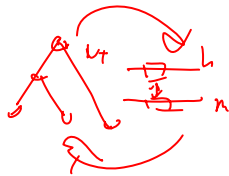


Computational Genomics

Molecular Evolution: nucleotide substitution models



Eric Xing

Lecture 13, February 27, 2007

Reading: DTW book, Chap 12
DEKM book, Chap 8



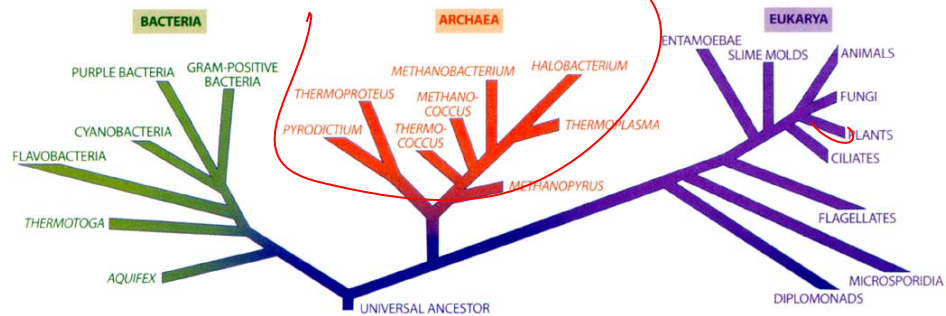
Some important dates in history (billions of years ago)

- | | |
|--------------------------------------|---------------------|
| • Origin of the universe | 15 ±4 |
| • Formation of the solar system | 4.6 |
| • First self-replicating system | 3.5 ±0.5 |
| • Prokaryotic-eukaryotic divergence | 1.8 ±0.3 |
| • Plant-animal divergence | 1.0 |
| • Invertebrate-vertebrate divergence | 0.5 |
| • Mammalian radiation beginning | 0.1 |

(86 CSH Doolittle et al.)



The three kingdoms

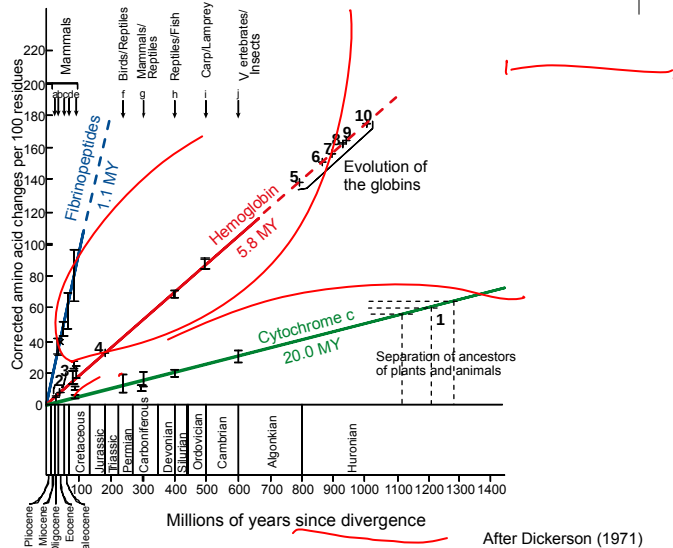


M. Madigan and B. Mairs, 1997

Two important early observations

- Different proteins evolve at different rates, and this seems more or less independent of the host organism, including its generation time.
- It is necessary to adjust the observed percent difference between two homologous proteins to get a distance more or less linearly related to the time since their common ancestor. (Later we offer a rational basis for doing this.)
- See next slide ...

Rates of macromolecular evolution



How does sequence variation arise?



- **Mutation:**
 - (a) Inherent: DNA replication errors are not always corrected.
 - (b) External: exposure to chemicals and radiation.
- **Selection:** Deleterious mutations are removed quickly. Neutral and rarely, advantageous mutations, are tolerated and stick around.
- **Fixation:** It takes time for a new variant to be established (having a stable frequency) in a population.

10^{-9} 10^{-13-15}

Modeling DNA base substitution



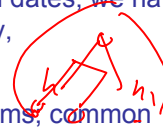
- Standard assumptions (sometimes weakened)
 - Site independence.
 - Site homogeneity.
 - Markovian: given current base, future substitutions independent of past.
 - Temporal homogeneity: stationary Markov chain.
- Strictly speaking, only applicable to regions undergoing little selection.



