

Computational Genomics

Molecular Evolution: Phylogenetic trees

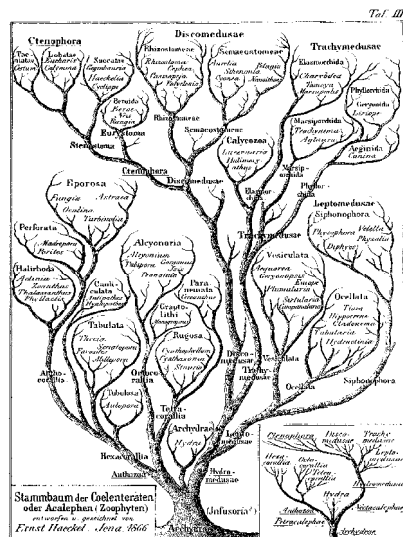
Eric Xing

Lecture 14, March 6, 2007

Reading: DTW book, Chap 12
DEKM book, Chap 7, 8



Phylogeny



In 1866, Ernst Haeckel coined the word “phylogeny” and presented phylogenetic trees for most known groups of living organisms.

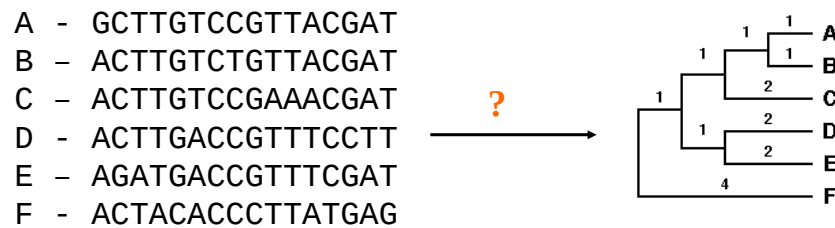


Ernst Haeckel (1834-1919)



Phylogenetic Inference

- Given a multiple alignment, how do we construct the tree?

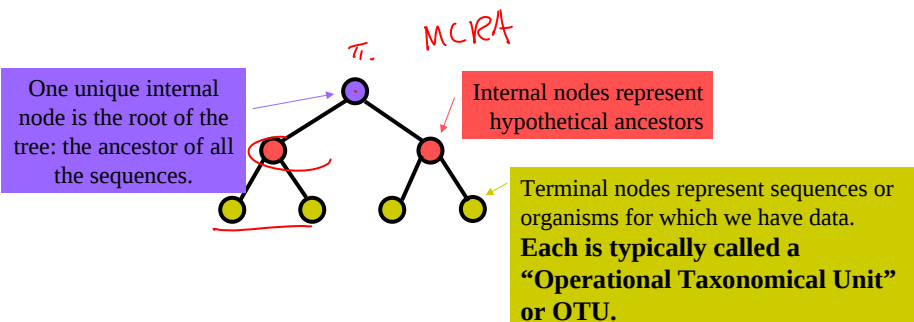


What Is a Tree?

- A tree is a mathematical structure which represents a model of an actual evolutionary history of a group of sequences or organisms.

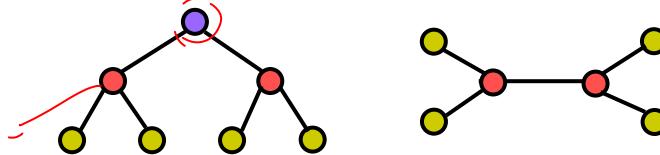
- In other words, it is an evolutionary hypothesis.

- A tree consists of nodes connected by branches.



Types of Trees

- Rooted vs. Unrooted



		Branches	Nodes
Rooted	Interior	$M - 2$	$M - 1$
	Total	$2M - 2$	$2M - 1$
Unrooted	Interior	$M - 3$	$M - 2$
	Total	$2M - 3$	$2M - 2$

● M is the number of OTU's

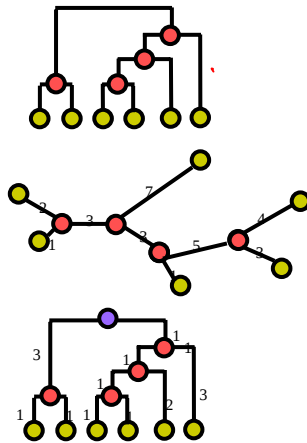
The number of rooted and unrooted trees:

