

Readings:
K&F: 4.5, 12.2, 12.3, 12.4

Kalman Filters Switching Kalman Filter

Graphical Models – 10708
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Carnegie Mellon University
November 20th, 2006

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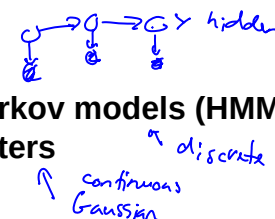
Adventures of our BN hero

- Compact representation for probability distributions
- Fast inference
- Fast learning
- Approximate inference

- But... Who are the most popular kids?

2 and 3.

Hidden Markov models (HMMs)
Kalman Filters



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The Kalman Filter

- An HMM with Gaussian distributions
- Has been around for at least 50 years
- Possibly the most used graphical model ever
- It's what
 - does your cruise control
 - tracks missiles
 - controls robots
 - ...
- And it's so simple...
 - Possibly explaining why it's so used
- Many interesting models build on it...
 - An example of a Gaussian BN (more on this later)

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Example of KF – SLAT Simultaneous Localization and Tracking

[Funiak, Guestrin, Paskin, Sukthankar '06]

- Place some cameras around an environment, don't know where they are
- Could measure all locations, but requires lots of grad. student (Stano) time
- Intuition:
 - A person walks around
 - If camera 1 sees person, then camera 2 sees person, learn about relative positions of cameras

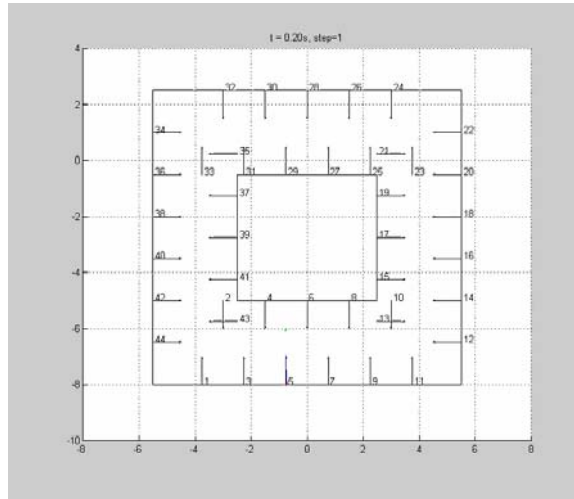


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Example of KF – SLAT

Simultaneous Localization and Tracking

[Funiak, Guestrin, Paskin, Sukthankar '06]



$x_t \leftarrow$ position
 $C_i \leftarrow$ camera location

$\vdots$

$P(C_i | m_{1:t})$

$P(x_t, C | m_{1:t})$

$\uparrow$ 2dim $\uparrow$ #cansx 3

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Multivariate Gaussian

$$p(\underline{X_1}, \dots, \underline{X_n}) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (\underline{x} - \underline{\mu})^T \Sigma^{-1} (\underline{x} - \underline{\mu}) \right\}$$

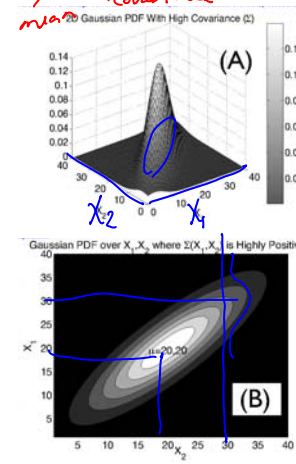
normalizer

Mean vector:

$$\underline{\mu} = \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_n \end{pmatrix}$$

Covariance matrix:

$$\Sigma = \begin{pmatrix} \sigma_{11} & \dots & \sigma_{1n} \\ \vdots & \ddots & \vdots \\ \sigma_{n1} & \dots & \sigma_{nn} \end{pmatrix}$$



Conditioning a Gaussian

Joint Gaussian:

$$p(X, Y) \sim N(\mu; \Sigma)$$

Conditional linear Gaussian:

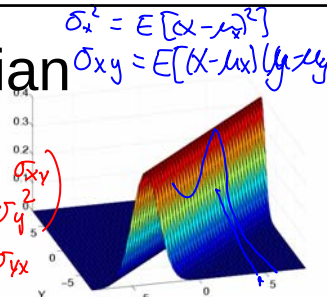
$$p(Y|X) \sim N(\mu_{Y|X}; \sigma_{Y|X}^2)$$

$$\mu_{Y|X} = \mu_Y + \frac{\sigma_{YX}}{\sigma_X^2}(x - \mu_X)$$

$$\sigma_{Y|X}^2 = \sigma_Y^2 - \frac{\sigma_{YX}^2}{\sigma_X^2}$$

$$\mu = \begin{pmatrix} \mu_X \\ \mu_Y \end{pmatrix} \quad \Sigma = \begin{pmatrix} \sigma_X^2 & \sigma_{XY} \\ \sigma_{YX} & \sigma_Y^2 \end{pmatrix}$$

$$\sigma_{XY} = \sigma_{YX}$$



x above mean & $\sigma_{XY} > 0 \Rightarrow$ positively correlated

$\mu_{Y|X}$ above average

$\mu_{Y|X}$ is a linear function of x

doesn't depend on observed x

observation make you more sure

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Gaussian is a "Linear Model"

Conditional linear Gaussian:

$$p(Y|X) \sim N(\beta_0 + \beta X; \sigma^2)$$

$$\mu_{Y|X} = \mu_Y + \frac{\sigma_{YX}}{\sigma_X^2}(x - \mu_X)$$

$$\sigma_{Y|X}^2 = \sigma_Y^2 - \frac{\sigma_{YX}^2}{\sigma_X^2}$$

$$\mu_{Y|X} = \underbrace{\mu_Y - \mu_X \frac{\sigma_{YX}}{\sigma_X^2}}_a + \underbrace{\frac{\sigma_{YX}}{\sigma_X^2}}_b x$$

$$= a + bx$$

doesn't depend on x



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Conditioning a Gaussian

Joint Gaussian:

$$\square p(\hat{X}, \hat{Y}) \sim N(\mu; \Sigma)$$

Conditional linear Gaussian:

$$\square p(Y|X) \sim N(\mu_{Y|X}; \Sigma_{YY|X})$$

$$\mu_{Y|X} = \mu_Y + \Sigma_{YX} \Sigma_{XX}^{-1} (x - \mu_X)$$

$$\Sigma_{YY|X} = \Sigma_{YY} - \Sigma_{YX} \Sigma_{XX}^{-1} \Sigma_{XY}$$

$$\Sigma_{XY} = \frac{1}{n} \begin{pmatrix} \dots & \sigma_{x_i y_j} \end{pmatrix}$$

$$\sigma_{x_i y_j} = E[(x_i - \mu_{x_i})(y_j - \mu_{y_j})]$$

$$\text{Marginal: } Y \sim N(\mu_Y, \Sigma_{YY})$$



Σ_{XX} cov. matrix for X
 Σ_{YY} " " Y
 Σ_{XY} is covar bet X, Y

Conditional Linear Gaussian (CLG) – general case

Conditional linear Gaussian:

$$\square p(Y|X) \sim N(\beta_0 + BX; \Sigma_{YY|X})$$

$$\mu_{Y|X} = \mu_Y + \Sigma_{YX} \Sigma_{XX}^{-1} (x - \mu_X)$$

$$\Sigma_{YY|X} = \Sigma_{YY} - \Sigma_{YX} \Sigma_{XX}^{-1} \Sigma_{XY}$$

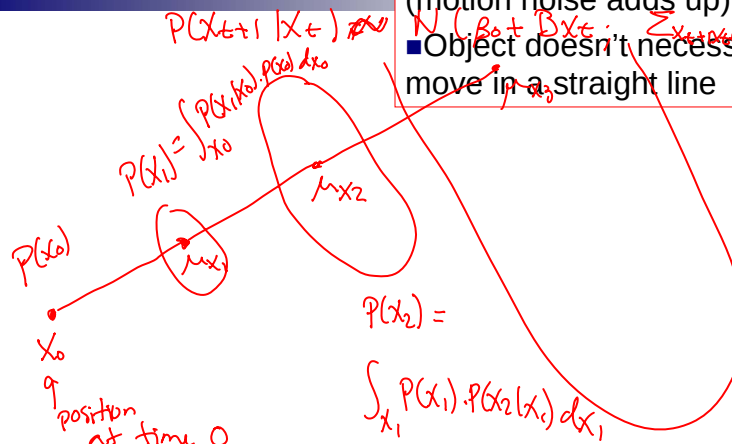
$$\mu_{Y|X} = \mu_Y + \Sigma_{YX} \Sigma_{XX}^{-1} (x - \mu_X) = \beta_0 + BX$$

