

# How to choose the right sketching matrix S? [S]

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- S is a Subsampled Randomized Hadamard Transform
  - $S = P^*H^*D$
  - D is a diagonal matrix with +1, -1 on diagonals
  - H is the Hadamard matrix:  $H_{i,j} = (-1)^{\langle i,j \rangle} / n^{.5}$
  - P just chooses a random (small) subset of rows of  $H^*D$
  - $S^*A$  can be computed in  $O(nd \log n)$  time

*Why does it work?*

# Why does this work?

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- We can again assume columns of  $A$  are orthonormal
- It suffices to show  $|SAx|_2^2 = |PHDAx|_2^2 = 1 \pm \epsilon$  for all  $x$
- $HD$  is a rotation matrix, so  $|HDAx|_2^2 = |Ax|_2^2 = 1$  for any  $x$ 
  - Notation: let  $y = Ax$
- Flattening Lemma: For any fixed  $y$ ,

$$\Pr [ |HDy|_\infty \geq C \frac{\log^{.5}(\frac{nd}{\delta})}{n^{.5}} ] \leq \frac{\delta}{2d}$$

# Proving the Flattening Lemma

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- **Flattening Lemma:**  $\Pr [|\text{HDy}|_\infty \geq C \frac{\log^5 nd/\delta}{n^5}] \leq \frac{\delta}{2d}$
- Let  $C > 0$  be a constant. We will show for a fixed  $i$  in  $[n]$ ,

$$\Pr [|(HDy)_i| \geq C \frac{\log^5 nd/\delta}{n^5}] \leq \frac{\delta}{2nd}$$

- If we show this, we can apply a union bound over all  $i$
- $|(HDy)_i| = \sum_j H_{i,j} D_{j,j} y_j$
- (Azuma-Hoeffding) For independent zero-mean random variables  $Z_j$ :

$$\Pr [|\sum_j Z_j| > t] \leq 2e^{-\frac{t^2}{2\sum_j \beta_j^2}}, \text{ where } |Z_j| \leq \beta_j \text{ with probability 1}$$

- $Z_j = H_{i,j} D_{j,j} y_j$  has 0 mean
- $|Z_j| \leq \frac{|y_j|}{n^5} = \beta_j$  with probability 1
- $\sum_j \beta_j^2 = \frac{1}{n}$

- $\Pr \left[ |\sum_j Z_j| > \frac{C \log^5(\frac{nd}{\delta})}{n^5} \right] \leq 2e^{-\frac{C^2 \log^2(\frac{nd}{\delta})}{2}} \leq \frac{\delta}{2nd}$

# Consequence of the Flattening Lemma

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- Recall columns of  $A$  are orthonormal
- HDA has orthonormal columns
- Flattening Lemma implies  $|HDAe_i|_\infty \leq C \frac{\log^5 nd/\delta}{n^{.5}}$  with probability  $1 - \frac{\delta}{2d}$  for a fixed  $i \in [d]$
- With probability  $1 - \frac{\delta}{2}$ ,  $|e_j HDAe_i| \leq C \frac{\log^5 nd/\delta}{n^{.5}}$  for all  $i, j$
- Given this,  $|e_j HDA|_2 \leq C \frac{d^{.5} \log^5 nd/\delta}{n^{.5}}$  for all  $j$

(Can be optimized further)

# Matrix Chernoff Bound

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- Let  $X_1, \dots, X_s$  be independent copies of a symmetric random matrix  $X \in \mathbb{R}^{d \times d}$  with  $E[X] = 0$ ,  $|X|_2 \leq \gamma$ , and  $|E[X^T X]|_2 \leq \sigma^2$ . Let  $W = \frac{1}{s} \sum_{i \in [s]} X_i$ . For any  $\epsilon > 0$ ,

$$\Pr[|W|_2 > \epsilon] \leq 2d \cdot e^{-s\epsilon^2 / (\sigma^2 + \frac{\gamma\epsilon}{3})}$$

(here  $|W|_2 = \sup |Wx|_2 / |x|_2$ )

