

Linear Regression

Aarti Singh

Machine Learning 10-701/15-781
Sept 27, 2010



MACHINE LEARNING DEPARTMENT



Discrete to Continuous Labels

Classification

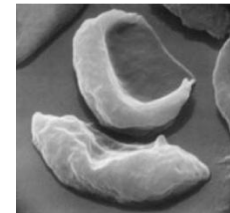


X = Document

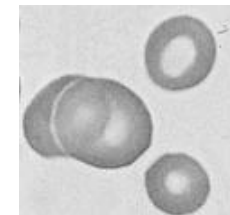


Sports
Science
News

Y = Topic



Anemic cell
Healthy cell



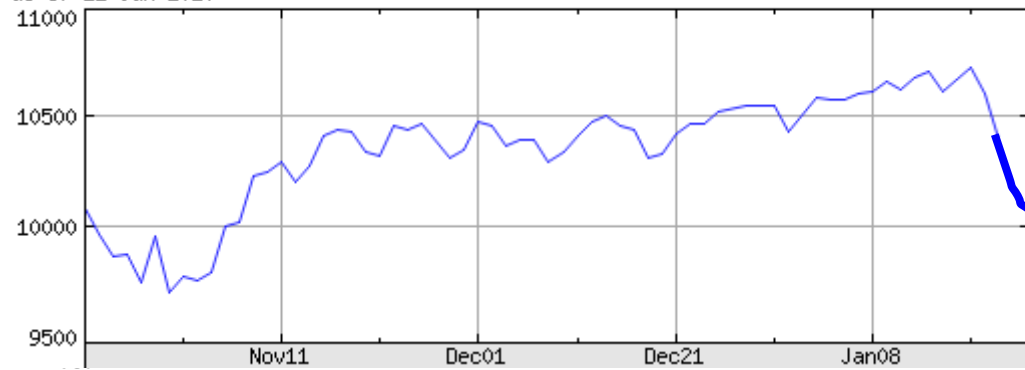
X = Cell Image

Y = Diagnosis

Regression

Stock Market
Prediction

DJ INDU AVERAGE (DOW JONES & CO
as of 22-Jan-2010



X = Feb01

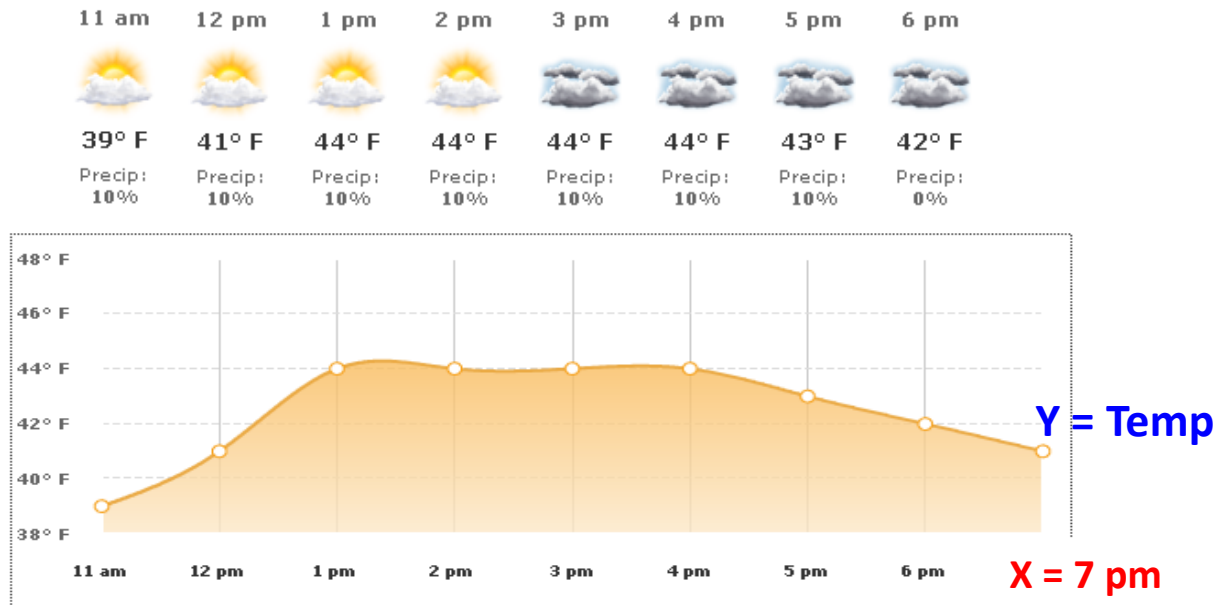
Y = ?

Copyright 2010 Yahoo! Inc.