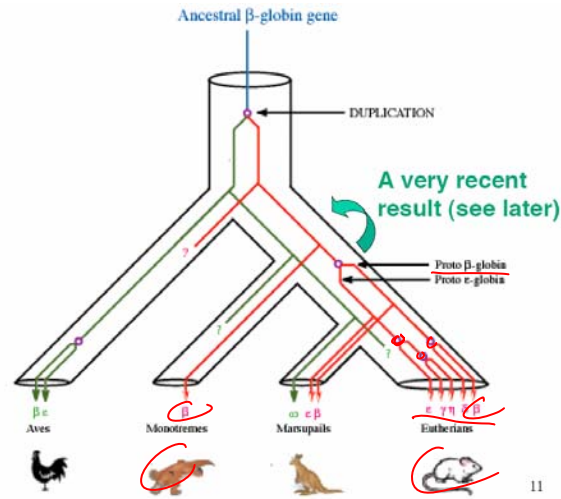
Some terminology



- In evolution, homology (here of proteins), means similarity due to common ancestry.
- A common mode of protein evolution is by duplication. Depending on the relations between duplication and speciation dates, we have two different types of homologous proteins. Loosely,
- Orthologues: the “same” gene in different organisms; common ancestry goes back to a speciation event.
- Paralogues: different genes in the same organism; common ancestry goes back to a gene duplication.
- Lateral gene transfer gives another form of homology.



Speciation vs. duplication



Beta-globins (orthologues)

		10	20	30	40
BG-human	MVHLT	E E K S A V T A L W G K V N V D E V G G E A L G R L L V V Y P W T Q			
BG-macaque	..	.. N .. T ..			
BG-bovine	.. M .. A ..	.. F .. K ..			
BG-platypus	.. S G G ..	.. N .. I N L ..			
BG-chicken	.. W .. Q L I G ..	.. A C A .. A .. I ..			
BG-shark	.. W S E V L H E I T T K S I D K H S L A K ..	.. A M F I .. T ..			
		50	60	70	80
BG-human	R F F E S F G D L S T P D A V M G N P K V K A H G K	V L G A F S D G L A H L D			
BG-macaque	..	.. S ..	..	.. N ..	
BG-bovine	..	.. A .. N ..	.. D S .. N M K ..		
BG-platypus	..	.. A .. S A G ..	.. A .. T S G A K N ..		
BG-chicken	..	.. A .. N .. S T I L .. M R ..	.. T S G A V K N ..		
BG-shark	..	.. Y G N L K E F T A C S Y G ..	.. E A .. T .. L G V A V T .. G		
		90	100	110	120
BG-human	N L K G T F A T L S E L H C D K L H V D P E N F R L L G N V L V C V L A H H F G				
BG-macaque	..	.. Q ..	.. K ..		
BG-bovine	..	.. A ..	.. K ..	.. V .. R N ..	
BG-platypus	..	.. K ..	.. N R ..	.. I V .. R .. S	
BG-chicken	..	.. I N .. S Q ..	..	.. D I .. I I .. A .. S	
BG-shark	..	.. D V S Q T D .. K K A E E ..	.. V S K .. A K C F V E G I L L K		
		130	140		
BG-human	K E F T P P V Q A A	Y Q K V V A G V A N A L A H K Y H			
BG-macaque	..	.. Q ..	..		
BG-bovine	..	.. V L .. F ..	.. R ..		
BG-platypus	..	.. D S E .. W .. L S .. H .. G ..			
BG-chicken	..	.. D .. E C .. W .. L R V .. H .. R ..			
BG-shark	..	.. D K A Q T .. W E Y F G V V D I S K E ..			

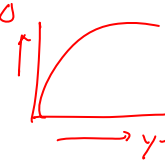
. means same as reference sequence
- means deletion

Beta-globins: uncorrected pairwise distances



- DISTANCES between protein sequences (calculated over: 1 to 147)
 - Below diagonal: observed number of differences
 - Above diagonal: number of differences per 100 amino acids

	hum	mac	bov	pla	chi	sha
hum	----	5	16	23	31	65
mac	7	----	17	23	30	62
bov	23	24	----	27	37	65
pla	34	34	39	----	29	64
chi	45	44	52	42	----	61
sha	91	88	91	90	87	----

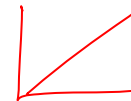


Beta-globins: corrected pairwise distances



- DISTANCES between protein sequences (calculated over: 1 to 147)
 - Below diagonal: observed number of differences
 - Above diagonal: number of differences per 100 amino acids
 - Correction method: Jukes-Cantor

	hum	mac	bov	pla	chi	sha
hum	----	5	17	27	37	108
mac	7	----	18	27	36	102
bov	23	24	----	32	46	110
pla	34	34	39	----	34	106
chi	45	44	52	42	----	98
sha	91	88	91	90	87	----