$$\beta_0 = \mu_Y - \Sigma_{YX} \Sigma_{XX}^{-1} \mu_X$$

$$B = \Sigma_{YX} \Sigma_{XX}^{-1}$$

Understanding a linear Gaussian – the 2d case

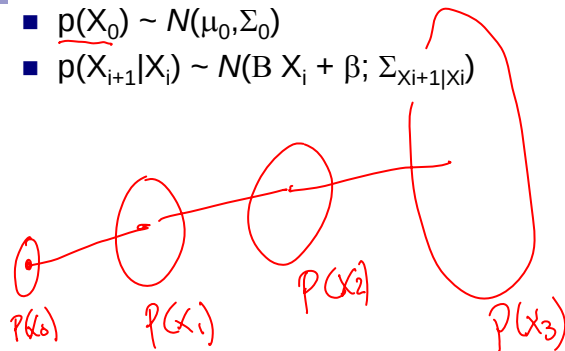
- Variance increases over time (motion noise adds up)
- Object doesn't necessarily move in a straight line



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Tracking with a Gaussian 1

- $p(X_0) \sim N(\mu_0, \Sigma_0)$
- $p(X_{i+1}|X_i) \sim N(B X_i + \beta; \Sigma_{X_{i+1}|X_i})$



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Tracking with Gaussians 2 – Making observations

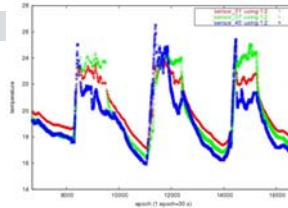
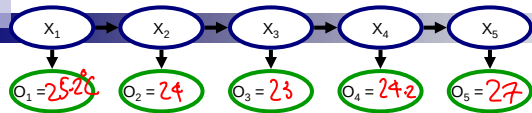
- We have $p(X_i) \sim \mathcal{N}(\mu_i; \Sigma_{X_i X_i})$
- Detector observes $O_i = o_i$
- Want to compute $p(X_i | O_i = o_i)$
- Use Bayes rule: $p(X_i | O_i = o_i) \propto \underbrace{p(X_i)}_{\text{prior}} \cdot \underbrace{P(O_i = o_i | X_i)}_{\text{likelihood}}$
- Require a CLG observation model
 - $p(O_i | X_i) \sim \mathcal{N}(W X_i + v; \Sigma_{O_i | X_i})$



T_i true temp
 M_i sensor
 $M_i = a T_i + b + \epsilon$
 $\epsilon \sim \mathcal{N}(0, \sigma^2)$
 $\uparrow$ calibration param.

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Operations in Kalman filter



- Compute $p(X_t | O_{1:t} = o_{1:t})$
- Start with $p(X_0)$ prior
- At each time step t :
 - **Condition** on observation $\leftarrow$ observations up to t
 $p(X_t | o_{1:t}) \propto p(X_t | o_{1:t-1}) p(o_t | X_t)$
 - **Prediction** (Multiply transition model) $\leftarrow$ observation at t
 $p(X_{t+1}, X_t | o_{1:t}) = p(X_{t+1} | X_t) p(X_t | o_{1:t})$
 - **Roll-up** (marginalize previous time step)
 $p(X_{t+1} | o_{1:t}) = \int_{X_t} p(X_{t+1}, x_t | o_{1:t}) dx_t$
- I'll describe one implementation of KF, there are others
 - Information filter

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Exponential family representation of Gaussian: Canonical Form $e^{a+b} = e^a \cdot e^b$

$$p(X_1, \dots, X_n) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \right\}$$

$$\exp \left\{ -\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \right\} = \exp \left[\underbrace{-\frac{1}{2} \mathbf{x}^T \Sigma^{-1} \mathbf{x}}_{\text{quadratic in } \mathbf{x}} + \underbrace{2 \mu^T \Sigma^{-1} \mathbf{x}}_{\text{linear in } \mathbf{x}} + \underbrace{\mu^T \Sigma^{-1} \mu}_{\text{constant (doesn't dep. on } \mathbf{x})} \right]$$

$$p(\mathbf{x}) = \frac{1}{Z} e^{-\frac{1}{2} \mathbf{x}^T \Sigma^{-1} \mathbf{x} + 2 \mu^T \Sigma^{-1} \mathbf{x} + \mu^T \Sigma^{-1} \mu}$$

