

## 1 $p$ -norm Estimation

Recall the sketching matrices  $P \cdot D$ , where  $P$  consists of a CountSketch matrix, and  $D$  consists of a diagonal matrix with diagonal elements  $1/E_i^{1/p}$ , with  $E_i$  being independent standard exponential random variables.

For arbitrary  $y$ ,  $\|Dy\|_\infty$  looks like

$$\|Dy\|_\infty^2 = \max_i \frac{\|y_i\|^p}{E_i} = \frac{1}{\min_i \frac{E_i}{\|y_i\|^p}} \equiv \frac{1}{E/\|y\|_p^p} = \frac{\|y\|_p^p}{E} \quad (1)$$

and the probability of a reasonable value of  $E$  is  $\Pr[E \in [1/10, 10]] = (1 - e^{-10}) - (1 - e^{-1/10}) > 4/5$  (this actually evaluates to just over  $9/10$ ).

As such,  $\|Dy\|_\infty^p$  is a good estimate for  $\|y\|_p^p$ , but  $Dy \in \mathbb{R}^n$  is a large vector, so sketching using matrix  $P \in \mathbb{R}^{s \times n}$  is needed to reduce computation cost.

Intuitively,  $P$  is hashing coordinates of  $Dy$  into buckets and taking a signed sum; most items cancel out and then  $\|PDy\|_\infty \simeq \|Dy\|_\infty$ . It is known previously that  $P$  is composed of hash functions  $h : [n] \rightarrow [s]$  and  $\sigma : [n] \rightarrow \{-1, 1\}$  (assuming they are truly random). Given that  $\|Dy\|_\infty/\|y\|_p \in [1/10^{1/p}, 10^{1/p}]$  with probability  $> 4/5$ , to achieve  $\|PDy\|_\infty \simeq \|Dy\|_\infty$  with good probability, it is necessary to have:

1. in each bucket  $i$  not containing the maximum value,  $|(PDy)_i| \leq \|y\|_p/100$
2. in each bucket  $i$  containing the maximum value,  $\left| |(PDy)_i| - \|Dy\|_\infty \right| \leq \|y\|_p/100$

Let  $\delta(\text{event}) = 1$  if a given event holds and  $\delta(\text{event}) = 0$  otherwise. It is then possible to define a given element of  $PDy$  as  $(PDy)_i = \sum_j \delta(h(j) = i) \cdot \sigma_j \cdot (Dy)_j$ . Due to  $\sigma$ , its expectation is  $\mathbb{E}[(PDy)_i] = 0$ . The evaluation of its variance as follows:

$$\mathbb{E}_P[(PDy)_i^2] = \sum_{j,j'} \mathbb{E}[\delta(h(j) = i) \cdot \delta(h(j') = i) \cdot \sigma_j \cdot \sigma_{j'}] (Dy)_j (Dy)_{j'} = \frac{1}{s} \|Dy\|_2^2 \quad (2)$$

$$\mathbb{E}_D[\|Dy\|_2^2] = \sum_i y_i^2 \mathbb{E}[D_{i,i}^2] \quad (3)$$

$$\mathbb{E}[D_{i,i}^2] = \int_0^\infty t^{\frac{2}{p}} e^{-t} dt = \int_0^1 t^{\frac{2}{p}} e^{-t} dt + \int_1^\infty t^{\frac{2}{p}} e^{-t} dt \quad (4)$$

$$\leq \int_0^1 t^{\frac{2}{p}} dt + \int_1^\infty e^{-t} dt = \left( \frac{1}{1 - \frac{2}{p}} \right) t^{-\frac{2}{p}} \Big|_0^1 - e^{-t} \Big|_1^\infty = O(1) \quad (5)$$

$$\mathbb{E}_P[(PDy)_i^2] = O\left(\frac{1}{s}\right) \|y\|_2^2 = O\left(\frac{1}{s}\right) (n^{1-\frac{2}{p}} \|y\|_p^2). \quad (6)$$

The last line holds due to Hölder's Inequality ( $\|y\|_2^2 = \sum_i^n y_i^2 \leq (\sum_i^n (y_i^2)^{p/2})^{2/p} (\sum_i^n 1^q)^{1/q} = \|y\|_p^2 \cdot n^{1-2/p}$ ).

**Definition** (Bernstein's Bound). Suppose independent random variables  $R_1, \dots, R_n$ , and for all  $j$ ,  $|R_j| \leq K$ , and  $\text{Var}[\sum_j R_j] = \sigma^2$ . Then, there exists constants  $c, C$  such that for all  $t > 0$ ,

$$\Pr \left[ \left| \sum_j R_j - \mathbb{E}[\sum_j R_j] \right| > t \right] \leq C \left( e^{-\frac{ct^2}{\sigma^2}} + e^{-\frac{ct}{K}} \right) \quad (7)$$

In order to get  $1/\text{poly}(n)$  error probability, set  $R_j = \delta(h(j) = i) \dot{\sigma}_j (Dy)_j$ ,  $t = \|y\|_p/100$ , and  $s = \Theta(n^{1-2/p} \log n)$  to handle all parameters required for Bernstein's bound other than  $K$ .