Human globins (paralogues)



		10	20	30	
alpha-human	-	V L S P A D K T N V K A A W G K V G A H A G E Y G A E A L E R M F L S F P T T			
beta-human	V H . T . E E . S A . T . L	N V D . V . G	G . L L V V Y . W .		
delta-human	V H . T . E E . . A . N . L	N V D A V . G	G . L L V V Y . W .		
epsilon-human	V H F T A E E . A A . T S L . S . M	N V E . A . G	G . L L V V Y . W .		
gamma-human	G H F T E E . . A T I T S L	N V E D A . G . T . G . L L V V Y . W .			
myo-human	- G . . D G E W Q L . L N V	E . D I P G H . Q . V . I . L . K G H . E .			

		40	50	60	70
alpha-human	K T Y F P H F - D L S H G S A - - - -	Q V K G H G K K V A D A L T N A V A H V			
beta-human	Q R F . E S . G	T P D . V M G N P K . . A	L G . F S D G L . . L		
delta-human	Q R F . E S . G	T P D . V M G N P K . . A	L G . F S D G L . . L		
epsilon-human	Q R F . D S . G N . . S P	I L G N P K . . A	L T S F G D . I K N M		
gamma-human	Q R F . D S . G N . . S A	I M G N P K . . A	L T S . G D . I K . L		
myo-human	L E K . D K . K H . K S E D E N K A S E D L . K	A T . L T . . G G I L K K K			

		80	90	100	110
alpha-human	D D M P N A L S A L S D L H A H K L R V D P V N F K L L S H C L L V T L A A H L				
beta-human	. N L K G T F A T . . E	C D . H	R . . G N V . V C V . . H . F		
delta-human	. N L K G T F . Q . . E	C D . H	R . . G N V . V C V . . R N F		
epsilon-human	. N L K P . F A K . . E	C D . H	R . . G N V M V I . . T . F		
gamma-human	. . L K G T F A Q . . E	C D . H	R . . G N V . V T V . . I . F		
myo-human	G H H E A E I K P . A Q S	T . H K I P V K Y L E F I . E . I I Q V . Q S K H			

		120	130	140	
alpha-human	P A E F T P A V H A S L D K F L A S V S T V L T S K Y R - - - -				
beta-human	G K	P . Q . A Y Q . V V . G . A N A . A H			
delta-human	G K	Q M Q . A Y Q . V V . G . A N A . A H			
epsilon-human	G K	E . Q . A W Q . L V S A . A I A . A H			
gamma-human	G K	E . Q	W V T A . A S A . S . R . H		
myo-human	. G D . G A D A Q G A M N . A . E L F R K D M A . N . K E L G F Q G				

Human globins: corrected pairwise distances

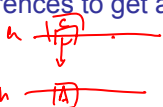


- DISTANCES between protein sequences (calculated over 1 to 141)
 - Below diagonal: observed number of differences
 - Above diagonal: estimated number of substitutions per 100 amino acids
 - Correction method: Jukes-Cantor

	alpha	beta	delta	epsil	gamma	myo
alpha	---	281	281	281	313	208
beta	82	---	7	30	31	1000
delta	82	10	---	34	33	470
epsil	89	35	39	---	21	402
gamma	85	39	42	29	---	470
myo	116	117	116	119	118	---

Correcting distances between DNA and protein sequences



- Why it is necessary to **adjust** observed percent differences to get a distance measure which scales linearly with time? 
- This is because we can have **multiple** and **back substitutions** at a given position along a lineage.
- All of the correction methods (with names like **Jukes-Cantor**, **2-parameter Kimura**, etc) are justified by simple probabilistic arguments involving Markov chains whose basis is worth mastering.
- The same molecular evolutionary models can be used in **scoring** sequence alignments.

Markov chain



- State space = {A, C, G, T}.

$$p(i,j) = \text{pr}(\text{next state } S_j \mid \text{current state } S_i)$$

- **Markov assumption:**

$$p(\text{next state } S_j \mid \text{current state } S_i \text{ \& any configuration of states before this}) = p(i,j)$$

Only the *present* state, not previous states, affects the probs of moving to next states.