- Let  $V = HDA$ , and recall  $V$  has orthonormal columns
- Suppose  $P$  in the  $S = \text{PHD}$  definition samples  $s$  rows uniformly with replacement. If row  $i$  is sampled in the  $j$ -th sample,  $P_{j,i} = \frac{\sqrt{n}}{\sqrt{s}}$ , and is 0 otherwise
- Let  $Y_i$  be the  $i$ -th sampled row of  $V = HDA$
- Let  $X_i = I_d - n \cdot Y_i^T Y_i$ 
  - $E[X_i] = I_d - n \cdot \sum_j \left(\frac{1}{n}\right) V_j^T V_j = I_d - V^T V = 0^{d \times d}$
  - $|X_i|_2 \leq |I_d|_2 + n \cdot \max |e_j HDA|_2^2 = 1 + n \cdot C^2 \log\left(\frac{nd}{\delta}\right) \cdot \frac{d}{n} = \Theta\left(d \log\left(\frac{nd}{\delta}\right)\right)$

# Matrix Chernoff Bound

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- Recall: let  $Y_i$  be the  $i$ -th sampled row of  $V = HDA$
- Let  $X_i = I_d - n \cdot Y_i^T Y_i$
- $$E[X^T X + I_d] = I_d + I_d - 2n E[Y_i^T Y_i] + n^2 E[Y_i^T Y_i Y_i^T Y_i]$$

$$= 2I_d - 2I_d + n^2 \sum_i \left(\frac{1}{n}\right) \cdot v_i^T v_i v_i^T v_i = n \sum_i v_i^T v_i \cdot |v_i|_2^2$$
- Define  $Z = n \sum_i v_i^T v_i C^2 \log\left(\frac{nd}{\delta}\right) \cdot \frac{d}{n} = C^2 d \log\left(\frac{nd}{\delta}\right) I_d$
- Note that  $E[X^T X + I_d]$  and  $Z$  are real symmetric, with non-negative eigenvalues
- Claim: for all vectors  $y$ , we have:  $y^T E[X^T X + I_d] y \leq y^T Z y$
- Proof:  $y^T E[X^T X + I_d] y = n \sum_i y^T v_i^T v_i y |v_i|_2^2 = n \sum_i \langle v_i, y \rangle^2 |v_i|_2^2$  and
 
$$y^T Z y = n \sum_i y^T v_i^T v_i y C^2 \log\left(\frac{nd}{\delta}\right) \cdot \frac{d}{n} = d \sum_i \langle v_i, y \rangle^2 C^2 \log\left(\frac{nd}{\delta}\right)$$
- Hence,  $|E[X^T X]|_2 \leq |E[X^T X] + I_d|_2 + |I_d|_2 = |E[X^T X + I_d]|_2 + 1$ 

$$\leq |Z|_2 + 1 \leq C^2 d \log\left(\frac{nd}{\delta}\right) + 1$$
- Hence,  $|E[X^T X]|_2 = O\left(d \log\left(\frac{nd}{\delta}\right)\right)$

# Matrix Chernoff Bound

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- Hence,  $|E[X^T X]|_2 = O\left(d \log\left(\frac{nd}{\delta}\right)\right)$
- Recall: (Matrix Chernoff) Let  $X_1, \dots, X_s$  be independent copies of a symmetric random matrix  $X \in \mathbb{R}^{d \times d}$  with  $E[X] = 0$ ,  $|X|_2 \leq \gamma$ , and  $|E[X^T X]|_2 \leq \sigma^2$ . Let  $W = \frac{1}{s} \sum_{i \in [s]} X_i$ . For any  $\epsilon > 0$ ,  $\Pr[|W|_2 > \epsilon] \leq 2d \cdot e^{-s\epsilon^2/(\sigma^2 + \frac{\gamma\epsilon}{3})}$

$$\Pr\left[|I_d - (\text{PHDA})^T(\text{PHDA})|_2 > \epsilon\right] \leq 2d \cdot e^{-s\epsilon^2/(\Theta(d \log(\frac{nd}{\delta})))}$$

- Set  $s = d \log\left(\frac{nd}{\delta}\right) \frac{\log(\frac{d}{\delta})}{\epsilon^2}$ , to make this probability less than  $\frac{\delta}{2}$