$\Lambda = \Sigma^{-1}$ $[\mu^T \Sigma^{-1}]^T = \eta$

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Canonical form

$\log \text{lin model} \doteq \frac{1}{Z} e^{\sum w_i f_i(\mathbf{x})}$

$$p(X_1, \dots, X_n) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \right\}$$

$$= K \exp \left\{ \underbrace{\eta^T \mathbf{x}}_{\text{vector } \mathbf{x}} - \underbrace{\frac{1}{2} \mathbf{x}^T \Lambda \mathbf{x}}_{\text{matrix } \mathbf{x}} \right\}$$

Param $\theta(\eta, \Lambda)$
a log linear model !!
 $e^{\sum \eta_i x_i - \frac{1}{2} \sum_{ij} \Lambda_{ij} x_i x_j}$

- Standard form and canonical forms are related: $f(\mathbf{x})$

$$\underline{\mu} = \underline{\Lambda}^{-1} \underline{\eta}$$

$$\underline{\Sigma} = \underline{\Lambda}^{-1}$$

$\underline{\eta} = \underline{\Sigma}^{-1} \underline{\mu} \Rightarrow \underline{\mu} = \underline{\Sigma} \underline{\eta} = \underline{\Lambda}^{-1} \underline{\eta}$

- Conditioning is easy in canonical form
- Marginalization easy in standard form

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Conditioning in canonical form

$$p(X_t | o_{1:t}) \propto p(X_t | o_{1:t-1})p(o_t | X_t)$$

- First multiply: $p(A, B) = p(A)p(B | A)$

$$p(A) : \mathcal{N}(\eta_1, \Lambda_1)$$

$$p(B | A) : \eta_2, \Lambda_2$$

$$p(A, B) : \eta_3 = \eta_1 + \eta_2, \Lambda_3 = \Lambda_1 + \Lambda_2$$

$$\begin{pmatrix} n(\eta_1) \\ m(0) \end{pmatrix} + \begin{pmatrix} n(\eta_2) \\ m(0) \end{pmatrix}$$

$$\begin{pmatrix} n(\Lambda_1) & 0 \\ 0 & 0 \end{pmatrix}$$

$$\begin{pmatrix} n(\Lambda_2) \\ m(0) \end{pmatrix}$$

$$\eta_3 = \begin{pmatrix} \eta_1 \\ \eta_2 \end{pmatrix}$$

$\eta_1 = n$ -vector
 $\Lambda_1 = n \times n$ matrix
 $\Lambda_2 = (n+m) \times (n+m)$
 $\eta_2 = n+m$ - vector
 $n+m$

- Then, condition on value $B = y$ $p(A | B = y)$

$$\eta_{A|B=y} = \eta_A - \Lambda_{AB} \cdot y$$

$$\Lambda_{AA|B=y} = \Lambda_{AA}$$

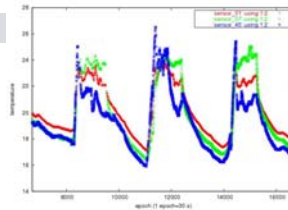
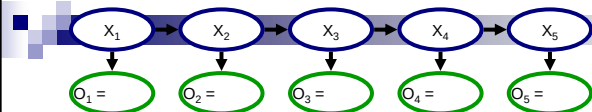
$$P(A, B) = \begin{pmatrix} n(\eta_A) \\ m(\eta_B) \end{pmatrix}$$

$$\Lambda = \begin{pmatrix} \Lambda_{AA} & \Lambda_{AB} \\ \Lambda_{BA} & \Lambda_{BB} \end{pmatrix}$$

linear function

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Operations in Kalman filter



- Compute $p(X_t | O_{1:t} = o_{1:t})$

- Start with $p(X_0)$

- At each time step t :