The multiplication rule

$$\begin{aligned}
 & \text{pr}(\text{state after next is } S_k \mid \text{current state is } S_i) \\
 &= \sum_j \text{pr}(\text{state after next is } S_k, \text{next state is } S_j \mid \text{current state is } S_i) \quad [\text{addition rule}] \\
 &= \sum_j \text{pr}(\text{next state is } S_j \mid \text{current state is } S_i) \times \text{pr}(\text{state after next is } S_k \mid \text{current state is } S_j, \text{next state is } S_j) \quad [\text{multiplication rule}] \\
 &= \sum_j \bar{p}_{ij} \times p_{jk} \quad [\text{Markov assumption}] \\
 &= (i,k)\text{-element of } P^2, \text{ where } P=(p_{ij}).
 \end{aligned}$$

Handwritten note: P (after next) (here)

More generally,

$$\text{pr}(\text{state } t \text{ steps from now is } S_k \mid \text{current state is } S_i) = i,k \text{ element of } P^t$$

Continuous-time version

- For any (s, t) :
 - Let $p_{ij}(t) = \text{pr}(S_j \text{ at time } t+s \mid S_i \text{ at time } s)$ denote the stationary (time-homogeneous) transition probabilities.
- Let $P(t) = (p_{ij}(t))$ denote the matrix of $p_{ij}(t)$'s.
 - Then for any (t, u) : $P(t+u) = P(t)P(u)$.
- It follows that $P(t) = \exp(Qt)$, where $Q = P'(0)$ (the derivative of $P(t)$ at $t = 0$).
- Q is called the infinitesimal matrix (transition rate matrix) of $P(t)$, and satisfies

$$P'(t) = QP(t) = P(t)Q.$$
- Important approximation: when t is small,

$$P(t) \approx I + Qt.$$

$$e^x = 1 + x$$

$$\boxed{r} = \boxed{r} + \frac{\boxed{r}}{t} \cdot t$$

Interpretation of Q

- Roughly, q_{ij} is the **rate** of transitions of i to j , while $q_{ii} = -\sum_{j \neq i} q_{ij}$, so each row sum is 0 (Why?).
- Now we have the short-time approximation:

$$p_{i \neq j}(t+h) = q_{ij}h + o(h)$$

$$p_{i=j}(t+h) = 1 + q_{ii}h + o(h)$$

where $p_{ij}(t+h)$ is the probability of transitioning from i at time t to j at time $t+h$

- Now consider the Chapman-Kolmogorov relation: (assuming we have a continuous-time Markov chain, and let $p_j(t) = \text{pr}(S_j \text{ at time } t)$):

$$\begin{aligned} p_j(t+h) &= \sum_i \text{pr}(S_i \text{ at } t, S_j \text{ at } t+h) \\ &= \sum_i \text{pr}(S_i \text{ at } t) \text{pr}(S_j \text{ at } t+h | S_i \text{ at } t) \\ &= p_j(t) \times (1 + q_{jj}h) + \sum_{i \neq j} p_i(t) \times h q_{ij} \end{aligned}$$

i.e., $h^{-1}(p_j(t+h) - p_j(t)) = p_j(t)q_{jj} + \sum_{i \neq j} p_i(t)q_{ij}$, which becomes: $p' = Qp$ as $h \rightarrow 0$.

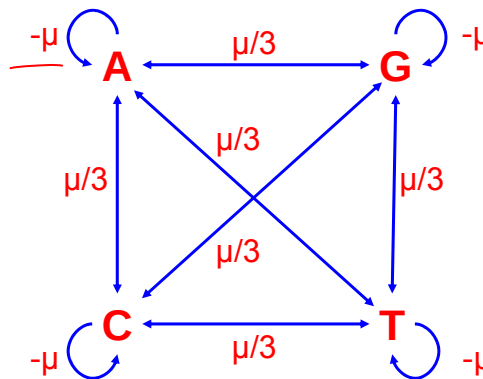
Probabilistic models for DNA changes

Orc:
Elf:
Dwarf:
Hobbit:
Human:

ACAGTGACGCCCCAAACGT
ACAGTGACGCTACAAACGT
CCTGTGACGTAACAAACGA
CCTGTGACGTAGCAAACGA
CCTGTGACGTAGCAAACGA

The Jukes-Cantor model (1969)

- Substitution rate:



$$P' = \alpha P$$

$$P_{t_0} = \exp(\alpha t)$$

the simplest symmetrical model for DNA evolution

Transition probabilities under the Jukes-Cantor model

- IID assumption:
 - All sites change independently
 - All sites have the same stochastic process working at them
- Equiprobability assumption:
 - Make up a fictional kind of event, such that when it happens the site changes to one of the 4 bases chosen at random equiprobably
- Equilibrium condition:
 - No matter how many of these fictional events occur, provided it is not zero, the chance of ending up at a particular base is 1/4.
- Solving differentially equation system $P' = QP$

$$\begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix}$$

$$\begin{bmatrix} 1/4 \\ 1/4 \\ 1/4 \\ 1/4 \end{bmatrix} \leftarrow$$

$$\pi = P_{\pi}$$

Transition probabilities under the Jukes-Cantor model (cont.)