# SRHT Wrapup

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- Have shown  $\|I_d - (\text{PHDA})^T(\text{PHDA})\|_2 < \epsilon$  using Matrix Chernoff Bound and with  $s = d \log\left(\frac{nd}{\delta}\right) \frac{\log\left(\frac{d}{\delta}\right)}{\epsilon^2}$  samples
- Implies for every unit vector  $x$ ,  
$$|1 - |\text{PHDA}x|_2^2| = |x^T x - x^T (\text{PHDA})^T (\text{PHDA}) x| < \epsilon,$$
so  $|\text{PHDA}x|_2^2 \in 1 \pm \epsilon$  for all unit vectors  $x$
- Considering the column span of  $A$  adjoined with  $b$ , we can again solve the regression problem
- The time for regression is now only  $O(nd \log n) + \text{poly}\left(\frac{d \log(n)}{\epsilon}\right)$ . Nearly optimal in matrix dimensions ( $n \gg d$ )

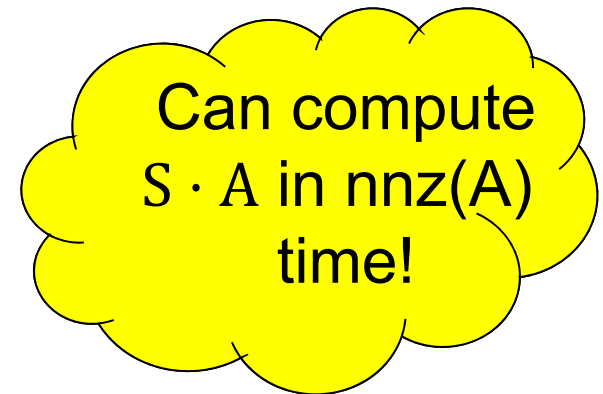


# Faster Subspace Embeddings S [CW,MM,NN]

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- CountSketch matrix
- Define  $k \times n$  matrix  $S$ , for  $k = O(d^2/\epsilon^2)$
- $S$  is really sparse: single randomly chosen non-zero entry per column

$$\begin{bmatrix} 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & -1 & 1 & 0 & -1 & 0 \\ 0 & -1 & 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix}$$



- $\text{nnz}(A)$  is number of non-zero entries of  $A$

# Simple Proof [Nguyen]

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- Can assume columns of  $A$  are orthonormal
- Suffices to show  $|SAx|_2 = 1 \pm \varepsilon$  for all unit  $x$ 
  - For regression, apply  $S$  to  $[A, b]$
- $SA$  is a  $6d^2/(\delta\varepsilon^2) \times d$  matrix
- Suffices to show  $\|A^T S^T SA - I\|_2 \leq \|A^T S^T SA - I\|_F \approx \varepsilon$
- Matrix product result shown below:  
$$\Pr[|CS^TSD - CD|_F^2 \leq [6/(\delta(\# \text{ rows of } S))] * |C|_F^2 |D|_F^2] \geq 1 - \delta$$
- Set  $C = A^T$  and  $D = A$ .
- Then  $|A|_F^2 = d$  and  $(\# \text{ rows of } S) = 6 d^2/(\delta\varepsilon^2)$

# Matrix Product Result [Kane, Nelson]

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- Show:  $\Pr[|CS^TSD - CD|_F^2 \leq [6/(\delta(\# \text{ rows of } S))] * |C|_F^2 |D|_F^2] \geq 1 - \delta$
- (JL Property) A distribution on matrices  $S \in \mathbb{R}^{k \times n}$  has the  $(\epsilon, \delta, \ell)$ -JL moment property if for all  $x \in \mathbb{R}^n$  with  $|x|_2 = 1$ ,
 
$$E_S \left| |Sx|_2^2 - 1 \right|^\ell \leq \epsilon^\ell \cdot \delta$$
- (From vectors to matrices) For  $\epsilon, \delta \in \left(0, \frac{1}{2}\right)$ , let  $D$  be a distribution on matrices  $S$  with  $k$  rows and  $n$  columns that satisfies the  $(\epsilon, \delta, \ell)$ -JL moment property for some  $\ell \geq 2$ . Then for  $A, B$  matrices with  $n$  rows,

$$\Pr_S \left[ |A^T S^T S B - A^T B|_F \geq 3 \epsilon |A|_F |B|_F \right] \leq \delta$$

# From Vectors to Matrices

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- (From vectors to matrices) For  $\epsilon, \delta \in \left(0, \frac{1}{2}\right)$ , let  $D$  be a distribution on matrices  $S$  with  $k$  rows and  $n$  columns that satisfies the  $(\epsilon, \delta, \ell)$ -JL moment property for some  $\ell \geq 2$ . Then for  $A, B$  matrices with  $n$  rows,

$$\Pr_S \left[ \left| A^T S^T S B - A^T B \right|_F \geq 3 \epsilon |A|_F |B|_F \right] \leq \delta$$