- Condition on observation

$$p(X_t | o_{1:t}) \propto p(X_t | o_{1:t-1})p(o_t | X_t)$$

- Prediction (Multiply transition model)

$$p(X_{t+1}, X_t | o_{1:t}) = p(X_{t+1} | X_t)p(X_t | o_{1:t})$$

- Roll-up (marginalize previous time step)

$$p(X_{t+1} | o_{1:t}) = \int_{X_t} p(X_{t+1}, x_t | o_{1:t}) dx_t$$

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Prediction & roll-up in canonical form

$$p(X_{t+1} | o_{1:t}) = \int_{X_t} p(X_{t+1} | x_t) p(x_t | o_{1:t}) dx_t$$

$N(\eta_{x_t}, \Lambda_{x_t})$

■ First multiply: $p(A, B) = p(A)p(B | A)$

$\text{sum } \eta\text{'s \& } \Lambda\text{'s}$

■ Then, marginalize X_t : $p(A) = \int_B p(A, b) db$

$$\eta_A^m = \eta_A - \Lambda_{AB} \Lambda_{BB}^{-1} \eta_B$$

$$\Lambda_{AA}^m = \Lambda_{AA} - \Lambda_{AB} \Lambda_{BB}^{-1} \Lambda_{BA}$$

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Announcements

■ Lectures the rest of the semester:

- Special time: Monday Nov 27 - 5:30-7pm, Wean 4615A:
Dynamic BNS
- Wed. 11/30, regular class time: Causality (Richard Scheines)
- Friday 12/1, regular class time: Finish Dynamic BNS & Overview of Advanced Topics

■ Deadlines & Presentations:

- Project Poster Presentations: Dec. 1st 3-6pm (NSH Atrium)
 - popular vote for best poster
- Project write up: Dec. 8th by 2pm by email
 - 8 pages – limit will be strictly enforced
- Final: Out Dec. 1st, Due Dec. 15th by 2pm (strict deadline)

What if observations are not CLG?

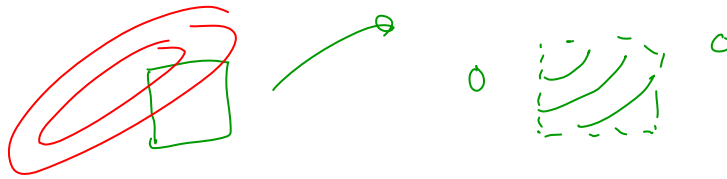
- Often observations are not CLG

- CLG if $O_i = B X_i + \beta_0 + \varepsilon$ ← $\mathcal{N}(\mu, \sigma^2)$

- Consider a motion detector

- $O_i = 1$ if person is likely to be in the region

- Posterior is not Gaussian



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Linearization: incorporating non-linear evidence

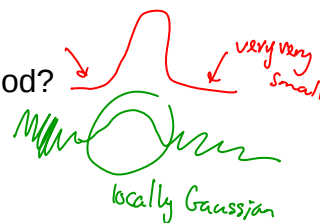
- $p(O_i|X_i)$ not CLG, but...

- Find a Gaussian approximation of $p(X_i, O_i) = p(X_i) p(O_i|X_i)$

- Instantiate evidence $O_i = o_i$ and obtain a Gaussian for $p(X_i|O_i = o_i)$

- Why do we hope this would be any good?

- Locally, Gaussian may be OK



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Linearization as integration

- Gaussian approximation of $p(X_i, O_i) = p(X_i) p(O_i | X_i)$

$\approx$ Gaussian μ, Σ

- Need to compute moments

$\circ E[X_i] = \mu_{x_i}$

$\square E[O_i] = \int_{o_i} o_i p(o_i) do_i = \int_{x_i} \int_{o_i} o_i p(x_i) p(o_i | x_i) do_i dx_i$

$\square E[O_i^2]$

$\square E[O_i X_i] = \int_{x_i} \int_{o_i} o_i x_i p(o_i | x_i) p(x_i) do_i dx_i$

$\uparrow$ arbitrary $\uparrow$ Gaussian

- Note: Integral is product of a Gaussian with an arbitrary function

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Linearization as numerical integration

- Product of a Gaussian with arbitrary function

- Effective numerical integration with Gaussian quadrature method

- $\square$ Approximate integral as **weighted sum over integration points**
- $\square$ Gaussian quadrature defines location of points and weights