- Prob transition matrix:

$$P(t) = \begin{matrix} & \begin{matrix} A & C & G & T \end{matrix} \\ \begin{matrix} A \\ C \\ G \\ T \end{matrix} & \begin{pmatrix} r(t) & s(t) & s(t) & s(t) \\ s(t) & r(t) & s(t) & s(t) \\ s(t) & s(t) & r(t) & s(t) \\ s(t) & s(t) & s(t) & r(t) \end{pmatrix} \end{matrix}$$



Where we can derive:

$$r(t) = \frac{1}{4} \left(1 + 3e^{-\frac{4}{3}\mu t} \right)$$

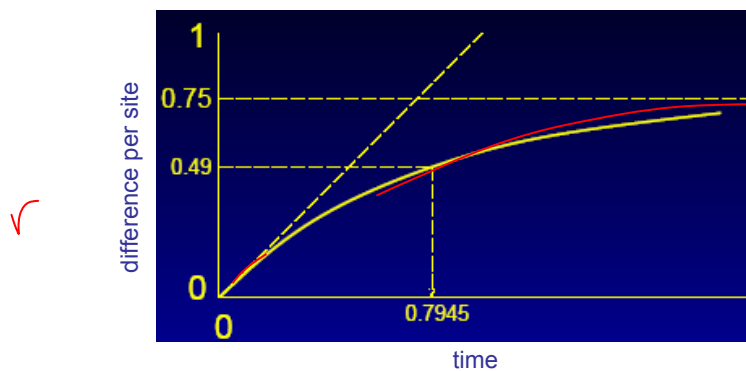
$$s(t) = \frac{1}{4} \left(1 - e^{-\frac{4}{3}\mu t} \right)$$

Homework!

Jukes-Cantor (cont.)

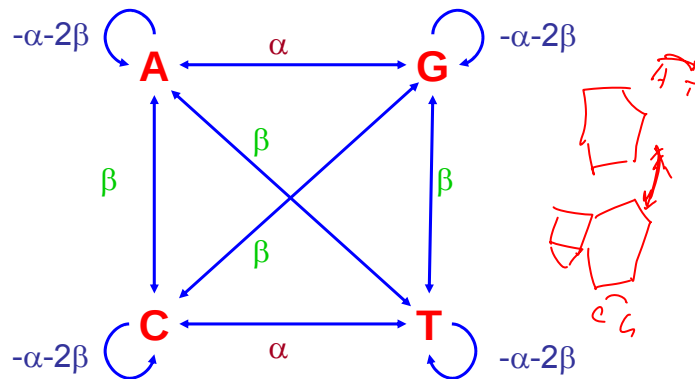


- Fraction of sites differences



Kimura's K2P model (1980)

- Substitution rate:



- which allows for different rates of transition and transversions.
- Transitions (rate α) are much more likely than transversions (rate β).

Kimura (cont.)

- Prob transition matrix:

$$P(t) = \begin{pmatrix} r(t) & s(t) & u(t) & s(t) \\ s(t) & r(t) & s(t) & u(t) \\ u(t) & s(t) & r(t) & s(t) \\ s(t) & u(t) & s(t) & r(t) \end{pmatrix}$$

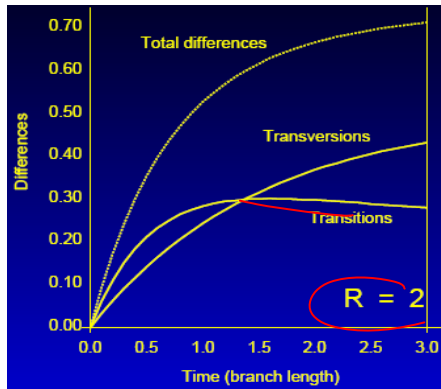
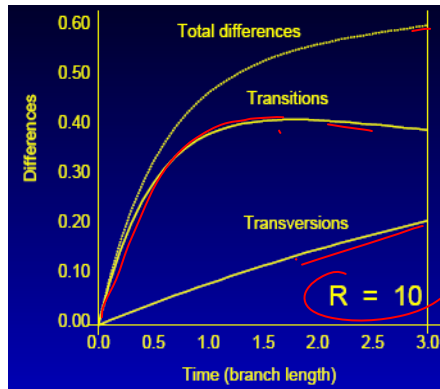
Where

$$\begin{aligned} s(t) &= \frac{1}{4} (1 - e^{-4\beta t}) \\ u(t) &= \frac{1}{4} (1 + e^{-4\beta t} - e^{-2(\alpha+\beta)t}) \\ r(t) &= 1 - 2s(t) - u(t) \end{aligned}$$

- By proper choice of α and β one can achieve the overall rate of change and Ts=Tn ratio R you want (warning: terminological tangle).