- Proof: For a random scalar  $X$ , let  $|X|_p = (E|X|^p)^{1/p}$ 
  - Sometimes consider  $X = |T|_F$  for a random matrix  $T$
  - $\left| |T|_F \right|_p = \left( E \left[ |T|_F^p \right] \right)^{1/p}$
- Can show  $|\cdot|_p$  is a norm if  $p \geq 1$ 
  - Minkowski's Inequality:  $|X + Y|_p \leq |X|_p + |Y|_p$
- For unit vectors  $x, y$ , we will bound  $|\langle Sx, Sy \rangle - \langle x, y \rangle|_\ell$

# Minkowski's Inequality

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- Minkowski's Inequality:  $|X + Y|_p \leq |X|_p + |Y|_p$

- Proof:

- If  $|X|_p, |Y|_p$  are finite, then so is  $|X + Y|_p$ . **Why?**

- $f(x) = x^p$  is convex for  $p \geq 1$ , so for any fixed  $x, y$ :

$$|.5x + .5y|^p \leq |.5x| + |.5y|^p \leq .5|x|^p + .5|y|^p, \text{ so}$$
$$|x + y|^p \leq 2^{p-1}(|x|^p + |y|^p)$$

- So,  $E[|X + Y|_p^p] \leq E[2^{p-1}(|X|_p^p + |Y|_p^p)]$

- $|X + Y|_p^p = \int |x + y|^p d\mu$

$$= \int |x + y| \cdot |x + y|^{p-1} d\mu$$

$$\leq \int (|x| + |y|) |x + y|^{p-1} d\mu$$

$$= \int |x| |x + y|^{p-1} d\mu + \int |y| |x + y|^{p-1} d\mu$$

$$\leq \left( \left( \int |x|^p d\mu \right)^{\frac{1}{p}} + \left( \int |y|^p d\mu \right)^{\frac{1}{p}} \right) \left( \int |x + y|^{(p-1) \left( \frac{p}{p-1} \right)} d\mu \right)^{\frac{p-1}{p}}$$

$$= (|X|_p + |Y|_p) |X + Y|_p^{p-1}$$

# From Vectors to Matrices

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- For unit vectors  $x, y$ ,  $|\langle Sx, Sy \rangle - \langle x, y \rangle|_\ell$

$$= \frac{1}{2} |(|Sx|_2^2 - 1) + (|Sy|_2^2 - 1) - (|S(x-y)|_2^2 - |x-y|_2^2)|_\ell$$

$$\leq \frac{1}{2} (||Sx|_2^2 - 1|_\ell + ||Sy|_2^2 - 1|_\ell + ||S(x-y)|_2^2 - |x-y|_2^2|_\ell)$$

$$\leq \frac{1}{2} (\epsilon \cdot \delta^\ell + \epsilon \cdot \delta^\ell + |x-y|_2^2 \epsilon \cdot \delta^\ell)$$

$$\leq 3 \epsilon \cdot \delta^\ell$$
- By linearity, for arbitrary  $x, y$ ,  $\frac{|\langle Sx, Sy \rangle - \langle x, y \rangle|_\ell}{|x|_2 |y|_2} \leq 3 \epsilon \cdot \delta^\ell$
- Suppose  $A$  has  $d$  columns and  $B$  has  $e$  columns. Let the columns of  $A$  be  $A_1, \dots, A_d$  and the columns of  $B$  be  $B_1, \dots, B_e$
- Define  $X_{i,j} = \frac{1}{|A_i|_2 |B_j|_2} \cdot (\langle SA_i, SB_j \rangle - \langle A_i, B_j \rangle)$
- $$|A^T S^T S B - A^T B|_F^2 = \sum_i \sum_j |A_i|_2^2 \cdot |B_j|_2^2 X_{i,j}^2$$