<http://finance.yahoo.com/>

Regression Tasks

Weather Prediction



Estimating Contamination

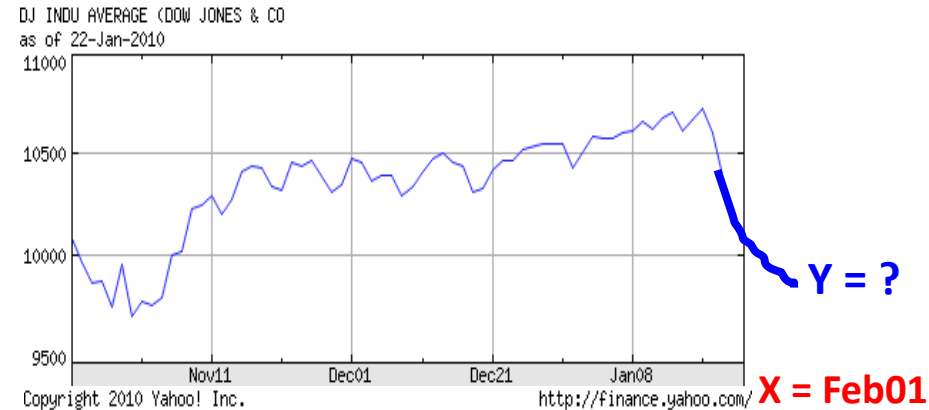


Supervised Learning

Goal: Construct a **predictor** $f : X \rightarrow Y$ to minimize a risk (performance measure) $R(f)$



Sports
Science
News



Classification:

$$R(f) = P(f(X) \neq Y)$$

Probability of Error

Regression:

$$R(f) = \mathbb{E}[(f(X) - Y)^2]$$

Mean Squared Error

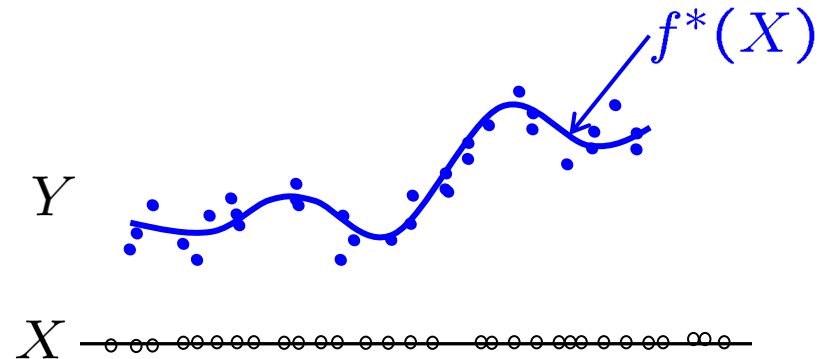
Regression

Optimal predictor:

$$f^* = \arg \min_f \mathbb{E}[(f(X) - Y)^2]$$
$$= \mathbb{E}[Y|X] \quad (\text{Conditional Mean})$$

Intuition: Signal plus (zero-mean) Noise model

$$Y = f^*(X) + \epsilon$$



Regression

Optimal predictor: $f^* = \arg \min_f \mathbb{E}[(f(X) - Y)^2] = \mathbb{E}[Y|X]$

Proof Strategy: $R(f) \geq R(f^*)$ for any prediction rule f

$$R(f) = \mathbb{E}_{XY}[(f(X) - Y)^2] = \mathbb{E}_X[\mathbb{E}_{Y|X}[(f(X) - Y)^2|X]]$$

Dropping subscripts
for notational convenience

$$= E \left[E \left[\underbrace{(f(X) - E[Y|X])}_{\text{red bracket}} + \underbrace{E[Y|X] - Y}_{\text{red bracket}} \right)^2 | X \right]$$

$$= E \left[\begin{aligned} &E[(f(X) - E[Y|X])^2|X] \\ &+ 2E[(f(X) - E[Y|X])(E[Y|X] - Y)|X] \\ &+ E[(E[Y|X] - Y)^2|X] \end{aligned} \right]$$

$$= E \left[\begin{aligned} &E[(f(X) - E[Y|X])^2|X] \\ &+ 2(f(X) - E[Y|X]) \times 0 \\ &+ E[(E[Y|X] - Y)^2|X] \end{aligned} \right]$$

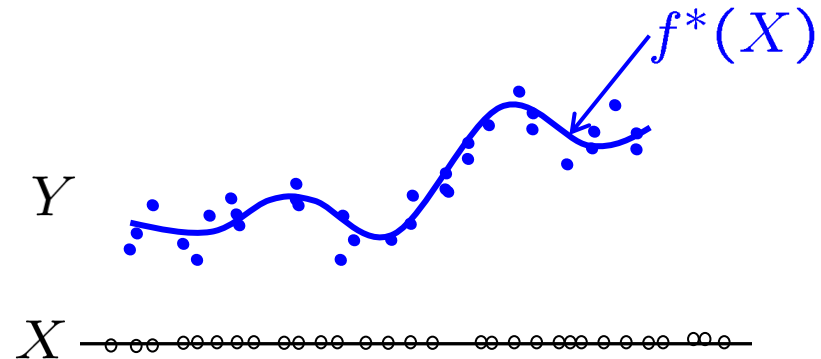
$$= \underbrace{E[(f(X) - E[Y|X])^2]}_{\geq 0} + R(f^*).$$

Regression

Optimal predictor: $f^* = \arg \min_f \mathbb{E}[(f(X) - Y)^2]$
 $= \mathbb{E}[Y|X]$ (Conditional Mean)