Kimura (cont.)

- Transitions, transversions expected under different R:



Other commonly used models

- Two models that specify the **equilibrium base frequencies** (you provide the frequencies A; C; G; T and they are set up to have an equilibrium which achieves them), and also let you control the transition/transversion ratio:
- The **Hasegawa-Kishino-Yano (1985) model**:

to : from :	A	G	C	T
A	—	$\alpha\pi_G + \beta\pi_G$	$\alpha\pi_C$	$\alpha\pi_T$
G	$\alpha\pi_A + \beta\pi_A$	—	$\alpha\pi_C$	$\alpha\pi_T$
C	$\alpha\pi_A$	$\alpha\pi_G$	—	$\alpha\pi_T + \beta\pi_T$
T	$\alpha\pi_A$	$\alpha\pi_G$	$\alpha\pi_C + \beta\pi_C$	—

Other commonly used models



- The **F84 model** (Felsenstein)

to : from :	A	G	C	T
A	—	$\alpha\pi_G + \beta\frac{\pi_G}{\pi_R}$	$\alpha\pi_C$	$\alpha\pi_T$
G	$\alpha\pi_A + \beta\frac{\pi_A}{\pi_R}$	—	$\alpha\pi_C$	$\alpha\pi_T$
C	$\alpha\pi_A$	$\alpha\pi_G$	—	$\alpha\pi_T + \frac{\beta\pi_T}{\pi_Y}$
T	$\alpha\pi_A$	$\alpha\pi_G$	$\alpha\pi_C + \beta\frac{\pi_C}{\pi_Y}$	—

- where $\pi_R = \pi_A + \pi_G$ and $\pi_Y = \pi_C + \pi_T$ (The equilibrium frequencies of purines and pyrimidines)

The general time-reversible model

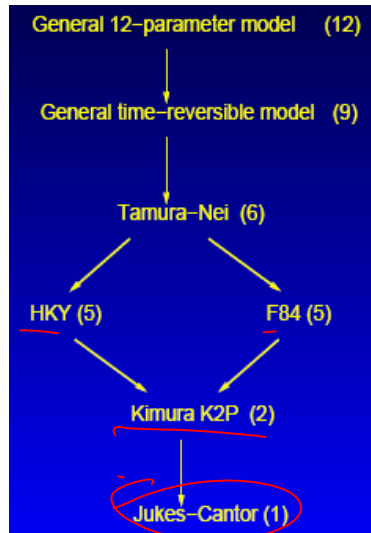


- It maintains "detailed balance" so that the probability of starting at (say) A and ending at (say) T in evolution is the same as the probability of starting at T and ending at A:

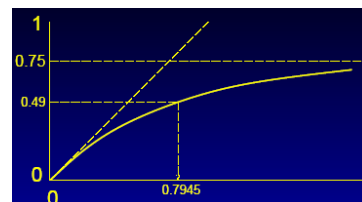
	A	C	G	T
A	—	$\alpha\pi_C$	$\beta\pi_G$	$\gamma\pi_T$
C	$\alpha\pi_A$	—	$\delta\pi_G$	$\epsilon\pi_T$
G	$\beta\pi_A$	$\delta\pi_C$	—	$\nu\pi_T$
T	$\gamma\pi_A$	$\epsilon\pi_C$	$\nu\pi_G$	—

- And there is of course the general 12-parameter model which has arbitrary rates for each of the 12 possible changes (from each of the 4 nucleotides to each of the 3 others).
- (Neither of these has formulas for the transition probabilities, but those can be done numerically.)

Relation between models



Adjusting evolutionary distance using base-substitution model



The Jukes-Cantor model



Common ancestor of human and orang

t time unit

Human (now)

$Q =$

$$\begin{bmatrix} -3\alpha & \alpha & \alpha & \alpha \\ \alpha & -3\alpha & \alpha & \alpha \\ \alpha & \alpha & -3\alpha & \alpha \\ \alpha & \alpha & \alpha & -3\alpha \end{bmatrix}$$

$P =$

$$\begin{bmatrix} r & s & s & s \\ s & r & s & s \\ s & s & r & s \\ s & s & s & r \end{bmatrix}$$

Consider e.g. the 2nd position in a-globin2 Alu1.

$$r = (1 + 3e^{-4\alpha t})/4, \quad s = (1 - e^{-4\alpha t})/4.$$

Definition of PAM



- Let $P(t) = \exp(Qt)$. Then the A,G element of $P(t)$ is

$$pr(G \text{ now} | A \text{ then}) = (1 - e^{-4\alpha t})/4.$$

- Same for all pairs of different nucleotides.
- Overall rate of change $k = 3\alpha t$.