Number of OTU's	Possible Number of	
	Rooted trees	Unrooted trees
2	1	1
3	3	1
4	15	3
5	105	15
6	945	105
7	10395	945
8	135135	10395
9	2027025	135135
10	34459425	2027025

More different kinds of trees



- Different kinds of trees can be used to depict different aspects of evolutionary history



1. Cladogram:
simply shows relative order of common ancestry
2. Additive trees:
a cladogram with branch lengths,
also called phylograms and metric trees
3. Ultrametric trees:
(dendograms) special kind of additive tree in which
the tips of the trees are all equidistant from the root

Phylogeny methods



Basic principles:

- Degree of sequence difference is proportional to length of independent sequence evolution
- Only use positions where alignment is pretty certain – avoid areas with (too many) gaps

Major methods:

Clustering methods	UPGMA	Neighbor-joining
↑	WPGMA	
	Single linkage	
	Complete linkage	
Objective criterion-based methods	Least-squares distance	
	Maximum parsimony	
	Minimum evolution	
	Maximum likelihood	

UPGMA



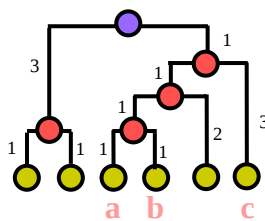
- Construction of a distance tree using clustering with the Unweighted Pair Group Method with Arithmetic Mean (UPGMA)
- First, construct a distance matrix:

A - GCTTGTCCGTTACGAT
 B - ACTTGTCTGTTACGAT
 C - ACTTGTCCGAAACGAT
 D - ACTTGACCGTTTCCTT
 E - AGATGACCGTTTCGAT
 F - ACTACACCCTTATGAG

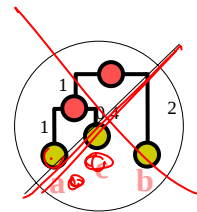
	A	B	C	D	E
B	2				
C	4	4			
D	6	6	6		
E	6	6	6	4	
F	8	8	8	8	8

From <http://www.icp.ucl.ac.be/~opperd/private/upgma.html>

Ultrametric Trees



- Metric distances (for additive trees) must obey 4 rules:
 - Non-negativity: $d(a,b) \geq 0$
 - Distinctness: $d(a,b) = 0$ if and only if $a = b$
 - Symmetry: $d(a,b) = d(b,a)$
 - Triangle Inequality: $d(a,c) \leq d(a,b) + d(b,c)$
- Ultrametric must obey one additional rule:
 - Three point condition: $d(a,b) \leq \max(d(a,c), d(b,c))$



UPGMA



- First round

	A	B	C	D	E
B	2				
C	4	4			
D	6	6	6		
E	6	6	6	4	
F	8	8	8	8	8

Choose the most similar pair, cluster them together and calculate the new distance matrix.



$$\begin{aligned} \text{dist}(A,B),C &= (\text{dist}AC + \text{dist}BC) / 2 = 4 \\ \text{dist}(A,B),D &= (\text{dist}AD + \text{dist}BD) / 2 = 6 \\ \text{dist}(A,B),E &= (\text{dist}AE + \text{dist}BE) / 2 = 6 \\ \text{dist}(A,B),F &= (\text{dist}AF + \text{dist}BF) / 2 = 8 \end{aligned}$$



	B,A	C	D	E
C	4			
D	6	6		
E	6	6	4	
F	8	8	8	8

UPGMA

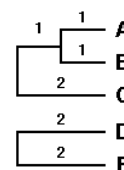
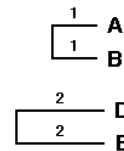