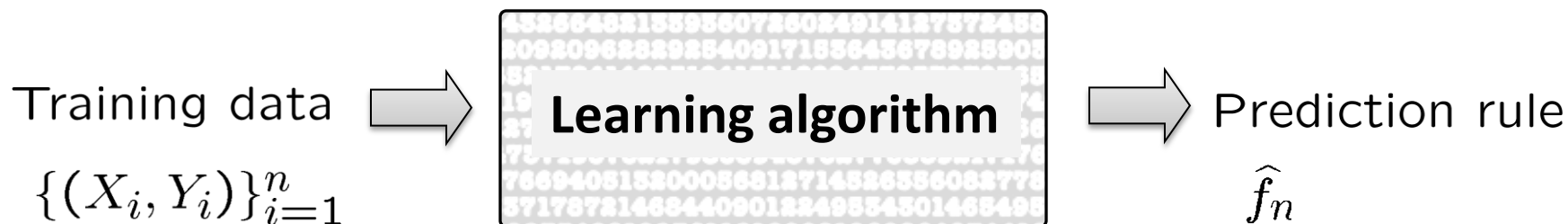
Intuition: Signal plus (zero-mean) Noise model

$$Y = f^*(X) + \epsilon$$



Depends on **unknown** distribution P_{XY}

Regression algorithms



Linear Regression

Lasso, Ridge regression (Regularized Linear Regression)

Nonlinear Regression

Kernel Regression

Regression Trees, Splines, Wavelet estimators, ...

Empirical Risk Minimization (ERM)

Optimal predictor: $f^* = \arg \min_f \mathbb{E}[(f(X) - Y)^2]$

Empirical Risk Minimizer: $\hat{f}_n = \arg \min_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2$

Class of predictors Empirical mean

$$\frac{1}{n} \sum_{i=1}^n [\text{loss}(Y_i, f(X_i))] \xrightarrow[\text{Numbers}]{\text{Law of Large Numbers}} \mathbb{E}_{XY} [\text{loss}(Y, f(X))]$$

More later...

ERM – you saw it before!

- Learning Distributions

Max likelihood = Min -ve log likelihood empirical risk

$$\max_{\theta} P(D|\theta) = \min_{\theta} \frac{1}{n} \sum_{i=1}^n \underbrace{-\log P(X_i|\theta)}_{\text{loss}(X_i, \theta)} \quad \begin{array}{l} \text{Negative log} \\ \text{Likelihood loss} \end{array}$$

What is the class \mathcal{F} ?

Class of parametric distributions

Bernoulli (θ)

Gaussian (μ, σ^2)

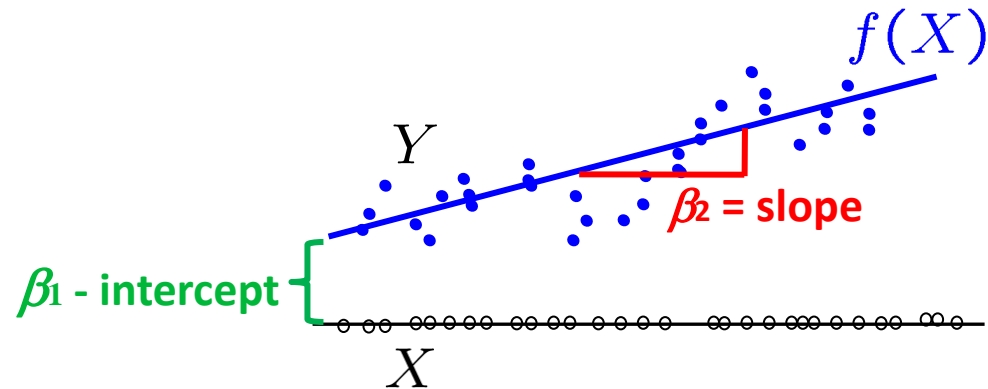
Linear Regression

$$\hat{f}_n^L = \arg \min_{f \in \mathcal{F}_L} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2 \quad \text{Least Squares Estimator}$$

\mathcal{F}_L - Class of Linear functions

Uni-variate case:

$$f(X) = \beta_1 + \beta_2 X$$



Multi-variate case:

$$f(X) = f(X^{(1)}, \dots, X^{(p)}) = \beta_1 X^{(1)} + \beta_2 X^{(2)} + \dots + \beta_p X^{(p)}$$

$$= X\beta \quad \text{where} \quad X = [X^{(1)} \dots X^{(p)}], \quad \beta = [\beta_1 \dots \beta_p]^T$$

Least Squares Estimator

$$\hat{f}_n^L = \arg \min_{f \in \mathcal{F}_L} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2$$



$$\hat{\beta} = \arg \min_{\beta} \frac{1}{n} \sum_{i=1}^n (X_i \beta - Y_i)^2$$

$$\hat{f}_n^L(X) = X \hat{\beta}$$

$$= \arg \min_{\beta} \frac{1}{n} (\mathbf{A} \beta - \mathbf{Y})^T (\mathbf{A} \beta - \mathbf{Y})$$

$$\mathbf{A} = \begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix} = \begin{bmatrix} X_1^{(1)} & \dots & X_1^{(p)} \\ \vdots & \ddots & \vdots \\ X_n^{(1)} & \dots & X_n^{(p)} \end{bmatrix} \quad \mathbf{Y} = \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix}$$

Least Squares Estimator

$$\hat{\beta} = \arg \min_{\beta} \frac{1}{n} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) = \arg \min_{\beta} J(\beta)$$

$$J(\beta) = (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y})$$

$$\left. \frac{\partial J(\beta)}{\partial \beta} \right|_{\hat{\beta}} = 0$$

Normal Equations

$$\underbrace{(\mathbf{A}^T \mathbf{A})}_{p \times p} \underbrace{\hat{\beta}}_{p \times 1} = \underbrace{\mathbf{A}^T \mathbf{Y}}_{p \times 1}$$

If $(\mathbf{A}^T \mathbf{A})$ is invertible,

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y} \qquad \hat{f}_n^L(X) = X \hat{\beta}$$

When is $(\mathbf{A}^T \mathbf{A})$ invertible ?

