

15-251

Great Theoretical Ideas in Computer Science

mr

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Setting Expectations for Computer Scientists

Probability Theory II

Lecture 12 (September 30, 2010)

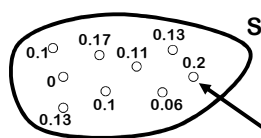


$$E[X+Y] = E[X] + E[Y]$$



A few things from last time...

Sample Space



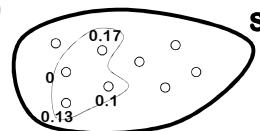
Sample space

weight or
probability of t
 $D(t) = p(t) = 0.2$

Events

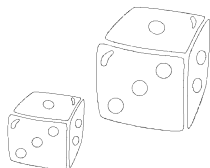
Any set $E \subseteq S$ is called an event

$$\Pr_D[E] = \sum_{t \in E} p(t)$$



$$\Pr_D[E] = 0.4$$

Binomials and Gaussians...



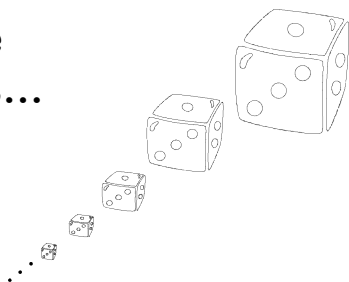
$$\Pr(\text{see } k \text{ heads in } n \text{ flips}) = \binom{n}{k} p^k (1-p)^{n-k}$$

Given a fair coin ($p = 1/2$)

$$\binom{n}{k} 2^{-n}$$

As $n \rightarrow \infty$, the plot for $\Pr(k \text{ heads})$ for $p = 1/2$ tends to “bell curve” or
“Gaussian/normal distribution”

Infinite sample spaces...



The “Geometric” Distribution

A bias- p coin is tossed until the first time that a head turns up.

sample space $S = \{H, TH, TTH, TTTH, \dots\}$
(shorthand $S = \{1, 2, 3, 4, \dots\}$)

$$\Pr(k) = (1-p)^{k-1} p$$

(sanity check) $\sum_{k \geq 1} \Pr(k) = \sum_{k \geq 1} (1-p)^{k-1} p$
 $= p * (1 + (1-p) + (1-p)^2 + \dots)$
 $= p * 1/(1-(1-p)) = 1$

The Geometric Distribution

A bias- p coin is tossed until the first time that a head turns up.

sample space $S = \{1, 2, 3, 4, \dots\}$

$E =$ “first heads at even numbered flip”

$= \{2, 4, 6, \dots\}$ “ $= \{TH, TTTH, TTTTTH, \dots\}$ ”

$$\begin{aligned} \Pr(E) &= \sum_{x \in E} \Pr(x) = \sum_{k \text{ even}} (1-p)^{k-1} p \\ &= p(1-p) * (1 + (1-p)^2 + (1-p)^4 + \dots) \\ &= p(1-p) * 1/(1-(1-p)^2) = (1-p)/(2-p) \end{aligned}$$

Expectations

$$E[\text{die}]$$

“Expectation = (weighted) average”

What is the average height of a 251 student?

Sample space S = all 251 students

(implicit) probability distribution = uniform

average = (sum of heights in S)/ $|S|$

$$= (\sum_{t \in S} \text{Height}(t)) / |S|$$

$$= \sum_{t \in S} \text{Pr}(t) \text{Height}(t)$$

Expectation

Given a probability distribution $D = (S, \text{Pr})$

And any function $f: S \rightarrow \text{reals}$

$$E[f] = \sum_{t \in S} f(t) \text{Pr}(t)$$

A function from the sample space to the reals is called a “random variable”

Random Variable

Let S be sample space in a probability distribution

A Random Variable is a function from S to reals

Examples:

X = value of white die in a two-dice roll

$$X(3,4) = 3, \quad X(1,6) = 1$$

Y = sum of values of the two dice

$$Y(3,4) = 7, \quad Y(1,6) = 7$$

Notational Conventions

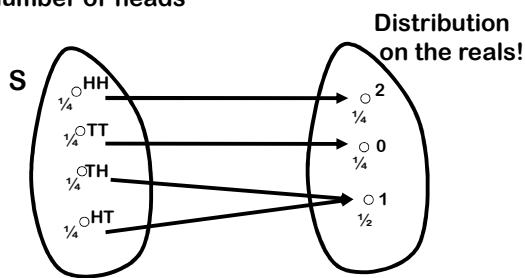
Use letters like A, B, C for events

Use letters like X, Y, f, g for R.V.'s

R.V. = random variable

Two Coins Tossed

$X: \{TT, TH, HT, HH\} \rightarrow \{0, 1, 2\}$ counts the number of heads



Two Views of Random Variables

Think of a R.V. as

Input to the function is random

A function from S to the reals R

Or think of the induced distribution on R

Randomness is “pushed” to the values of the function

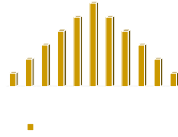
Given a distribution, a random variable transforms it into a distribution on reals