- Second round

	B,A	C	D	E
C	4			
D	6	6		
E	6	6	4	
F	8	8	8	8

- Third round

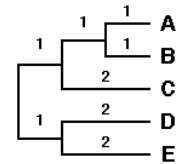
	B,A	C	,D E
C	4		
,D E	6	6	
F	8	8	8



UPGMA

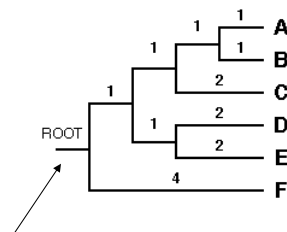
- Forth round

	C,AB	E,D
E,D	6	
F	8	8



- Fifth round

	DE,ABC
F	8

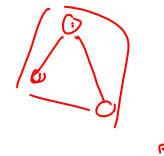
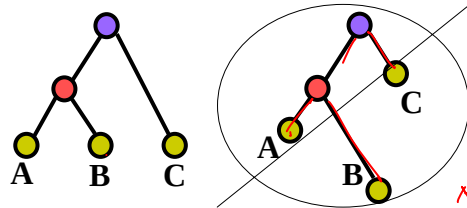


Note the this method identifies the root of the tree.

UPGMA assumes a molecular clock

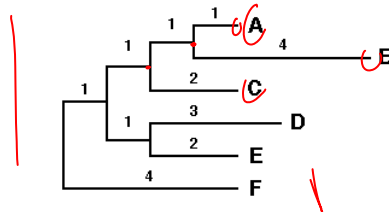
- The UPGMA clustering method is very sensitive to unequal evolutionary rates (assumes that the evolutionary rate is the same for all branches).
- Clustering works only if the data are ultrametric
- Ultrametric distances are defined by the satisfaction of the 'three-point condition'.

The three-point condition:

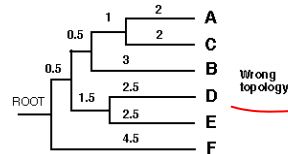
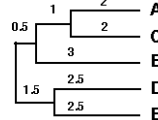
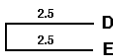
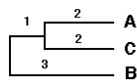
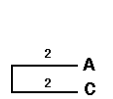


For any three taxa, the two greatest distances are equal.

UPGMA fails when rates of evolution are not constant



	A	B	C	D	E
B	5				
C	4	7			
D	7	10	7		
E	6	9	6	5	
F	8	11	8	9	8



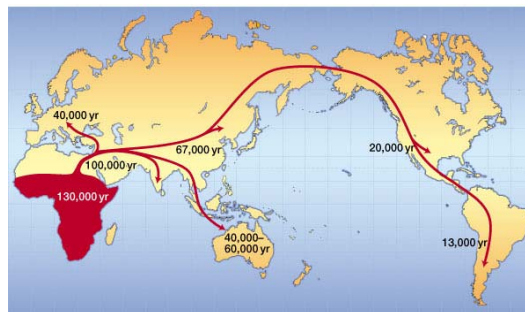
(Neighbor joining will get the right tree in this case.)

From <http://www.icp.ucl.ac.be/~opperd/private/upgma.html>

Phylogeny based upon the molecular clock

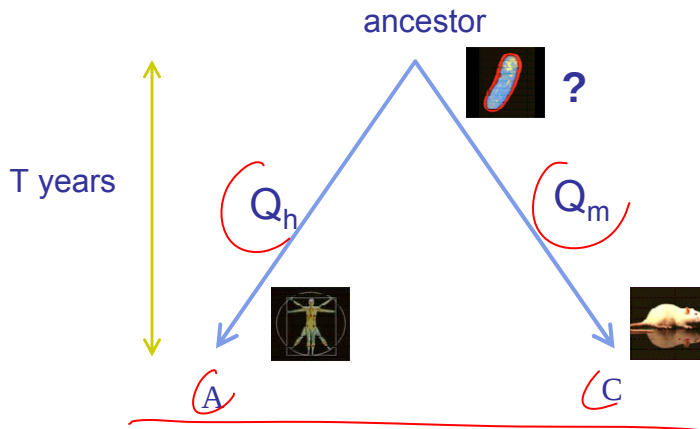


- Evidence for a human mitochondrial origin in Africa: African sequence diversity is twice as large as that of non-African
- Gyllensten and colleagues estimate that the divergence of Africans and non-Africans occurred 52,000 to 28,000 years ago.



Ingman, M., Kaessmann, H., Pääbo, S. & Gyllensten, U. (2000) *Nature* 408: 708-713.

A pair of homologous bases



Typically, the ancestor is unknown.

How does sequence variation arise?