Recall: Full rank matrices are invertible. What is rank of $(\mathbf{A}^T \mathbf{A})$?

What if $(\mathbf{A}^T \mathbf{A})$ is not invertible ?

Regularization (later)

Geometric Interpretation

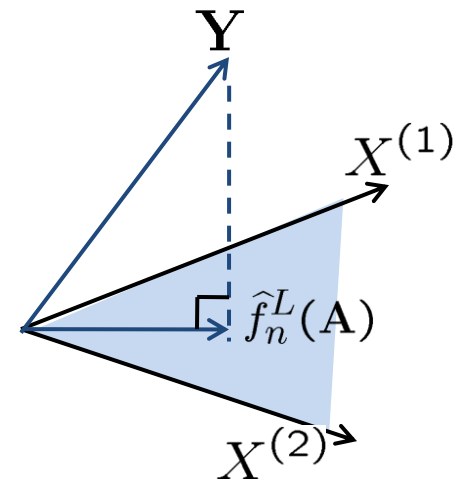
$$\hat{f}_n^L(X) = X\hat{\beta} = X(\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y}$$

Difference in prediction on training set:

$$\hat{f}_n^L(\mathbf{A}) - \mathbf{Y} =$$

$$\mathbf{A}^T (\hat{f}_n^L(\mathbf{A}) - \mathbf{Y}) = 0$$

$\hat{f}_n^L(\mathbf{A})$ is the orthogonal projection of \mathbf{Y} onto the linear subspace spanned by the columns of \mathbf{A}



Revisiting Gradient Descent

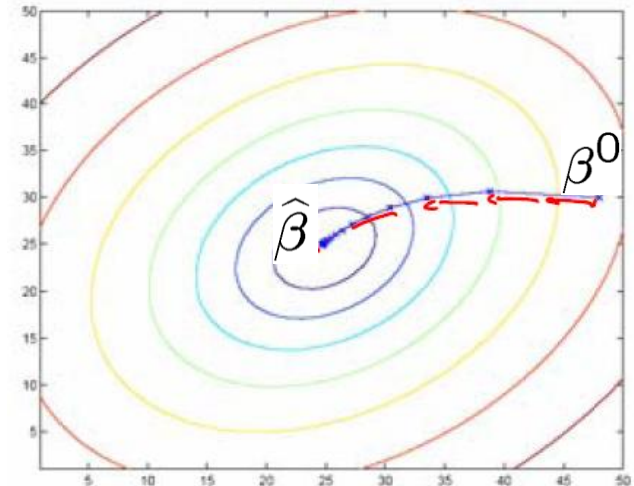
Even when $(\mathbf{A}^T \mathbf{A})$ is invertible, might be computationally expensive if \mathbf{A} is huge.

$$\hat{\beta} = \arg \min_{\beta} \frac{1}{n} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) = \arg \min_{\beta} J(\beta)$$