- PAM = accepted point mutation**

- When $k = .01$ described as 1 PAM
- Put $t = .01/3\alpha = 1/300\alpha$. Then the resulting $P = P(1/300\alpha)$ is called the PAM(1) matrix.

$$P(t),$$

- Why use PAMs?

Evolutionary time, PAM



- Since sequences evolve at different rates, it is convenient to rescale time so that 1 PAM of evolutionary time corresponds to 1% expected substitutions.
- For Jukes-Cantor, $k = 3\alpha t$ is the expected number of substitutions in $[0, t]$, so is a distance. (Show this.)
 - Set $3\alpha t = 1/100$, or $t = 1/300\alpha$, so 1 PAM = $1/300\alpha$ years.

Distance adjustment



- For a pair of sequences, $k = 3\alpha t$ is the desired metric, but not observable. Instead, $pr(\text{different})$ is observed. So we use a model to convert $pr(\text{different})$ to k .

- This is completely analogous to the conversion of

$$\theta = pr(\text{recombination})$$

to genetic (map) distance (= expected number of crossovers) using the Haldane map function

$$\theta = 1/2 \times (1 - e^{-2d}),$$

assuming the no-interference (Poisson) model.

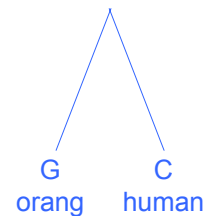


Towards Jukes-Cantor adjustment



- E.g., 2nd position in a-globin Alu 1
- Assume that the common ancestor has A, G, C or T with probability 1/4.

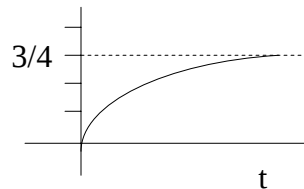
common ancestor



- Then the chance of the nt differing

$$p_{\neq} = \frac{3}{4} \times (1 - e^{-8\alpha t})$$

$$= \frac{3}{4} \times (1 - e^{-4k/3}), \text{ since } k = 2 \times 3\alpha t$$



Jukes-Cantor adjustment



- If we suppose all **nucleotide** positions behave identically and independently, and $n_{\neq}$ differ out of n , we can invert this, obtaining

$$\hat{k} = -\frac{3}{4} \times \log\left(1 - \frac{4}{3} \frac{n_{\neq}}{n}\right)$$

- This is the **corrected** or **adjusted** fraction of differences (under this simple model). $\times 100$ to get PAMs
- The analogous simple model for **amino** acid sequences has

$$\hat{k} = -\frac{19}{20} \times \log\left(1 - \frac{20}{19} \frac{n_{\neq}}{n}\right)$$

$\times 100$ for PAM.

Illustration



1. **Human** and bovine beta-globins are aligned with no deletions at 145 out of 147 sites. They differ at 23 of these sites. Thus $n_{\neq}/n = 23/145$, and the corrected distance using the Jukes-Cantor formula is (natural logs)

$$- 19/20 \times \log(1 - 20/19 \times 23/145) = 17.3 \times 10^{-2}.$$

2. The **human** and **gorilla** sequences are aligned without gaps across all 300 bp, and differ at 14 sites. Thus $n_{\neq}/n = 14/300$, and the corrected distance using the Jukes-Cantor formula is

$$- 3/4 \times \log(1 - 4/3 \times 14/300) = 4.8 \times 10^{-2}.$$

Correspondence between observed a.a. differences and the evolutionary distance (Dayhoff et al., 1978)

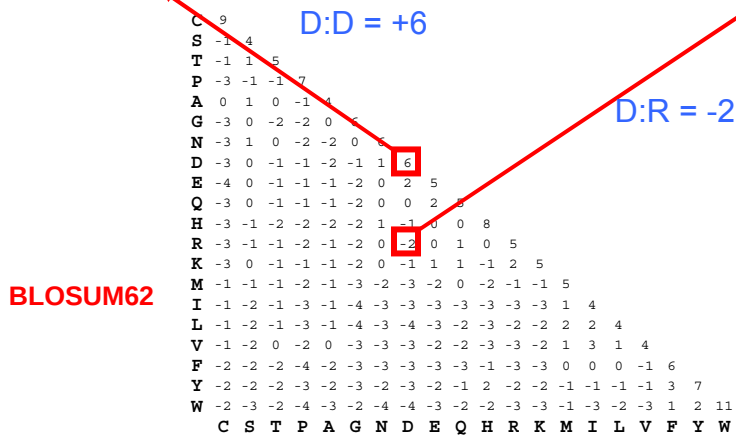


Observed Percent Difference	Evolutionary Distance in PAMs
1	1
5	5
10	11
15	17
20	23
25	30
30	38
35	47
40	56
45	67
50	80
55	94
60	112
65	133
70	159
75	195
80	246
85	328