- **Mutation:**
 - (a) Inherent: DNA replication errors are not always corrected.
 - (b) External: exposure to chemicals and radiation.
- **Selection:** Deleterious mutations are removed quickly. Neutral and rarely, advantageous mutations, are tolerated and stick around.
- **Fixation:** It takes time for a new variant to be established (having a stable frequency) in a population.

Modeling DNA base substitution



- Strictly speaking, only applicable to regions undergoing little selection.
- Standard assumptions (sometimes weakened)
 1. Site independence.
 2. Site homogeneity.
 3. Markovian: given current base, future substitutions independent of past.
 4. Temporal homogeneity: stationary Markov chain.

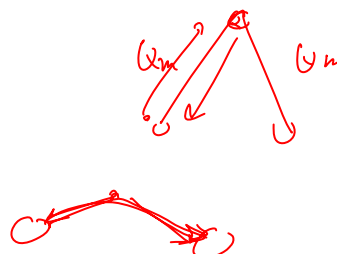
More assumptions



- $Q_h = s_h Q$ and $Q_m = s_m Q$, for some positive s_h, s_m , and a rate matrix Q .
- The ancestor is sampled from the stationary distribution π of Q .
- Q is **reversible**: for $a, b, t \geq 0$

$$\pi(a)P(t, a, b) = P(t, b, a)\pi(b),$$
 (detailed balance).

$$p = e^{tQ}$$



The stationary distribution



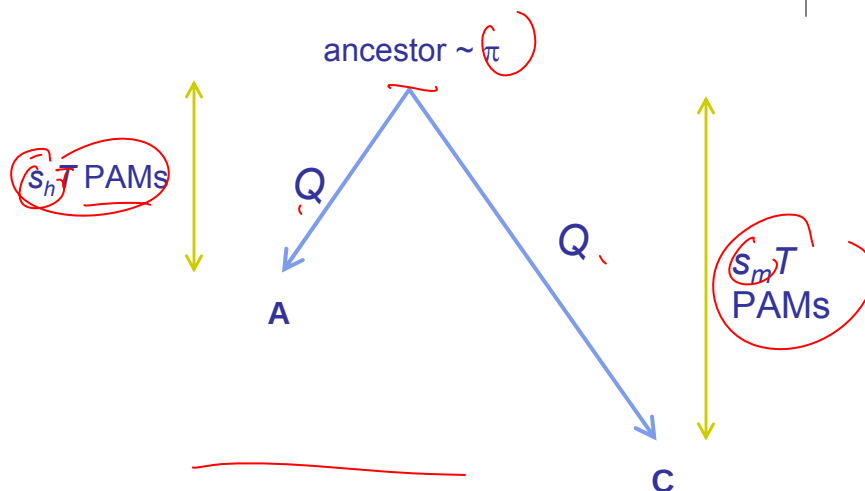
- A probability distribution π on $\{A, C, G, T\}$ is a **stationary distribution** of the Markov chain with transition probability matrix $P = P(i, j)$, if for all j ,

$$\sum_i \pi(i) P(i, j) = \pi(j).$$

- Exercise.** Given any initial distribution, the distribution at time t of a chain with transition matrix P converges to π as $t \rightarrow \infty$. Thus, π is also called an **equilibrium** distribution.
- Exercise.** For the Jukes-Cantor and Kimura models, the uniform distribution is stationary. (Hint: diagonalize their infinitesimal rate matrices.)

We often assume that the ancestor sequence is i.i.d π .

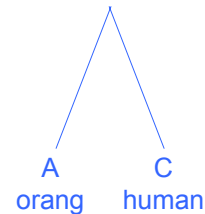
New picture from a statistical evolution model



The Jukes-Cantor adjustment

- Assume that the common ancestor has A, G, C or T with probability 1/4.

common ancestor

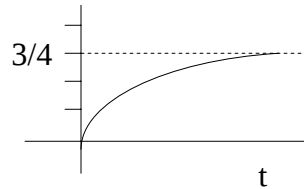


- Then the chance of the nt differing

$$p_{\neq} = \frac{3}{4} \times (1 - e^{-8\alpha t})$$

$$= \frac{3}{4} \times (1 - e^{-4k/3}), \text{ since } k = 2 \times 3\alpha t$$

When $k = .01$, described as 1 PAM



Joint probability of A and C

- Under the model in the previous slides, the joint probability is

$$p(A,C) = \sum_a \pi(a) p(A | s_h T, Q, a) p(C | s_m T, Q, a)$$

$$= \sum_a \pi(A) p(a | s_h T, Q, A) p(C | s_m T, Q, a)$$

$$= \pi(A) p(C | s_h T + s_m T, Q, A)$$

$$= (t, A, C)$$



- where $t = s_h T + s_m T$ is the (evolutionary) distance between A and C. Note that s_h , s_m and T are not identifiable.
- The matrix $F(t)$ is symmetric. It is equally valid to view A as the ancestor of C or vice versa.