Gradient Descent since $J(\beta)$ is convex

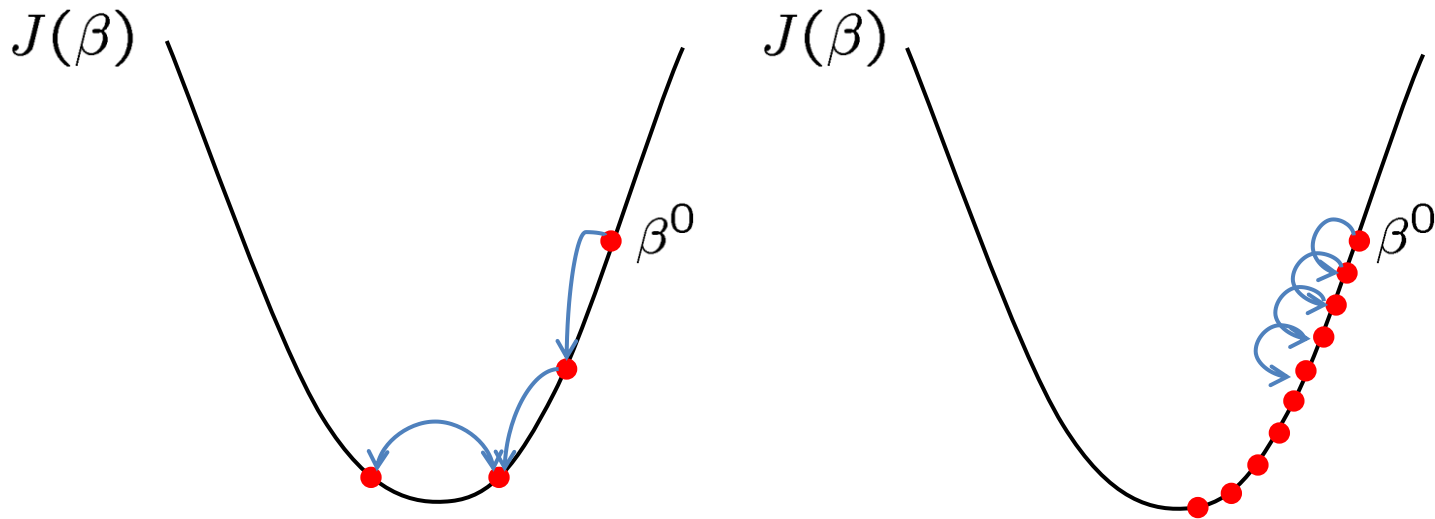
Initialize: β^0

$$\begin{aligned} \text{Update: } \beta^{t+1} &= \beta^t - \frac{\alpha}{2} \frac{\partial J(\beta)}{\partial \beta} \bigg|_t \\ &= \beta^t - \alpha \underbrace{\mathbf{A}^T (\mathbf{A}\beta^t - \mathbf{Y})}_{0 \text{ if } \beta^t = \hat{\beta}} \end{aligned}$$



Stop: when some criterion met e.g. fixed # iterations, or $\left. \frac{\partial J(\beta)}{\partial \beta} \right|_{\beta^t} < \epsilon$.

Effect of step-size α



Large $\alpha \Rightarrow$ Fast convergence but larger residual error
Also possible oscillations

Small $\alpha \Rightarrow$ Slow convergence but small residual error

Least Squares and MLE

Intuition: Signal plus (zero-mean) Noise model

$$Y = f^*(X) + \epsilon = X\beta^* + \epsilon \quad \epsilon \sim \mathcal{N}(0, \sigma^2 \mathbf{I})$$

$$Y \sim \mathcal{N}(X\beta^*, \sigma^2 \mathbf{I})$$

$$\hat{\beta}_{\text{MLE}} = \arg \max_{\beta} \underbrace{\log p(\{(X_i, Y_i)\}_{i=1}^n | \beta, \sigma^2)}_{\text{log likelihood}}$$

$$= \arg \min_{\beta} \sum_{i=1}^n (X_i \beta - Y_i)^2 = \hat{\beta}$$

Least Square Estimate is same as Maximum Likelihood Estimate under a Gaussian model !

Regularized Least Squares and MAP

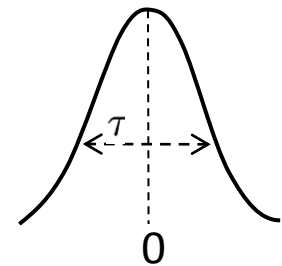
What if $(\mathbf{A}^T \mathbf{A})$ is not invertible ?

$$\hat{\beta}_{\text{MAP}} = \arg \max_{\beta} \underbrace{\log p(\{(X_i, Y_i)\}_{i=1}^n | \beta, \sigma^2)}_{\text{log likelihood}} + \underbrace{\log p(\beta)}_{\text{log prior}}$$

I) Gaussian Prior

$$\beta \sim \mathcal{N}(0, \tau^2 \mathbf{I})$$

$$p(\beta) \propto e^{-\beta^T \beta / 2\tau^2}$$



$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \lambda \|\beta\|_2^2$$

Ridge Regression

Closed form: HW

\downarrow
constant(σ^2, τ^2)

Prior belief that β is Gaussian with zero-mean biases solution to “small” β

Regularized Least Squares and MAP

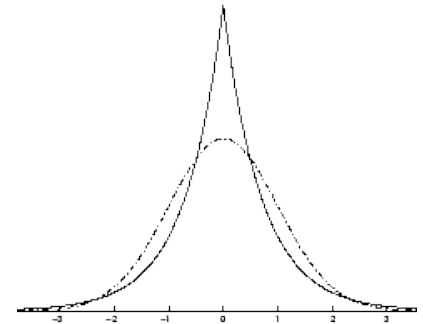
What if $(\mathbf{A}^T \mathbf{A})$ is not invertible ?

$$\hat{\beta}_{\text{MAP}} = \arg \max_{\beta} \underbrace{\log p(\{(X_i, Y_i)\}_{i=1}^n | \beta, \sigma^2)}_{\text{log likelihood}} + \underbrace{\log p(\beta)}_{\text{log prior}}$$

II) Laplace Prior

$$\beta_i \stackrel{iid}{\sim} \text{Laplace}(0, t)$$

$$p(\beta_i) \propto e^{-|\beta_i|/t}$$



$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \underbrace{\lambda \|\beta\|_1}_{\text{constant}(\sigma^2, t)}$$

Lasso

Prior belief that β is Laplace with zero-mean biases solution to “small” β

Ridge Regression vs Lasso

$$\min_{\beta} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) + \lambda \text{pen}(\beta) = \min_{\beta} J(\beta) + \lambda \text{pen}(\beta)$$

Ridge Regression:

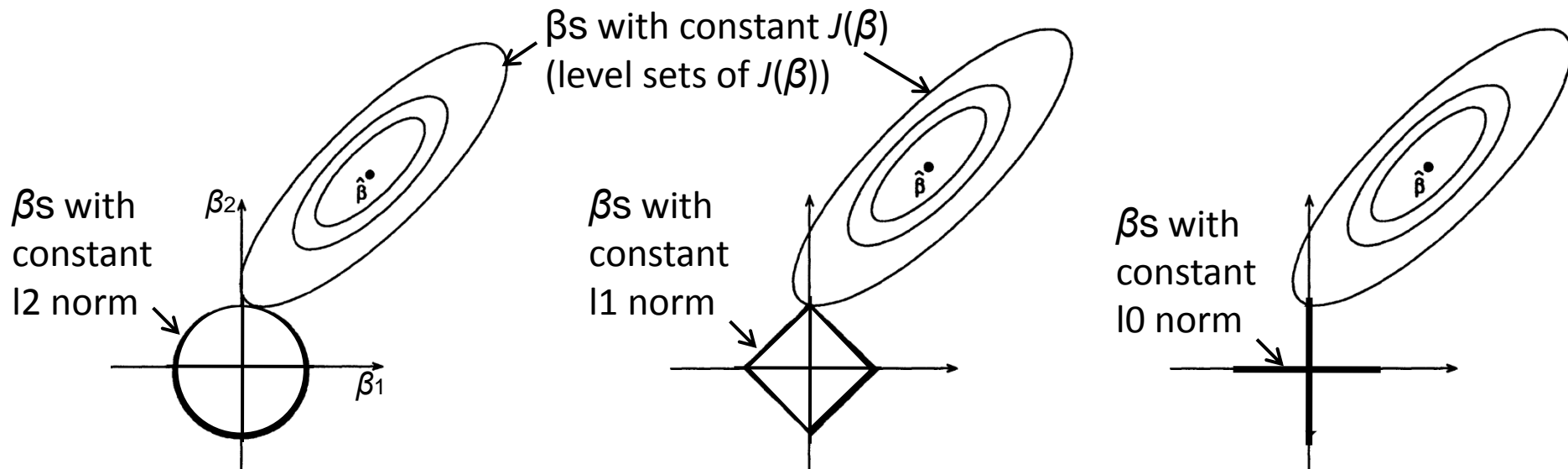
$$\text{pen}(\beta) = \|\beta\|_2^2$$

Lasso:

$$\text{pen}(\beta) = \|\beta\|_1$$

HOT!

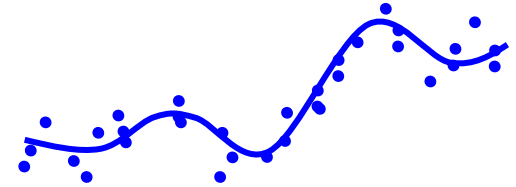
Ideally l0 penalty,
but optimization
becomes non-convex



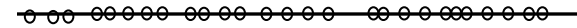
Lasso (l1 penalty) results in sparse solutions – vector with more zero coordinates
Good for high-dimensional problems – don't have to store all coordinates!

Beyond Linear Regression

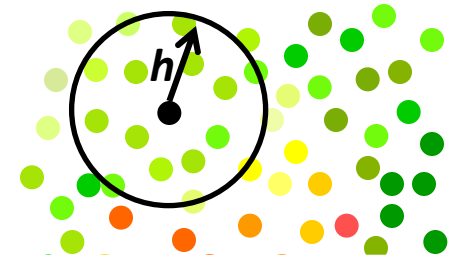
Polynomial regression



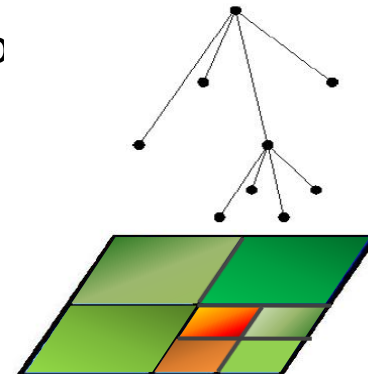
Regression with nonlinear features/basis functions



Kernel regression - Local/Weighted regression



Regression trees – Spatially adaptive regression



Polynomial Regression

Univariate (1-d) case: $f(X) = \beta_0 + \beta_1 X + \beta_2 X^2 + \dots + \beta_m X^m = \mathbf{X}\beta$

where $\mathbf{X} = [1 \ X \ X^2 \ \dots \ X^m]$, $\beta = [\beta_1 \ \dots \ \beta_m]^T$