It is possible to treat large  $R_j$  separately, where  $R_j > \frac{\alpha \|y\|_p}{\log n}$  for a sufficiently small  $\alpha > 0$ . If  $|R_j| > \frac{\alpha \|y\|_p}{\log n}$ , then necessarily  $(Dy)_j \geq \frac{\alpha \|y\|_p}{\log n}$  (define  $j$  as "large" if this is the case, "small" otherwise). Then, any  $j$  may be large with probability and expectation

$$\Pr[j \text{ is large}] = \Pr \left[ \frac{|y_j|}{E_j^{1/p}} \leq \frac{\alpha \|y\|_p}{\log n} \right] = \Pr \left[ \frac{|y_j|^p}{\alpha^p \|y\|_p^p} \log^p n \leq E_j \right] \quad (8)$$

$$= 1 - e^{-\frac{|y_j|^p \log^p n}{\alpha^p \|y\|_p^p}} \leq \frac{|y_j|^p \log^p n}{\alpha^p \|y\|_p^p} \quad (9)$$

$$\mathbb{E}[R_j \text{ for large } j] \leq \sum_j \frac{|y_j|^p \log^p n}{\alpha^p \|y\|_p^p} = \frac{\log^p n}{\alpha^p} \quad (10)$$

There are  $s = O(n^{1-2/p} \log n)$  buckets and  $\frac{\log^p n}{\alpha^p}$  items. By Markov bound, there are  $O(\log^p n)$  large  $j$  with constant probability.  $D$  is conditioned on the above as well as  $\|Dy\|_\infty \in [\|y\|_p/10^{1/p}, \|y\|_p \cdot 10^{1/p}]$  (which happens with probability  $> 4/5$ ). All the large  $j$  should then be perfectly hashed into separate buckets by  $P$ . (If there are  $b$  balls and  $Cb$  bins,  $\Pr[\text{collision}] \leq \binom{b}{2} 1/Cb \leq 1/2C$ )

Bernstein's bound can then be applied separately for the small indices  $j$  for each hash bucket.  $\mathbb{E}[(PDy)_i] = 0$  for each hash bucket  $i$ , and  $\mathbb{E}[(PDy)_i^2] = O(1/s)(n^{1-2/p} \|y\|_p^2)$ . Assuming  $K = \max_j |R_j| \leq \alpha \|y\|_p / \log n$  for small  $j$  in a bucket (it can be shown that  $\text{Var}[R_j]$  is  $O(1/s)(n^{1-2/p} \|y\|_p^2)$ )

even if no  $j$  is large. Setting  $t = \|y\|_p/100$  and  $s = \Theta(n^{1-2/p}\log n)$  in Bernstein's bound, for a bucket  $(PDy)_i$

$$\Pr \left[ \left| \sum_{\text{small } j} \delta(h(j) = i) \cdot \sigma_j \cdot (Dy)_i \right| > \frac{\|y\|_p}{100} \right] \leq C \left( e^{-\Theta(\log n)} + e^{-c \frac{\log n}{100\alpha}} \right) \leq \frac{1}{n^2} \quad (11)$$

By union bound over all  $s$  buckets, the signed sum of all small  $j$  in every bucket will be at most  $\|y\|_p/100$ . Therefore, for all  $i$ ,

1. in each bucket  $i$  without large indices  $j$ ,  $|(PDy)_i| \leq \|y\|_p/100$
2. in each bucket  $i$  with one large index  $j$ ,  $|(PDy)_i| = |\sigma_j(Dy)_j| \pm \|y\|_p/100$

and no bucket has more than one large  $j$  as shown in the perfect hashing assumption above. Conditioning on  $\|Dy\|_\infty \in [\|y\|_p/10^{1/p}, \|y\|_p \cdot 10^{1/p}]$ ,

$$\frac{\|y\|_p}{10^{1/p}} - \frac{\|y\|_p}{100} \leq \|PDy\|_\infty \leq 10^{1/p} \cdot \|y\|_p + \frac{\|y\|_p}{100} \quad (12)$$

Therefore, it is reasonable to use  $\|PDy\|_\infty$  as an estimate for  $\|y\|_p$ . The total space used is  $s = O(n^{1-2/p}\log n)$ , i.e.  $O(n^{1-2/p}\log^2 n)$  bits. This space complexity still holds even when considering the pseudorandom generation of matrix  $P$ , see [1].

## 2 Heavy Hitters

$l_1$  **guarantee**: output a set containing all items  $j$  for which  $|x_j| \geq \phi\|x\|_1$ , and the set should not contain any  $j$  with  $|x_j| \leq (\phi - \varepsilon)\|x\|_1$ .

$l_2$  **guarantee**: output a set containing all items  $j$  for which  $x_j^2 \geq \phi\|x\|_2^2$ , and the set should not contain any  $j$  with  $x_j^2 \leq (\phi - \varepsilon)\|x\|_2^2$ . This guarantee is much stronger: suppose  $x = [\sqrt{n}, 1, \dots, 1]$ ,  $\sqrt{n}$  is an  $l_2$ -heavy hitter for constant  $\phi$  and  $\varepsilon$ , but not an  $l_1$ -heavy hitter. Also, if  $|x_j| \geq \phi\|x\|_1$ , it means that  $x_j^2 \geq \phi^2\|x\|_1^2 \geq \phi^2\|x\|_2^2$  as well.

## References

- [1] N. Nisan. Pseudorandom generators for space-bounded computations. In *Proceedings of the Twenty-Second Annual ACM Symposium on Theory of Computing*, STOC '90, page 204â212, New York, NY, USA, 1990. Association for Computing Machinery. doi:10.1145/100216.100242.