Estimating the evolutionary distance between two sequences



- Suppose two aligned protein sequences $a_1 \dots a_n$ and $b_1 \dots b_n$ are separated by t PAMs.
- Under a **reversible substitution model** that is IID across sites, the likelihood of t is

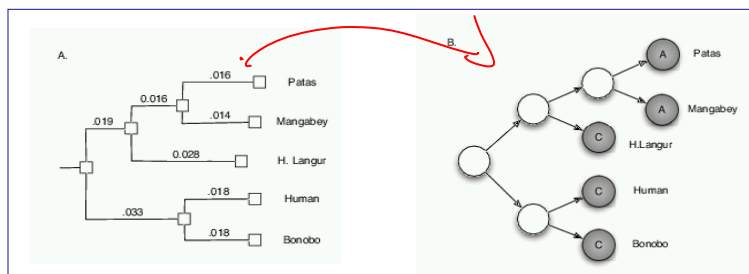
$$\begin{aligned} L(t) &= p(a_1 \dots a_n, b_1 \dots b_n | \text{model}) \\ &= \prod_k (t, a_k, b_k) \\ &= \prod_{a,b} (t, a, b)^{c(a,b)} \end{aligned}$$



- where $c(a,b) = \# \{k : a_k = a, b_k = b\}$.

- Maximizing this quantity gives the maximum likelihood estimate of t . This generalizes the distance correction with Jukes-Cantor.

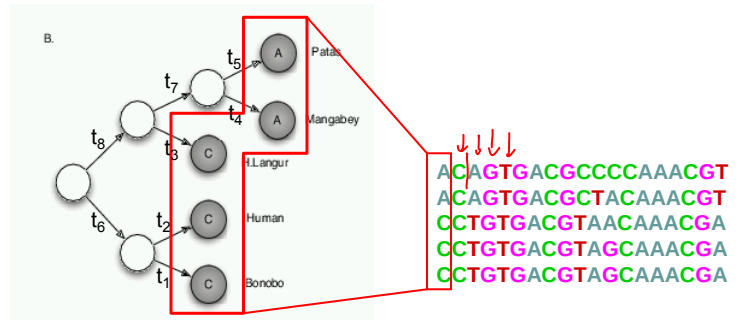
Phylogeny



- The shaded nodes represent the observed nucleotides at a given site for a set of organisms
- The unshaded nodes represent putative ancestral nucleotides
- Transitions between nodes capture the dynamic of evolution

Likelihood methods

- A tree, with branch lengths, and the data at a single site.



- Since the sites evolve independently on the same tree,

$$L = P(D|T) = \prod_{i=1}^m P(D^{(i)} | T)$$

Likelihood at one site on a tree

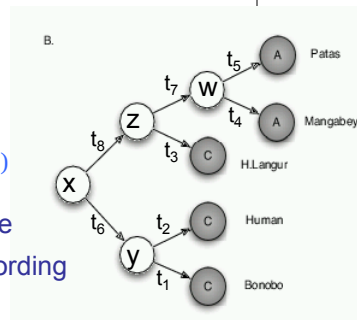
- We can compute this by summing over all assignments of states x, y, z and w to the interior nodes:

$$P(D^{(i)} | T) = \sum_x \sum_y \sum_z \sum_w P(A, A, C, C, C, x, y, z, w | T)$$

- Due to the Markov property of the tree, we can factorize the complete likelihood according to the tree topology:

$$P(A, A, C, C, C, x, y, z, w | T) = P(x) P(y|x, t_6) P(C|y, t_1) P(C|y, t_2) P(z|x, t_8) P(C|z, t_3) P(w|z, t_7) P(A|w, t_4) P(A|w, t_5)$$

- Summing this up, there are 256 terms in this case!