Two dice

I throw a white die and a black die.

Sample space $S =$

$\{(1,1), (1,2), (1,3), (1,4), (1,5), (1,6),$
 $(2,1), (2,2), (2,3), (2,4), (2,5), (2,6),$
 $(3,1), (3,2), (3,3), (3,4), (3,5), (3,6),$
 $(4,1), (4,2), (4,3), (4,4), (4,5), (4,6),$
 $(5,1), (5,2), (5,3), (5,4), (5,5), (5,6),$
 $(6,1), (6,2), (6,3), (6,4), (6,5), (6,6)\}$



$X = \text{sum of both dice}$

function with $X(1,1) = 2, X(1,2) = 3, \dots, X(6,6) = 12$

It's a Floor Wax And a Dessert Topping



It's a function on the sample space S

It's a variable with a probability distribution on its values



You should be comfortable with both views

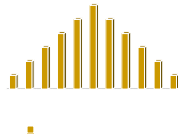


Two dice

I throw a white die and a black die.

Sample space $S =$

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$X = \text{sum of both dice}$

function with $X(1,1) = 2, X(1,2) = 3, \dots, X(6,6) = 12$

From Random Variables to Events

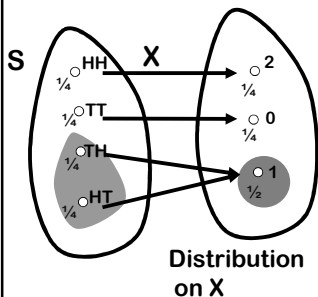
Note that each event in the induced distribution corresponds to some event in the original one.

For any random variable X and value a , we can define the event A that $X = a$

$$\Pr(A) = \Pr(X=a) = \Pr(\{t \in S \mid X(t)=a\})$$

Two Coins Tossed

$X: \{TT, TH, HT, HH\} \rightarrow \{0, 1, 2\}$ counts # of heads



$$\Pr(X=a) = \Pr(\{t \in S \mid X(t)=a\})$$

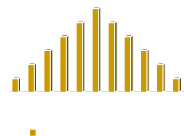
$$\begin{aligned} \Pr(X=1) &= \Pr(\{t \in S \mid X(t)=1\}) \\ &= \Pr(\{TH, HT\}) = \frac{1}{2} \end{aligned}$$

Two dice

I throw a white die and a black die.

Sample space $S =$

$\{(1,1), (1,2), (1,3), (1,4), (1,5), (1,6),$
 $(2,1), (2,2), (2,3), (2,4), (2,5), (2,6),$
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 $(6,1), (6,2), (6,3), (6,4), (6,5), (6,6)\}$



$X = \text{sum of both dice}$

$$\Pr(X=7) = \frac{1}{6} = \frac{6}{36}$$

Definition: Expectation

The expectation, or expected value of a random variable X is written as $E[X]$, and is

$$E[X] = \sum_{t \in S} \Pr(t) X(t) = \sum_k k \Pr[X = k]$$

X is a function
on the sample space S

X has a prob.
distribution on
its values

(assuming X takes values in the naturals)

Quick check: $\sum_{t \in S} \Pr(t) X(t) = \sum_k k \Pr[X = k]$

A Quick Calculation...

What if I flip a coin 3 times? What is the expected number of heads?

$$E[X] = \frac{1}{8} \times 0 + \frac{3}{8} \times 1 + \frac{3}{8} \times 2 + \frac{1}{8} \times 3 = 1.5$$

But $\Pr[X = 1.5] = 0$

Moral: don't always expect the expected.
 $\Pr[X = E[X]]$ may be 0!