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y}$$

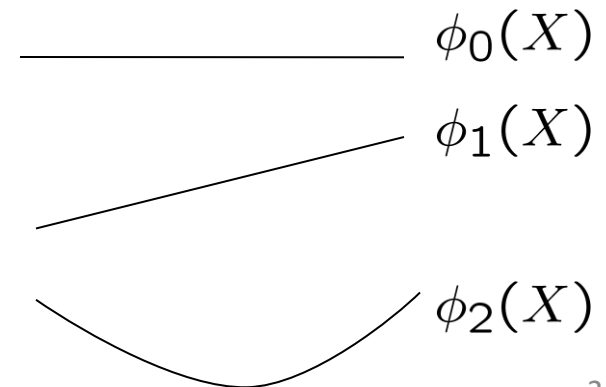
$$\mathbf{A} = \begin{bmatrix} 1 & X_1 & X_1^2 & \dots & X_1^m \\ \vdots & & & \ddots & \vdots \\ 1 & X_n & X_n^2 & \dots & X_n^m \end{bmatrix}$$

$$\hat{f}_n(X) = \mathbf{X}\hat{\beta}$$

$$f(X) = \sum_{j=0}^m \beta_j X^j = \sum_{j=0}^m \beta_j \phi_j(X)$$

Weight of
each feature

Nonlinear
features



Polynomial Regression

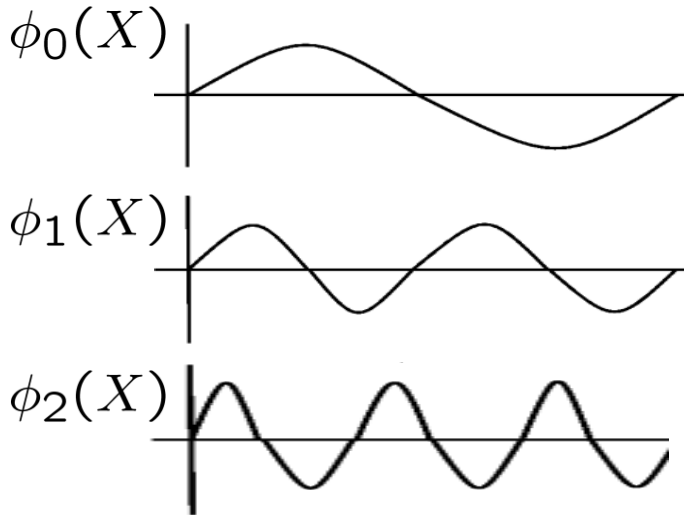
<http://mste.illinois.edu/users/exner/java.f/leastquares/>

Nonlinear Regression

$$f(X) = \sum_{j=0}^m \beta_j \phi_j(X)$$

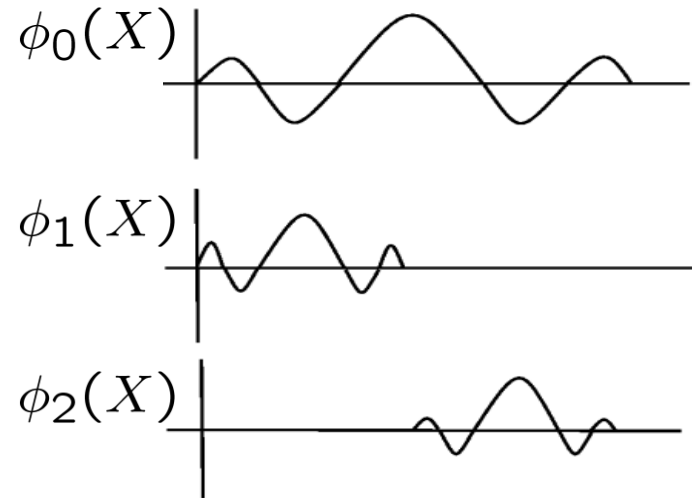
Basis coefficients \leftarrow Nonlinear features/basis functions