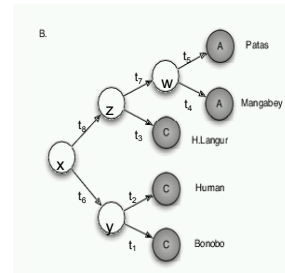
Getting a recursive algorithm

- when we move the summation signs as far right as possible:

$$P(D^{(i)} | T) = \sum_x \sum_y \sum_z \sum_w P(A, A, C, C, C, x, y, z, w | T) =$$

$$\sum_x P(x) \left(\sum_y P(y | x, t_6) P(C | y, t_1) P(C | y, t_2) \right)$$

$$\left(\sum_z P(z | x, t_8) P(C | z, t_3) \left(\sum_w P(w | z, t_7) P(A | w, t_4) P(A | w, t_5) \right) \right)$$



Felsenstein's Pruning Algorithm

- To calculate $P(x_1, x_2, \dots, x_N | T, t)$

Initialization:

Set $k = 2N - 1$

Recursion:

Compute $P(L_k | a)$ for all $a \in \Sigma$

If k is a leaf node:

Set $P(L_k | a) = 1(a = x_k)$

If k is not a leaf node:

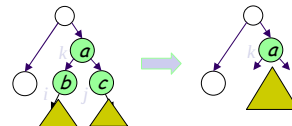
1. Compute $P(L_i | b)$, $P(L_j | c)$ for all b and c , for daughter nodes i, j

2. Set $P(L_k | a) = \sum_{b, c} P(b | a, t_i) P(L_i | b) P(c | a, t_j) P(L_j | c)$

Termination:

Likelihood at this column = $P(x_1, x_2, \dots, x_N | T, t) = \sum_a P(L_{2N-1} | a) P(a)$

- This algorithm can easily handle Ambiguity and error in the sequences (how?)



Finding the ML tree



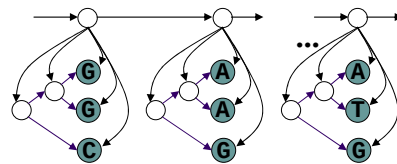
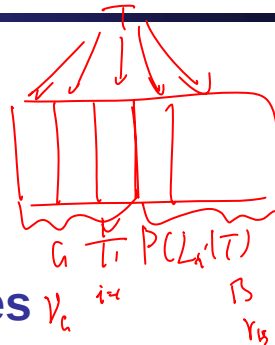
- So far I have just talked about the computation of the likelihood for one tree with branch lengths known.
- To find a ML tree, we must search the space of tree topologies, and for each one examined, we need to optimize the branch lengths to maximize the likelihood.

Bayesian phylogeny methods



- Bayesian inference has been applied to inferring phylogenies (Rannala and Yang, 1996; Mau and Larget, 1997; Li, Pearl and Doss, 2000).
 - All use a prior distribution on trees. The prior has enough influence on the result that its reasonableness should be a major concern. In particular, the depth of the tree may be seriously affected by the distribution of depths in the prior.
 - All use Markov Chain Monte Carlo (MCMC) methods. They sample from the posterior distribution.
 - When these methods make sense they not only get you a point estimate of the phylogeny, they get you a distribution of possible phylogenies.

Modeling rate variation among sites

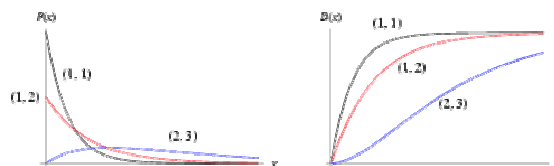


A model of variation in evolutionary rates among sites

- The basic idea is that the rate at each site is drawn independently from a distribution of rates. The most widely used choice is the Gamma distribution, which has density function:

$$f(r) = \frac{\lambda^\alpha r^{\alpha-1} e^{-\lambda r}}{\Gamma(\alpha)} = \frac{r^{\alpha-1} e^{-r/\theta}}{\Gamma(\alpha) \theta^\alpha}$$

- Gamma distributions (α, θ)



Unrealistic aspects of the model:



- There is no reason, aside from mathematical convenience, to assume that the Gamma is the right distribution.
- A common variation is to assume there is a separate probability f_0 of having rate 0.
- Rates at different sites appear to be correlated, which this model does not allow.
- Rates are not constant throughout evolution, they change with time.