- Exact if arbitrary function is polynomial of bounded degree

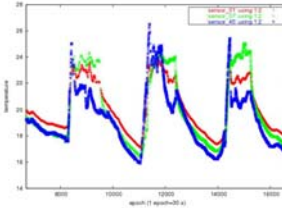
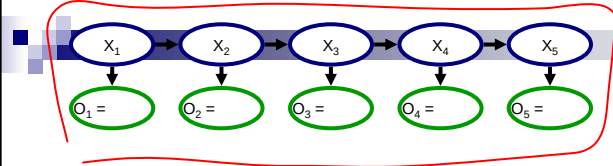
- Number of integration points exponential in number of dimensions d

- Exact monomials requires exponentially fewer points

- $\square$ For $2d+1$ points, this method is equivalent to effective Unscented Kalman filter
- $\square$ Generalizes to many more points

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Operations in non-linear Kalman filter



- Compute $p(X_t | O_{1:t} = o_{1:t})$
- Start with $p(X_0)$
- At each time step t :
 - **Condition** on observation (use numerical integration)

$$p(X_t | o_{1:t}) \propto p(X_t | o_{1:t-1})p(o_t | X_t)$$
 - **Prediction** (Multiply transition model, use numerical integration)

$$p(X_{t+1}, X_t | o_{1:t}) = p(X_{t+1} | X_t)p(X_t | o_{1:t})$$
 - **Roll-up** (marginalize previous time step)

$$p(X_{t+1} | o_{1:t}) = \int_{X_t} p(X_{t+1}, x_t | o_{1:t}) dx_t$$

non-linear KF

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What you need to know about Kalman Filters

- **Kalman filter**
 - Probably most used BN
 - Assumes Gaussian distributions
 - Equivalent to linear system
 - Simple matrix operations for computations
- **Non-linear Kalman filter**
 - Usually, observation or motion model not CLG
 - Use numerical integration to find Gaussian approximation

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What if the person chooses different motion models?

- With probability θ , move more or less straight
- With probability $1-\theta$, do the “moonwalk”

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The moonwalk



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What if the person chooses different motion models?

- With probability θ , move more or less straight
- With probability $1-\theta$, do the “moonwalk”

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Switching Kalman filter

- At each time step, choose one of k motion models:
 - You never know which one!
- $p(X_{i+1}|X_i, Z_{i+1})$
 - CLG indexed by Z_i
 - $p(X_{i+1}|X_i, Z_{i+1}=j) \sim N(\beta^j_0 + B^j X_i, \Sigma^j_{X_{i+1}|X_i})$

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Inference in switching KF – one step

- Suppose
 - $p(X_0)$ is Gaussian
 - Z_1 takes one of two values
 - $p(X_1|X_0, Z_1)$ is CLG
- Marginalize X_0
- Marginalize Z_1
- Obtain mixture of two Gaussians!

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Multi-step inference

- Suppose
 - $p(X_i)$ is a mixture of m Gaussians
 - Z_{i+1} takes one of two values
 - $p(X_{i+1}|X_i, Z_{i+1})$ is CLG
- Marginalize X_i
- Marginalize Z_i
- Obtain mixture of $2m$ Gaussians!
 - Number of Gaussians grows exponentially!!!

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Visualizing growth in number of Gaussians

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Computational complexity of inference in switching Kalman filters

- Switching Kalman Filter with (only) 2 motion models
- Query:
- **Problem is NP-hard!!!** [Lerner & Parr '01]
 - Why “!!!”?
 - Graphical model is a tree:
 - Inference efficient if all are discrete
 - Inference efficient if all are Gaussian
 - But not with hybrid model (combination of discrete and continuous)

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Bounding number of Gaussians

- $P(X_i)$ has 2^m Gaussians, but...
- usually, most are bumps have low probability and overlap:

- **Intuitive approximate inference:**

- Generate $k.m$ Gaussians
- Approximate with m Gaussians

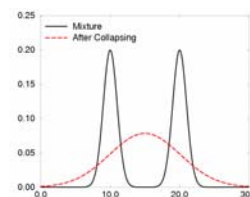
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Collapsing Gaussians – Single Gaussian from a mixture

- Given mixture $P <w_i; N(\mu_i, \Sigma_i)>$
- Obtain approximation $Q \sim N(\mu, \Sigma)$ as:

$$\mu = \sum_i w_i \mu_i$$

$$\Sigma = \sum_i w_i \Sigma_i + \sum_i w_i (\mu_i - \mu)(\mu_i - \mu)^T$$



- **Theorem:**
 - P and Q have same first and second moments
 - **KL projection:** Q is single Gaussian with lowest KL divergence from P

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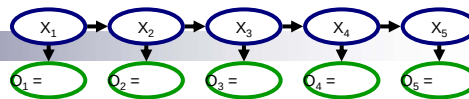
Collapsing mixture of Gaussians into smaller mixture of Gaussians

- Hard problem!
 - Akin to clustering problem...