# From Vectors to Matrices

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- Have shown: for arbitrary  $x, y$ ,  $\frac{|\langle Sx, Sy \rangle - \langle x, y \rangle|}{|x|_2 |y|_2} \leq 3\epsilon \cdot \delta^{\frac{1}{\ell}}$
- For  $X_{i,j} = \frac{1}{|A_i|_2 |B_j|_2} \cdot (\langle SA_i, SB_j \rangle - \langle A_i, B_j \rangle)$ :  $|A^T S^T S B - A^T B|_F^2 = \sum_i \sum_j |A_i|_2^2 \cdot |B_j|_2^2 X_{i,j}^2$
- $$\begin{aligned} ||A^T S^T S B - A^T B|_F^2|_{\ell/2} &= \left| \sum_i \sum_j |A_i|_2^2 \cdot |B_j|_2^2 X_{i,j}^2 \right|_{\ell/2} \\ &\leq \sum_i \sum_j |A_i|_2^2 \cdot |B_j|_2^2 |X_{i,j}^2|_{\ell/2} \\ &= \sum_i \sum_j |A_i|_2^2 \cdot |B_j|_2^2 |X_{i,j}|_{\ell} \\ &\leq \left(3\epsilon\delta^{\frac{1}{\ell}}\right)^2 \sum_i \sum_j |A_i|_2^2 |B_j|_2^2 \\ &= \left(3\epsilon\delta^{\frac{1}{\ell}}\right)^2 |A|_F^2 |B|_F^2 \end{aligned}$$
- Since  $E \left[ |A^T S^T S B - A^T B|_F^{\ell} \right] = \left| |A^T S^T S B - A^T B|_F^2 \right|_{\frac{\ell}{2}}^{\ell/2}$ , by Markov's inequality,
- $\Pr \left[ |A^T S^T S B - A^T B|_F > 3\epsilon |A|_F |B|_F \right] \leq \left( \frac{1}{3\epsilon |A|_F |B|_F} \right)^{\ell} E \left[ |A^T S^T S B - A^T B|_F^{\ell} \right] \leq \delta$

# Result for Vectors

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- Show:  $\Pr[|CS^TSD - CD|_F^2 \leq [6/(\delta(\# \text{ rows of } S))] * |C|_F^2 |D|_F^2] \geq 1 - \delta$
- (JL Property) A distribution on matrices  $S \in \mathbb{R}^{k \times n}$  has the  $(\epsilon, \delta, \ell)$ -JL moment property if for all  $x \in \mathbb{R}^n$  with  $|x|_2 = 1$ ,

$$E_S | |Sx|_2^2 - 1 |^\ell \leq \epsilon^\ell \cdot \delta$$

- (From vectors to matrices) For  $\epsilon, \delta \in (0, \frac{1}{2})$ , let  $D$  be a distribution on matrices  $S$  with  $k$  rows and  $n$  columns that satisfies the  $(\epsilon, \delta, \ell)$ -JL moment property for some  $\ell \geq 2$ . Then for  $A, B$  matrices with  $n$  rows,

$$\Pr_S \left[ |A^T S^T S B - A^T B|_F \geq 3 \epsilon |A|_F |B|_F \right] \leq \delta$$

- Just need to show that the CountSketch matrix  $S$  satisfies JL property and bound the number  $k$  of rows



# CountSketch Satisfies the JL Property

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- (JL Property) A distribution on matrices  $S \in \mathbb{R}^{k \times n}$  has the  $(\epsilon, \delta, \ell)$ -JL moment property if for all  $x \in \mathbb{R}^n$  with  $\|x\|_2 = 1$ ,

$$\mathbb{E}_S \left| \|Sx\|_2^2 - 1 \right|^\ell \leq \epsilon^\ell \cdot \delta$$

- We show this property holds with  $\ell = 2$ . First, let us consider  $\ell = 1$

- For CountSketch matrix  $S$ , let

- $h: [n] \rightarrow [k]$  be a 2-wise independent hash function
- $\sigma: [n] \rightarrow \{-1, 1\}$  be a 4-wise independent hash function

- Let  $\delta(E) = 1$  if event  $E$  holds, and  $\delta(E) = 0$  otherwise

- $$\begin{aligned} \mathbb{E}[\|Sx\|_2^2] &= \sum_{j \in [k]} \mathbb{E} \left[ \left( \sum_{i \in [n]} \delta(h(i) = j) \sigma_i x_i \right)^2 \right] \\ &= \sum_{j \in [k]} \sum_{i_1, i_2 \in [n]} \mathbb{E} [\delta(h(i_1) = j) \delta(h(i_2) = j) \sigma_{i_1} \sigma_{i_2}] x_{i_1} x_{i_2} \\ &= \sum_{j \in [k]} \sum_{i \in [n]} \mathbb{E} [\delta(h(i) = j)^2] x_i^2 \\ &= \left( \frac{1}{k} \right) \sum_{j \in [k]} \sum_{i \in [n]} x_i^2 = \|x\|_2^2 \end{aligned}$$