Type Checking



$\downarrow$
 $P(B)$ B must be an event

$E(X)$ X must be a R.V.

cannot do $P(\text{R.V.})$ or $E(\text{event})$

Operations on R.V.s

You can sum them, take differences,
or do most other math operations...

$$\text{E.g., } (X + Y)(t) = X(t) + Y(t)$$

$$(X * Y)(t) = X(t) * Y(t)$$

$$(X^Y)(t) = X(t)^{Y(t)}$$

**Random variables
and expectations
allow us to give elegant
solutions to
problems that seem
really really messy...**

If I randomly put 100 letters into 100 addressed envelopes, on average how many letters will end up in their correct envelopes?

On average, in class of size m , how many pairs of people will have the same birthday?



Pretty messy with direct counting...

The new tool is called
“Linearity of Expectation”

Linearity of Expectation

If $Z = X + Y$, then

$$E[Z] = E[X] + E[Y]$$

Even if X and Y are not independent



$$\begin{aligned} E[Z] &= \sum_{t \in S} \Pr[t] Z(t) \\ &= \sum_{t \in S} \Pr[t] (X(t) + Y(t)) \\ &= \sum_{t \in S} \Pr[t] X(t) + \sum_{t \in S} \Pr[t] Y(t) \\ &= E[X] + E[Y] \end{aligned}$$

By Induction

$$E[X_1 + X_2 + \dots + X_n] =$$

$$E[X_1] + E[X_2] + \dots + E[X_n]$$

The expectation
of the sum
=
The sum of the
expectations



Expectation of a Sum of r.v.s

= Sum of their Expectations

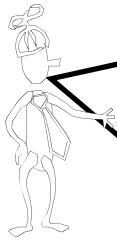
even when r.v.s are not independent!

Expectation of a Product of r.v.s

False in general = Product of their Expectations

only when r.v.s are independent!

Independence for r.v.s



Two random variables X and Y are independent if for every a, b , the events $X=a$ and $Y=b$ are independent

How about the case of
 $X=1\text{st die}$, $Y=2\text{nd die}$?
 $X = \text{left arm}$, $Y=\text{right arm}$?

Let's test our
 Linearity of Expectation
 chops...

If I randomly put 100 letters
 into 100 addressed
 envelopes, on average how
 many letters will end up in
 their correct envelopes?



Hmm...

$\sum_k k \Pr(\text{exactly } k \text{ letters end up in correct envelopes})$

$= \sum_k k (\dots \text{aargh!!} \dots)$



Use Linearity of Expectation

Let A_i be the event the i^{th} letter ends up in its correct envelope

Let X_i be the "indicator" R.V. for A_i

$$X_i = \begin{cases} 1 & \text{if } A_i \text{ occurs} \\ 0 & \text{otherwise} \end{cases}$$

Let $Z = X_1 + \dots + X_{100}$

We are asking for $E[Z] = \sum_i E[X_i]$

$E[X_i] = \Pr(A_i) = 1/100$

So $E[Z] = 1$



So, in expectation, 1 letter will be
 in the same correct envelope

Pretty neat: it doesn't depend on
 how many letters!

Question: were the X_i independent?

No! E.g., think of $n=2$



Use Linearity of Expectation

General approach:

View thing you care about as
 expected value of some RV

Write this RV as sum of
 simpler RVs

Solve for their expectations
 and add them up!



Example #2

We flip n coins of bias p . What is the expected number of heads?

We could do this by summing

$$\sum_k k \Pr(X = k) = \sum_k k \binom{n}{k} p^k (1-p)^{n-k} = n \cdot p$$

But now we know a better way!



Linearity of Expectation!

Let X = number of heads when n independent coins of bias p are flipped

Break X into n simpler RVs:

$X = \sum_i X_i$

$$X_i = \begin{cases} 1 & \text{if the } i^{\text{th}} \text{ coin is } \text{heads} \\ 0 & \text{if the } i^{\text{th}} \text{ coin is } \text{tails} \end{cases}$$

$$E[X] = E[\sum_i X_i] = \sum_i E[X_i] = \sum_i p = np$$

On average, in class of size m , how many pairs of people will have the same birthday?

$$\sum_k k \Pr(\text{exactly } k \text{ collisions}) = \sum_k k (\dots \text{aargh!!!!} \dots)$$



