i} | X_{n+1}, \dots, X_{n+2-i})
 \end{aligned}$$

Substituting this expression in the main equality:

$$H(X_1, \dots, X_{n+1}) = \sum_{i \geq 0} H(X_{n+1-i} | X_{n+1}, \dots, X_{n+2-i})$$

Which completes the proof. Equality holds when X_1, \dots, X_n are jointly independent, that is, any two disjoint subsets of (X_1, \dots, X_n) are independent.

1.2.3 2.3

$$\begin{aligned}
 I(X_1, \dots, X_n; Y) &= H(X_1, \dots, X_n) - H(X_1, \dots, X_n | Y) = \\
 &= H(X_1, \dots, X_n) - H(X_1, \dots, X_n, Y) - H(Y) =
 \end{aligned}$$

Applying item (b) twice:

$$\begin{aligned}
&= \sum_{i=1}^n (H(X_i|X_{i-1}, \dots, X_1) - H(X_i|X_{i-1}, \dots, X_1, Y)) + H(Y) - H(Y) = \\
&= \sum_{i=1}^n I(X_i; Y|X_{i-1}, \dots, X_1)
\end{aligned}$$

1.2.4 2.4

Note: many of you did not prove the first part—that $D(\cdot || \cdot)$ is convex. Rafael's solution is very slick and I recommend that you take a look if you hadn't proved it yourself.

$$\begin{aligned}
D_{KL}(\lambda p_1 + (1-\lambda)p_2 || \lambda q_1 + (1-\lambda)q_2) &\leq \lambda D_{KL}(p_1 || q_1) + (1-\lambda) D_{KL}(p_2 || q_2) \\
D_{KL}(\lambda p_1 + (1-\lambda)p_2 || \lambda q_1 + (1-\lambda)q_2) &= \int (\lambda p_1 + (1-\lambda)p_2) \log\left(\frac{\lambda p_1 + (1-\lambda)p_2}{\lambda q_1 + (1-\lambda)q_2}\right) dx = \\
&= - \int (\lambda p_1 + (1-\lambda)p_2) \log\left(\frac{\lambda p_1}{\lambda p_1 + (1-\lambda)p_2} \cdot \frac{\lambda q_1}{\lambda p_1} + \frac{\lambda p_2}{\lambda p_1 + (1-\lambda)p_2} \cdot \frac{(1-\lambda)q_2}{(1-\lambda)p_2}\right) dx \leq \\
&\leq - \int \lambda p_1 \log\left(\frac{q_1}{p_1}\right) + (1-\lambda)p_2 \log\left(\frac{q_2}{p_2}\right) dx = \lambda D_{KL}(p_1 || q_1) + (1-\lambda) D_{KL}(p_2 || q_2)
\end{aligned}$$

We wish to prove that $H(\lambda p_1 + (1-\lambda)p_2) \geq \lambda H(p_1) + (1-\lambda)H(p_2)$. Take $X \sim p_1$ and $Y \sim p_2$. Consider $\xi \sim \text{Ber}(\lambda)$ independent of X and Y . Define $Z = \xi X + (1-\xi)Y$. Observe that $Z \sim \lambda p_1 + (1-\lambda)p_2$. Hence the left hand side of the inequality equals $H(Z)$. Observe that the right hand side equals $H(Z|B)$ and, thus, by Exercise 2a, the proof is complete.

1.3 Problem 3

1.3.1 3.1

Using the data processing inequality, we know that $I(U; V) \leq I(X; V)$. Hence, we only need to prove that $I(X; V) \leq I(X; Y)$. By symmetry of mutual information, this is equivalent to show that $I(V, X) \leq I(Y; X)$. If we could invert the graph as in the hint, we could establish this inequality by using the data processing inequality once again. Hence, it only remains to prove that $X \rightarrow Y \rightarrow Z$ implies $Z \rightarrow Y \rightarrow X$.

Assume $X \rightarrow Y \rightarrow Z$. Hence, $p(x, y, z) = p(x)p(y|x)p(z|y)$. Using Bayes Rule, conclude that: $p(x, y, z) = p(x|y)p(y)p(z|y)$. Finally, using Bayes Rule again: $p(x, y, z) = p(x|y)p(y|z)p(z)$. Since x, y and z were arbitrary, we have established that $Z \rightarrow Y \rightarrow X$, which completes the proof.

1.3.2 3.2

$$f(x_1, \dots, x_n | \theta) = \prod_{i=1}^n I_{[\theta, \infty]}(x_i) I_{[-\infty, \theta+1]}(x_i) =$$

$$= \prod_{i=1}^n I_{[\theta, \infty]}(x_i) \prod_{i=1}^n I_{[-\infty, \theta+1]}(x_i) = I_{[\theta, \infty]}(x_{(1)}) I_{[-\infty, \theta+1]}(x_{(n)})$$

Hence, by Fisher's Factorization Theorem, T is a sufficient statistic for θ in the statistical sense. Hence, by the result seen during Recitation 1, it is also sufficient in the information theoretic sense.

Figure 1.1: Solution to Problem 1.3