Fourier Basis



Good representation for oscillatory functions

Wavelet Basis

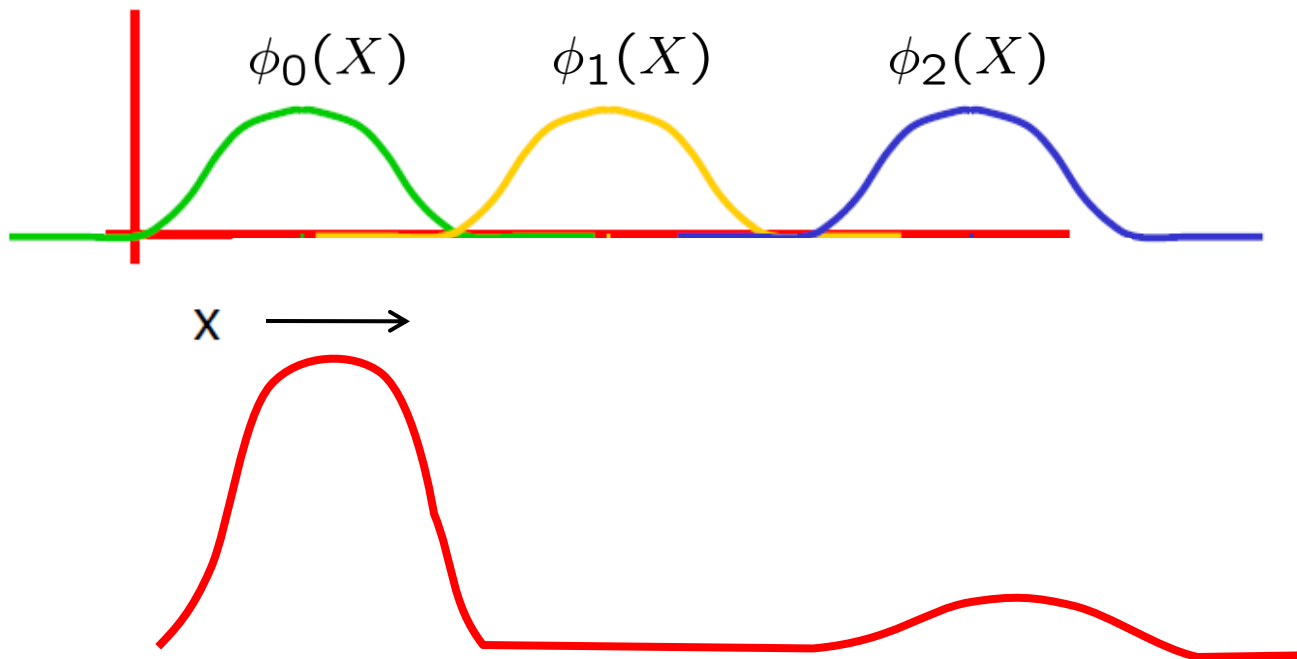


Good representation for functions localized at multiple scales

Local Regression

$$f(X) = \sum_{j=0}^m \beta_j \phi_j(X)$$

Basis coefficients \leftarrow Nonlinear features/basis functions

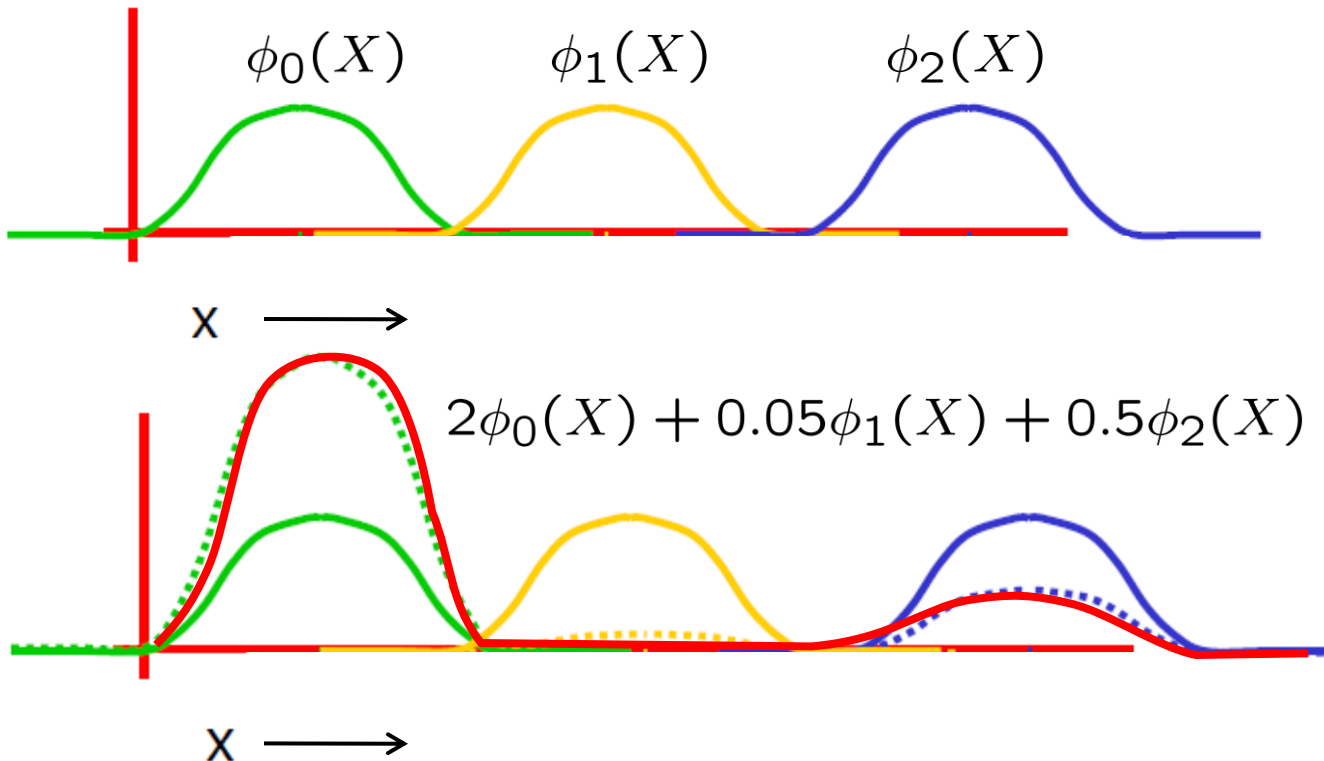


Globally supported
basis functions
(polynomial, fourier)
will not yield a good
representation

Local Regression

$$f(X) = \sum_{j=0}^m \beta_j \phi_j(X)$$

Basis coefficients \leftarrow Nonlinear features/basis functions



Globally supported basis functions (polynomial, fourier) will not yield a good representation

What you should know

Linear Regression

- Least Squares Estimator

- Normal Equations

- Gradient Descent

- Geometric and Probabilistic Interpretation (connection to MLE)

Regularized Linear Regression (connection to MAP)

- Ridge Regression, Lasso

Polynomial Regression, Basis (Fourier, Wavelet) Estimators

Next time

- Kernel Regression (Localized)
- Regression Trees