Rates varying among sites



- If $L^{(i)}(r_i)$ is the likelihood of the tree for site i given that the rate of evolution at site i is r_i , we can integrate this over a gamma density:

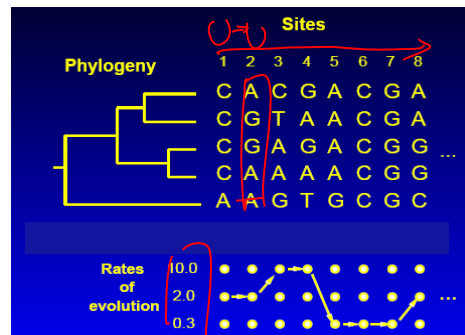
$$L^{(i)} = \int_0^\infty f(r_i; \alpha) L^{(i)}(r_i) dr_i$$

- so that the overall likelihood is

$$L = \prod_{i=1}^m \left[\int_0^\infty f(r_i; \alpha) L^{(i)}(r_i) dr_i \right]$$

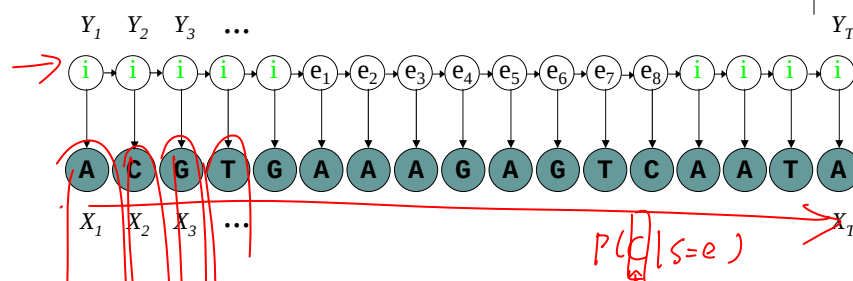
- Unfortunately these integrals cannot be evaluated for trees with more than a few tips as the quantities $L^{(i)}(r_i)$ becomes complicated.

Modeling rate variation among sites



- There are a finite number of rates (denote rate i as r_i).
- There are probabilities p_i of a site having rate i .
- A process not visible to us ("hidden") assigns rates to sites.
- The probability of our seeing some data are to be obtained by summing over all possible combinations of rates, weighting appropriately by their probabilities of occurrence.

Rocall the HMM



- The shaded nodes represent the observed nucleotides at particular sites of an organism's genome
- For discrete Y_i , widely used in computational biology to represent segments of sequences
 - gene finders and motif finders
 - profile models of protein domains
 - models of secondary structure

Definition (of HMM)

- **Observation space**

Alphabetic set: $C = \{c_1, c_2, \dots, c_K\}$

Euclidean space: $\mathbb{R}^d$

- **Index set of hidden states**

$$I = \{1, 2, \dots, M\}$$

- **Transition probabilities** between any two states

$$p(y_t^j = 1 | y_{t-1}^i = 1) = a_{i,j},$$

or $p(y_t | y_{t-1}^i = 1) \sim \text{Multinomial}(a_{i,1}, a_{i,2}, \dots, a_{i,M}), \forall i \in I.$

- **Start probabilities**

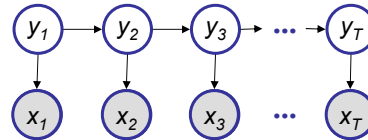
$$p(y_1) \sim \text{Multinomial}(\pi_1, \pi_2, \dots, \pi_M).$$