- Several heuristics exist
 - *c.f.*, K&F book

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Operations in non-linear switching Kalman filter



- Compute mixture of Gaussians for $p(X_t | O_{1:t} = o_{1:t})$
- Start with $p(X_0)$
- At each time step t :
 - For each of the m Gaussians in $p(X_t | O_{1:t})$:
 - **Condition** on observation (use **numerical integration**)
 - **Prediction** (Multiply transition model, use **numerical integration**)
 - Obtain k Gaussians
 - **Roll-up** (marginalize previous time step)
 - **Project** $k \cdot m$ Gaussians into m' Gaussians $p(X_t | O_{1:t+1})$

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Assumed density filtering

- Examples of very important **assumed density filtering**:

- ☐ Non-linear KF
- ☐ Approximate inference in switching KF

- General picture:

- ☐ Select an **assumed density**
 - e.g., single Gaussian, mixture of m Gaussians, ...
- ☐ After conditioning, prediction, or roll-up, **distribution no-longer representable with assumed density**
 - e.g., non-linear, mixture of $k.m$ Gaussians,...
- ☐ **Project** back into assumed density
 - e.g., numerical integration, collapsing,...

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When non-linear KF is not good enough

- Sometimes, distribution in non-linear KF is not approximated well as a single Gaussian

- ☐ e.g., a banana-like distribution

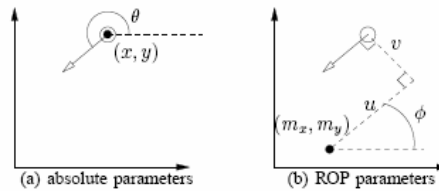
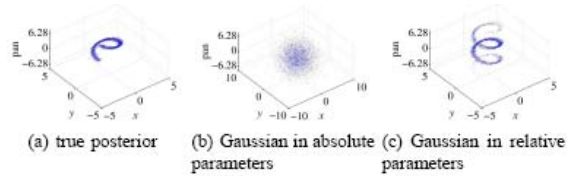
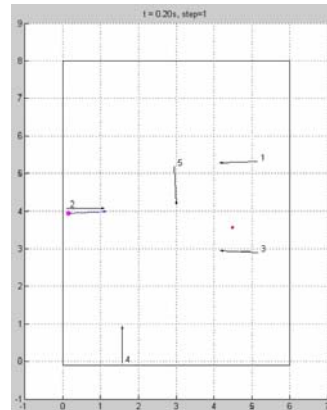
- Assumed density filtering:

- ☐ Solution 1: **reparameterize problem** and solve as a **single Gaussian**
- ☐ Solution 2: more typically, **approximate as a mixture of Gaussians**

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Reparameterized KF for SLAT

[Funiak, Guestrin, Paskin, Sukthankar '05]



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When a single Gaussian ain't good enough



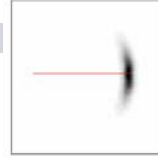
[Fox et al.]

- Sometimes, smart parameterization is not enough
 - Distribution has multiple hypothesis
- Possible solutions
 - Sampling – particle filtering
 - Mixture of Gaussians
 - ...
- Quick overview of one such solution...

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Approximating non-linear KF with mixture of Gaussians

- Robot example:



- $P(X_i)$ is a Gaussian, $P(X_{i+1})$ is a banana
- Approximate $P(X_{i+1})$ as a mixture of m Gaussians
 - e.g., using discretization, sampling,...
- Problem:
 - $P(X_{i+1})$ as a mixture of m Gaussians
 - $P(X_{i+2})$ is m bananas
- One solution:
 - Apply collapsing algorithm to project m bananas in m' Gaussians

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What you need to know

- **Switching Kalman filter**
 - Hybrid model – discrete and continuous vars.
 - Represent belief as mixture of Gaussians
 - Number of mixture components grows exponentially in time
 - Approximate each time step with fewer components
- **Assumed density filtering**
 - Fundamental abstraction of most algorithms for dynamical systems
 - Assume representation for density
 - Every time density not representable, project into representation

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