Use linearity of expectation

Suppose we have m people each with a uniformly chosen birthday from 1 to 366

X = number of pairs of people with the same birthday

$$E[X] = ?$$



X = number of pairs of people with the same birthday

$$E[X] = ?$$

Use $m(m-1)/2$ indicator variables, one for each pair of people

$X_{jk} = 1$ if person j and person k have the same birthday; else 0

$$E[X_{jk}] = (1/366) \cdot 1 + (1 - 1/366) \cdot 0 = 1/366$$



X = number of pairs of people with the same birthday

$X_{jk} = 1$ if person j and person k have the same birthday; else 0

$$E[X_{jk}] = (1/366) \cdot 1 + (1 - 1/366) \cdot 0 = 1/366$$

$$E[X] = E[\sum_{1 \leq j < k \leq m} X_{jk}]$$

$$= \sum_{1 \leq j < k \leq m} E[X_{jk}]$$

$$= m(m-1)/2 \times 1/366$$



E.g., setting $m = 23$, we get
 $E[\# \text{ pairs with same birthday }]$
 $= 0.691\dots$

How does this compare to
 $\Pr[\text{at least one pair has same birthday}]?$

You pick a number $n \in [1..6]$.

You roll 3 dice.

If any of them match n , you win \$1.

Else you pay me \$1. Want to play?

$X_i = 1$ if i -th die matches, 0 otherwise

Expected number of dice that match:

$E[X_1 + X_2 + X_3] = 1/6 + 1/6 + 1/6 = 1/2.$

BUT not same as $\Pr(\text{at least one die matches}).$

$\Pr(\text{at least one die matches}) = 1 - \Pr(\text{none match})$
 $= 1 - (5/6)^3 = 0.416.$

Back to Lin. of Exp.

Using linearity of expectations
in unexpected places...

A Puzzle

10% of the surface of a sphere is colored green, and the rest is colored blue. Show that no matter how the colors are arranged, it is possible to inscribe a cube in the sphere so that all of its vertices are blue.



Solution

Pick a random cube. (Note: any particular vertex is uniformly distributed over surface of sphere).

Let $X_i = 1$ if i th vertex is blue, 0 otherwise

Let $X = X_1 + X_2 + \dots + X_8$

$E[X_i] = P(X_i=1) = 9/10$

$E[X] = 8 \cdot 9/10 > 7$

So, must have some cubes where $X = 8$.

The general principle we used here

Show the expected value of some random variable is “high”

Hence, there must be an outcome in the sample space where the random variable takes on a “high” value.

(Not everyone can be below the average.)

called “the probabilistic method”

Some supplementary material

linearity of expectation:
it holds even when the
random variables are
not independent

Linearity of Expectation

E.g., 2 fair flips:

X = heads in 1st coin,

Y = heads in 2nd coin

Z = X+Y = total # heads

What is $E[X]$? $E[Y]$? $E[Z]$?



	1,1,2	
1,0,1	HH	0,1,1
HT	0,0,0	TH
	TT	

Linearity of Expectation

E.g., 2 fair flips:

X = at least one coin is heads

Y = both coins are heads, Z = X+Y

Are X and Y independent?

What is $E[X]$? $E[Y]$? $E[Z]$?



	1,1,2	
1,0,1	HH	1,0,1
HT	0,0,0	TH
	TT	

from random variables
to events
and vice versa

From Random Variables to Events

For any random variable X and value a,
we can define the event A that $X = a$

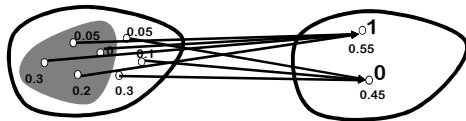
$$\Pr(A) = \Pr(X=a) = \Pr(\{t \in S \mid X(t)=a\})$$

From Events to Random Variables

For any event A, can define the
“indicator random variable” for event A:

$$X_A(t) = \begin{cases} 1 & \text{if } t \in A \\ 0 & \text{if } t \notin A \end{cases}$$

$$E[X_A] = 1 \times \Pr(X_A = 1) = \Pr(A)$$



How about multiplying expectations?

Multiplication of Expectations

A coin is tossed twice.

$X_i = 1$ if the i^{th} toss is heads and 0 otherwise.

$$E[X_i] = 1/2$$

$$E[X_1 X_2] = E[X_1] E[X_2] = 1/4$$

$$E[XY] = E[X]E[Y] \text{ if they are independent}$$

Multiplication of Expectations

Consider a single toss of a coin.

$X = 1$ if heads turns up and 0 otherwise.

Set $Y = 1 - X$

$$E[X] = E[Y] = 1/2$$

X and Y are
not
independent

$$E[XY] \neq E[X] E[Y]$$

$$\text{since } XY = 0$$