- **Emission probabilities** associated with each state

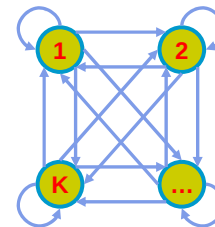
$$p(x_t | y_t^i = 1) \sim \text{Multinomial}(b_{i,1}, b_{i,2}, \dots, b_{i,K}), \forall i \in I.$$

or in general:

$$p(x_t | y_t^i = 1) \sim f(\cdot | \theta_i), \forall i \in I.$$

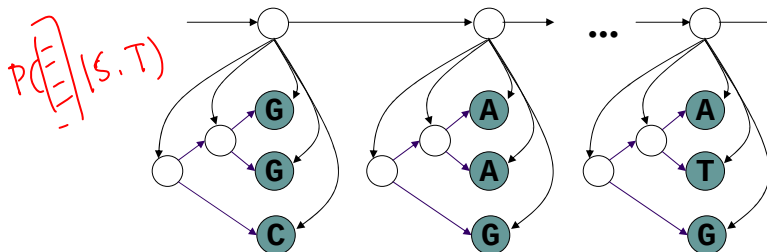


Graphical model



State automata

Hidden Markov Phylogeny



- Replacing the standard emission model with a tree

- A process not visible to us (.hidden") assigns rates to sites. It is a Markov process working along the sequence.
- For example it might have transition probability $\text{Prob}(j|i)$ of changing to rate j in the next site, given that it is at rate i in this site.

- These are the most widely used models allowing rate variation to be correlated along the sequence.



The Forward Algorithm

- We can compute α_t^k for all k, t , using dynamic programming!

Initialization:

$$\alpha_1^k = P(x_1 | y_1^k = 1) \pi_k$$

$$\begin{aligned} \alpha_1^k &= P(x_1, y_1^k = 1) \\ &= P(x_1 | y_1^k = 1) P(y_1^k = 1) \\ &= P(x_1 | y_1^k = 1) \pi_k \end{aligned}$$

Iteration:

$$\alpha_t^k = P(x_t | y_t^k = 1) \sum_i \alpha_{t-1}^i a_{i,k}$$

Handwritten note: $P(\prod_{i=1}^S l_{i,T})$

Termination:

$$P(x) = \sum_k \alpha_T^k$$



The Backward Algorithm

- We can compute β_t^k for all k, t , using dynamic programming!

Initialization:

$$\beta_T^k = 1, \forall k$$

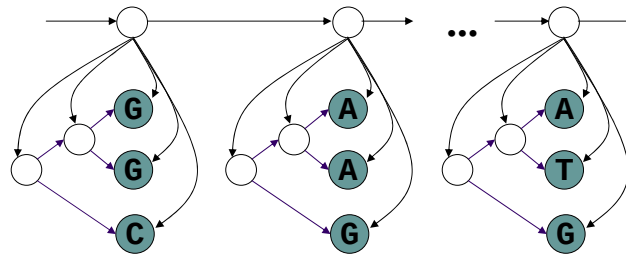
Iteration:

$$\beta_t^k = \sum_i a_{k,i} P(x_{t+1} | y_{t+1}^k = 1) \beta_{t+1}^i$$

Termination:

$$P(x) = \sum_k \alpha_1^k \beta_1^k$$

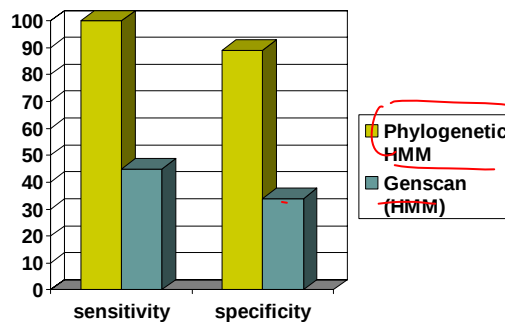
Hidden Markov Phylogeny



- this yields a gene finder that exploits evolutionary constraints

A Comparison of comparative genomic gene-finding and isolated gene-finding

- Based on sequence data from 12-15 primate species, McAuliffe et al (2003) obtained sensitivity of 100%, with a specificity of 89%.
 - Genscan (state-of-the-art gene finder) yield a sensitivity of 45%, with a specificity of 34%.



Open questions (philosophical)



Observation:

- Finding a good phylogeny will help in finding the genes.
- Finding the genes will help to find biologically meaningful phylogenetic trees

Which came first, the chicken or the egg?

Open questions (technical)



- How to learn a phylogeny (topology and transition prob.)?
- Should different site use the same phylogeny? Function-specific phylogeny?
- Other evolutionary events: duplication, rearrangement, lateral transfer, etc.

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