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60	112
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70	159
75	195
80	246
85	328



134 LQQGELDLVMTSDILPRSELHYSPMFD FEVRLVLAPDHPLASKTQITPEDLASETLLI
137 LDSNSVDLVLMGVPPRNVEVEAEAFMDNPLVVIAPDPHPLAGERAISLARLAEETFVM



From Henikoff 1996

Statistical motivation for alignment scores



Alignment: AGCTGATCA...
AACCGGTTA... Hypotheses: H = homologous (indep. sites, Jukes-Cantor)
R = random (indep. sites, equal freq.)

$$\text{pr}(\text{data} | H) = \text{pr}(AA | H)\text{pr}(GA | H)\text{pr}(CC | H)...$$

$$= (1-p)^a p^d, \text{ where } a = \# \text{agreements}, d = \# \text{disagreements}, p = \frac{3}{4}(1 - e^{-8\alpha t}).$$

$$\text{pr}(\text{data} | R) = \text{pr}(AA | R)\text{pr}(GA | R)\text{pr}(CC | R)...$$

$$= \left(\frac{1}{4}\right)^a \left(\frac{3}{4}\right)^d$$

$$\Rightarrow \log \left\{ \frac{\text{pr}(\text{data} | H)}{\text{pr}(\text{data} | R)} \right\} = a \log \frac{1-p}{1/4} + d \log \frac{p}{3/4} = a \times \sigma + d \times (-\mu).$$

- Since $p < 3/4$, $\sigma = \log((1-p)/(1/4)) > 0$, while $-\mu = \log(p/(3/4)) < 0$.
- Thus the alignment score = $a \times \sigma + d \times (-\mu)$, where the match score $\sigma > 0$, and the mismatch penalty is $-\mu < 0$.

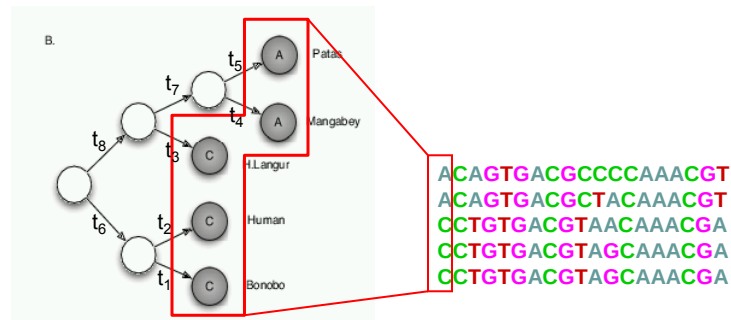
Large and small evolutionary distances



- Recall that
 - $p = (3/4)(1 - e^{-8\alpha t})$,
 - $\sigma = \log((1-p)/(1/4))$,
 - $-\mu = \log(p/(3/4))$.
- Now note that if $\alpha t \approx 0$,
 - then $p \approx 6\alpha t$, and $1-p \approx 1$, and so $\sigma \approx \log 4$, while $-\mu \approx \log 8\alpha t$ is large and negative.
 - That is, we see a big difference in the two values of σ and μ for small distances.
- Conversely, if αt is large,
 - $p = (3/4)(1 - \epsilon)$, hence $p/(3/4) = 1 - \epsilon$, giving $\mu = -\log(1 - \epsilon) \approx \epsilon$, while $1-p = (1+3\epsilon)/4$, $(1-p)/(1/4) = 1+3\epsilon$, and so $\sigma = \log(1+3\epsilon) \approx 3\epsilon$.
 - Thus the scores are about 3 (for a match) to 1 (for a mismatch) for large distances. This makes sense, as mismatches will on average be about 3 times more frequent than matches.
- the matrix which performs best will be the matrix that reflects the evolutionary separation of the sequences being aligned.

What about multiple alignment

- Phylogenetic methods: a tree, with branch lengths, and the data at a single site.



- See next lecture for how to compute likelihood under this hypothesis

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