# CountSketch Satisfies the JL Property

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- $$E[|Sx|_2^4] = E[\sum_{j \in [k]} \sum_{j' \in [k]} (\sum_{i \in [n]} \delta(h(i) = j) \sigma_i x_i)^2 (\sum_{i' \in [n]} \delta(h(i') = j') \sigma_{i'} x_{i'})^2] =$$

$$\sum_{j_1, j_2, i_1, i_2, i_3, i_4} E[\sigma_{i_1} \sigma_{i_2} \sigma_{i_3} \sigma_{i_4} \delta(h(i_1) = j_1) \delta(h(i_2) = j_1) \delta(h(i_3) = j_2) \delta(h(i_4) = j_2)] x_{i_1} x_{i_2} x_{i_3} x_{i_4}$$
- We must be able to partition  $\{i_1, i_2, i_3, i_4\}$  into equal pairs
- Suppose  $i_1 = i_2 = i_3 = i_4$ . Then necessarily  $j_1 = j_2$ . Obtain  $\sum_j \frac{1}{k} \sum_i x_i^4 = |x|_4^4$
- Suppose  $i_1 = i_2$  and  $i_3 = i_4$  but  $i_1 \neq i_3$ . Then get  $\sum_{j_1, j_2, i_1, i_3} \frac{1}{k^2} x_{i_1}^2 x_{i_3}^2 = |x|_2^4 - |x|_4^4$
- Suppose  $i_1 = i_3$  and  $i_2 = i_4$  but  $i_1 \neq i_2$ . Then necessarily  $j_1 = j_2$ . Obtain  $\sum_j \frac{1}{k^2} \sum_{i_1, i_2} x_{i_1}^2 x_{i_2}^2 \leq \frac{1}{k} |x|_2^4$ . Obtain same bound if  $i_1 = i_4$  and  $i_2 = i_3$ .
- Hence,  $E[|Sx|_2^4] \in [|x|_2^4, |x|_2^4(1 + \frac{2}{k})] = [1, 1 + \frac{2}{k}]$
- So,  $E_S ||Sx|_2^2 - 1|^2 \leq (1 + \frac{2}{k}) - 2 + 1 = \frac{2}{k}$ . Setting  $k = \frac{2}{\epsilon^2 \delta}$  finishes the proof

# Where are we?

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- (JL Property) A distribution on matrices  $S \in \mathbb{R}^{k \times n}$  has the  $(\epsilon, \delta, \ell)$ -JL moment property if for all  $x \in \mathbb{R}^n$  with  $\|x\|_2 = 1$ ,

$$\mathbb{E}_S \left| \|Sx\|_2^2 - 1 \right|^\ell \leq \epsilon^\ell \cdot \delta$$

- (From vectors to matrices) For  $\epsilon, \delta \in \left(0, \frac{1}{2}\right)$ , let  $D$  be a distribution on matrices  $S$  with  $k$  rows and  $n$  columns that satisfies the  $(\epsilon, \delta, \ell)$ -JL moment property for some  $\ell \geq 2$ . Then for  $A, B$  matrices with  $n$  rows,

$$\Pr_S \left[ \left| \|A^T S^T S B\|_F^2 - \|A^T B\|_F^2 \right| \geq 3 \epsilon^2 \|A\|_F^2 \|B\|_F^2 \right] \leq \delta$$

- We showed CountSketch has the JL property with  $\ell = 2$ , and  $k = \frac{2}{\epsilon^2 \delta}$

- Matrix product result we wanted was:

$$\Pr \left[ \|CS^TSD - CD\|_F^2 \leq \frac{6}{\delta k} \|C\|_F^2 \|D\|_F^2 \right] \geq 1 - \delta$$

- We are now done with the proof CountSketch is a subspace embedding

# Course Outline

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- Subspace embeddings and least squares regression
  - Gaussian matrices
  - Subsampled Randomized Hadamard Transform
  - CountSketch
- **Affine embeddings**
  - Application to low rank approximation
- High precision regression
- Leverage score sampling
- Distributed low rank approximation
- L1 Regression
- M-Estimator regression